



- 1 Ensemble-based data assimilation improves hyperresolution 2 snowpack simulations in forests.
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17 Abstract

18 Snowpack dynamics play a key role in controlling hydrological and ecological 19 processes at various scales, but snow monitoring remains challenging. Data 20 assimilation techniques are emerging as promising tools to improve uncertain 21 snowpack simulations by fusing state-of-the-art numerical models with 22 information rich, but noisy observations. However, the occlusion of the ground 23 below the forest canopy limits the retrieval of snowpack information from remote 24 sensing tools. Remote sensing observations in these environments are spatially 25 incomplete, impeding the implementation of fully distributed data assimilation 26 techniques. Here we propose different experiments to propagate the information 27 obtained in forest clearings, where it is possible to retrieve observations, towards 28 the sub-canopy, where the point of view of remote sensors is occluded. The 29 experiments were conducted in forests within Sagehen Creek watershed 30 (California, USA), by updating simulations conducted with the Flexible Snow Model 31 (FSM2) using airborne lidar snow data using the Multiple Snow data Assimilation





32 system (MuSA). The successful experiments improved the reference simulations 33 significantly both in terms of validation metrics (correlation coefficient from R=0.1 34 to R ~0.8 on average) and spatial patterns. Data assimilation configurations using 35 geographical distances and space of topographical dimensions, improved the 36 reference run. However, those creating a space of synthetic coordinates by 37 combining the spatiotemporal data assimilation with a principal components 38 analysis did not show any improvement, even degrading some validation metrics. 39 Future data assimilation initiatives would benefit from building specific localization 40 functions that are able to model the spatial snowpack relationships at different 41 resolutions.

42 1 Introduction

The seasonal snowpack is a crucial component in various ecological and 44 hydrological processes in mountain areas and cold regions (Han et al., 2024; 45 Slatyer et al., 2022), covering over 47 million square kilometers of the northern 46 hemisphere (Robinson & Frei, 2000) and 45% of global mountain areas (Gascoin et 47 al., 2024),. It has significant implications for both the economy and ecology of 48 these areas, as well as for downstream regions (Barnett et al., 2005; Qin et al., 49 2020; Sturm et al., 2017). However, accurately estimating the spatiotemporal 50 dynamics of the snowpack, in particular the snow water equivalent (SWE), remains 51 a challenging and unresolved issue (Tsang et al., 2022). These difficulties are only 52 increased in forested terrain, due to the complex relationships between snowpack 53 and canopy cover (Mazzotti, Essery, Moeser, et al., 2020).

54 The overlapping area between the snowpack and forested areas is estimated in at 55 least 19% of the terrain in the northern hemisphere, only accounting for the boreal 56 forest (Rutter et al., 2009). This estimation can only be higher considering the 57 overlapping area in alpine forests. Snow beneath the canopy behaves differently 58 than in open terrain (Dickerson-Lange et al., 2023; Safa et al., 2021; Varhola et al., 59 2010). One major process is the interception of snowfall in the forest canopy, 60 limiting under canopy snow accumulation compared to clearings (Essery et al., 61 2003). The intercepted snow will either sublimate, drip as liquid water or unload as 62 snow (Lundquist et al., 2021). In addition, the canopy cover changes the net 63 radiation available to melt the snowpack, both by shading the snow surface and 64 increasing the incoming longwave radiation (Lundquist et al., 2013). Generally, 65 this leads to increased ablation under the canopy in warmer environments from 66 longwave radiation compared to colder environments where shading from solar 67 radiation causes less ablation under canopy (Lundquist et al., 2013). This 68 relationship leads to differences between under canopy and open clearing





69 snowpack in most environments (Dickerson-Lange et al., 2017) that are 70 challenging to observe across complex terrain (Safa et al., 2021).

71 Direct observations of the snowpack under the forest are rare and challenging to 72 obtain (Kinar & Pomeroy, 2015). Deploying field based monitoring networks is a 73 complicated and expensive task. The harsh weather conditions of the remote areas 74 where the snowpack is present complicate the installation of monitoring networks, 75 with the number of automatic weather stations declining dramatically with higher 76 elevation (Matthews et al., 2020). Given the considerable complexity of the spatial 77 patterns of the seasonal snowpack, monitoring networks often suffer from a lack 78 of representativity (Herbert et al., 2024). In addition, monitoring SWE, which is the 79 key snow hydrological variable, is significantly more uncertain and costly to 80 measure than other variables such as snow depth, and remains an active research 81 topic (e.g., Gugerli et al., 2022; Orio-Alonso et al., 2023). The extensive spatial 82 extent of the seasonal snowpack and its temporal variability make monitoring 83 based on manual field campaigns challenging to deploy.

84 Remote sensing techniques are well established as snow cover monitoring tools 85 (Gascoin et al., 2024). Due to different remote sensing initiatives, it is possible to 86 monitor the dynamics of the snowpack even at continental scales at frequencies 87 approaching real time. Despite being traditionally restricted to the measurement 88 of snow cover properties such as the snow cover extent, fractional snow cover or 89 snow surface temperature, it is now possible to retrieve the snow depth, from 90 photogrammetry and lidar based sensors installed in traditional or remotely 91 piloted airborne platforms, or orbital sensors (Deschamps-Berger et al., 2020; 92 Harder et al., 2020; Painter et al., 2016). Unfortunately, most of these retrievals are 93 limited to observations in open terrain or clearings in forested areas, being limited 94 either spatially or temporally. Recent experiments based on airborne campaigns 95 have proven the potential of X- and Ku-band SAR technology to retrieve SWE 96 (Montpetit et al., 2024; Singh et al., 2024), a technology that is expected to be 97 implemented in the next generation of satellites in the near future (Derksen et al., 98 2021). Unfortunately obtaining SWE observations in dense forest areas will remain 99 problematic (Tsang et al., 2022). One partial solution to observing snow under the 100 canopy is with airborne lidar systems that can partially penetrate the canopy and 101 retrive the snow surface elevation. Recent work has processed lidar point clouds 102 to resolve under canopy snowpack and validated the results against field 103 observations (Kostadinov et al., 2019; Safa et al., 2021). Refinements to this 104 method offer promise for better resolving lidar returns from low canopy with the 105 snow surface (Piske et al., 2024) and creating datasets that can be used to train 106 models and improve other remote sensing snow products.





107 Numerical modeling of the snowpack allows simulating the complete state of the 108 snowpack, including the SWE, at any spatiotemporal resolutions. Modern 109 snowpack models of increasing complexity even represent the horizontal transport 110 of the snow caused by wind and avalanches, and the interactions with forests 111 (Mazzotti et al., 2020; Vionnet et al., 2021). However, numerical models often rely 112 on adjustable parameters to represent different physical processes, whose 113 transferability between different areas and model resolutions is usually complex, 114 leading to uncertain simulations (Essery et al., 2013). In addition, these models rely 115 on high resolution meteorological forcings, that are very challenging to generate 116 and constrain, in part due to the lack of dense in situ observations. An alternative 117 is to use meteorological downscaling techniques based on limited-area 118 atmospheric models (Alonso-González et al., 2021; Sharma et al., 2023). However, 119 the computational cost of regional atmospheric models increases significantly with 120 finer resolution, with the current state of the art at the kilometer scale (Rasmussen 121 et al., 2023). As such, dynamical downscaling is not yet a tractable option to couple 122 with high and hyper resolution snowpack simulations. A partial, and very 123 widespread, solution to this problem is to use simplified downscaling models that 124 rely on different assumptions and/or empirical approximations to generate high 125 resolution meteorological forcing fields. These may be predefined temperature 126 and precipitation lapse-rates, or using empirical relations between the 127 atmospheric variables and the underlying terrain (Fiddes & Gruber, 2014; Liston & 128 Elder, 2006; Reynolds et al., 2023). Despite their simplicity, these more heuristic 129 approaches may lead to a performance comparable with dynamically downscaled 130 meteorological products (Alonso-González et al., 2023; Gutmann et al., 2012; Kruyt 131 et al., 2022). Nonetheless, any (often considerable) remaining uncertainty in the 132 forcing will, together with the uncertainty in the snow model structure and 133 parameters, be propagated to the snowpack simulations, typically leading to 134 simulations that differ significantly from reality (Krinner et al., 2018).

135 Data assimilation (DA) is the exercise of merging noisy observations with uncertain 136 numerical models to exploit the strengths of both worlds (Evensen et al., 2022). 137 Thanks to DA, It is possible to constrain model uncertainty using partial information from 138 snowpack observations (Largeron et al., 2020). Although DA may not be as 139 widespread in the snow sciences as in other disciplines, its use is becoming more 140 common with a number of operational and experimental initiatives (e.g. Girotto et 141 al., 2024; Mott et al., 2023). Using DA, it is possible to infer uncertain parameters to 142 improve the simulations so as to better match the observations, providing an 143 estimation of the model uncertainty. However, snow DA is still rarely used in 144 forested areas due to the lack of reliable remote sensing observations of the 145 snowpack under the canopy.





146 Canopy cover impedes the direct observation of the snowpack from space or 147 airborne sensors, which collaterally hampers the use of DA, and may even degrade 148 simulation outputs if implemented in its simplest form (Yatheendradas et al., 149 2012). This is probably the reason that the majority of snow DA experiments have 150 been limited to arctic or alpine areas above the treeline, with only some 151 experiments approaching specifically the topic of snow DA in forested areas. 152 Smyth et al. (2022) tested the potential of a particle filter DA algorithm to improve 153 snowpack simulations generated by the Flexible Snow Model (FSM2) in the 154 presence of observations beneath the canopy. The results show that simulations 155 can be improved by assimilating data in snow models that consider canopy 156 interactions. However, the question of how to improve simulations of the 157 snowpack in case of a total occlusion of the snow view in certain regions of the 158 simulation domain (i.e. lack of local observations) remains unanswered. Pflug et 159 al., (2024) proposed a simplified three dimensional DA scheme to update the SWE 160 state variable at unobserved locations from remote observations in forest gaps 161 and tested their approach with a synthetic observing system simulation 162 experiment (OSSE). First, they used a one dimensional (purely temporal) Ensemble 163 Kalman Filter (EnKF) to update the cells where observations exist. In a second step 164 they updated the local unobserved pixels SWE using the ratio of the average 165 observations and average modeled SWE within a spatial window, generating a new 166 observation to be assimilated. Due to its simplicity, this heuristic procedure 167 succeeded in performing a promising synthetic assimilation experiment over a 168 very large area of North America at an affordable computational cost. Cho et al. 169 (2023) assimilated spatially coarse (5km2) airborne gamma ray based SWE 170 retrievals in forested environments, using a three-dimensional EnKF. These recent 171 works lay the foundations of snow data assimilation in forests, with great potential 172 to (i) improve snowpack simulations in forested watersheds, (ii) better understand 173 snow-forest processes, and (iii) identify shortcomings in snow-forest model 174 parameterizations. However, these previous works are based on necessarily 175 simplified approximations to limit the computational cost, synthetic experiments 176 or very coarse resolutions unable to capture the spatial variability present in 177 montane forests (Safa et al., 2021; Tennant et al., 2017). The emergence of new 178 technologies that allow the acquisition of snowpack observations at high and 179 hyper resolutions (Gascoin et al., 2024), make it necessary to adapt classical DA 180 techniques to maximize the value of the available information.

181 The interactions between the canopy and the snowpack behavior pose challenges 182 for inferring the snow mass beneath the canopy directly from nearby observed 183 locations in forest clearings, preventing simple interpolation techniques 184 (Dharmadasa et al., 2024) or DA techniques designed to update the model states





185 directly from the information obtained in nearby cells to work efficiently in this 186 context (Pflug et al., 2024). It is necessary to explore how the available information 187 can be transferred from the available observations in forest clearings to beneath 188 the canopy, where observations are typically either missing or highly uncertain. In 189 this work, we test a recently developed spatio-temporal snow DA methodology et al., 2023), specifically designed to update distributed 190 (Alonso-González 191 snowpack simulations from spatially incomplete observations such as in a forest 192 environment where the information from remote sensors is mostly available in 193 forest clearings. We combine that information with a unique post-processed lidar 194 dataset that resolves the under-canopy snowpack explicitly (Kostadinov et al., 195 2019; Piske et al., 2024) to validate the model. The objective of this work is (i) to 196 explore the potential of lidar-derived real observations to update distributed 197 snowpack simulations at hyperresolution (10 m) scales in forest environments, and 198 (ii) to test different spatiotemporal DA configurations for estimating snow under 199 the canopy when only observations in forest gaps are available. Here we propose 200 different spatio-temporal DA configurations to propagate information under the 201 canopy where the observations are often not available.





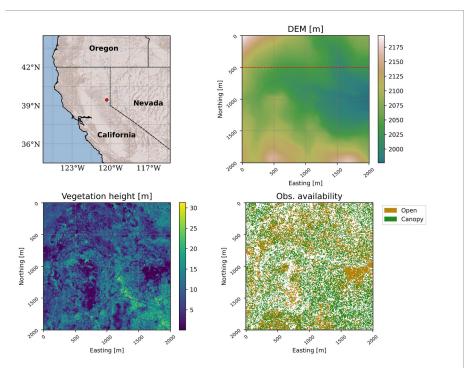
203 2 Data and Methods

204 2.1 Observed snow depth maps, vegetation parameters and 205 meteorological forcing

206 The experiments proposed in this work were developed in the Sagehen Creek 207 forest (California, USA, Fig. 1). The observations consist of one airborne LiDAR 208 derived snow depth map collected by the National Center for Airborne Laser 209 Mapping on 21 March 2022 (Piske, 2022) and a snow off flight (Graup, 2021). From 210 all the available areas, we have manually selected a domain of ~2x2 km that 211 maximizes canopy heterogeneity and the observed snowpack data that are 212 incomplete due to dense canopy cover, since pixels with potential vegetation-213 snowpack conflict were removed to increase the confidence in the snow depth 214 data. The Sagehen Creek site was used to develop and test a new method of under 215 canopy snow depth detection from airborne lidar (Kostadinov et al., 2019) that 216 resolves a considerable amount of snow information beneath the canopy (Fig. 1). 217 We use a slightly improved method to extract vegetation from the snow surface 218 described in Piske (2022). Based on nearby SNOTEL at a similar elevation (SNOTEL 219 Site: Independence https://wcc.sc.egov.usda.gov/nwcc/site? 539, Camp, 220 sitenum=539, last accessed: 11-Nov-2024), the SWE was 43 cm on 21 March when 221 lidar was collected compared to a maximum annual SWE of 48 cm on 9 March, 222 2022. The native spatial resolution of the lidar dataset was 1 m which was 223 resampled to 10m for use in the DA analysis. The error variance of the 224 observations was assumed to be σ^2 = 0.01 m² at 10m resolution based on previous 225 airborne LiDAR snow experiences that reported similar error metrics (Currier et al., 226 2019; Harpold et al., 2014; Mazzotti et al., 2019; Painter et al., 2016). Future 227 initiatives may benefit from more sophisticated error models. In addition to the 228 snow depth observations, different vegetation parameters were computed from 229 the three-dimensional lidar data, including vegetation height, the Vegetation area 230 index and the in forest sky view factor based on methods described in Broxton et 231 al., (2015) and Broxton et al. (2021). This dataset was segmented into grid cells in 232 forest clearings (to be assimilated) and canopy-covered cells (to be used as 233 independent validation) based on this vegetation information. The meteorological 234 forcing was generated using MicroMet (Liston & Elder, 2006) forced by the ERA5 235 atmospheric reanalysis (Hersbach et al., 2020). The meteorological fields were 236 downscaled to the same geometry of the observations using a LiDAR based digital 237 elevation model (Sourp et al., 2024). The precipitation partitioning was estimated 238 using the psychrometric parameterization scheme proposed by Harder & Pomeroy 239 (2013).







240 Figure 1. Localization map, Digital elevation model, vegetation height and available observations 241 (with its segmentation between canopy covered data used for validation or forest gaps to be 242 assimilated). The red transect in the digital elevation map indicates the location of the profile used 243 later for validation.

244 2.2 Data assimilation and computational setup





245 All DA experiments presented in this work were developed using the Multiple 246 Snow data Assimilation (MuSA) system (Alonso-González et al., 2022). MuSA is an 247 open-source DA toolbox designed primarily as a python wrapper around the 248 Flexible Snow Model (FSM2, Essery et al. (2024), but now providing support for 249 other numerical models as well while not necessarily being limited to snowpack 250 models. MuSA provides support to different DA algorithms, and simplifies the 251 implementation of new ones thanks to its modular design. In this work, the FSM2 252 model was chosen due to its already coupled canopy module that required only 253 minimal modifications of the original MuSA code to be activated. MuSA, and 254 therefore FSM2, was forced by the Micromet outputs, and provided with the 255 aforementioned lidar-derived vegetation parameter maps. It should be noted that 256 although in this work we focus on the MuSA snow depth outputs (as this is what 257 we can validate) posterior simulations include the full state vector of FSM2, 258 including SWE.

259 The spatio-temporal DA scheme is described in Alonso-González et al. (2023), and 260 therefore we only briefly introduce some key points here, its configuration, and 261 the new modifications implemented to improve its performance for the new 262 problem at hand. We only infer meteorological correction parameters and not 263 model states, leading to physically consistent (in terms of FSM2) simulations of the 264 modeled snow state across the snow season. As mentioned above, snowpack in 265 the forest gaps shows a different behavior than beneath the trees (Varhola et al., 266 2010), so trying to infer canopy-occluded states directly from the information we 267 can obtain in the gaps could also introduce artifacts in the simulations. Crucially 268 the forcing perturbations will also be modified by the canopy scheme in FSM2, so 269 even if the above canopy forcing is constrained to be similar for neighboring cells 270 the forcing that the below canopy snowpack experiences will be different thanks to 271 the model physics.





272 As the first step in our workflow, we generated an ensemble of 100 FSM2 273 simulations by randomly drawing stationary (i.e. constant across the water year) 274 spatially correlated prior parameters to perturb the meteorological forcing, 275 particularly the precipitation and 2m air fields. The choice of perturbing only 276 precipitation and temperature was motivated by previous successful experiments 277 with a similar setup, albeit in non-forested environments (Alonso-González et al., 278 2023; Alonso-González et al., 2022). Herein, the prior probability distributions that 279 we sampled using a random number generator were: a normal (additive 280 parameter) for the temperature bias and lognormal (multiplicative parameter) for 281 the precipitation scaling. These prior distributions were defined by: its mean ($\mu = 0$) 282 and standard deviation (σ = 1) in the case of the temperature and by the mean and 283 standard deviation ($\mu = 0$, $\sigma = 0.4$) of the *underlying* normal distribution in the case 284 of the precipitation. The latter results in log-normally distributed prior 285 multiplicative precipitation scaling parameters in the physical space whose median 286 is ~1. The objective of the algorithm is to update these parameters by assimilating 287 observations to directly correct the temperature and precipitation fields and 288 indirectly update the corresponding snowpack states

289 For this purpose we have used a deterministic ensemble Kalman filter (DEnKF)-290 based algorithm in iterative smoother mode, namely the Deterministic Ensemble 291 Smoother (DES, Sakov & Oke, 2008) with multiple data assimilation (DES-MDA, 292 Emerick, 2018). In this DES-MDA scheme, the update proceeds in two steps for 293 each grid cell $i=1,\dots,N_g$ and $\ell=0,\dots,(N_a-1)$ MDA iteration. Firstly, the $N_p\times 1$ 294 updated ensemble mean parameter column vector $\overline{\boldsymbol{\theta}}_{\ell+1}^{(i)}$ is obtained using a Kalman 295 analysis equation of the form

$$\overline{\boldsymbol{\theta}}_{\ell+1}^{(i)} = \overline{\boldsymbol{\theta}}_{\ell}^{(i)} + \mathbf{K}_{\ell}^{(i)} \left[\mathbf{y}^{(i)} - \overline{\mathbf{y}}_{\ell}^{(i)} \right]$$





296 where $\overline{\theta}_{\ell}^{(i)}$ is the $N_p \times 1$ ensemble mean parameter column vector from the current 297 (prior for $\ell=0$) iteration, the $N_p \times N_o^{(i)}$ matrix $\mathbf{K}_{\ell}^{(i)}$ is a localized and inflated 298 ensemble Kalman gain computed using ensemble covariances and the observation 299 error covariance, $\mathbf{y}^{(i)}$ is the $N_o^{(i)} \times 1$ local observation vector containing available 300 local observations that are within a yet to be defined distance-based neighborhood 301 d < 2c (see the GC function below) of grid cell i, and the $N_o^{(i)} \times 1$ vector $\overline{\mathbf{y}}_{\ell}^{(i)}$ contains 302 the corresponding local ensemble mean predicted (i.e. modeled) observations 303 from the last iteration obtained at neighboring grid cells. We refer to Alonso-304 Gonzalez et al. (2022) for the full form of the ensemble Kalman gain matrix in 305 particular and a more detailed overview of the implementation of spatio-temporal 306 DA using the DES-MDA in MuSA in general. Secondly, the $N_p \times N_e$ matrix $\Theta_{\ell+1}^{(i)\prime}$ 307 containing the updated ensemble of parameter vector anomalies (from the mean) 308 is obtained a modified Kalman analysis equation of the form

$$\boldsymbol{\Theta}_{\ell+1}^{(i)'} = \boldsymbol{\Theta}_{\ell}^{(i)'} - 0.5 \mathbf{K}_{\ell}^{(i)} \left[\mathbf{y}^{(i)} \mathbf{1}_{N_o^{(i)}}^{\mathrm{T}} - \widehat{\mathbf{Y}}_{\ell}^{(i)} \right]$$

309 where $\Theta_\ell^{(i)\prime}$ is the $N_p \times N_e$ matrix containing the ensemble of parameter vector 310 anomalies from the current iteration, $\mathbf{1}_{N_o^{(i)}}^{\mathbf{T}}$ is a $1 \times N_o^{(i)}$ row vector of ones, and $\widehat{\mathbf{Y}}_\ell^{(i)}$ 311 is the $N_o^{(i)} \times N_e$ matrix of predicted observations from the current iteration. Once 312 the mean and anomaly update steps have been carried out, the $N_p \times N_e$ matrix 313 $\Theta_{\ell+1}^{(i)}$ (without the prime) containing the updated ensemble of parameter vectors is 314 obtained through the matrix sum

$$\mathbf{\Theta}_{\ell+1}^{(i)} = \overline{m{ heta}}_{\ell+1}^{(i)} \mathbf{1}_{N_p}^{\mathrm{T}} + \mathbf{\Theta}_{\ell+1}^{(i)\prime}$$

315 where ${}^1{}^1{}^N{}_p$ is a ${}^1 imes N_p$ row vector of ones. Unlike the classic stochastic (perturbed 316 observation) ensemble Kalman scheme, this deterministic ensemble Kalman 317 scheme is less overconfident thanks to built-in model covariance inflation and also 318 avoids the need to factorize the observation error covariance that can be costly in 319 spatio-temporal problems (Emerick, 2018). In the loop over iterations above we 320 implicitly rerun the forward model, FSM2 in this case, with the updated parameter 321 values to generate an updated ensemble of hidden snowpack states including the 322 predicted snow depth observations to be assimilated.





323 The Gaussian assumptions inherent in this ensemble Kalman method make it 324 more robust against ensemble collapse (where a single member carries all the 325 posterior probability) than particle methods which are more widely used for snow 326 DA (Alonso-Gonzalez et al., 2022). In particular, we have used an iterative version 327 of DES, that performs the update of the parameters in multiple data assimilation 328 (MDA) steps, creating the DES-MDA used here (Emerick, 2018). The MDA is a form 329 of likelihood tempering (Murphy, 2023)that helps relax the undesirable effects of 330 the linear assumption inherent to EnKF based algorithms. In nonlinear DA 331 problems such as the one tackled here, previous work has shown that these MDA 332 iterations lead to significant improvement of the results compared with non-333 iterative versions of the algorithm(Aalstad et al., 2018; Alonso-González et al., 334 2022). In this work, based on previous studies (Alonso-González et al., 2022 and 335 references therein), the number of iterations was fixed to 4. To accommodate the 336 Gaussian assumption, we employed analytical Gaussian anamorphosis (Bertino et 337 al., 2003) to log transform the precipitation parameter distribution to a normal 338 distribution and perform the update in Gaussian space. After the update, the 339 parameters are mapped back to the model space using the exponential function 340 before generating the new ensembles.





341 The spatial propagation of information may happen through two main 342 mechanisms in the DES-MDA: observation error correlations or prior correlations 343 (van Leeuwen 2019). Since observation error correlations are more challenging to 344 specify and arguably less general than prior correlations, we will focus only on the 345 latter. A key component of the scheme is to draw random prior parameters for 346 each cell that are correlated with other cells in the domain, reflecting similarities 347 among different regions of the simulation domain. In the general DA literature, 348 this is typically done by computing the pairwise geographic (Euclidean) distance to 349 map the proximity of the cells. The pairwise distance matrix is then used to 350 generate a covariance matrix. In this work we have used the 5th-order piecewise 351 rational function proposed by Gaspari and Cohn (GC) (Gaspari & Cohn, 1999), as is 352 often done in DA to generate and localize the covariance matrix. The GC 353 localization function depends on an important hyperparameter, the correlation 354 length scale, that in practice controls how far information can be transfered 355 spatially. Crucially, this length scale willaffect both the posterior results and the 356 computational cost since a larger length scale results in a greater number of 357 neighbors with non-zero correlation. The GC function defines a distance-based 358 correlation

$$\rho\left(d,c\right) = \begin{cases} -\frac{1}{4} \left(\frac{d}{c}\right)^5 + \frac{1}{2} \left(\frac{d}{c}\right)^4 + \frac{5}{8} \left(\frac{d}{c}\right)^3 - \frac{5}{3} \left(\frac{d}{c}\right)^2 + 1\,, & \text{for } 0 \leq \left(\frac{d}{c}\right) \leq 1\,, \\ \frac{1}{12} \left(\frac{d}{c}\right)^5 - \frac{1}{2} \left(\frac{d}{c}\right)^4 + \frac{5}{8} \left(\frac{d}{c}\right)^3 + \frac{5}{3} \left(\frac{d}{c}\right)^2 - 5 \left(\frac{d}{c}\right) + 4 - \frac{2}{3} \left(\frac{d}{c}\right)^{-1}\,, & \text{for } 1 \leq \frac{d}{c} \leq 2\,, \\ 0\,, & \text{for } \frac{d}{c} > 2\,. \end{cases}$$





359 where, d is the pairwise distance between cells and c is the correlation length 360 scale. This function is used for localization, with two important roles: first, it 361 reduces spurious long range correlations that arise due to the limited size of the 362 ensemble (Morzfeld & Hodyss, 2023), and second, to save considerable 363 computational costs since relatively distant locations can be ignored when 364 updating a particular cell. . Note that without localization, the spatio-temporal DA 365 problem would essentially be intractable, especially in this context with a relatively 366 large domain and a high spatial density of observations. In addition to ensemble 367 collapse, this is another motivation for using the ensemble Kalman method over 368 particle techniques here, since more developed localization methods exist for the 369 former (Evensen 2022). Despite being the typical spatial snow DA configuration 370 (e.g. De Lannoy et al., 2012; Magnusson et al., 2014) and references herein), there 371 is no reason to restrict the distance mapping to the geographic (northing and 372 easting dimensions) space, since an arbitrary number of dimensions can be used 373 to define a feature space and generate the distance matrix. It is widely 374 acknowledged that snowpack redistribution is strongly dominated by the 375 topographic characteristics of the terrain, such as concavity, slope, and elevation 376 as well as vegetation parameters (e.g. Dharmadasa et al., 2023; Essery & Pomeroy, 377 2004; Revuelto et al., 2014; Zheng et al., 2019). In the context of snow DA, it is 378 possible to map the similarities between cells using a multidimensional feature 379 space of topographical (or any other) dimensions. The only two considerations to 380 be taken into account are that these feature dimensions may have different units, and 381 that they can be potentially correlated. This may generate a space of non-382 orthogonal dimensions where using the Euclidean distance directly may lead to a 383 spurious similarity mapping (Curriero, 2006). It is possible to overcome these 384 issues by using the Mahalanobis distance, a generalization of the Euclidean 385 distance that includes a covariance-based normalization attempting to address 386 these two problems in a single step. Alternatively, it may be possible to generate 387 other spaces using synthetic transformed orthogonal dimensions in a potentially 388 lower dimensional space from the previously scaled topographical dimensions 389 using a principal components analysis or multidimensional scaling approaches 390 (e.g. Aversano et al., 2019; Murphy et al., 2015), and compute the pairwise 391 Euclidean distance matrix in the new synthetic space.





392 Whichever approach is used to define the space that enables information to be 393 spread, it is necessary to generate a pairwise distance matrix to compute a prior 394 covariance matrix. The previous version of MuSA generated the complete distance 395 matrix, which is highly memory and time intensive with poor scalability. The 396 reason for this is that the computational cost and the size of the matrix scales 397 quadratically with the number of cells, further complicating subsequent linear 398 algebra operations. However, it is not necessary to compute the full distance 399 matrix since localization ensures that long distances will be ignored in the analysis 400 as the corresponding elements in the covariance matrix will be 0 beyond a certain 401 distance that is controlled by the GC length scale hyperparameter. This makes the 402 distance and the subsequent covariance matrix very sparse, opening new 403 possibilities to make the otherwise expensive prior sampling more tractable. As 404 such, in MuSA we have now implemented the capability of mapping the distances 405 using a k-dimensional tree (k-d tree) space-partitioning data structure, as 406 implemented in the SciPy python module (Virtanen et al., 2020). This allows MuSA 407 to ignore all distances beyond the GC hyperparameter value, generating a sparse 408 distance matrix. Unfortunately only Minkowski metrics (which includes the 409 Euclidean distance) are available so far with the k-d tree implementation. As such, 410 this method is not compatible with Mahalanobis spaces in the current MuSA 411 version, and therefore it was not used for all the experiments proposed here. In 412 addition, we have implemented the capability of computing the distance matrix 413 cell by cell, which has proven to be very memory efficient with a very manageable 414 loss of efficiency that is compatible with Mahalanobis, or any other, distance 415 metric. Since the distance matrix, and the generated prior covariance matrix, are 416 very sparse, we have now migrated most MuSA linear algebra routines to the 417 SciPy.sparse module. This allows for the use of sparse linear algebra, enabling us 418 to sample even in very large domains while, depending on the GC 419 hyperparameter, maintaining an affordable computational cost. All these 420 modifications are included in a new MuSA version (v2.2), compatible with the use 421 of arbitrary masks, even non-contiguous ones within the same simulation domain, 422 indicating over which cells to perform the analysis. This allows simulations to be 423 performed only in the areas of interest such as. above a certain elevation or within 424 a certain complex basin geometry), while still performing spatio-temporal 425 assimilation by propagating the information between the selected cells at a 426 considerably reduced computational cost.





427 The last step of the prior sampling requires approximating the square root of the 428 covariance matrix via Cholesky factorization. As noticed by previous research 429 (Alonso-González et al., 2023; Curriero, 2006), the use of non-Euclidean distances 430 (e.g. using the Mahalanobis distance) leads easily to non-positive definite 431 covariance matrices, making it impossible to compute the Cholesky factor. We 432 have increased the numerical stability of the prior sampling in MuSA by 433 regularizing the prior covariance matrix, adding small values to the elements of its 434 diagonal. These diagonal elements are increased iteratively up to a limit defined by 435 the user (from 1e-6 to a maximum of 0.1 in this study), following a technique 436 known as jitter as is commonly done in the Gaussian Process machine learning 437 community (Neal, 1999; Rasmussen & Williams, 2005). The remaining steps, 438 including the DES-MDA update itself, remain the same as in the previous version of 439 MuSA, despite a few minor updates with the intention of improving the I/O 440 performance by optimizing the compression routines. All these modifications are 441 packed as a new version, whose code has been released together with this work 442 (Alonso-González et al., 2024).

443 2.3 Experimental design

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444 We propose different experiments to evaluate the potential of ensemble-based 445 data assimilation techniques to update hyperresolution simulations in forest 446 environments. First, as a reference, we generated a deterministic reference run 447 without any DA for comparison with the updated simulations. Then, different 448 experiments were developed in an effort to find a MuSA configuration that is able 449 to exploit dispersed hyperresolution information in forested terrain. Here we are 450 not aiming to find a generalistic optimal configuration, since each specific case will 451 require a different configuration, depending on the resolution of the simulations, 452 the spatial density of the observations, the domain, and the availability of 453 computational resources. We propose 3 different information propagation 454 schemes, and two different GC hyperparameters for each, leading to 6 different 455 simulations:

• Using Euclidean distances in the geographical space. We developed two different simulations where the Euclidean distance over the northing and easting dimensions is used to map the similarities among cells, using the values of 50 (Eu50) and 100 (Eu100) m for the GC hyperparameter.



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- Using the Mahalanobis distance in a topographical space. Here, we propose two experiments where in addition to northing and easting, we included elevation, the Topographic Position index, the Diurnal Anisotropic Heat index and the slope to define a topographical space. Since we have separated the data beneath the canopy and in the forest gaps, using them for assimilation and validation data, it is not instructive to include dimensions based on vegetation parameters. In fact, due to the GC function, it might even prevent the information transfer towards the canopy covered cells. The open cells that are geographically (or topographically) distant, and nearby geographically (or topographically) cells under the canopy, would be equally far away in Mahalanobis distance from a given open observed cell in that hypothetical space including vegetation parameters. The distances were computed using the Mahalanobis distance (Ma), and the GC hyperparameters tested were 0.5 (Ma0.5) and 1 (Ma1).
- Using Euclidean distances in a synthetic topographical space. Here we 474 included a PCA (after z-score standardization) analysis over the 475 topographical space to generate an orthogonal space that ensures a 476 positive definite covariance matrix by sorting the cells prior to computing 477 478 the covariance matrix. This saves significant computational cost since it allows for distance mapping using the new k-d tree implementation. The 479 number of principal components was selected automatically using the 480 algorithm proposed by Minka (2000), which in practice resulted in 5 481 components. The GC hyperparameters tested were 0.5 (PCA0.5) and 1 482 483 (PCA1).

484 For each of the experiments, we have computed the cell wise Continuous Ranked 485 Probability Score (CRPS, Hersbach, 2000), a generalization of the mean absolute 486 error for probabilistic simulations:

$$CRPS(F, x^*) = \int_{\mathbb{R}} [F(x) - H(x - x^*)]^2 dx$$

487 Where F(x) is the predicted cumulative distribution function of the snowpack state 488 variable x to be evaluated, x^\star is the reference (ground truth) value for the state 489 obtained from observations, and $H(x-x^\star)$ is the Heaviside function resulting in 1 if 490 $x \geq x^\star$ and 0 otherwise. We have used a normal approximation of the posterior 491 snow depth distribution defined from the posterior mean and standard deviation 492 derived from the ensemble together with the observations to compute the cell by 493 cell mean CRPS and standard deviation (SD). We have also computed the spatial





494 bias, which is the mean error of all cells used for validation, where error is the 495 difference between the posterior mean and the observations. In addition, we 496 computed the correlation (R) and root mean square error (RMSE) between the 497 posterior mean and observations across the domain. To evaluate the spatial 498 patterns of each of the experiments, we calculated the variograms of each 499 simulation. To quantify how far the variogram curves are from the one obtained 500 from the observations under the forest canopy, we used the discrete Frechet 501 distance (FrDist) as an indicator of similarity between the variogram curves.





502 3 Results

503 3.1 Validation metrics of the reference run and DA experiments

504 Compared with the deterministic reference simulation, both the Euclidean (Eu) and 505 Mahalanobis (Ma) experiments improved the quantitative error metrics 506 considerably (Table 1). The marked improvement in R (from R = 0.1 to R \sim 0.8 on 507 average for all the Eu and Ma experiments) is especially notable, and, combined 508 with the lower Frechet distance values (FrDist = 0.29 for the reference, while FrDist 509 = 0.005 on average for the Eu and Ma experiments), indicates a significant 510 improvement of the spatial patterns of the simulation. RMSE values also improved 511 significantly (RMSE improvement ~30%). The bias remained lower and close to zero 512 (bias mean = -0.07 m) for the reference simulation compared with the Eu and Ma 513 experiments (bias mean ~ 0.13 m), suggesting a slight overestimation of the snow 514 mass in the updated simulations. However, the RMSE in the reference run (RMSE = 515 0.32) compared with the Ma and Eu experiments (RMSE = 0.2) suggest many cells 516 in the reference run exhibit higher errors than the ones of the Eu and Ma 517 experiments. The CRPS, which is the only uncertainty-aware metric considered that 518 accounts for both the precision and accuracy of the ensemble, showed lower 519 values for the Eu50 (CRPS = 0.12), but followed closely by the other experiments, 520 except the PCA0.5 and PCA1.

521 Unfortunately, despite the convenience of using a PCA preprocessing step, the 522 experiments using PCA exhibited only a slight improvement in some metrics while 523 degrading other indicators. In particular, they exhibited a slight improvement in 524 the correlation values (R=0.20 and 0.46 for PCA0.5 and PCA1 respectively), while all 525 other metrics were similar to the reference (e.g. bias), with a FrDist being 526 equivalent or significantly degraded relative to the reference for the PCA0.5 (FrDist 527 = 0.021) and PCA1 (FrDist = 0.046), respectively. This suggests not only that 528 absolute error metrics were not improved, but even that spatial patterns were not 529 adequately simulated with the PCA approach.

530 Table 1: Validation metrics of the experiments

Ехр.	RMSE	R	Bias	CRPS [mean(+/- SD)]	FrDist
Ref.	0.32	0.10	-0.07	-	0.029
Eu50	0.20	0.84	0.12	0.12 (+/-	0.006



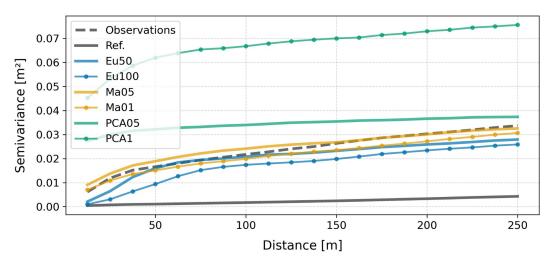


				0.09)	
Eu100	0.22	0.85	0.15	0.13 (+/- 0.12)	0.009
Ma0.5	0.22	0.76	0.10	0.14 (+/- 0.10)	0.003
Ma1	0.24	0.81	0.16	0.15 (+/- 0.12)	0.003
PCA0.5	0.33	0.20	-0.03	0.19 (+/- 0.13)	0.021
PCA1	0.33	0.46	0.08	0.21 (+/- 0.18)	0.046





531 Among the Eu experiments, Eu50 exhibited slightly better or similar error metrics 532 than the Eu100. However, the differences were minimal, suggesting there is 533 flexibility in choosing the GC hyperparameters, in this case at least, in terms of 534 validation metrics. A similar conclusion can be drawn from the validation metrics 535 of the Ma experiments, where there was not a clearly superior simulation. 536 Similarly, Eu and Ma yielded comparable performance according to these error 537 metrics. However, the FrDist metric was consistently lower in the Ma experiments 538 compared with the Eu experiments, suggesting a better representation of the 539 spatial patterns, while the remaining error metrics were slightly better or similar 540 for the Eu experiments. This superior performance in representing the spatial 541 patterns was evident in the snow depth semivariograms of the experiments (Fig.2), 542 where Ma experiments exhibited a semivariance much closer to the observations, 543 even reproducing accurately the nugget effect exhibited by the observations, 544 suggesting a better representation of the small scale patterns. In any case, the 545 variograms of the Eu and Ma experiments exhibit a closer shape to the one 546 obtained from the observations, compared with the one obtained in the reference 547 run, which is nearly flat.

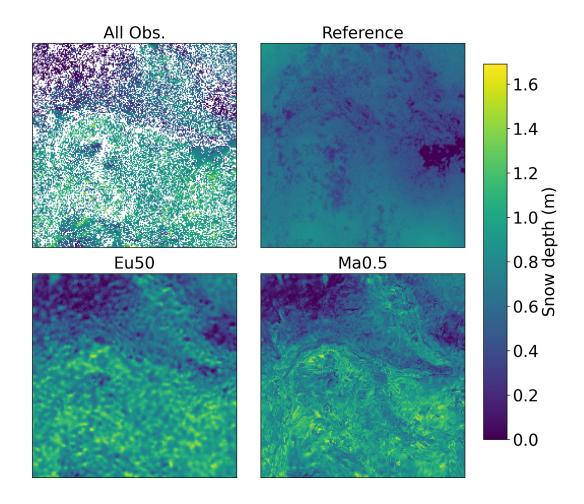


548 Figure 2: Snow depth spatial semivariance derived from the lidar-derived observations, the 549 reference run and different experiments.





When examining the distributed posterior mean simulations, these considerations for about the spatial patterns become evident (Fig. 3). First, there was a very limited spatial variability in the deterministic reference run, as reflected quantitatively by the Frechet distance and qualitatively by the variograms. Among the Eu50 and Ma0.5 posterior maps, there is a clear difference in its snow depth spatial patterns. While the large scale patterns were similar in both simulations, and close to the observations, the small scale patterns were different. In Eu50 small scale patterns for the posterior mean were clearly affected by the shape of the GC function, since the blurrier horizontal patterns are reminiscent of the Gaussian-like shape of this function. On the other hand, Ma0.5 small scale patterns, which do not depend solely on geographic distance, are considerably more intricate, which also explains the lower FrDist error metric.





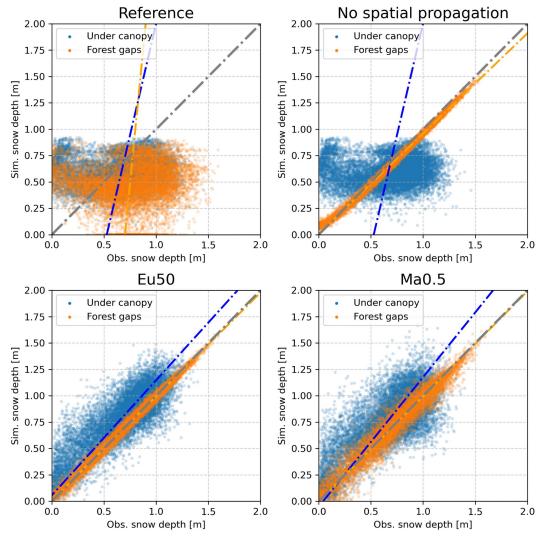


562 Figure 3: Distributed snow depth observations, reference simulation and posterior mean 563 simulations of the Eu50 and Ma0.5 experiments

564 While both in Ma0.5 and Eu50 point scale comparison with observations show a 565 similar overall R metric and distribution, it is worth noting the differences shown in 566 Fig.4. In Ma0.5, the cells with local observations (i.e. the cells in the forest gaps, 567 which include assimilated information) exhibit slightly larger residuals (R = 0.99 568 and R=0.97 for Eu50 and Ma0.5 respectively). These differences suggest that the 569 influence of the GC hyperparameter makes both schemes not fully comparable. 570 This is a consequence of the varying number of observations used to update the 571 parameters of each cell that differ for each experiment, depending on how much 572 space falls within the correlation length scale of the GC function in each case.





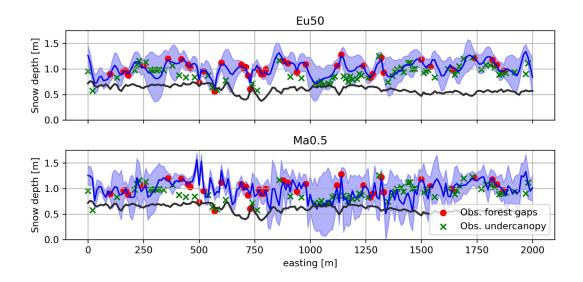


573 Figure 4: Scatterplot based comparison of the under canopy (withheld) and forest gaps 574 (assimilated) observations, with the reference simulation, and the posterior mean of a data 575 assimilation experiment without and with the spatial propagation of the information enabled 576 (experiments Eu50 and Ma0.5)





577 However, these error metrics should be taken with care. Most of them (except 578 CRPS) used the posterior mean as an optimal point estimate of the updated 579 simulation. This assumption was adopted for simplicity but may compromise the 580 interpretation of the results. Posterior simulations are not deterministic 581 simulations and come with an uncertainty estimate inherent in the posterior 582 ensemble. To investigate this issue, we extracted a longitudinal profile along the 583 easting dimension, including both the deterministic reference simulation and the 584 posterior mean, but for the latter we now included the associated uncertainty 585 represented by +/-1 posterior standard deviation (which accounts for 586 approximately 68% of the posterior probability, Fig. 5). In addition, we included a 587 representation of the observations obtained both beneath the canopy and in 588 forest gaps. The profile highlights the differences of using the GC function in the 589 Euclidean or topographic space, with Eu exhibiting a much smoother surface 590 compared with the sharper Ma profile. Both profiles exhibited a similar 591 performance if we account for the uncertainty. In terms of the posterior mean, 592 Ma0.5 was able to accurately capture snow depth in large areas beneath the 593 canopy (e.g. Fig.5 from 1000 to 1250), while maintaining most of the observations 594 in at least the range of its standard deviation. Both Eu50 and Ma0.5 improved the 595 reference run, which exhibited an evident underestimation and lack of 596 heterogeneity along this profile, with only a few observations approaching the 597 simulated reference values.

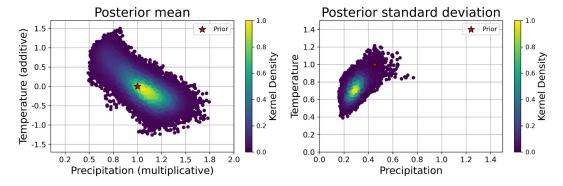






598 Figure 5: Snow depth profile showing the match between the reference run (black line), the Eu50 599 and Ma0.5 experiments and the observations for the horizontal profile delineated by the red line 600 shown in Figure 1. The dark blue line is the posterior mean and the shaded area the posterior 601 standard deviation.

602 Although the aim of the present work is to explore how to propagate the 603 information spatially, it is tempting to analyze the posterior distribution of the 604 parameters (Fig. 6). On average for all cells, using the experiment Ma0.5 as a 605 reference, the multiplicative precipitation parameter was 1.06 (+/- 0.30) and the 606 additive temperature parameters was -0.04 (+/-0.73).



607 Figure 6: Posterior distributions of perturbation parameters in the model space for the Ma0.5 608 experiment, each point represents a grid cell.

609 4 Discussion





610 The results shown here demonstrate the potential of ensemble-based DA 611 experiments to improve hyper resolution snowpack simulations in forested terrain, 612 by updating the canopy covered cells from information retrieved in clearings. 613 Recall that the DA schemes proposed herein are theoretically independent of the 614 underlying numerical model, meteorological forcing or site. As such, in practice any 615 other snow or land surface model forced by meteorological data generated by any 616 downscaling tool at any geographical location may benefit from the proposed 617 techniques. The aim of this work is not to perform the best possible simulation, 618 but to explore whether it is possible to improve snowpack simulations in forested 619 areas by means of DA. Future initiatives may choose to explore the added value of 620 including additional forcing corrections or internal model parameters in the 621 parameter vector since there is, in theory, not any particular limitation on this 622 provided that a large enough ensemble is computationally feasible.

623 All experiments were performed using the Centre National D'Etudes Spatiales 624 (CNES) supercomputing infrastructure. For reference, the Ma0.5 experiment took 625 one day and eight hours to complete, using 6 nodes with 10 CPUs each to solve the 626 40401 cells (201 cells in each geographical direction) that compose the domain 627 using the aforementioned DA scheme. This estimate of computational cost, which 628 could be considered very affordable, especially given the iterative nature of the 629 assimilation algorithm and the relatively low number of processors involved, 630 should be treated with some caution. The computational time varied significantly 631 between experiments, as in practice the I/O increases with the GC 632 hyperparameter, which effectively defines a search radius. In addition, MuSA 633 benefits from distributed systems that share I/O bottlenecks among their nodes, 634 so the computational scheme can also be very relevant. On the other hand, other 635 DA experiments with a lower density of observations will see their computational 636 cost dramatically reduced, independent of the GC hyperparameter.





637 Most of the DA configurations managed to improve the posterior simulations 638 compared with the deterministic reference simulation, with 639 configurations showing similar error metrics. However, the PCA based 640 experiments, despite their desirability given the orthogonal properties of the 641 synthetic coordinate system, did not perform as expected. We hypothesize that the 642 limitations found may come from the fact that the new set of coordinates do not 643 explicitly preserve the Cartesian northing and easting information by mixing them 644 with other dimensions, relaxing the relations between nearby cells in the Euclidean 645 space (Davis & Curriero, 2019). However, the same could be said when using the 646 Mahalanobis distance, but the performance of the Ma experiments was clearly 647 superior compared to the PCA ones. A potential reason may be the fact that, to 648 ease the positive-definiteness of the PCA-based covariance matrix by sorting the 649 cells in a lower dimensional space, we used the Minka algorithm to reduce the 650 dimensionality of the synthetic coordinate system. This dimension reduction 651 comes in practice with a loss of information. However, this is very unlikely, since in 652 practice it resulted in only one dimension being removed, which represented a 653 very low percentage of the total variance of the system. This requires further 654 research to fully understand how the information can be effectively propagated in 655 different spaces. A potential future approach may be the use of multidimensional 656 scaling techniques, instead of PCA, that have shown previous success in 657 geostatistics (R. R. Murphy et al., 2015). The challenges previously encountered in 658 generating non-positive definite covariance matrices have been substantially 659 mitigated. Previous research has proposed to enforce positive definiteness in 660 covariance matrices by using (potentially iterative) methods based on 661 eigendecomposition, to make any negative eigenvalues of the covariance matrix 662 become nonnegative (e.g. Davis & Curriero, 2019 and references herein), which 663 imposed a considerable computational burden, particularly for large matrices. 664 However, regularizing the covariance matrix with the introduction of the jitter 665 technique (where small values are iteratively added to the diagonal) has proven to 666 be both highly effective and computationally efficient. Whether the results of prior 667 sampling differ significantly between these two approaches to regularize the 668 covariance matrix remains an open question for future investigation.





669 The fact that in these experiments we update meteorological correction 670 parameters only, and not snowpack states, allows the numerical model to resolve 671 the snow-canopy interactions. This prevents the posterior simulations to be 672 degraded by the fact that in reality the snowpack beneath the canopy behaves 673 differently than in open terrain (Pflug et al., 2024; Varhola et al., 2010), by updating 674 only parts of the simulation that we assume to be similar independently of the 675 canopy cover (such as the precipitation or temperature), and letting the model to 676 resolve the parts that can't be constrained (such as snow states), due to the lack of 677 information. Since the main objective of this experiment was to explore how the 678 information can be propagated effectively from clearings towards the canopy 679 covered cells, we split the observation dataset in two, keeping the cells beneath 680 the canopy for validation. This has not allowed us to include vegetation 681 parameters in the distance mapping of the Ma experiments, as the cells inside and 682 outside the forest would have been too far away in Mahalanobis space, and 683 therefore due to the localization, the information would not have been transmitted 684 from the clearings towards the sub-canopy. Some vegetation model parameters 685 could have been included in the inference, but because the information is located 686 in the forest gaps, they could not have been constrained. However, given the 687 success of the experiments, future research would benefit from assimilating data 688 also in canopy-covered cells, if a proper error model is developed. State of the art 689 remote sensing techniques are able to retrieve at least a partial information of the 690 snowpack in forested terrain (Mazzotti et al., 2019), or even snow cover 691 information from satellites (Xiao et al., 2022). This may be used not just to further 692 improve the posterior simulations but as a tool to infer internal model parameters 693 spotting weakness in canopy/snow models or their parameters. It should be noted 694 that these spatio-temporal techniques are compatible with joint DA initiatives, 695 where more than one type of observation is assimilated into the same simulation, 696 potentially only spatially spreading some of them (Mazzolini et al., 2024). This may 697 include the ingestion of under canopy in situ observations jointly with remotely 698 sensed retrievals of any kind. It is worth noting that, due to the assimilation of only 699 a single incomplete snow distribution map, the posterior simulations exhibit 700 equifinality (Beven & Freer, 2001), which prevents us from exploring in detail which 701 of these components is more dominant over the other since they are correlated 702 (Fig. 6). Adding other data sources and using more varied information could help 703 address this issue in future studies. In any case, the mean posterior values 704 obtained were close to unity for precipitation (in the physical space) and close to 705 zero for temperature, suggesting that it is not the total amount of precipitation 706 that is biased, but rather the small-scale redistribution of the meteorological 707 forcing.





708 Among the experiments that improved the simulations compared with the 709 deterministic reference run, there was not a clearly superior experiment 710 depending on the GC correlation length scale hyperparameter. Similar conclusions 711 could be drawn from the findings in Cho et al. (2023), who tested different 712 correlation length scales for their Gaussian decay-based localization function, 713 showing that the differences were always lower than the improvement compared 714 with their reference simulation. This suggests some flexibility in the choice of this 715 hyperparameter, which may be complex especially when using non-Euclidean 716 distances, and often limited by the availability of considerable computing 717 resources. When comparing the Eu and Ma experiments, it was also difficult to 718 spot differences if considering only quantitative error metrics. However, the spatial 719 patterns at smaller scales seem more realistic when using the Ma configuration, as 720 also found in Alonso-González et al. (2023). This is based on the fact that the snow 721 spatial patterns are correlated with the characteristics of the terrain, since it 722 controls its distribution by modulating accumulation and melt processes in both 723 open and forested terrain (Geissler et al., 2024; Revuelto et al., 2014). It should be 724 noted that the proposed domain is relatively small exhibiting a limited 725 topographical complexity. Other experiments over larger areas of increasing 726 topographical complexity may benefit from the increasing topographical 727 variability. A potential limitation of this method will be found in non-complex 728 terrain, as is typical in high latitude areas, where the topographical control of the 729 snowpack dynamics may be less clear, although still very relevant (Bennett et al., 730 2022). In any case, snowpack in these areas exhibits less spatial variability, so we 731 hypothesize that the use of Euclidean distance to map cell similarity is likely to be 732 sufficient in these environments and/or at coarser resolutions. Alternatively, it is 733 possible to use snow climatologies or observations to perform a more direct cell 734 similarity mapping based on the persistence of the spatial patterns of the snow 735 (Alonso-González, et al., 2023; Mazzolini et al., 2024). Despite the fact that 736 developing snow cover climatologies in forest environments is significantly more 737 challenging than in open terrain due to the aforementioned limitations of satellites 738 to retrieve information beneath the canopy, it is possible to generate maps of the 739 snow distribution in forested terrain by combining different techniques such as 740 ground observations, lidar and field campaigns (Geissler et al., 2023). The 741 generation of such products requires a significant effort in logistics that prevent its 742 operational exploitation as a real time monitoring tool. In addition, such field 743 methods will not be able to retrieve information at other times that the 744 observation time itself. A promising application of the assimilation scheme 745 presented here is to exploit such products to map the similarity between cells in 746 forested terrain, allowing the significant effort needed for these initiatives to be





747 exploited to generate gap-free re-analyses or near real time updated simulations.

748 In this work, we have explored the effect of using the GC function to create a prior 749 covariance matrix in different spaces. However, what remains to be investigated is 750 the potential benefit of using different covariance (or kernel) functions. It is 751 possible that other functions may offer a more accurate representation of 752 snowpack correlograms across various spatial scales and resolutions, especially in 753 topographical Mahalanobis spaces. One obvious source of inspiration is to take 754 advantage of the extensive literature on kernels developed by the Gaussian 755 process community (Rasmussen & Williams, 2005). In particular, kernels with 756 compact support—those that become zero beyond a certain boundary— (Barber, 757 2020) could be of special interest since they will behave similarly to the GC 758 function, helping in limiting the computational cost and preventing spurious 759 correlations among the ensembles. Given the increasing availability of snow depth 760 information over large domains (Magnusson et al., 2024; Painter et al., 2016), it 761 will be beneficial for the snow DA community to explore which kernel functions 762 better approximate the empirical snowpack spatial variability in different spaces 763 and resolutions. Given that snowpack exhibits persistent spatial patterns in both 764 forest and open terrain (Geissler et al., 2024; Helfricht et al., 2014), there is 765 potential to find a single flexible kernel configuration, ideally depending on a very 766 limited number of parameters, to be widely used in both spatiotemporal DA and 767 observation interpolation initiatives.

768 **5 Conclusions**

769 In this work, we have explored the potential of the observations obtained in forest 770 clearings to be used to update spatially complete snow simulations in forest 771 environments by means of spatio-temporal ensemble-based data assimilation. Six 772 different experiments were conducted in the Sagehen Creek (California, USA) using 773 different data assimilation configurations, demonstrating the potential obvious 774 benefits of spatiotemporal DA in forest environments. While most of the 775 experiments greatly improved the reference snow simulations, those relying on a 776 set of synthetic dimensions generated by a PCA were clearly inferior. Future 777 research may benefit from exploring other dimension reduction techniques such 778 as multidimensional scaling.

779 Among the remaining successful experiments, there was not a clearly superior 780 configuration, in that the differences among them were significantly lower than 781 the improvement compared with the reference run. This suggests some flexibility 782 on the selection of the critical hyperparameters of the DA. However, we found that





783 in terms of both qualitative and quantitative error metrics, those experiments built 784 on a cell similarity mapping based on the Euclidean distance were slightly more 785 accurate in terms of absolute validation metrics, but with a more realistic 786 representation of the spatial variance when using the Mahalanobis distance in a 787 topographical space. This suggests that this latter technique is better suited for 788 preserving spatial relationships in complex terrain. The critical differences found in 789 the implementation of a prior covariance function in different spaces, suggests the 790 importance of future research investing effort in development of specific kernels 791 with the aim of improving distributed snowpack simulations from spatially 792 incomplete observations in forested and/or complex terrain.

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802 Open Research

803 MuSA (v2.2) is open source and can be found at Alonso-González et al. (2024). Future 804 versions of MuSA will be submitted to https://github.com/ealonsogzl/MuSA. The 805 assimilated airborne lidar snow depth data can be found at Piske (2022).

806 Author Contribution

807 Conceptualization was by EAG, AH, SG and JL. Methodology was by EAG and KA. Software 808 was by EAG, KA and LS. Validation and formal analysis was by EAG. Investigation was by 809 EAG and KA. Resources were provided by AH. Data Curation and visualization was by EAG. 810 Writing the original draft was led by EAG with key contributions from all authors. All 811 authors contributed to the review & editing of the original draft.

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