



# Consistency of seismic hazard estimates from a physics-based earthquake simulator: a case study in south-eastern Spain

Octavi Gómez-Novell<sup>1</sup>, Francesco Visini<sup>2</sup>, Paula Herrero-Barbero<sup>1</sup>, José A. Álvarez-Gómez<sup>3</sup>, Julián García-Mayordomo<sup>1</sup>

<sup>1</sup>CN Instituto Geológico y Minero de España, CSIC, Madrid, Spain

<sup>2</sup>Istituto Nazionale di Geofisica e Vulcanologia, Chieti, Italy

<sup>3</sup>Universidad Complutense de Madrid, Madrid, Spain

Correspondence to: [ogomez.novell@igme.es](mailto:ogomez.novell@igme.es)

## Abstract

Estimating seismic hazard is crucial for enhancing societal resilience and risk mitigation strategies. Probabilistic Seismic Hazard Analysis (PSHA) is the current standard framework, traditionally relying on empirical earthquake rupture forecasts (ERFs) and ground-motion models. In this framework, physics-based earthquake cycle simulators are emerging as powerful tools in PSHA, capable of replicating observed seismicity and seismic hazard statistics. Here, we present a quantitative consistency evaluation of physics-based PSHA models at the Eastern Betic Shear Zone in south-eastern Spain against both historical macroseismic intensity data and instrumental ground-shaking records. We use synthetic catalogues from RSQSim earthquake-cycle simulations to construct two physics-based ERFs that we pipeline into a PSHA calculation. Results indicate that the physics-based ERFs derived from the best-performing simulation model, previously ranked against empirical benchmarks, achieve the best overall agreement with observed macroseismic intensities and acceleration records at 10 sites, outperforming both the lower-performing simulation and a traditional area-source model. Our findings highlight that the incorporation of physics-based models into a PSHA framework is reliable, enabling the inclusion of rupture behaviour and fault-system interactions, key challenges in PSHA. We also advocate for the complementarity of physics-based models with traditional approaches in PSHA to better capture epistemic uncertainties in the hazard representation.

## 1 Introduction

Estimating seismic hazard is essential for enhancing societal resilience, guiding building codes, informing insurance companies and supporting risk assessments and public policies aimed at mitigating earthquake impacts. Probabilistic seismic hazard analysis (PSHA), formalized by Cornell (1968), remains the standard framework. It combines an earthquake rupture forecast (ERF, sometimes also called seismicity model) with empirical ground-motion models (GMMs) to estimate the probability of exceeding different shaking intensity measures.



Commonly, an ERF is made of two components: a distributed seismicity model, or area source model, describing the occurrence of earthquakes not associated with well-identified fault sources, and, when sufficient information for fault characterization is available, a fault source model. (e.g., Gerstenberger et al. 2020, Visini et al., 2021; Danciu et al., 2024). The use of active faults in seismic hazard assessments is well-established in high-strain regions such as California (Field et al., 2014), the Marmara Region in Turkey (for example, Chartier et al., 2021), and New Zealand (Gerstenberger et al., 2024). Fault-based models have also been applied successfully in moderate-to-low strain areas like, for example, southeast Spain, Italy, France, and Malawi (e.g., García-Mayordomo et al., 2007; Scotti et al., 2014; Gómez-Novell et al., 2020a; Valentini et al., 2017; Rivas-Medina et al., 2018; Williams et al., 2023). One of the advantages of considering fault sources is the incorporation of long-term fault-slip rates, which provide insights into cumulative earthquake displacements over multiple seismic cycles and help avoid biases arising from short-term variability. These rates are directly linked to earthquake recurrence through geological moment accumulation (Brune, 1968). Conventional models often simplify this relationship by assuming complete seismic release of the accumulated moment, leading to uniform slip-rate estimates on faults, while also considering faults as independent sources. Geological observations, however, reveal pronounced spatial variability in fault displacement (e.g., Hergert & Heidbach, 2010; Reitman et al., 2022; Roberts, 2007), controlled by factors such as fault geometry, interaction or off-fault damage (e.g., Cappa et al., 2014; Iezzi et al., 2018; Nicol et al., 1996) that strongly influence earthquake recurrence and ground-shaking estimates (e.g., Cowie et al., 2012; Faure Walker et al., 2019; Wang & Barbot, 2024; Rodríguez-Picada et al., 2025). Moreover, observations show that faults do not always rupture independently; instead, multi-fault ruptures frequently occur in large earthquakes, as exemplified by the 1992 Mw 7.1 Landers, 2002 Mw 7.9 Denali, and 2016 Mw 7.8 Kaikōura earthquakes.

Modern fault-based PSHA models now accommodate multi-fault ruptures, as their consideration allows for more accurate representations of the seismic hazard. The Uniform California Earthquake Rupture Forecast (UCERF3; Field et al., 2014), followed by the 2022 Aotearoa New Zealand National Seismic Hazard Model (Gerstenberger et al., 2024), pioneered a new methodology by exploring all plausible fault section-to-section ruptures, moving beyond more classical rigid segmentation approaches and allowing for more accurate representations of fault system behavior in PSHA. In this line, Visini et al. (2020) compared alternative approaches for modeling multi-fault ruptures in PSHA and demonstrated that incorporating finite-to-finite ruptures and spatial slip-rate variability significantly affects hazard estimates, further emphasizing the importance of fault-system complexity.

Beyond fault-based models, physics-based earthquake cycle modeling is emerging as a powerful complementary tool for seismic hazard assessment (e.g., Rafiei et al., 2022 in Alborz, Khodaverdian et al., 2016 in NW Iran, or the work by Shaw et al., 2018 in the US). These earthquake simulators inherently account for fault friction and stress interactions to model the physical processes governing earthquake rupture. This allows them to reproduce the spatiotemporal patterns of slip variability within fault systems and multi-fault interactions without requiring these behaviours to be predefined as inputs. Published



studies demonstrate that in many regions of active faulting the magnitude-frequency distributions of simulated earthquakes are comparable to the ones obtained with historical and instrumental earthquakes (e.g., Herrero-Barbero et al., 2021; Shaw et al., 2022), while simulated spatiotemporal rupture patterns are consistent with paleoseismic observations (e.g., Howarth et al., 2021). These similarities support the use of earthquake simulators for either estimating seismicity models or for constraining parameter inputs to seismic hazard models. Shaw et al. (2018) found that RSQSim simulations for the Californian fault system can replicate the seismic hazard statistics from the long-term UCERF3. Similarly, Milner et al. (2021) employed RSQSim to establish a nonergodic framework for probabilistic seismic-hazard analysis, while Chartier et al. (2021) utilized RSQSim as an auxiliary tool for weighting a logic tree that explores seismicity uncertainties in the Sea of Marmara. Despite this, efforts to validate physics-based models are ongoing as their use for real-world hazard applications requires the ability to reproduce natural behaviour and relations observed during earthquakes. Recently, Gómez-Novell et al. (2025a) proposed a benchmarking approach using empirical scaling laws, magnitude-frequency distributions, and paleoearthquake data to rank simulations.

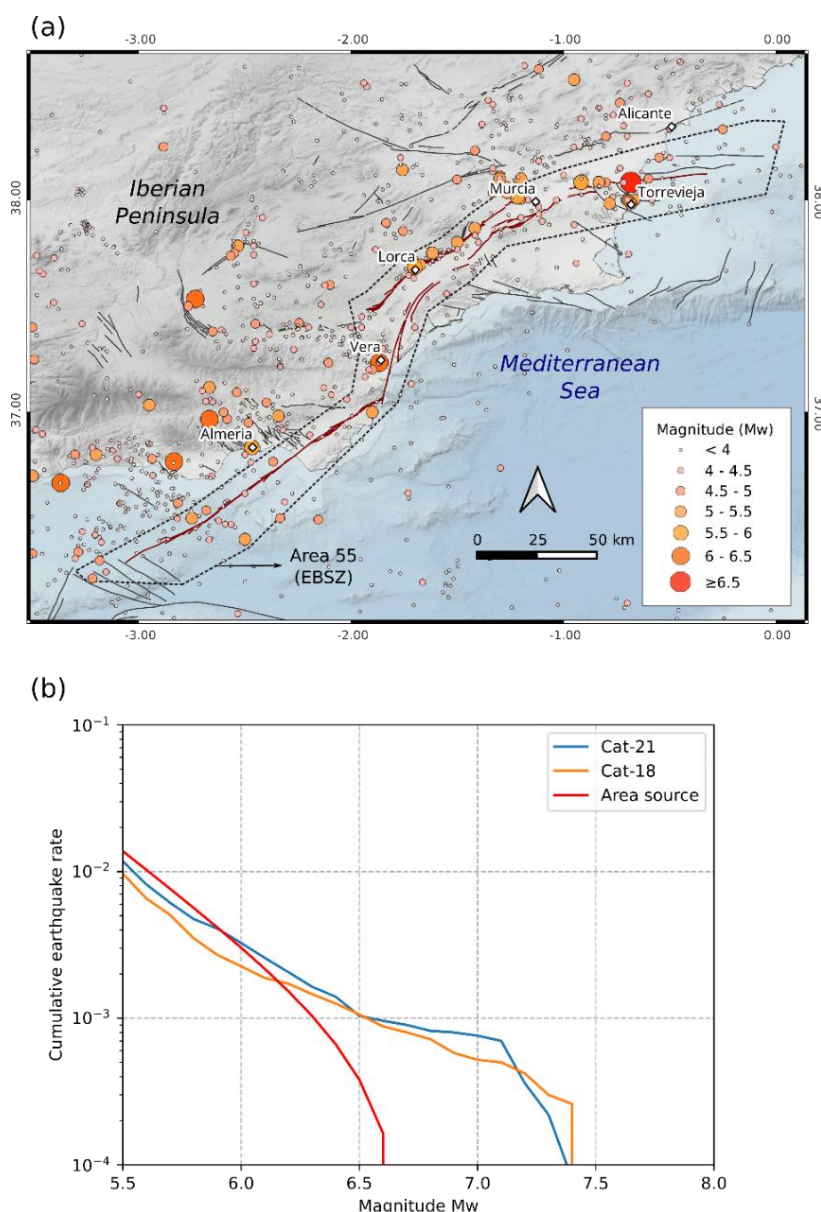
Here we use synthetic earthquake catalogues from multi-cycle simulations by Herrero-Barbero et al. (2021) for the Eastern Betic Shear Zone (EBSZ), a low-to-moderate strain active fault system in southern Spain (Figure 1). We pipeline simulated ruptures into ground shaking using GMMs, which are then used to produce PSHA models for the region. We evaluate the predictive performance of these physics-based PSHA models against both historical macroseismic intensity data and instrumental ground-shaking records. Figure 2 illustrates the workflow from simulated rupture sets to PSHA and model testing: a pipeline of simulated earthquake ruptures into PSHA and a consistency testing approach.

To our knowledge, this study represents one of the first attempts to evaluate consistency of physics-based PSHA against observations in Europe, offering a potential step forward for the integration of simulation-based approaches into the future of seismic hazard assessment, not only in Europe, but in a worldwide context. The results presented in this study are encouraging and highlight the potential of physics-based earthquake cycle models as a robust and viable alternative in the suite of fault-based approaches for PSHA. In this line, physics-based based models can also be relevant to explore the high-impact, low-probability events, as they can have severe consequences in a region and require efficient decision-making strategies under uncertainty (Ellingwood & Wen, 2005). While the modeling framework could be further enhanced by incorporating additional components, such as novel simulators like MCQSim (Zielke and Mai, 2023) or advanced physics-based ground motion models like CyberShake (Graves et al., 2011), this work represents a first systematic attempt to evaluate consistency of PSHA outcomes from an earthquake simulator against empirical observations, supporting the continued development and application of such approaches in seismic hazard analysis.

## 2. Data and methods

The methodological core of this work consists of two steps (Figure 2). First, we develop a method that pipelines RSQSim earthquake cycle simulation outputs into seismic hazard calculations. Second, we perform a quantitative consistency analysis

of the simulation-based seismic hazard statistics with macroseismic and instrumental observations. We deploy this approach  
 95 in southern Spain to test the applicability and consistency of physics-based models for PSHA in any adapted framework. The  
 method is generalizable to any output of physics-based earthquake cycle simulators, not only RSQSim.



**Figure 1:** (a) Map location of the Eastern Betic Shear Zone and seismicity from the IGN (2025) seismic catalogue, comprising from  
 year 1048 to 2025. Fault traces are from the QAFI v4 database (IGME, 2022) and area source 55 from the ZESIS database (IGME,  
 2015). (b) Magnitude-frequency distributions for the two synthetic physics-based earthquake catalogues explored in this study (Cat-  
 21 and Cat-18 from Herrero-Barbero et al., 2021) and for the area source 55 posterior to re-cluster correction (see text for details).



## 2.1 The RSQSim simulator

RSQSim is a boundary element model developed by Richards-Dinger and Dieterich (2012) and based on the rate-and-state friction law introduced by Dieterich (1979). RSQSim incorporates four primary approximations to elasto-dynamics: (1) the seismic cycle is defined by three states: healing, nucleation and, seismic rupture; (2) stress interactions between elements are quasi-static and elastic, neglecting dynamic stress transfer; (3) seismic sliding occurs at a fixed slip rate; and (4) the seismic rupture is resolved by using a quasi-dynamic approximation. These simplifications enable a computationally efficient simulation of long seismic sequences by allowing adaptive time stepping only during state changes, significantly reducing computational cost. Although RSQSim is not equivalent to fully dynamic models, previous studies (e.g., Richards-Dinger and Dieterich, 2012) show that it can reproduce key features of earthquake physics, including clustering and delayed nucleation, yielding realistic emergent seismic sequences.

## 2.2 The Eastern Betic Shear Zone fault model and RSQSim catalogues

For our study, we use the synthetic RSQSim catalogues of the Eastern Betic Shear Zone (EBSZ) from Herrero-Barbero et al. (2021). The EBSZ (Figure 1) is an active fault system located in the Eastern Betic Cordillera of the southeastern Iberian Peninsula. The fault system represents the onshore continuation of the Trans-Alboran Shear Zone (De Larouzière et al., 1988) and it is primarily composed of strike-slip faults, many of which also exhibit reverse slip component.

### 2.2.1 Fault model

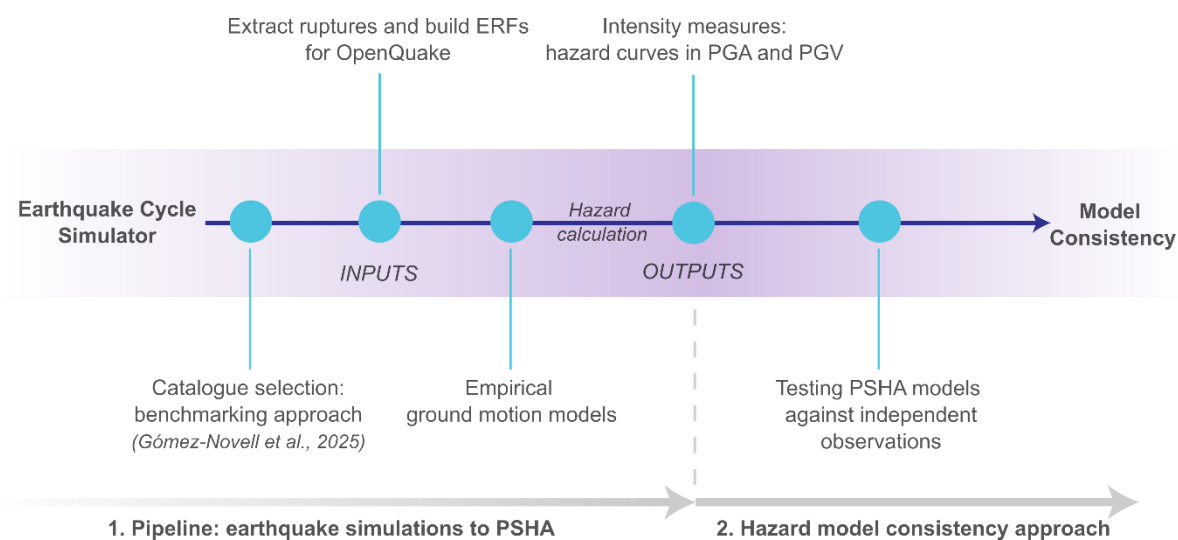
The EBSZ, as modeled by Herrero-Barbero et al. (2021), consists of six main seismogenic faults. This fault system extends over approximately 400 km and comprises 16 oblique-reverse, left-lateral strike-slip fault segments, with seismogenic depths ranging between 8 and 12 km (Figure 3). The Herrero-Barbero et al. (2021) model reproduces the geometric characteristics of the EBSZ faults both along strike and in the down-dip direction, based on the main geometrical and segmentation characteristics compiled in the Quaternary Active Faults Database of Iberia - QAFI - database (IGME, 2022; García-Mayordomo et al., 2017). Fault segmentation within the EBSZ is mainly constrained by fault zone geometry, seismicity patterns, and variations in the hanging wall topography, while seismogenic crustal depths are constrained between 8 and 12 km, consistent with seismotectonic studies from southeastern Spain (see Herrero-Barbero et al., 2021).

For modeling purposes, the faults were simplified from the QAFI database to a single fault trace, smoothing out small-scale complexities such as secondary branches, relay zones, and horsetail splays, while preserving first-order along-strike bends and variations in dip angle at depth, as these are relevant for the overall fault geometry at the scale of the simulation. Fault surfaces are discretized into triangular elements of  $\sim 1 \text{ km}^2$ . Geological fault slip rates and kinematics (rake angle) are assigned based on published literature data available for each one of the faults. Detailed information on fault geometries, slip rates, and rake assigned to each fault can be found in Herrero-Barbero et al. (2021).

## 2.2.2 Synthetic catalogues and model selection

The first step of our approach is model selection (Figure 2). For our analysis we use two of the eleven RSQSim synthetic earthquake catalogues computed by Herrero-Barbero et al. (2021) as part of a sensitivity analysis on input parameters: Cat-21 and Cat-18. The models share the same fault model but have different frictional input parameters, namely initial normal and shear stresses, and a-b coefficients of the rate and state law (Table 1). All catalogues are 100 kyr long with earthquakes of  $M_w > 4$ . We select Cat-21 and Cat-18 (Figure 1b) as the best and worst performing catalogues, respectively, according to the benchmarking method developed by Gómez-Novell et al. (2025a), which proposes a ranking scheme based on how well each catalogue matches four reference metrics: rupture area–magnitude and average displacement–magnitude scaling laws, agreement with a target magnitude-frequency distribution (magnitude-frequency distribution, with  $b = 1.0$ ), and consistency with paleoearthquake data. We select these end-member models of the EBSZ benchmarking ranks to evaluate the sensitivity of seismic hazard results to catalogue quality.

For each catalogue we discard the first 10 kyr to ensure that the modeling has evolved from its initial conditions and that the seismic cycle is stable. We then analyse the following 50 kyr of ruptures from each catalogue. The catalogue duration is determined after a sensitivity analysis on earthquake recurrence variability (see figure S1), which demonstrates that 50 kyr-long sub-catalogues are enough to produce stable earthquake rates and inter-event times compared to the whole 100 kyr catalogue. Such a shorter rupture catalogue also reduces computational demand on the subsequent hazard models without impacting the hazard estimates.



**Figure 2: Workflow (pipeline) of this study. This study combines a multicycle earthquake rupture simulator with empirical ground motion models. Probabilistic seismic hazard outputs are compared against macroseismic and instrumental observations.**





**Table 1: Input parameters of the selected RSQSim models from Herrero-Barbero et al. (2021) at the Eastern Betic Shear Zone.**

Model	Initial normal stress	Initial shear stress	a	b
Cat-21	20 MPa/km	60 MPa	0.001	0.006
Cat-18	140 MPa	60 MPa	0.005	0.01

### 2.3 Pipeline to integrate simulated earthquake ruptures into probabilistic seismic hazard assessment

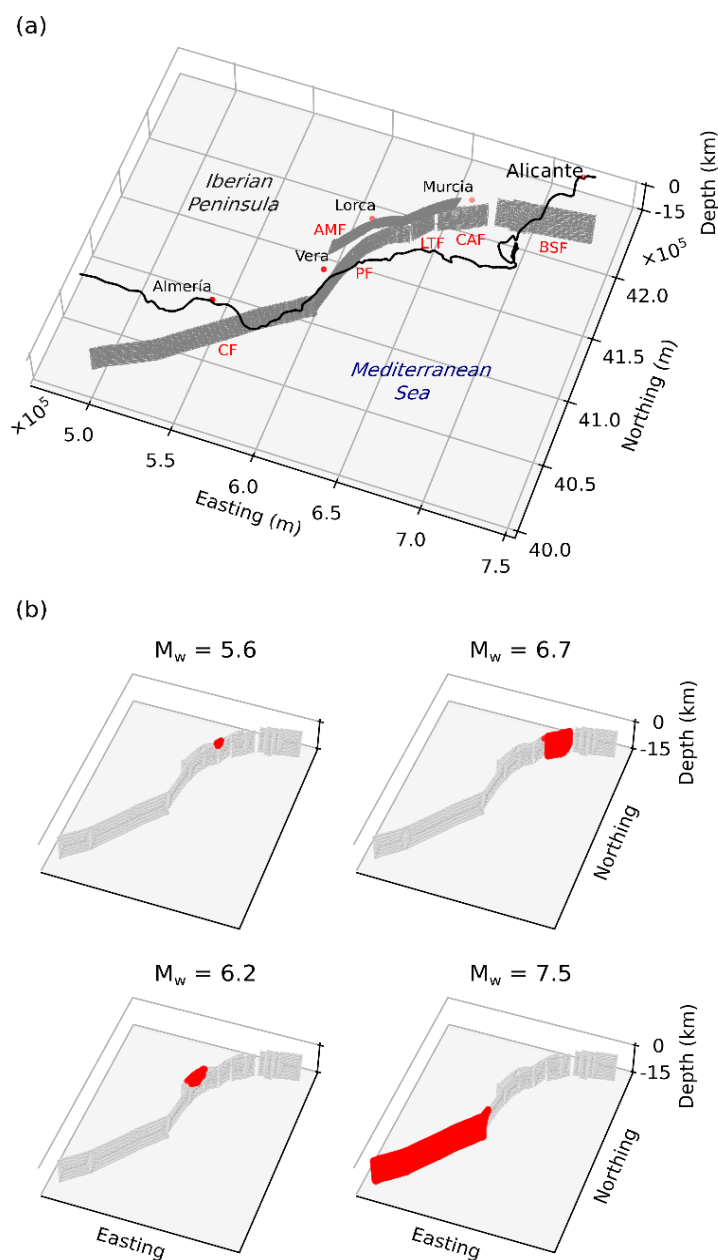
For each simulated catalogue (Cat-21 and Cat-18), we consider synthetic earthquakes of  $M_w \geq 5.5$  consistent with regional practices and to ensure that we capture only fault-related seismicity relevant to the hazard modeling, as smaller magnitudes are more likely to happen also outside the faults considered here. It is a common practice in PSHA to adopt a magnitude threshold to distinguish earthquakes associated with identified faults from those attributed to background (or distributed) seismogenic sources. This approach aims to ensure the completeness and consistency of the fault-related seismic catalogue. For instance, Gómez-Novell et al. (2020a) adopt a threshold magnitude of  $M_w 4$  for the EBSZ, while Valentini et al. (2017) apply a  $M_w 5.5$  threshold for the Italian territory. In Europe, Danciu et al. (2024) use a higher threshold of  $M_w 6.0$ , while in models such as UCERF and those applied in New Zealand (e.g., Stirling et al., 2012), this threshold is often set to  $M_w 7.0$ .

Instead of using the simulated earthquake ruptures to assign occurrence rates to specific faults as in classical seismic hazard, we consider the three-dimensional geometry of each one of the ruptures in the simulated catalogues to construct the source models for the earthquake rupture forecasts (ERFs). This facilitates to capture the inherent spatial complexities and patterns of earthquake ruptures within the EBSZ system and propagate them to the hazard. Our approach follows that introduced by Shaw et al. (2018) to integrate RSQSim ruptures into PSHA calculations in California.

The rupture outputs generated by RSQSim, as well as by other boundary-element simulators such as MCQSim, consist of a list of fault elements, often triangular, involved in individual earthquake events. As such, for each simulated rupture we extract the grid of barycentres of the triangular fault elements activated. In Figure 3b, we show some examples of ruptures for the Cat-21 model, whereas videos showing the rupture occurrence along the EBSZ in the 50 kyr-long simulations for both Cat-21 and Cat-18 are available at the Zenodo repository (Gómez-Novell et al., 2025b). Each rupture grid is incorporated into a source model as a characteristic source and following the “Gridded Surface” representation in OpenQuake (Pagani et al., 2014; OpenQuake Documentation), where individual fault ruptures are represented as a mesh of geo-spatially located points forming a discrete surface. Each rupture is assigned an annual occurrence rate of the inverse of the temporal length of the catalogue (50 kyr), hence  $1/50,000$ . The full stack of Gridded Surfaces for both catalogues (Cat-21 and Cat-18), considering their



corresponding spatial extensions, magnitudes, rake and annual occurrence rates, is used to build the models' ERFs in OpenQuake.



180 **Figure 3: (a) Scheme of the Eastern Betic Shear Zone (EBSZ) fault model. AMF, Alhama de Murcia fault; BSF, Bajo Segura fault; CAF, Carrascoy fault; CF, Carboneras fault; LTF, Los Tornos fault; PF, Palomares fault. See Herrero-Barbero et al. (2021) for fault**





**parameter details. (b) Four examples of earthquake ruptures for four different magnitudes. The red areas are the barycentres of the patches of the triangular mesh involved in each rupture.**

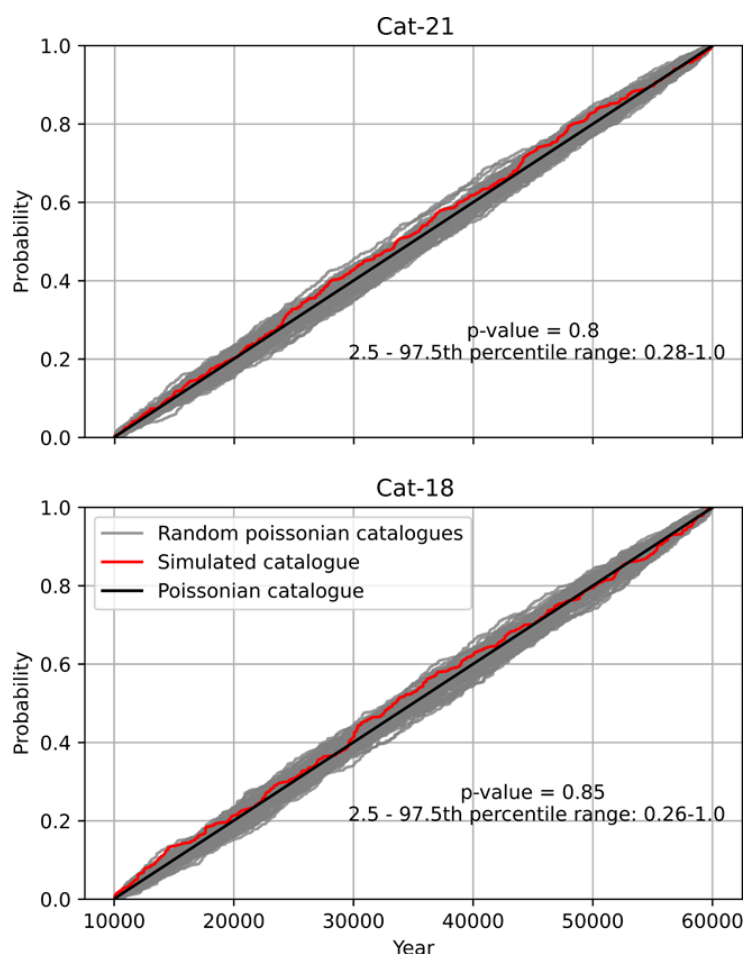
Because the temporal distribution of ruptures directly affects the fitness of time-independency assumptions in PSHA, we first  
185 evaluated whether the simulated catalogues can be treated as time-independent. Given that the simulated catalogues used are  
not declustered, we investigated the plausibility of constructing time-independent ERFs by testing the temporal distribution of  
simulated ruptures against the null hypothesis of a Poisson process. To this aim, we employed the Kolmogorov–Smirnov (KS)  
test to evaluate the statistical agreement between the empirical distribution of rupture occurrences and the theoretical  
cumulative distribution function (CDF) of a Poisson process, following the formulation adopted in the European Seismic  
190 Hazard Model 2020 (Danciu et al., 2021; Danciu et al., 2024). Specifically, we computed the maximum distance between the  
empirical CDF of simulated rupture times and the idealized Poisson CDF expected from a uniformly distributed dataset. This  
distance was then compared to a reference distribution of distances generated from synthetic random datasets drawn from a  
uniform distribution over the same time window of the simulated ruptures (Figure 4). The null hypothesis of temporal  
uniformity is accepted or rejected based on a 95% confidence level.

195 It is important to highlight that the distance metric defined by Danciu et al. (2024) represents a modified version of the classical  
KS distance. In particular, the distance between empirical and idealized Poisson CDFs is compared to a reference distribution  
of distances obtained from random synthetic datasets generated from perturbed realizations of the null hypothesis. This  
modification accounts for limitations in standard declustering techniques, which may fail to remove aftershocks from  
earthquake catalogues completely. Such residual clustering often manifests as a concave deviation in the normalized  
200 cumulative number time series relative to the theoretical expectation derived from a well-declustered Poissonian catalogue  
(see figure 4 and Danciu et al., 2021). Since our simulated catalogues do not exclude aftershocks or foreshocks, the use of this  
modified KS distance is particularly appropriate. It inherently accommodates the presence of such clustering by avoiding  
penalization of these deviations in the cumulative number time series, thus preventing them from biasing the goodness-of-fit  
assessment. In Figure 4, we show the comparison of the CDFs obtained with the aforementioned process for Cat-21 and Cat-  
205 18. The Poisson process cannot be rejected for both catalogues given that the p-value is within the 95% confidence level of  
the synthetic datasets.

We compute ground shaking, expressed in terms of peak ground acceleration and velocity (PGA, in units of g; PGV in cm/s)  
using the OpenQuake software (Pagani et al., 2014) within a “Classical” PSHA framework, according to the own OpenQuake  
nomenclature. This approach follows the standard methodology for hazard integration as formulated by Cornell (1968),  
210 McGuire (1976), and later refined by Field et al. (2003), allowing for the computation of hazard curves and maps by combining  
the simulated ERF of finite ruptures and GMMs. For ground motion prediction, we adopt the model by Bindi et al. (2014).  
This GMM estimates the geometric mean of horizontal ground motion components and uses the Joyner–Boore distance ( $R_{jb}$ )



as the distance metric. The model is calibrated on a comprehensive dataset of shallow crustal earthquakes in Europe, making it suitable for application in the study area.



215 **Figure 4: Kolmogorov–Smirnov (KS) test results for the temporal distribution of simulated ruptures in the Cat-21 (top) and Cat-18 (bottom) catalogues. The empirical cumulative distribution functions (CDFs) of rupture occurrence times (red curves) are compared against the theoretical CDFs expected from a Poisson process (black dashed lines). The grey lines represent CDFs of rupture occurrence times generated by random sampling from a uniform distribution over the same time window of the simulated ruptures.**

## 2.4 Model consistency against observations

220 In general, a model is considered reliable when its predictions agree with observations. If the observations are independent of the model, this reliability is referred to as validation (Marzocchi and Jordan, 2014; Schorlemmer et al., 2018). However, when observations have been used, directly or indirectly, in building the model, the appropriate term is consistency. This distinction is far from semantic: while a flawed model typically fails a validation test, it may still pass a consistency test by overfitting



the data. Therefore, in practice, only the rejection of a consistency test is meaningful, as it highlights structural errors in the  
 225 model, often termed "unknown unknowns" (Marzocchi and Jordan, 2014).

In PSHA, especially in low-deformation regions, observations are typically limited and rarely independent of the model. As a  
 result, PSHAs are generally assessed for consistency rather than validation. This is typically done by comparing model outputs  
 to a set of observations (Stein et al., 2015). The most straightforward choice for such observations is ground-shaking  
 observations at specific sites.

230 A general framework for testing PSHA reliability can be expressed as:

$$y \sim g(Y) \tag{1}$$

where  $y$  is ground shaking and  $g(Y)$  is the statistical distribution of the random variable  $Y$ , derived from the PSHA. The forecast-  
 observation comparison is made through pairs of the form  $(g, y)$ , where  $g$  is the forecast from the mean hazard model  $f(x)$  and  
 $y$  is the observed value (Gneiting and Katzfuss, 2014).

235 A commonly used consistency/validation test in PSHA involves comparing the number of observed exceedances of ground  
 motion thresholds to the expected number. In regions with low strain rates, such as Italy, D'Amico et al. (2024) proposed a  
 scoring procedure that compares hazard curves to time series of macroseismic observations (or "seismic histories") at selected  
 sites. In this study, we test the expected number of exceedances against both macroseismic observations and recorded levels  
 of ground shaking. The general procedure is the same for the two types of observations; however, given the treatment necessary  
 240 in the use of macroseismic data, we describe separately the two tests.

In the second part of our approach (Figure 2), we test the two physics-based hazard models for Cat-21 and Cat-18, for which  
 the use of the term "validation" is appropriate since the observational data were not used in model construction, and an area  
 source model, for which the term "consistency check" is more appropriate. This is because the Gutenberg-Richter parameters  
 of this area source model are derived from earthquake catalogues significantly based on macroseismic intensities (see details  
 245 on the ZESIS database; IGME, 2015). The area source hazard model is computed using the Gutenberg-Richter parameters of  
 area "55" from the ZESIS model (García-Mayordomo, 2015; IGME, 2015) and considering the GMM as for Cat-21 and Cat-  
 18 models. Area "55" is an area defined to encompass the geological and seismogenic features of the EBSZ fault system, and  
 hence is appropriate for comparison with the fault-based seismic hazard we perform for the simulated catalogues. Despite the  
 mentioned terminology, for better clarity we use the term consistency to describe all analyses against observations

250 To be consistent with the physics-based hazard models, we apply the cluster correction proposed by Marzocchi and Taroni  
 (2014) to the magnitude-frequency distribution of the area source. Marzocchi and Taroni (2014) define multiplicative  
 coefficients to calculate non-declustered (complete) frequencies for magnitude values starting from their declustered values.



The adoption of an area source–based ERF for PSHA plays an important role in both the qualitative and quantitative evaluation of fault-related impacts, particularly in terms of the spatial distribution of seismic hazard. Statistical tests are primarily used to assess the relative performance of competing models and to identify those that cannot be rejected; therefore, an area source–based approach offers complementary value. By producing a spatially smoothed representation of seismic hazard, it serves as an effective tool for visualizing and interpreting the regional influence and performance of physics-based earthquake models.

#### 2.4.1 Macroseismic consistency test

The first step in the procedure for using macroseismic data involves selecting a set of localities to evaluate the consistency of PSHA models with available macroseismic observations. We select here five sites that correspond to prominent cities along the EBSZ (Lorca, Murcia, Almería, Torrevieja and Vera). All sites are located along or at a distance shorter than 10 km, except Almería, located at approximately 20 km from the faults of the EBSZ. This close distance criterion is used to ensure that the macroseismic records are related to the faults within the EBSZ fault system rather than seismicity from distant sources.

For each site, we built two datasets: (i) A set of observed macroseismic intensity (MI) values (see Table S1) and (ii) a set of expected MI values derived from the hazard models. To compare these datasets, it is first necessary to estimate the completeness time intervals for each intensity degree in the observed MI catalogue, that is, the periods in which it is reasonable to assume that all earthquake effects above a certain intensity were reported. These intervals are computed using the statistical procedure of Albarello et al. (2001), which returns a probability density function of the completeness time for each desired intensity threshold. To be coherent with the ERFs, we retain intensity data related to earthquakes with magnitudes  $M_w \geq 5.5$ .

According to D’Amico et al. (2024), macroseismic observations with uncertain intensity values (e.g., MI 6-7) are treated in two hypotheses: Option 1 - assigned to the lower MI value (e.g., MI 6); Option 2 - assigned to the higher MI value (e.g., MI 7). For each MI threshold we compute their completeness functions applying the Albarello et al. (2021) procedure. Then, as described in D’Amico et al. (2024), we use the medians and 75th percentiles of the completeness functions to investigate the uncertainty in the assessment of the starting year of the completeness period. By combining the treatment of the uncertainty of the half-degree of the macroseismic intensities with the starting years of the completeness function ( $T_c$ ), we obtain four estimates of the pairs “starting time of completeness – number of observed data” for each intensity level at each site:

- Option 1 and the median of the completeness functions (median  $T_c$  value)
- Option 1 and the 75th percentile  $T_c$
- Option 2 and the median  $T_c$
- Option 2 and the 75th percentile  $T_c$

To compute the number of expected intensity data at each site, we start from the RSQSim-derived hazard curves in PGV computed considering site-specific  $V_{s30}$  values (see table S2) taken from the soil classification from Herrero-Barbero et al.



(2023), which encompasses the whole EBSZ. To each city we assign the upper Vs30 value from the range corresponding to the soil class type. While this geology-based approach provides a practical framework for assigning Vs30 values across the region, it may be affected by local uncertainties when applied to specific sites, such as lateral variations in stiffness within a given geological unit. Nevertheless, it remains a common method for regional-scale applications where in-situ measurements are lacking. Hazard maps are computed considering a uniform Vs30 of 600 m/s to facilitate recognition of spatial patterns of hazard estimates at a regional scale. Also, the selected Vs30 value corresponds to class B soils ( $360 \leq V_{s30} < 800$  m/s), which is the average of the Vs30 values across the entire EBSZ according to the map by Herrero-Barbero et al. (2023). We use PGV as a hazard measure as it is the most likely linked to damage (Gomez-Capera et al., 2020; Cataldi et al., 2021).

In Figure 5, we show the computational flow to derive the number of expected intensity data from the hazard curves. From the hazard curve expressed as the annual frequency of exceedance (AFoE) of PGV levels, we first calculate the annual frequency of occurrence (AFoO; Figure 5a), which corresponds to the incremental counterpart of the AFoE. For each PGV level, we build the conditional probability of being an MI level (Figure 5b) using the Ground Motion Intensity Conversion Equation (GMICE) by Gomez Capera et al. (2020). These equations establish a probability distribution to each MI level, which is related to a ground shaking measure (i.e., PGV). Then, we multiply the annual occurrence rate of each PGV level by the corresponding probability of MI levels from the GMICE to obtain the annual rate of occurrence of MI ( $\lambda(MI)$ ) as a function of PGV (Figure 5c):

$$\lambda(MI) = \sum_{j=1}^n \lambda(x_j) P(MI|x_j) \quad (2)$$

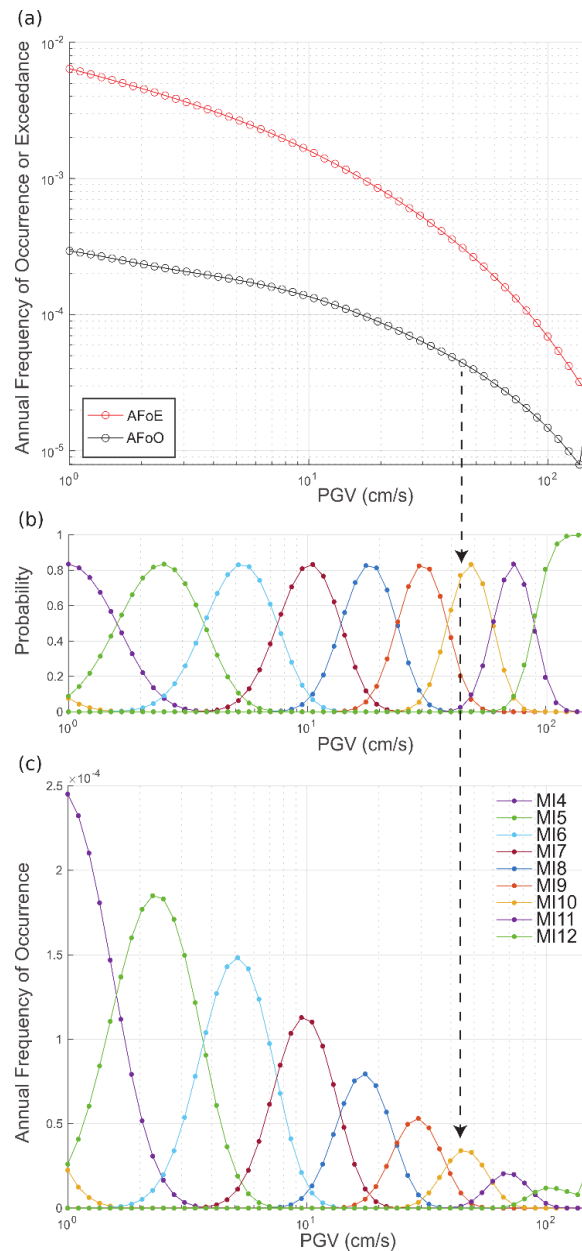
Here,  $\lambda(x_j)$  is the annual occurrence rate for the  $j$ -th level of PGV, and  $P(MI|x_j)$  is the conditional probability of a PGV level being an MI level from the GMICE.

The occurrence rates of each MI are summed and multiplied by the corresponding length of completeness periods (from the four completeness estimates above) to determine the expected number of intensity data for each MI threshold. Importantly, the number of expected intensities per MI threshold follows a Poisson distribution.

For each considered testing locality, four pairs of observed and expected numbers of intensity data are obtained for each MI threshold, corresponding to the four different estimates of the completeness starting years ( $T_c$ ), i.e., for the median value and the 75th percentile of the completeness function, and for the two options of uncertain intensity degrees (Options 1 and 2).

We perform the statistical test by computing the p-value from the tails of the Poisson distribution, as described by Zechar et al. (2010). Here,  $F$  is the right-continuous Poisson cumulative distribution function and, if the number of observations ( $N_{obs}$ ) exceeds the expected count ( $N_{exp}$ ), the p-value is calculated as:

$$p = 1 - F(N_{obs} - 1; N_{exp}) \quad (3)$$



**Figure 5.** a) Example of a hazard curve for a site, expressed in terms of annual frequency of exceedance and occurrence of PGV (AFoE and AFoO). b) PDF of the relationship between PGV and MI as proposed in Gomez-Capera et al. (2020). c) Annual frequency of occurrence of MI from 4 to 12 as a function of PGV for a site. The arrow illustrates an example of the procedure: the annual frequency of occurrence value (AFoO) in terms of PGV (black circle in panel a) is multiplied by the probability of corresponding to Modified Intensity levels 9, 10, or 11 (panel b), to obtain the annual frequency of occurrence of MI as a function of PGV as shown in panel c. This procedure is repeated for each PGV value in the hazard curve. Then, all the corresponding annual frequencies of occurrence for each MI are summed to obtain the final annual frequency of occurrence per MI level.



Conversely, if the  $Nobs$  is less than or equal to the  $Nexp$ , the p-value is given by:

$$320 \quad p = F(Nobs; Nexp) \quad (4)$$

These two p-value calculations address specific questions: the p-value from the former case indicates whether the forecast is too low compared to observations, while the latter assesses if it's too high.

For each MI, we compute the average p-value across the four cases. Subsequently, to compare the performance of the tested catalogues, we calculated the logarithm of the averaged p-value ( $LogP$ ).

### 325 **2.4.2 Ground shaking consistency test**

For sites with ground-shaking data recorded at seismic stations, the computational workflow is more straightforward. We consider a set of five accelerometer stations located near the fault system (EXLO, ENIJ, M04, EXTO and EXVE2; see figure 6), continuously operational over the last decades. These stations correspond to the Spanish Instituto Geográfico Nacional (IGN) Seismic Alert Network. Data from these stations, including station coordinates and records, have been obtained from  
 330 the IGN Seismic Alert Network Stations database (IGN 2025, last access October 2025).

We compute hazard curves for peak ground acceleration (PGA) at the mentioned sites, incorporating site-specific  $Vs30$  values provided by the Instituto Geográfico Nacional (IGN) and available in table S2 (see Supplements). We obtain the expected number of exceedances by multiplying the annual frequency of exceedance by each station operational period: 1989-2025 for EXLO, 2002-2025 for EXNIJ, 2008-2025 for M04, 1993-2025 for EXTO and 2012-2025 for EXVE2. We then compute the  
 335 p-values between both the observed and expected exceedances using equations (3) and (4), and we calculate their logarithms ( $LogP$ ) to compare model performance.

## **3. Results**

### **3.1 Seismic hazard estimates**

Figures 6 and 7 show the hazard results at the EBSZ obtained from the physics-based earthquake simulations on Cat-21 and  
 340 Cat-18, as well as those from an area source model. Hazard maps are for PGV at 2% probability of exceedance in 50 years for soil B conditions.

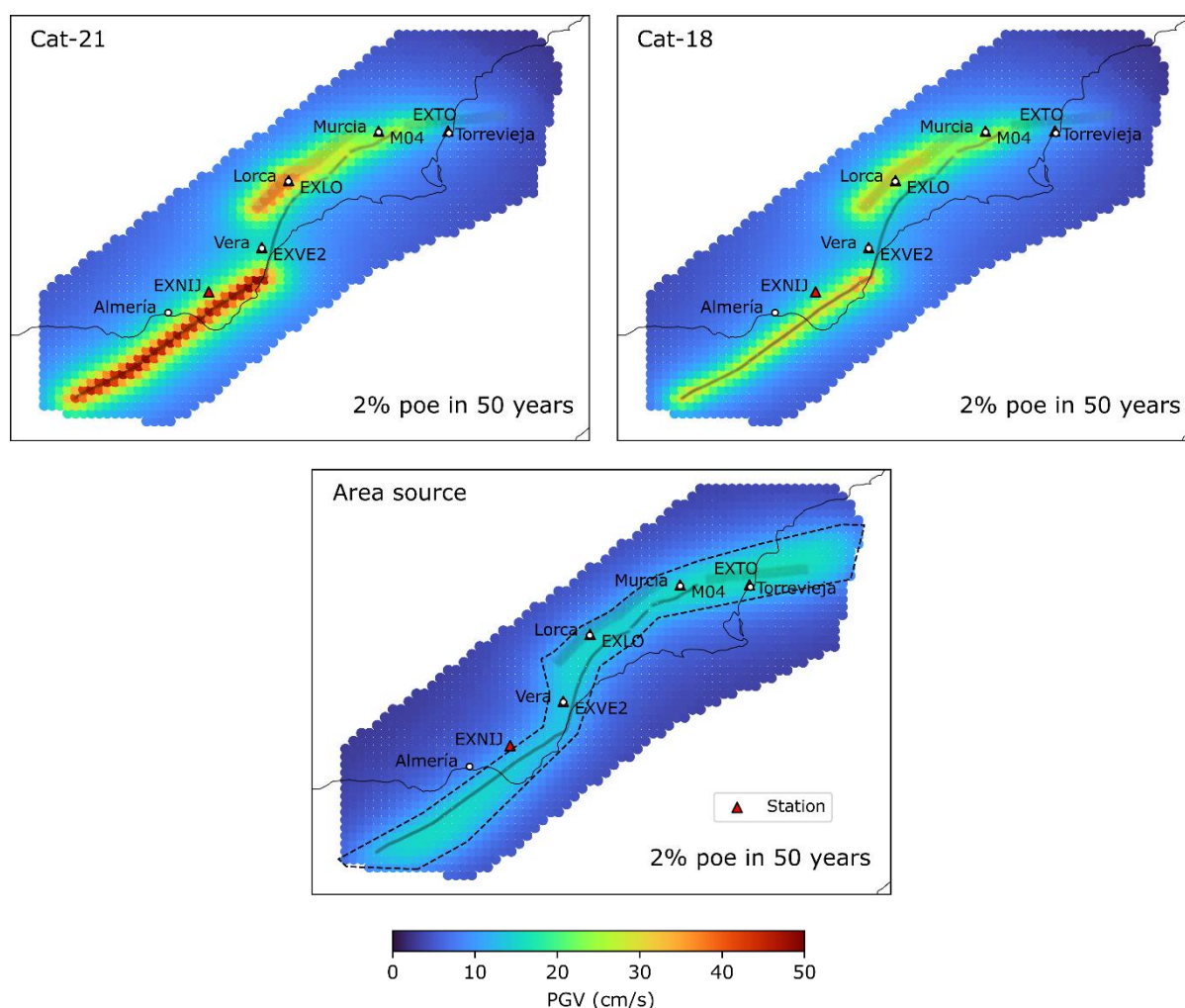
Hazard maps on Cat-21 and Cat-18 (Figure 6) show generally higher PGV values than the area source model, exceeding approximately 20 cm/s in the regions along the fastest slipping major faults of the EBSZ: the Alhama de Murcia and Carboneras faults (see map in figure 3a for location). Conversely, the map obtained using the area source model (Figure 6) results in a





345 smoother spatial distribution of PGV within the area source (area 55 of the ZESIS model), with the highest values below 20  
 cm/s inside the modelled area.

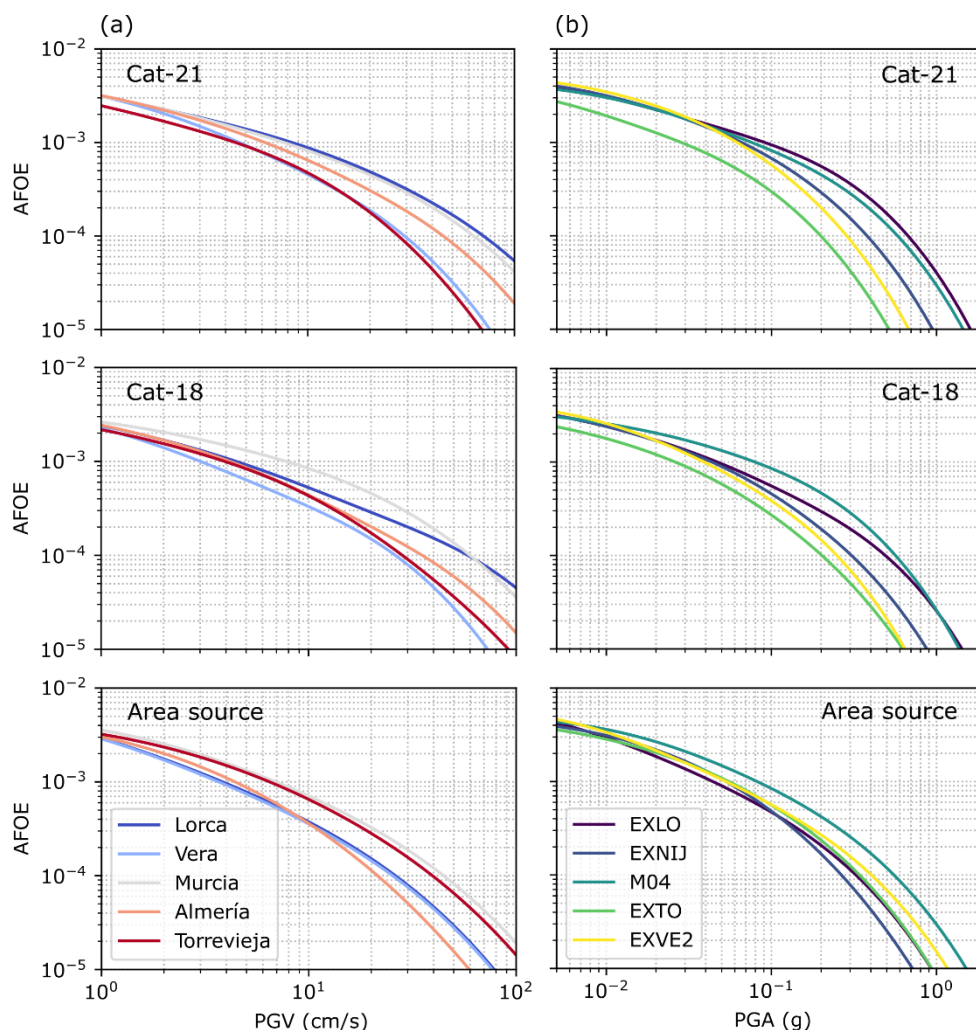
Remarkably, out of the physics-based seismic hazard maps, Cat-21 is the one returning highest PGV values. This has to do  
 with the fact that Cat-21 returns higher earthquake rates than Cat-18 for magnitudes above 5.5 (see figure 1b). PGV peaks at  
 around 50 cm/s along the Carboneras fault, which is the fault with higher estimated slip rate (at 1.1-1.3 mm/yr; see Herrero-  
 350 Barbero et al., 2021). Such discrepancy in earthquake rates lies on the differences in the frictional parameters and initial stresses  
 of both models, as fault system geometry, slip rates and hazard calculations remain constant across models.



**Figure 6.** PGV hazard maps for 2% probability of exceedance in 50 years (homogeneous  $V_{s30}=600\text{m/s}$ ) computed using the OpenQuake software for both the simulator-based catalogues (Cat-21 and Cat-18 from Herrero-Barbero et al., 2021) and an area



355 source model: Zone 55 from ZESIS model (García-Mayordomo, 2015; IGME, 2015). White circles and triangles show, respectively, selected localities and accelerometer station sites for consistency tests.



**Figure 7. (a) Hazard curves for PGV for the five selected sites. (b) Hazard curves for PGA for the five selected stations. AFOE: Annual frequency of exceedance. Vs30 for each site and station are provided in table S2.**

360 In Figure 7, we show the hazard curves for PGV and PGA at the selected localities and station sites. For Cat-21 and Cat-18, the highest hazard estimates are found in Lorca, Murcia and Almería, all three located near some prominent, higher slipping, faults of the EBSZ (Alhama de Murcia, Carrascosy and Carboneras faults, respectively; see Herrero-Barbero et al., 2021). In detail, Lorca is the site with highest seismic hazard in Cat-21 due to the influence of the Alhama de Murcia fault, similarly to Almería with the Carboneras fault. However, in Murcia, the higher hazard values are likely due to the amplification effect on



365 PGV generated by the low Vs30 value (150 m/s) because the Carrascoy fault slips at a rate about a half of the Alhama de Murcia and Carboneras faults. This is further emphasized in Cat-18, where Murcia overcomes Lorca in the hazard estimates. In this case, this is likely due to the combined effect of the low Vs30 and the decreased activity rates in the Alhama de Murcia Fault in that catalogue, a feature that is recognizable by the lower hazard estimates in the map of figure 6. In Vera and Torrevieja, seismic hazard is consistently lower for both simulated models as the faults near both cities have activity rates that  
 370 have a small influence in the seismic hazard (Fig. 6).

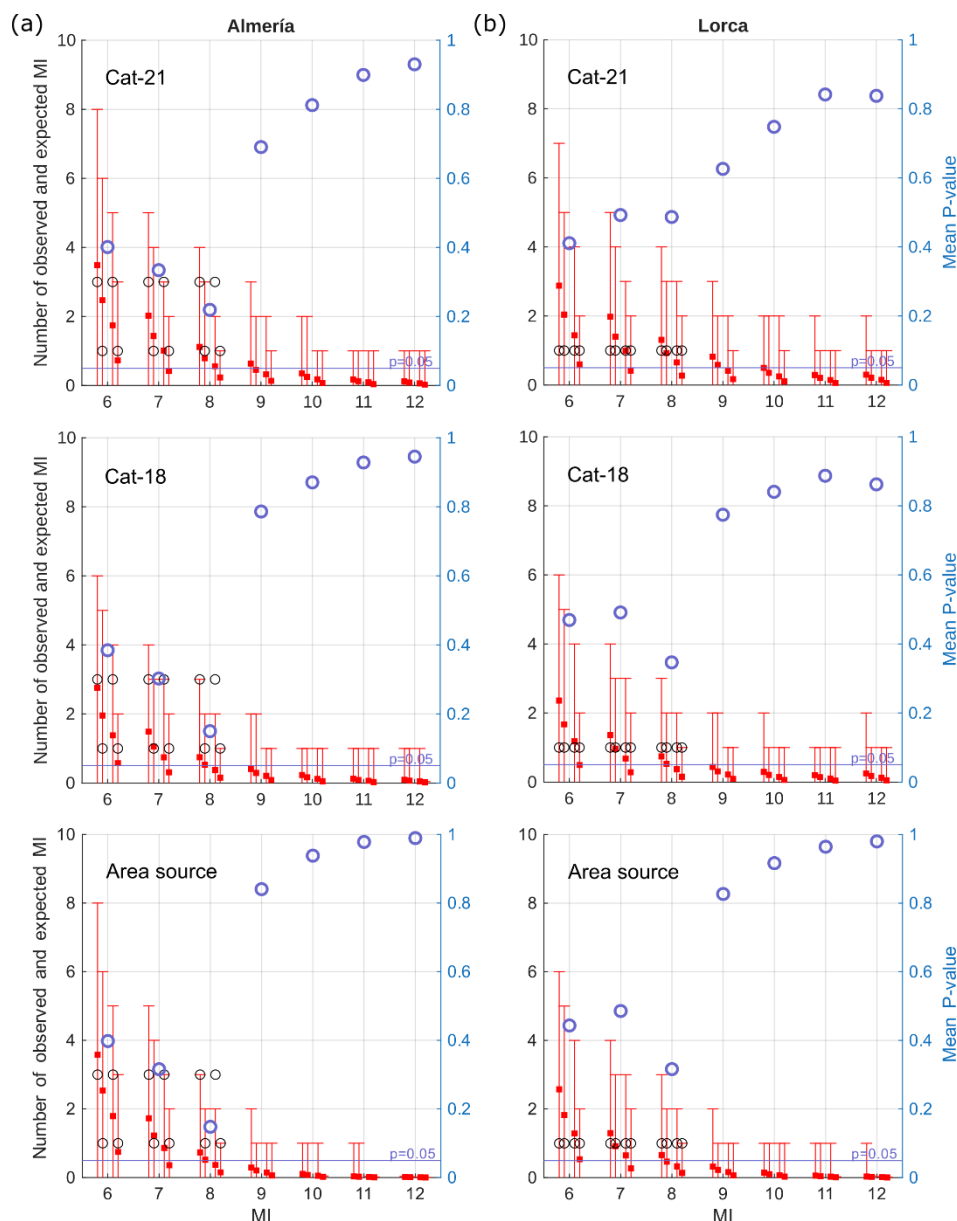
In contrast, the area source model provides hazard curves that are generally lower for most sites and with less dispersion due to the smoother spatial distribution of the hazard in this model (see figure 6). Among these, the curve for Almería shows the lowest values of AFoE in this model, especially for larger PGV values, because it is located outside the area source. The higher  
 375 hazard curves for Torrevieja and Murcia are due to the lower Vs30 values in these cities compared to other cities within the area source (Lorca and Vera, both at 360 m/s; see table S2).

Figure 7b, we show the hazard curves for PGA at the seismic stations. The results of the hazard curves for these station locations are interchangeable with those made for the cites, as most stations are located within them or in very similar settings  
 380 (e.g., EXNIJ is also outside the area source limit; Fig. 6).

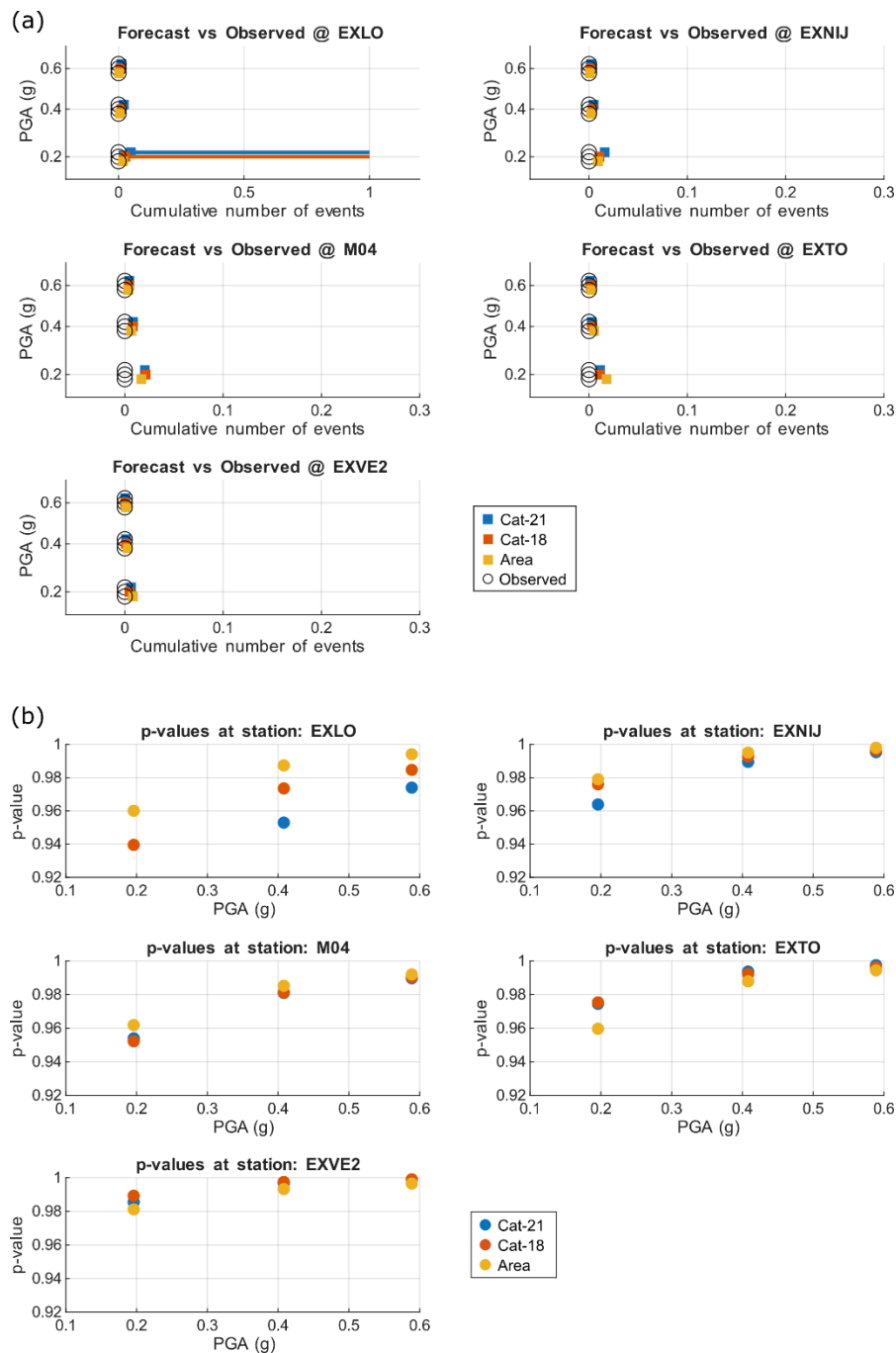
### 3.2 Consistency tests

In Figure 8, we show an example of the comparison between the four pairs of observed and expected numbers of intensity data obtained for each MI threshold, while figure 9 presents the analysis for PGA at accelerometric stations. In figure 8, we recall that the four pairs correspond to different estimates of the completeness periods (median value and the 75th percentile of the  
 385 completeness function) and the two options for treating uncertain half intensity degrees. The examples are for the localities of Almería and Lorca for the Cat-21, Cat-18, and area source models, whereas the complete set of comparisons is given in the Supplements. We represent the observed number of exceedances with black edge circles and the expected numbers with the red squares. In addition, we plot the resulting mean p-value for each MI, with a blue circle (values on the right-side y-axis). The horizontal dashed line marks the confidence level of 0.05 of p-value, above which we cannot reject the null hypothesis.

390 All models show p-values above the 0.05 confidence level, highlighting overall good agreement between models and observations. We observe a first-order shift of p-values between MI 8 and MI 9 due to the absence of MI greater than 8 in the catalogue, but all models tend to furnish a consistent number of expected values of MI for MI 6-MI 8, which basically means that no models over-predict the observations. A slight tendency to under-predict the observations is instead noticeable with the area source model for the MI 8 at the Almeria site, with p-values closer to the 0.05 threshold.



395 **Figure 8. Comparison, for different MI thresholds, between the number of observed macroseismic data and the number of expected intensities from Cat-21, Cat-18, and area source models in a) Almería and b) Lorca. Key: black edge circles are the observed number of exceedances; the red squares are the expected numbers from the hazard models, with vertical bars representing the 2.5th and 97.5th percentiles of the Poisson distribution of the expected MI; blue edge circles are the mean p-values (values on the right-side y-axis). Comparison graphs for the rest of the sites are available in the Supplements of the paper.**



400 **Figure 9. (a) Observed vs forecasted number of exceedances for the different models at different PGA thresholds and at each station location due to  $M_w \geq 5.5$  events. (b) p-values at different PGA thresholds for each station representing closeness between the number of recorded instrumental PGA exceedances corresponding to  $M_w \geq 5.5$  and the PGA exceedances from the hazard models Cat-21, Cat-18 and area source.**



Regarding the PGA thresholds at four seismic stations located in the EBSZ (Figure 9), we present the p-value analysis between  
 405 the observed and expected number of exceedances. A particularly informative case occurs when zero exceedances are observed  
 at a station, as it is the case of the EBSZ for  $M_w \geq 5.5$ . In such cases, the comparison with model expectations is especially  
 relevant, as it allows us to exclude models that predict unrealistically high hazard levels. When no exceedances are recorded  
 during the observational period, any model that forecasts a significantly non-zero number of exceedances is likely  
 overestimating the hazard at that location. Under these circumstances, the model yielding the lowest expected number of  
 410 exceedances is statistically the most consistent with observations. This is a direct consequence of Poisson-based testing,  
 wherein the probability of observing zero events decreases rapidly as the expected number increases. Hence, if all models  
 predict non-zero exceedances, the one with the lowest expected rate will return the highest p-value and, thus, cannot be rejected  
 based on observational evidence (i.e., it is the one with better performance). This approach provides a straightforward yet  
 robust method to discriminate between competing hazard models in data-scarce regions, where even the absence of observed  
 415 exceedances carries meaningful information.

Figure 9b shows the p-value analysis between the observed and expected PGA exceedances in the different stations and for  
 each model. For all stations, except EXTO and EXVE2, the area source model provides the highest p-values, inferring superior  
 performance over the simulated catalogues. However, this agreement might reflect model self-consistency, since the low  
 number of expected exceedances of  $M_w \geq 5.5$  in area source 55 for short-term periods are largely inherited from the lack of  
 420  $M_w \geq 5.5$  instrumental records compared to the physics-based models. Conversely, EXTO and EXVE2 shows superior  
 performance of simulation-based hazard models over the area source model. In this case, the smoothing effect of the area  
 source model leads to overprediction in regions where the influence of nearby active faults in the hazard is less prominent.  
 This is further evidence that simulated catalogues might improve the spatial distribution of the hazard at near-fault locations.  
 Figure 9a depicts this effect, with physics-based models predicting larger number of events for all PGA thresholds in all  
 425 stations except EXTO and EXVE2.

The consistency analysis demonstrates how this testing strategy is particularly effective for identifying models that tend to  
 overpredict shaking intensity, allowing us to refine hazard estimates and increase model reliability at the site scale.

### 3.3 Model performance assessment

In Figure 10, we present the  $\text{Log}P$  scores obtained for the three tested hazard models, Cat-21, Cat-18, and the area source  
 430 model, across five localities and four instrumental stations. These scores, which are the logarithm of the p-values from the  
 previous tests, quantify the statistical agreement between observed and expected exceedances of ground-motion thresholds,  
 both in terms of MI and PGA. We recall here that higher (less negative)  $\text{Log}P$  values indicate a better agreement with the  
 observed data, as they correspond to higher p-values under the Poisson-based statistical framework used for model evaluation.





At the site-specific level, the Cat-21 model generally outperforms both Cat-18 and the area source model across the tested MI thresholds ( $MI \geq 6$ ,  $MI \geq 7$ , and  $MI \geq 8$ ). Specifically: (i) for  $MI \geq 6$ , Cat-21 yields the best (least negative) log-likelihood values at Almería; (ii) for  $MI \geq 7$ , Cat-21 performs best at Lorca, Vera, and Almería and; (iii) for  $MI \geq 8$ , Cat-21 returns the highest likelihood at Lorca, Vera, and Almería. In contrast, the Cat-18 model performs slightly better at  $MI \geq 6$  thresholds in all sites <10km from active faults (Lorca, Vera Murcia and Torrevecija) due to slight overprediction in Cat-21 (Fig S2). Also, the area source model performs better in Murcia and Torrevecija for  $MI \geq 7$ , as the physics-based models underestimate the exceedances in these locations. In these locations, where nearby active faults have less influence on the hazard (Fig. 6), the observations showcase a potential bias toward higher values because of the short completeness period of the different macroseismic intensity classes. That is, the completeness periods are shorter (few hundreds of years) than full earthquake cycles on those individual faults (thousands of years; see analysis of the catalogues by Herrero-Barbero et al., 2021).

In terms of instrumental data (PGA in Figure 10), the  $LogP$  values indicate that the area source model outperforms the other models in most stations. However, differences are minor compared to the MI test as evidenced by the generally lower (closer to 0) p-values. In these stations, the simulated catalogues slightly overpredict the observations relative to the area source model. This may be explained by the fact that the station records span, at best, 36 years, a considerably short time window to properly capture the  $M_w \geq 5.5$  seismicity in a region where recurrence intervals average around 100 years (see Figure S1 and the MFDs in Figure 1b). Given that earthquake activity rates are highly variable over time, as illustrated by Herrero-Barbero et al. (2021) in the simulated catalogues, short observational periods are more likely to be affected by such variability, difficulting the accurate characterization of longer-term seismicity patterns. Conversely, the simulated catalogues perform better in Torrevecija and Vera, where active faults have low activity rates that do not influence the hazard significantly (Fig. 6). In these cases, the smoothing effect of the area source model distributes the hazard more evenly across the area source, leading to overestimated forecasts.

The overall model performance is assessed by computing the sum of  $LogP$  values across all macroseismic and instrumental thresholds. This analysis returns a final summed  $LogP$  score of -13.363 for Cat-21, -14.598 for Cat-18 and -14.513 for the area source model. These results highlight that Cat-21 overall offers the most consistent predictions when both macroseismic intensities and acceleration records are considered, underscoring its potential in reproducing both short and long-term shaking patterns and damage, respectively.





		Macroseismic intensity		
Site	Model	MI 6	MI 7	MI 8
Lorca	Cat-21	-0.890	-0.709	-0.720
	Cat-18	-0.756	-0.711	-1.059
	Area	-0.813	-0.722	-1.150
Vera	Cat-21	-0.883	-0.867	-0.924
	Cat-18	-0.790	-1.212	-1.160
	Area	-0.846	-0.936	-1.078
Murcia	Cat-21	-0.900	-0.709	-0.676
	Cat-18	-0.807	-0.707	-0.712
	Area	-1.075	-0.686	-0.628
Almería	Cat-21	-0.914	-1.096	-1.518
	Cat-18	-0.956	-1.196	-1.897
	Area	-0.921	-1.152	-1.911
Torrevieja	Cat-21	-0.774	-0.673	-0.950
	Cat-18	-0.734	-0.725	-1.060
	Area	-1.037	-0.672	-0.782

		PGA		
Station	Model	0.2g	0.4g	0.6g
EXLO	Cat-21	-0.049	-0.021	-0.011
	Cat-18	-0.027	-0.012	-0.007
	Area	-0.018	-0.006	-0.003
EXNIJ	Cat-21	-0.016	-0.005	-0.002
	Cat-18	-0.011	-0.003	-0.001
	Area	-0.009	-0.002	-0.001
M04	Cat-21	-0.020	-0.008	-0.004
	Cat-18	-0.021	-0.008	-0.004
	Area	-0.017	-0.006	-0.003
EXTO	Cat-21	-0.011	-0.003	-0.001
	Cat-18	-0.011	-0.003	-0.002
	Area	-0.018	-0.005	-0.002
EXVE2	Cat-21	-0.006	-0.001	-0.000
	Cat-18	-0.005	-0.001	-0.000
	Area	-0.008	-0.003	-0.002
<b>TOTAL</b>	Cat-21	<b>-13.363</b>		
	Cat-18	<b>-14.598</b>		
	Area	<b>-14.513</b>		

**Figure 10.** *LogP* scores obtained for the three tested hazard models, Cat-21, Cat-18, and the area source model, across five localities and five instrumental stations. With green background, we mark the best model for each situation.



## 465 4. Discussion

The results presented in this study demonstrate the viability of physics-based earthquake simulators for PSHA in low-to-moderate strain regions. On the one hand, we develop a pipeline approach to accurately include simulation-based earthquake ruptures, including complex rupture geometries, into PSHA calculations. On the other hand, we compute the probabilistic seismic hazard from two best and worst performing synthetic catalogues by Herrero-Barbero et al. (2021) at the EBSZ. Upon  
470 this we assess whether these models adequately reproduce the spatial variability in hazard by comparing them to macroseismic and ground shaking observations.

### 4.1 Pipeline of simulations to PSHA

Earthquake cycle simulators allow the generation of large earthquake rupture datasets, including complex multi-fault rupture geometries. However, such complexities are generally disregarded for PSHA as the most straightforward way to work with  
475 the simulator outputs is to use synthetic catalogues to assign earthquake rates to each specific fault source (e.g., Khodaverdian et al., 2016; Rafiei et al., 2022). This simplification deprecates earthquake rupture complexity (including multi-fault ruptures) and downgrades the accurate spatial representation of seismic hazard, as most engines, like OpenQuake, distribute earthquake ruptures uniformly over fault surfaces based only on the earthquake occurrence rates. Instead, the treatment of the simulated ruptures individually as finite characteristic sources allows to preserve physical features of ruptures within the fault system.  
480 As such, our formulation inherently allows the representation of multi-fault rupture complexity, preferent rupture concentration zones linked to fault geometry and rupture asymmetry, among others, onto the spatial distribution of the hazard. This physics-based approach can overcome difficulties frequently found in the definition and inclusion of multi-fault ruptures in fault-based PSHA, which require establishing pre-defined segmentation rules (e.g., Field et al., 2015; Chartier et al., 2017; Gómez-Novell et al., 2020b), usually based on empirical statistics on earthquake ruptures rather than on fully physics-based criteria.

485 The inclusion of detailed rupture geometries can be a critical benefit in some hazard assessments as spatial rupture representations govern the source-to-site distances, a key parameter in GMMs. Depending on the specific GMM, distances may be defined in terms of hypocentral, rupture, or Joyner-Boore distances, with those relying on hypocentral or rupture distances being particularly sensitive to the initiation location and final geometry of the rupture, respectively.

### 4.2 Seismic hazard results

490 Our results demonstrate that earthquake cycle simulations do not only reproduce the large-scale hazard patterns observed in the EBSZ, but they also capture the localized differences seen in empirical data. The agreement is not only quantitative (in terms of likelihood and statistical tests) but also spatially coherent across a range of return periods and intensities.

The overall better performance of the physics-based model Cat-21 compared to the area source model in the combined  $\text{Log}P$   
495 scoring for macroseismic intensities and ground shaking, highlights the simulator ability to generate realistic hazard statistics.



Its ability to incorporate detailed fault geometry, long-term slip rates, and realistic rupture behaviour into the ERF overperforming Cat-18 (see Gómez-Novell et al., 2025a) contributes to capturing key aspects of seismic hazard in regions with mapped active faults. Notably, the model's performance is robust in an area with sparse historical seismicity like the EBSZ, addressing a critical gap in conventional PSHA for low-strain regions.

500 The physics-based model outperforms the area-source model against macroseismic empirical data, while the opposite happens for instrumental ground shaking observations. On the one hand, the hazard models from synthetic catalogues are influenced by larger magnitudes than the area source model (see MFDs in figure 1). These magnitudes reflect long-term behaviour of the fault system, which might explain the better agreement of the physics-based model to longer-term observations of the catalogue (macroseismic data). On the other hand, the more accurate spatial distribution of hazard along fault in the simulator-based  
 505 models, likely resolves better the expected exceedances from macroseismic data compared to the area source model, which has a smoothing effect on hazard estimates. In terms of ground shaking, the area source model generally returns lower expected PGA exceedances for the short-term (station operational period) investigation time compared to the physics-based models at the station locations. In this case, even though the performance gap is less pronounced, the area source model performs slightly better than the physics-based models. This effect is likely a reflect of the self-consistency between the seismic catalogue -  
 510 scarce in large magnitudes compared to the physics-based models - and the area source model, which is partly derived from the same catalogue (see section 3).

The good performance of the RSQSim model Cat-21 underscores the simulator's ability to reproduce seismic hazard statistics, a result that is aligned with previous findings in high seismicity regions like California (Shaw et al., 2018). However, the choice of simulation model is one crucial step in PSHA as it directly impacts the hazard estimates. In fact, in some cases,  
 515 physics-based simulations can perform worse than statistical models like area sources, a feature that we observe for the synthetic catalogue Cat-18 (Figure 10). This has to do with a central aspect of physics-based simulations, the parametrization, which can significantly influence the resulting simulated earthquake statistics. These parameters include the initial stress conditions, namely normal and shear stresses, frictional parameters of the fault such as the rate (a) and state (b) coefficients and other parameters like slip rate variability or fault geometry. In this study, we applied a benchmarking method that uses a  
 520 proper scoring algorithm (Gómez-Novell et al., 2025a) to select the two synthetic catalogues that perform best and worst from the eleven computed by Herrero-Barbero et al. (2021). The method scores the catalogues based on their fit to a widely accepted scaling relationship, to a standard Gutenberg-Richter magnitude frequency distribution with a b-value of 1, and the rates of paleoearthquakes at the EBSZ. This ensures that the consistency results for physics-based model Cat-21 are suitable for seismic hazard applications in the study area, while the results for Cat-18 highlight the relevance of choosing a proper model design  
 525 in a physics-based framework.



## 4.3 Limitations

### 4.3.1. Model setup and parameters

Besides the advantages, there are limitations in our physics-based approach. In terms of the RSQSim models, it is important to highlight that these assume important simplifications in the driving friction parameters, as well as fault geometries and properties. First, initial friction and stressing conditions (variables of the rate and state law, shear stress) are considered uniform throughout the whole EBSZ faults. Despite this common practice in earthquake cycle simulations allows a reliable and realistic approximation to real catalogue statistics (e.g., Shaw et al., 2018; 2022; Delogkos et al., 2023), recent research has shown that the inclusion of friction and stressing variability (e.g., depth dependent) is able to produce more realistic earthquake rupture patterns as well as improving statistical features of the simulated MFDs and hypocentral depth distributions (Mika Liao et al., 2024).

Second, fault geometry and prescribed slip rates are also simplified in the EBSZ model. Faults are simplified into single surfaces, while slip rate is mostly assumed constant in each fault throughout its whole fault surface. Similar to initial stress and friction conditions, simplified geometries and slip rate profiles can lead to less realistic features in the simulated catalogues. For instance, Delogkos et al. (2023) demonstrated that the inclusion of complexity in fault geometries and slip rate profiles along-strike and down dip, allowed to improve the simulated catalogues by removing characteristic behaviour and improving depth distribution. Similarly, Allam et al. (2019), identified that fault roughness, enhanced predictability in earthquake rupture location, which can be a critical point in earthquake-rupture dependent approaches like the present one.

The EBSZ is a complex fault system, where faults are generally composed by several fault strands with complex geometric links at depth, frictional properties, and slip rate partitioning (e.g., Herrero-Barbero et al., 2020; Gómez-Novell et al., 2022; Martín-Banda et al., 2022; Yazdi and García-Mayordomo, 2024). Considering complexity in fault friction, geometry and slip rate might help further refining the simulated catalogues and improve the accuracy of seismic hazard estimates, especially in a physics-based approach where the spatial variability component is key to the hazard. These epistemic uncertainties could be treated in a systematic way in the framework of a PSHA through the use of a logic tree approach in which the weight of each branch could be assigned by the fit to selected validation benchmarks (e.g., macroseismic intensities and ground shaking).

### 4.3.2. Model consistency

Validation of synthetic rupture sets through comparison with independent observational datasets, considered the gold standard in testing procedures, provides a robust mean to assess the predictive capability of region-specific physics-based models. In our study, we performed proper validation tests against both macroseismic and instrumental datasets for the simulated catalogues because validation data is not used for the physics-based modeling. Fault geometries, friction and slip rates come from geological datasets in the region (see Herrero-Barbero et al., 2021 for details), while model benchmarking is driven by their agreement with global empirical relationships and regional paleoseismic data not used for the validation step. In contrast,



the performance of the area source model can only be assessed through consistency because the EBSZ seismic catalogue was the basis to develop the seismic parameters of the area source in the ZESIS database. In the consistency case, only the poor performance of the area source model against the Cat-21 model in the macroseismic data tests is meaningful as it highlights the inadequacy of the model compared to the physics-based approach. Conversely, the superior performance in the ground shaking analysis reflects self-consistency rather than true predictive adequacy.

Incorporating additional constraints from paleoseismic records or geological indicators into model evaluation would help extend the applicability and robustness of the models for hazard curve testing, and would also introduce independence in area source consistency checks. For instance, in the EBSZ case, geological data is not used to define area source parameters, hence a consistency test with paleoseismic data would be fully independent.

### 4.3.3. Magnitude threshold

A limitation of our work, like in most fault-based PSHA studies, is the selection of the  $M_w$  5.5 threshold above which ruptures are considered for the analysis, because with such selection comes the assumption that all earthquakes above that threshold are caused by the faults in the model. At the EBSZ, especially within the area source (IGME, 2015), there is consensus that most moderate-to-large earthquakes are caused by the main known fault systems, supported by decades of seismicity and paleoseismic research (e.g., Sanz de Galdeano, et al., 2020). As such, most  $M_w \geq 5.5$  earthquakes at the EBSZ have been attributed to the faults considered in this study (e.g., Yazdi and García-Mayordomo, 2024). Despite this, our magnitude threshold might also be debatable because  $M_w < 5.5$  earthquakes have also been caused by EBSZ faults, the most recent one being the 2011  $M_w$  5.1 Lorca earthquake (Martínez-Díaz et al., 2012). Leaving such earthquakes out removes the opportunity to test the models against non-zero observations at different stations. To tackle this limitation, we performed an alternative analysis considering a lower  $M_w$  5.0 threshold. This implies extracting all ruptures of  $M_w \geq 5.0$  from the simulated catalogues and re-computing hazard calculations as well as the consistency tests. The main difference in this case is that the macroseismic data includes intensity records from the 2011  $M_w$  5.1 Lorca earthquake and station observations are not all zero, but include the PGA values from this same event. Macroseismic and station records are extracted from the IGN (2025) database for the different stations and PGA for the 2011 earthquake is computed as the mean between N-S and E-W accelerometer components (Table S3). All analyses in figures from this secondary test are available in the Supplements (Figures S3 to S6).

Considering a lower magnitude threshold rejects the poissonian process for Cat-18 (Fig. S3), meaning that earthquake occurrence in this catalogue cannot be treated as time-independent. Consequently, the subsequent hazard analysis (based on the Poisson assumption) is less meaningful. Despite this, the analysis raises some insightful points as we discuss.

In this test, the area source model outperforms both simulated catalogues for the combined macroseismic and PGA consistency analyses, although performances are very similar between Cat-21 and the areas model. In general discrepancies between



observations and models are larger, evidenced by larger  $\text{Log}P$  values (Fig. S6), but smaller between models. Although the score rank patterns observed for the Mw 5.5 assumption are similar (i.e., overforecasting and underforecasting of observations; Figs. S4 and S5), here the areas model performs better in more intensity measure level classes, particularly for macroseismic data, while Cat-21 steps down as the best-performing model in some locations. As such, for instance, cities like Lorca, Murcia and Torre Vieja all perform better in the area source model for the MI 6 class, while Cat-21 is not the best performing model anymore in Lorca for  $\text{MI} \geq 7$  (Fig. S6).

The larger (worse)  $\text{Log}P$  scores and the overall better performance of the area source model suggests that the consideration of magnitudes below Mw 5.5 on faults (simulated catalogues) is less appropriate. That is, the magnitudes below this threshold might be related to other fault sources not considered in the current study (e.g., secondary faults) and, therefore, they do not necessarily follow the earthquake rates constrained by the activity parameters of the faults considered. In addition, the overall increase in the  $\text{Log}P$  scores for the area source model compared to the previous models suggests that Mw < 5.5 are less adequate to explain observations within the area and might evidence issues in the seismicity assignment to the EBSZ area source 55. Smaller magnitudes increase epicentral location uncertainties for historical events, which compose most of the Mw > 5 seismicity assigned to area source 55. In fact, 67% of the Mw  $\geq 4$  earthquakes at the EBSZ are previous to 1962 with epicentral errors larger than 10-20 km (García-Mayordomo, 2005; Gómez-Novell et al., 2020a). Thus, lowering the magnitude threshold increases the chances of inaccuracies in the seismicity assignment to the area source, that might explain the lower performance. Cat-18 remains the model with the poorest performance, further highlighting the relevance of a proper benchmarking of the simulations prior to the hazard assessment.

Accounting for the magnitude threshold limitations depicted in this analysis is important for PSHA, and underscores the importance of combining fault-based approaches with distributed source methods to better capture the whole range of epistemic uncertainties in the hazard, especially for regulatory hazard evaluations.

#### 4.4 Implications for low-to-moderate strain regions

We evaluate the performance of seismic hazard estimates from physics-based earthquake rupture forecasts in southern Spain, compared to a traditional area source model. To do so, we compare the hazard models to empirical observations on macroseismic intensity and ground shaking in the region. The physics-based models inherently incorporate fault rupture complexity including multi-fault ruptures, and long-term fault interaction across seismic cycles in 50-kyr long catalogues. When physics-based models properly match empirical relationships, we find that the physics-based models outperform the area source model for the macroseismic data, while the area source model shows better consistency for the ground shaking tests. These results suggest that the physics-based approach can estimate long-term hazard better than area sources, a fact that can be of high value in regions like the EBSZ, where long-term fault behaviour is largely unknown, and hazard estimates are usually restricted to the available seismic catalogues. Moreover, the physics-based models allow a better spatial representation



of the hazard along the fault zones, which can help to prioritize hazard mitigation strategies and territorial planning in the region.

As opposed to other fault-based approaches, such as analytical developments based on seismicity rate modeling, earthquake cycle simulators inherently incorporate physical constraints on fault rupturing into the simulated earthquake catalogues. On the one hand, this minimises potential biases in expert judgment introduced during the definition of earthquake rupture complexity and fault interaction. On the other hand, these simulators allow to model earthquake catalogues beyond the recorded seismicity. This feature is particularly useful at the EBSZ, where the historical earthquake catalogue is scarce and imprecise compared to the long recurrence periods of active faults (Herrero-Barbero et al., 2021).

Nevertheless, area source models remain a valuable component of probabilistic seismic hazard analysis, even if they are incomplete for large events in the instrumental records. In this study, the area source model shows good consistency with PGA exceedances at short observational periods and second-to-best performance in the macroseismic intensity tests (above Cat-18). Their flexibility, broad applicability, and robustness, especially at low shaking thresholds or in regions where detailed fault data are lacking, make them essential tools in regional-scale assessments. In particular, their tendency to smooth the hazard over space can provide conservative estimates in data-scarce zones and complement the more localized/detailed outputs of fault-based approaches. This would be particularly relevant for a larger-scale regional approach, since outside the EBSZ most faults in southern Spain lack the detailed information required to support the type of modeling presented here.

The overall good agreement between models and observations, indicates that physics-based and area source models should be viewed as complementary, rather than competing. A combined use of both, for instance, within a logic-tree framework, may offer a more complete representation of epistemic uncertainties, highlighting the strengths of each modeling approach. In this context, the integration of physics-based earthquake rupture forecasts into broader hazard modeling strategies can enhance the resolution and accuracy of hazard estimates, without compromising the stability offered by more traditional models.

We recommend integrating physics-based PSHA models to better account for large lower-frequency events combined with traditional approaches that model historical and instrumental data for background low-to-moderate magnitude events, within an ensemble or logic tree framework. This would enable the quantification of epistemic uncertainties linked to data and modeling approaches, while retaining the complexity of rupture processes inherent to earthquake cycle models. In addition to that, physics-based models can explore further important sources of uncertainty, such as the temporal variability in the hazard, and therefore address known limitations of current approaches, as noted by Gerstenberger et al. (2020). This can be achieved by comparing simulated catalogues of variable durations (e.g., 10 kyr vs. 50 kyr), providing insights into the time dependence of hazard. Finally, coupling RSQSim with dynamic ground-motion models (e.g., Milner et al., 2021) may improve shaking estimates, particularly for near-fault sites.





## 5. Conclusions

This study presents the development and evaluation of physics-based probabilistic seismic hazard models in the Eastern Betic Shear Zone (EBSZ), a low-to-moderate strain fault system in southern Spain. Using long-term RSQSim earthquake cycle simulation models by Herrero-Barbero et al. (2021), we construct two earthquake rupture forecasts (ERF) corresponding to the best and worst performing catalogues according to a benchmark analysis: Cat-21 and Cat-18, respectively. We pipeline them into a PSHA calculation using the OpenQuake Engine and compare their hazard outputs with those from a classical area source model (ZESIS model). The performance of the three models is evaluated with consistency tests against both macroseismic intensity observations over the last millennium and ground motion records.

Our results show that the Cat-21 model provides the best agreement out of all three, outperforming both the lower-ranked Cat-18 and the area source model, particularly at sites located near the faults with faster rates, and across the range of intensity thresholds and selected ground-motion levels. The performance of Cat-21, especially to long-term macroseismic data, highlights the value of incorporating complex fault rupture geometries and long-term recurrence in PSHA, both inherent products of earthquake cycle simulators. These features are especially relevant in regions where empirical data are sparse or heterogeneously distributed, and therefore fault behaviour characterization becomes particularly challenging. Despite this, the classical area source approach also shows good agreement with observed macroseismic intensities and accelerometric exceedances, highlighting the complementarity of both approaches to better capture epistemic uncertainties in seismic hazard.

Physics-based models, when properly calibrated, produce credible large-scale hazard patterns and also capture the localized variability in seismic hazard associated with near-fault effects. Thus, our findings support that the integration of physics-based ERFs within more standardized PSHA workflows is feasible and reliable. First, earthquake simulators are able to introduce the physical features of fault behaviour into seismic hazard assessment, including complex fault ruptures and fault system scale long-term fault interaction. Second, these simulators allow to extend earthquake catalogues beyond instrumental and historical coverages, overcoming completeness limitations of seismicity catalogues and offering a complementary approach to purely empirical methods (e.g., Gutenberg-Richter distributions or spatial smoothing). Future developments should aim at combining simulator-based ERFs with advanced ground motion simulation tools to deploy fully physics-based PSHA. Additionally, paleoseismological data can be used for further validating model outputs.

Earthquake cycle simulators represent robust tools ready to be implemented in next-generation PSHA, not only for low-strain, tectonically complex regions, but also across fault systems worldwide. Importantly, our study pinpoints that physics-based models require proper benchmarking against empirical seismic and geological observations before being pipelined in a PSHA calculation.



## 685 **Code availability**

The RSQSim code, developed by Richards-Dinger and Dieterich (2012), can be obtained from the authors upon request. The code to convert earthquake simulator ruptures to OpenQuake format inputs is stored at a GitHub repository: [https://github.com/octavigomez/simulator\\_to\\_openquake](https://github.com/octavigomez/simulator_to_openquake)

## **Data availability**

690 Seismic hazard models, including input files and hazard results, are available in a Zenodo repository under the CC-BY 4.0 license (Gómez-Novell et al., 2025b).

## **Author contribution**

**OGN:** Conceptualization, Methodology, Software, Validation, Formal Analysis, Investigation, Data Curation, Writing (original Draft Preparation), Visualization; **FV:** Conceptualization, Methodology, Software, Validation, Formal Analysis, Investigation, Data Curation, Writing (original Draft Preparation), Visualization; **PHB:** Investigation, Data Curation, Writing (review and editing); **JAAG:** Investigation, Resources, Funding acquisition, Writing (review and editing); **JGM:** Investigation, Resources, Funding acquisition, Writing (review and editing).

## **Competing interests**

The authors declare that they have no conflict of interest.

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