

PaleoSTeHM v1.0-~~re~~: a modern, scalable spatio-temporal hierarchical modeling framework for paleo-environmental data

Yucheng Lin¹, Robert E. Kopp^{1,2}, Alexander Reedy^{1,2}, Matteo Turilli^{3,4}, Shantenu Jha^{2,3,5}, and Erica L. Ashe¹

¹Department of Earth & Planetary Sciences, Rutgers University, Piscataway, NJ, USA

²Rutgers Climate and Energy Institute, Rutgers University, New Brunswick, NJ, USA

³Department of Electrical and Computer Engineering, Rutgers University, Piscataway, NJ, USA

⁴Computational Science Initiative, Brookhaven National Laboratory, Upton, NY, USA

⁵Computational Science Department, Princeton Plasma Physics Laboratory, Princeton, NJ, USA

Correspondence: Yucheng Lin (yc.lin@rutgers.edu) and Robert E. Kopp (robert.kopp@rutgers.edu)

Abstract. Geological records of past environmental change provide crucial ~~information for assessing~~ insights into long-term climate variability, trends, non-stationarity, and ~~nonlinearities~~ nonlinear feedback mechanisms. However, reconstructing spatio-temporal fields from these records is statistically challenging due to their sparse, indirect, and noisy nature. Here, we present PaleoSTeHM, a scalable and modern framework for spatio-temporal hierarchical modeling of paleo-environmental data. This framework enables the implementation of flexible statistical models that rigorously quantify spatial and temporal variability from geological data ~~with clear distinguishing between~~ while clearly distinguishing measurement and inferential uncertainty from process variability. We illustrate its application by reconstructing temporal and spatio-temporal paleo sea-level changes across multiple locations. Using various modeling and analysis choices, PaleoSTeHM demonstrates the impact of different methods on inference results and computational efficiency. Our results highlight the critical role of model selection in addressing specific paleo-environmental questions, showcasing the PaleoSTeHM framework’s potential to enhance the robustness and transparency of paleo-environmental reconstructions.

1 Introduction

As humans push the planet’s climate and biosphere increasingly far outside the range of our species’ experience, the ~~geological record provides~~ environmental reconstructions derived from the geological record provide critical out-of-sample ~~data against~~ which information to test the physical models used to project future environmental change. ~~Yet, as an environmental record, the geological data is quite sparse and often noisy and indirect~~ However, as environmental records, geological data are sparse, often noisy, and indirect (PAGES2k Consortium, 2017; Shennan, 2015). Reconstructing paleo-environmental fields is thus a critical and challenging statistical task (Tingley et al., 2012).

From ~~an analytical~~ a modeling perspective, spatio-temporal hierarchical statistical models provide a natural, conceptually straightforward framework for reconstructing ~~paleo-environment~~ paleo-environmental signals (Ashe et al., 2019; Cressie and Wikle, 2015; Tingley et al., 2012). Hierarchical statistical models, often employed within a Bayesian framework, decompose

the various sources of random variation contributing to individual observations into distinct levels, thereby providing a ~~clearer~~ clear articulation of the assumptions underlying the statistical analysis. ~~It has~~ They have been increasingly used to model paleo-climate fields from geological proxies ~~(e.g., which are naturally occurring physical characteristics or chemical markers that can~~ be used to reconstruct past climate and environmental conditions, such as temperature and precipitation ~~from, from sources like~~ tree rings and corals ~~)(Walter et al., 2022; PAGES2k Consortium, 2017)~~. These applications have proven crucial in assessing the robustness of scientific knowledge of past climate and placing changes in the modern, instrumentally observed period in the context of longer-term variability. For example, they have shown an increasing influence of ice melt and thermal expansion on ~~GMSL~~ global mean sea level (GMSL) since 1860 CE (Walker et al., 2021), that GMSL rise over the 20th century was faster than during any century in at least 3000 years (Kemp et al., 2018; Kopp et al., 2016) ~~;~~ and that several early 21st century Arctic summers exhibited warmth unprecedented in at least 600 years (Tingley and Huybers, 2013). ~~There is substantial community demand for the use of such techniques~~

Hierarchical models are in high demand within the paleoenvironmental research community. For example, in the past few years, numerous papers have used temporal or ~~spatiotemporal~~ spatio-temporal hierarchical models with ~~Gaussian Process (GP)~~ Gaussian Process (GP) priors to interpret paleo sea-level proxies ~~(Tan et al., 2023; Khan et al., 2022; Vacchi et al., 2021)-~~ (e.g., Tan et al., 2023; Khan et al., 2022; Vacchi et al., 2021). To meet the demand of the ~~paleo-environment~~ paleo-environmental research community, this paper describes PaleoSTeHM v1.0, which is designed to support the flexible and high-performance implementation of ~~spatiotemporal~~ spatio-temporal hierarchical modeling for paleo-environmental data. PaleoSTeHM (<https://github.com/radical-collaboration/PaleoSTeHM>) is a framework built ~~using modern machine learning architecture and~~ in the spirit of open science and utilizes modern machine learning architecture (e.g., Pollack et al., 2024). It is designed so users can select not only various modeling choices, such as change-point models for temporal analysis or GP for ~~spatiotemporal~~ spatio-temporal analysis, but also ~~analysis choices, including fully Bayesian, empirical Bayesian, and variational Bayesian analysis~~ analysis choices, including fully Bayesian, empirical Bayesian and variational Bayesian analysis (more details in section 2), to investigate different research questions ~~asked~~, with different types of data and ~~spatiotemporal~~ spatio-temporal scales (e.g., local to global, years to millennia) considered. In this paper, some key terms and phrases are defined in Table 1 and are italicized upon their first occurrence for clarity.

2 Hierarchical statistical modeling

Hierarchical modeling is a statistical approach that separates multiple sources of variability contributing to individual observations into distinct levels, enabling a clear understanding and quantification of uncertainties. This section briefly describes ~~a basic~~ theory about hierarchical modeling of basic theory of hierarchical modeling in the paleo-environment, using paleo sea level as an illustrative example. For more systematic introductions ~~on to~~ hierarchical statistical modeling of paleo sea-level and paleo-climate, readers can refer to Ashe et al. (2019) and Tingley et al. (2012).

Bayesian statistics denotes a statistical theory that uses Bayes' theorem to update probabilities conditioned on data and prior knowledge. Based on Bayes' theorem ~~;~~ the conditional probability (see definition in Table 1) (Laplace, 1810), the conditional

Table 1. Definitions of relevant terms in this study. This paper employs terminology based on Ashe et al. (2019).

Term	Meaning
analysis choices	decisions in how to implement a specific model structure, including the selection of deterministic or probabilistic methods (e.g., Bayesian analysis) and approaches to incorporate temporal uncertainty, such as errors-in-variables frameworks or noisy-input methods
auto-differentiation	automatic differentiation, a set of computational techniques to automatically evaluate the partial derivative of a function
Bayesian statistics	a statistical theory that uses Bayes' theorem to update probabilities conditioned on data and prior knowledge
conditional probability	a measure of the probability of an event occurring, given that another event is already known to have occurred
continuous core	near-continuous records from a single core of sediment or a single coral reef
covariance function	defines prior beliefs about the relationship or correlation between variables
data level	model representations of the relationship between the phenomenon and observed data
errors-in-variable (EIV)	a fully-Bayesian framework that accounts for measurement uncertainty in independent variables
error	difference between a measurement and the true value
empirical Bayesian analysis	an analysis choice to estimate the prior from data by fixing higher-level parameters at their most likely values, typically determined using maximum likelihood estimation
fully Bayesian analysis	an analysis choice that assigns prior distributions to all model parameters, combining prior knowledge and observed data to shape the posterior distribution. Since the posterior is often analytically intractable, MCMC methods are used to approximate it
Gaussian Process (GP)	a stochastic process that generalizes the multivariate Gaussian distribution to continuous time and space, defined by mean and covariance functions
GPU acceleration	a computational technique that leverages Graphics Processing Unit (GPU) to significantly accelerate machine learning computational processes
glacial isostatic adjustment (GIA)	Earth surface water mass redistribution induced solid Earth deformation and gravitational and rotational fields variation
hierarchical model	a statistical framework that partitions the multiple random effects that lead to individual observations into different levels
hyperparameter	parameter of a prior distribution
isotropy	a property of having identical statistical characteristics in all directions
likelihood	the probability of observing the given data as a function of the model parameters
Markov Chain Monte Carlo (MCMC)	techniques used to generate random variables, perform complicated calculations and simulate complicated distributions through random sampling in Bayesian models
noisy-input	a framework that applies a first-order Taylor series approximation to account for errors in the independent variable (e.g., time) thereby translating these into equivalent errors in the dependent variable
non-parametric model	a model that does not involve any assumptions about the functional form of the relationship between variables
optimization	the process of iteratively improving the accuracy of a machine learning model, lowering the degree of error
parameter	a quantity used in mathematical equations, computer programs, physical models, and other scientific applications to describe the characteristics of a system or process
parameter level	model representations of prior beliefs about parameters used to control the behavior of a statistical and/or physical model at different levels of the hierarchy
physical model	a class of model based on physical principles to describe natural phenomena, typically using mathematical representations of a system or process that uses numbers and equations to describe physical conditions
probabilistic programming	a programming paradigm that integrates probabilistic models and inference algorithms into standard programming languages, enabling users to define complex statistical models and automatically perform inference on uncertain variables
parametric model	a model that explicitly assumes a specific functional form or mathematical relationship between variables, defined by a fixed set of parameters that summarize the underlying process
process level	model representations of the underlying processes responsible for the data generation
posterior distribution	a type of conditional probability that results from updating the prior probability with observational information summarized by the likelihood
prior distribution	the assumed probability distribution before any observational evidence is taken into account, which can be uninformative or subjective based on a priori knowledge
residual	the difference between an observed and a modeled or predicted value
relative sea level (RSL)	vertical distance between the solid Earth surface and the ocean surface. In the paleo sea-level context, RSL is typically measured relative to present day, where a positive value indicates a higher RSL and a negative value indicates a lower RSL
smoothness	the characteristic of a process that reflects the gradualness of its variations over time or space, often controlled by the kernel's differentiability in Gaussian Process models
sampling covariance function	a function that describes the covariance structure of a random process, derived from the variability observed in a set of model ensembles or sampled data
space-time separability	a property of processes where the spatial and temporal components of the covariance function are treated as independent, so the covariance is expressed as a product of purely spatial and purely temporal functions
stationarity	a property of processes or signals where their statistical properties, like mean and variance, remain consistent over time or space
uncertainty	a parameter defining the range within which a measured value is likely to fall, given a specified probability.
teleconnection	a causal connection or correlation between meteorological or other environmental phenomena which occur a long distance apart
variational Bayesian analysis	an analysis choice that approximates the full posterior distribution with a simpler parametric distribution, transforming Bayesian inference into an optimization problem to reduce computational expense while estimating uncertainty
white noise	serially uncorrelated random variation (zero mean and finite variance)

55 probability of the observed data (y) can be ~~inverted~~ derived from the conditional probability of unknown ~~parameter(s)~~ parameter(s) or process(es) (θ):

$$p(\theta|y) = \frac{p(y|\theta)p(\theta)}{p(y)} \quad (1)$$

where p denotes ‘probability’ and $|$ represents ‘given’. The ~~likelihood~~ likelihood function, $p(y|\theta)$ represents the probability of observing the data y given the parameter(s) or process(es) θ of the model. The ~~prior distribution~~ prior distribution, $p(\theta)$, captures
60 *a priori* beliefs about the unknown parameter(s) or process(es) before any data ~~is~~ are observed. The term $p(y)$, known as the marginal likelihood (or evidence), is the probability of the observed data averaged over all possible parameters or processes:-

$$p(y) = \int p(y | \theta) p(\theta) d\theta \quad (2)$$

Given the observations, the ~~posterior distribution~~ posterior distribution, $p(\theta|y)$, reflects the updated beliefs about the parameter(s) or process(es) after considering the data. Since the marginal likelihood $p(y)$ is often intractable and remains constant
65 for ~~static observations~~ a given data set, we use the simplified form of Bayes’ theorem:- where the posterior distribution is proportional to the product of the likelihood and the prior:

$$p(\theta|y) \propto p(y|\theta)p(\theta) \quad (3)$$

where \propto indicates ‘is proportional to.’

A basic hierarchical statistical model ~~for paleo sea level distinguishes the fundamental RSL change~~ distinguishes the change
70 in observations from both its inherent variability and the observational noise. ~~Hierarchical~~ These models achieve probabilistic uncertainty estimation for time series and/or spatial fields by ~~inverting conditional~~ treating observed data as conditional on a
latent (unobserved) process and unknown parameters, enabling separate quantification of uncertainties at each level through the
application of Bayesian conditional probabilities. Each level of the model quantifies uncertainties independently, necessitating careful evaluation of their respective sources. Generally, three levels are defined: ~~the data level, the process level, and~~ the
75 ~~parameter level~~ data level, the process level and the parameter level:

$$p(f, \theta_s, \theta_d | y) \propto \underbrace{p(y | f, \theta_d)}_{\text{data model}} \cdot \underbrace{p(f | \theta_s)}_{\text{process model}} \cdot \underbrace{p(\theta_d, \theta_s)}_{\text{parameter model}} \quad (4)$$

~~The data level~~ Taking paleo relative sea-level (RSL) change as an example, the data level model defines the relationship between the latent (unobserved) RSL process (f) and the observed RSL data (instrumental and/or proxy), y , while accounting for measurement, inferential (e.g., ~~uncertainties~~ arising from converting a proxy’s elevation to a distribution of RSL) ~~and~~
80 dating uncertainties (often inherited from geochronology techniques, Reimer et al., 2020; Wright et al., 2017). This level represents the probability distribution of observing a particular sea-level height at a given age, conditioned on the underlying latent process and the associated ~~measurement, inferential, and dating~~ uncertainties, encapsulated by the data level parameters, θ_d .

The process level distinguishes the underlying phenomenon of interest and its inherent variability, from the noisy observation captured at the data level. This model integrates scientific understanding and associated uncertainties into the estimation of the true RSL process ~~using conditional~~ conditioned on model parameters, θ_s . These parameters may represent unobserved ~~physical model~~ physical model parameters (e.g., Earth's rheology in a ~~glacial isostatic adjustment model~~; Table 1 ~~glacial isostatic adjustment (GIA) model~~), statistical model parameters (such as the ~~linear rate in a linear rate of change in a sea-level model~~), or ~~hyperparameters~~ hyperparameters (parameters of a prior distribution, such as length scale and variance in a ~~model with a Gaussian Process prior~~ GP model). At the foundational level, the parameter model specifies the prior distribution for all unknown parameters, effectively capturing the essential characteristics of both the data and process levels through the unobserved parameters.

In addition to constructing models at the data, process, and parameter levels, often referred to as modeling choices (Ashe et al., 2019), it is essential to choose an appropriate analysis choice for a specific model (Table 1). This involves decisions regarding the implementation of a model structure, such as deterministic methods ~~including or~~ probabilistic methods like Bayesian analysis. Deterministic methods, such as least-squares analysis (Wilks, 1938) and likelihood maximization (~~Wilks, 1938; Aitken, 1936~~) ~~or probabilistic methods~~ (Aitken, 1936), rely on fixed relationships between states and events without incorporating randomness into the modeling process. In contrast, probabilistic methods, like Bayesian analysis (Hastings, 1970), account for uncertainty explicitly by representing model parameters and outputs as probability distributions, enabling flexible and robust uncertainty quantification. Analysis choices are also integral to addressing how measurement uncertainties, particularly those arising from geochronological techniques ~~or~~ (i.e., input uncertainty), are incorporated and managed within the model. This ensures that the uncertainty is properly quantified and reflected in the final analysis outputs (Ashe et al., 2019). ~~Several factors, including the complexity of the problem, the~~ The selection of modeling and analytical choices should consider the problem's complexity, data size and resolution of the data available, the computational resources at hand, and the extent of prior knowledge applicable to the modeling effort, should guide the selection of modeling and analytical choices, computational resources, and prior knowledge.

105 3 ~~Model description~~ Software architecture

This section provides a comprehensive overview of PaleoSTeHM, detailing its foundational model implementation (section 3.1), the basic architecture for a typical PaleoSTeHM experiment (section 3.2) and the development of PaleoSTeHM modules (sections 3.3, 3.4 and 3.5).

3.1 Model implementation

110 PaleoSTeHM is designed to be a functionally extensible and high-performing toolkit for modeling paleo data. It is fully open-source and developed under a four-layer structure to maintain a flexible and generic design that is agile to future development (Figure 1). The ~~core toolkit and development reside in L3, comprising modules that integrate existing capabilities from L2, enabling PaleoSTeHM to function across various computing platforms defined in four layers, shown from the bottom to top in Figure 1~~ are: (L1 ~~or~~) computing platforms, (L2 ~~employs Python as the user interface language and utilizes a high-performance~~

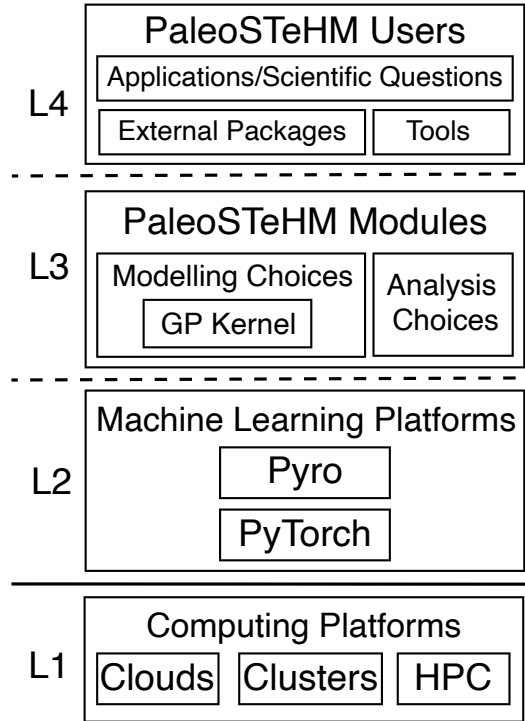


Figure 1. Schematic illustration of the four-layer structure of PaleoSTeHM. L1 specifies various computing platforms (clouds, clusters and HPC), L2 comprises machine learning platforms (Pyro, Bingham et al., 2019, and PyTorch, Paszke et al., 2019), L3 includes PaleoSTeHM modules (Modeling Choices, GP kernel and Analysis Choices, see Figure 2) and L4 consists of the user layer, facilitating interaction with external packages and tools for practical applications and scientific inquiries.

115 ~~machine learning platform as the execution back-end. Atop~~ machine learning platforms, ~~(L3 is the PaleoSTeHM User layer)~~
PaleoSTeHM modules and ~~(L4) PaleoSTeHM users~~. At the fundamental level, PaleoSTeHM utilizes computational power from
various platforms (L1), such as clouds, clusters and high-performance computing systems, to ensure scalability and flexibility
for diverse applications. Built upon L1, L2 employs Python as the user interface language, ~~facilitating interaction with external~~
~~packages and tools, thereby supporting practical applications and the resolution of scientific inquiries.~~ ~~leveraging PyTorch~~
120 ~~(Paszke et al., 2019) and Pyro (Bingham et al., 2019) as its high-performance machine learning platforms~~. The fast-evolving
ecosystem of these popular machine learning platforms enables PaleoSTeHM to support *probabilistic programming*, *auto-differentiation*,
GPU acceleration, and state-of-the-art *optimization* algorithms, making it highly efficient and adaptable to a variety of paleoenvironmental
statistical tasks.

Schematic illustration of the four-layer structure of PaleoSTeHM. L1 specifies various computing platform (clouds, clusters
125 and HPC), L2 comprises machine learning platforms (Pyro and PyTorch), L3 includes PaleoSTeHM modules (Modeling

Choices, GP kernel, and Analysis Choices, see Figure 2), and L4 consists of the user layer, facilitating interaction with external packages and tools for practical applications and scientific inquiries.

PaleoSTeHM modules were built upon Pyro (Bingham et al., 2019), a universal probabilistic programming language supported by PyTorch (Paszke et al., 2017), a popular machine learning library for artificial intelligence applications. Therefore, PaleoSTeHM not only supports probabilistic programming but also leverages an ecosystem of existing machine learning capabilities, including auto-differentiation, GPU acceleration, and modern optimization algorithms. By utilizing such advanced machine learning platforms, we have developed three core PaleoSTeHM modules in the The core toolkit and development reside in L3, which comprises modules that integrate existing machine learning capabilities from L2 layer. This layer includes three primary components: (1) the Modeling Choices module, which ~~incorporates multiple~~ provides options for data, process, and ~~parameter~~ level-modeling-parameter-level modeling (section 3.3); (2) the Gaussian Process kernel module, a sub-module of the Modeling Choices module that supports kernel construction ~~of a process model~~ using GP priors (section 3.4); and (3) the Analysis Choices module, which incorporates ~~multiple methods to consider temporal uncertainty and perform methods to propagate temporal uncertainty into inference results (i.e., temporal uncertainty treatment)~~ and Bayesian inference (~~Figure 1~~ section 3.5). These modules ~~provide enable~~ flexible and efficient ~~options for~~ spatio-temporal hierarchical modeling for a wide range of paleo-environmental applications.

We anticipate PaleoSTeHM interacting with external packages and/or tools for practical applications and addressing scientific questions on the PaleoSTeHM User layer (~~L3~~ L4, Figure 1). Here, ‘External Packages’ refer to external Python libraries, which provide various pre-processing and post-processing data functions. For example, in PaleoSTeHM tutorials (see section 4), we use Scipy (Virtanen et al., 2020) for interpolation and Matplotlib (Hunter, 2007) for visualization. ‘Tools’ represent frameworks and services adapted by other developers to integrate PaleoSTeHM capabilities into their toolkits (e.g., Framework for Assessing Changes To Sea level (FACTS); Kopp et al., 2023). Such plug-in implementations will make it easy for users drawn from any of the PaleoSTeHM categories to use, extend, or contribute to core capabilities for various scientific applications.

3.2 PaleoSTeHM ~~experiment architecture~~ modeling workflow

~~Constructing and training~~ Constructing and optimizing a hierarchical model within PaleoSTeHM ~~consists~~ involves a workflow consisting of five sequential selection steps ~~.Each step corresponds to different~~ (outlined in Figure 2), with a focus on modeling and analysis choices ~~introduced in section 2 in layer L3 as depicted in Figure 1~~. Typical PaleoSTeHM experiment steps include: (1) selecting data-level models for paleo-environmental data; (2) choosing an appropriate process-level model to describe the latent process; (3) defining prior distributions for each model parameter; (4) selecting a temporal uncertainty treatment method; and (5) choosing a Bayesian inference method (Figure 2). These five steps reflect core functionalities ~~developed within three PaleoSTeHM modules, of PaleoSTeHM Modules (layer L3~~ shown in Figure 1). To support the effective selection of modeling and analytical choices provided by PaleoSTeHM for various paleo-environmental applications, ~~The, the~~ the fundamental theories and example applications for each modeling ~~option~~ choices will be introduced in section 3.3.

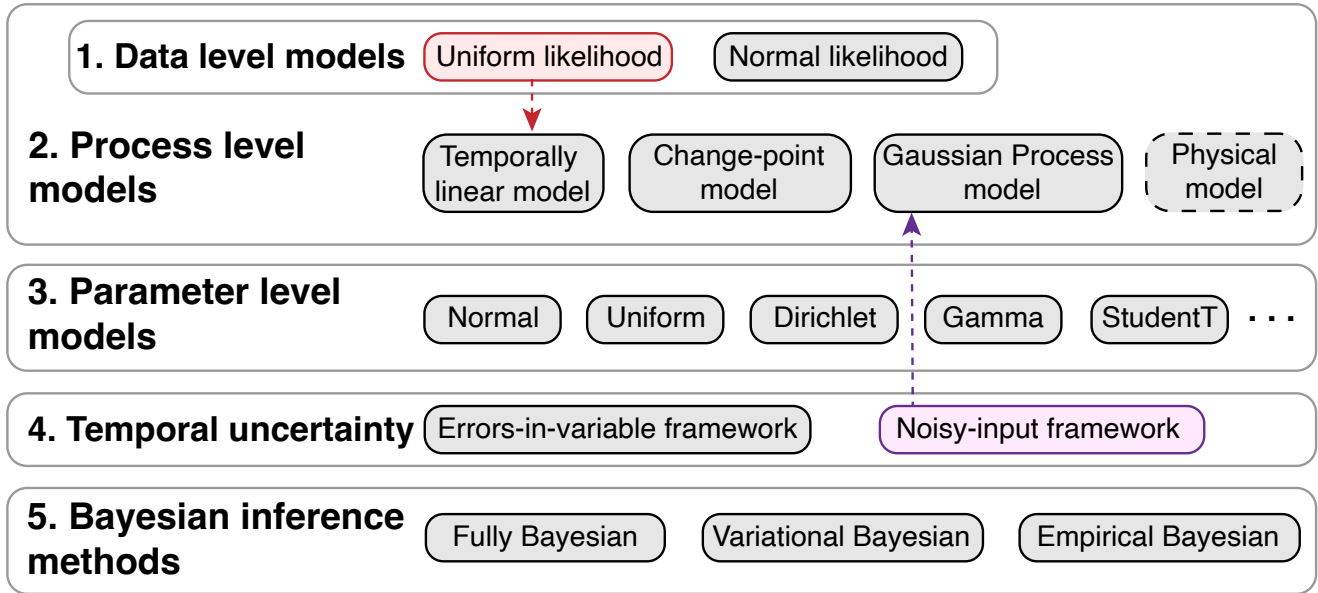


Figure 2. ~~Schematic~~ A schematic illustration of the PaleoSTeHM ~~experiment architecture described modeling workflow, providing more detailed information about layer L3 in this paper~~ Figure 1. The large numbered boxes represent five steps to build a hierarchical model, and it should be noted that the data-level model is specified within each process-level model in PaleoSTeHM v1.0. The smaller boxes indicate different modeling choices within each step. Grey boxes denote available choices that apply to other grey boxes in different steps. Red and purple boxes represent a specific data level model and temporal uncertainty treatment method corresponding to a specific process level model (e.g., temporally linear and Gaussian Process models), as indicated by colored arrows. The dashed grey box (physical model) highlights that no specific physical model is implemented in PaleoSTeHM. Instead, PaleoSTeHM utilizes outputs from other physical models (see section 3.3.2.)

3.3 Modeling Choices module

160 As mentioned above, spatio-temporal hierarchical modeling experiments begin with selecting an appropriate modeling choice for a specific problem. This module ~~provides multiple~~ offers a variety of commonly used temporal ~~or~~ and spatio-temporal modeling ~~choices used in options for~~ paleo-environmental studies (Figure 2). ~~We will briefly introduce the fundamental theories for each modeling choice and provide examples of paleo-environment studies that adopted such a model. While we do~~ While this paper does not include a ~~specific dedicated~~ section for parameter-level modeling, ~~leveraging the ecosystem of Pyro and~~ PyTorch ~~the integration of Pyro~~ (Bingham et al., 2019) and PyTorch (Paszke et al., 2019) enables users to ~~easily~~ define prior probabilities for data and process-level model parameters using ~~most of the~~ a wide range of commonly used probability distributions ~~(Figure 2)~~. This functionality allows users to customize priors as needed for their specific modeling requirements.

165

3.3.1 Data level modeling

170 The data level of a hierarchical statistical model characterizes the relationship between true (unobserved) target signals and uncertain observations due to multiple ~~error~~error sources. For example, in reconstructing past sea-level changes, the data level addresses uncertainties arising from elevation measurements, indicative range ~~̄~~and leveling errors (Khan et al., 2017). Additionally, proxy data are often subject to inherent temporal uncertainties stemming from various geochronological methods ~~(e.g., radiocarbon Reimer et al., 2020; Heaton et al., 2020).~~ This hierarchical structure (e.g., radiocarbon dating, Reimer et al., 2020; Heaton et al., 2020). This relationship between observed data and latent process can be formally expressed as:

$$y_i = f(x_i, t_i) + \epsilon_i^y \quad (5)$$

$$t_i = \hat{t}_i + \epsilon_i^t \quad (6)$$

where y_i is the observed data, x_i is the noise-free spatial location of i -th observation, t_i is its true age, \hat{t}_i is the mean observational age ~~̄~~and ϵ_i^t and ϵ_i^y are uncertainties in the age measurement and target signal reconstruction. For paleo-environmental studies, a commonly made assumption is that both ϵ_i^t and ϵ_i^y are multivariate normally distributed with zero mean and heteroscedastic covariance, so ϵ^y can be expressed as:

$$\epsilon^y \sim \mathcal{N}(0, \Sigma_y) \quad (7)$$

$$185 \quad \Sigma_y = \begin{pmatrix} \text{var}(y_1) & \text{cov}(y_1, y_2) & \cdots & \text{cov}(y_1, y_n) \\ \text{cov}(y_2, y_1) & \text{var}(y_2) & \cdots & \text{cov}(y_2, y_n) \\ \vdots & \vdots & \ddots & \vdots \\ \text{cov}(y_n, y_1) & \text{cov}(y_n, y_2) & \cdots & \text{var}(y_n) \end{pmatrix} \quad (8)$$

where n indicates the number of observations available, $\text{var}(\cdot)$ represents the variance of specific data ~~̄~~and $\text{cov}(\cdot, \cdot)$ stands for covariance between two data points, which is often assumed to be 0 when all data are assumed to be independently distributed. ~~Whereas strong covariance could potentially exist for some~~ However, in practice, strong correlations in paleo-environmental data can emerge from shared processes or dependencies, such as sedimentary records ~~collected~~ from the same core or data dated using age-depth modeling ~~technique (Cahill et al., 2015; Blaauw, 2010).~~ techniques, where shared depositional history introduces correlated uncertainties (Cahill et al., 2015; Blaauw, 2010). Ignoring these correlations can lead to biased estimates and reduced model reliability. Adapting the likelihood structure to account for covariance, for example, by using a structured covariance model from age-depth modeling, allows for more accurate and robust inference (Cahill et al., 2015).

In PaleoSTeHM v1.0, the data-level model is specified within each process-level model, which is assumed to be normally and independently distributed (Figure 2). For illustrative purposes, PaleoSTeHM v1.0 also includes an implementation of uniform

likelihood together with a temporally linear model (see Figure 2 and section 4.1). For specific problems requiring different likelihood structures, users can replace the likelihood sampling code ~~within each process-level model with~~ (a probabilistic random sampling operation in Pyro) to utilize most of the standard probability distributions supported by Pyro, such as multivariate Normal distributions with covariance structures mentioned above.

200 3.3.2 Process level modeling

The process level is a hierarchical layer where the variability of the paleo-environment signal is modeled and, in certain cases, decomposed. The process level reflects a scientific understanding of environmental change processes. PaleoSTeHM v1.0 offers multiple process-level models for temporal or spatio-temporal data analysis.

3.3.3 ~~Temporally linear models~~

205 Temporally linear models. Starting with temporal data analysis, probably the most straightforward method for estimating linear trends and the average rate of paleo-environmental change is to fit a linear model to the observed data over time (i.e., straight line model). For example, Engelhart et al. (2009) and ~~Lin et al. (2021)~~ Islam et al. (2021) applied linear regression to discrete ~~paleo-sea-level-paleoenvironment~~ data to estimate respectively the average rate of ~~RSL change during the Common Era and Meltwater Pulse 1A. Over that period~~ sea-level, rainfall and temperature change during specific time intervals. Over
 210 those periods, the observations were ~~qualitatively assessed-assumed~~ to be well represented by a linear trend. A temporally linear model can be expressed as:

$$f(t) = \alpha + \beta t \quad (9)$$

where $f(t)$ is the modeled true RSL, β is the constant rate of change in ~~RSL~~, paleo-environmental variable and α is the intercept (Ashe et al., 2019).

215 3.3.4 ~~Change-point models~~

Change-point models. Change-point models describe a single time series by partitioning it into distinct, contiguous segments, each characterized by a linear trend over time (Carlin et al., 1992). These models are widely used to identify the timing of abrupt changes in past climate conditions. For instance, Caesar et al. (2021) and Kemp et al. (2015) ~~employed a~~ respectively employed change-point ~~model-models~~ to determine the onset of reduced strength in the Atlantic Meridional Overturning
 220 Circulation and the commencement of modern sea-level rise in Connecticut. With m change points, ~~it the~~ change-point model can be written as:

$$f(t) = \begin{cases} \alpha_1 + \beta_1(t - \gamma_1), & \text{when } t < \gamma_1 \\ \alpha_{j-1} + \beta_j(t - \gamma_{j-1}), & \text{when } \gamma_{j-1} < t < \gamma_j \\ \alpha_m + \beta_{m-1}(t - \gamma_m), & \text{when } \gamma_m \leq t \end{cases} \quad (10)$$

where γ_k represents the ~~a~~ change point, α_k denotes the expected value of RSL at that change point ~~and~~ β_j indicates the rate of RSL change for each of the $m + 1$ segments. This model incorporates a continuity constraint ensuring that α_k equals α_{k-1} plus the product of β_{k-1} and the difference between γ_k and γ_{k-1} . In PaleoSTeHM, the change-point model is implemented to allow users to specify any number of change points (i.e., m in equation 10) in the model.

3.3.5 Gaussian Process models

~~Gaussian Process (GP)~~ Gaussian Process models. GP modeling is a ~~nonparametric and non-parametric~~ Bayesian approach that has been frequently used to infer temporal (or spatio-temporal) variation of paleo-environmental change, including magnitude and rate (Ashe et al., 2019). In models with GP priors, the relationships among any set of points (e.g., over time or across both space and time) are described by a multivariate normal distribution, fully characterized by a mean function and a ~~covariance function~~ covariance function (or kernel). Unlike ~~parametric models~~ parametric models, such as linear or change-point models used for spatio-temporal analysis, GP models offer greater flexibility because the shape of the curve is determined by the covariance matrix, which ~~is inferred based on the data~~ reflects the relationship between data points and is inferred directly from the data, rather than being ~~constrained~~ restricted by a predefined functional form.

~~The GP model has~~ GP models have gained considerable traction in paleo-environmental science, largely owing to ~~its~~ their proficiency in extracting meaningful insights from relatively small datasets. ~~It utilizes a nonparametric~~ They utilize a non-parametric framework to interpret intricate data patterns effectively. For example, Kay et al. (2021) utilized a GP model to assess herbivore richness for different latitudes in Argentina. Apart from that, Walker et al. (2021) estimated the trend and rate of RSL change across the US Atlantic coast with a GP model. A spatio-temporal GP model, which is defined by its mean function, ~~$\mu(t)$~~ $\mu(X)$ and covariance function (i.e., kernel) $K(X, X')$, can be expressed as:

$$f(X) \sim GP(\mu(X), K(X, X')) \quad (11)$$

~~here where~~ X indicates spatio-temporal information of a specific date, for which spatial information is frequently represented by latitude and longitude values with no uncertainty location. A popular choice for many paleo-environmental studies is using the zero-mean function, indicating $\mu(X) = 0$ everywhere. In this case, the predictions are only determined by covariance function $K(X, X')$, which defines prior expectations about how information is shared between points in different time and space, which typically decays as the time and space differences increase (Rasmussen and Williams, 2006).

Constructing the covariance function is a pivotal and challenging step in a GP model, as it significantly influences the outcome of the inference results. Yet, ~~its justification can sometimes be complex (Stein, 2012)~~ justifying the form of the covariance function in Gaussian processes for paleoenvironmental studies can be challenging because the processes being modeled are influenced by a wide range of spatial and temporal dependencies, many of which are complex, nonstationary, and not well understood (Tingley et al., 2012; Stein, 2012, 2005a). PaleoSTeHM addresses this by incorporating a ‘GP kernel’ module under the Modeling Choices module, designed to offer more flexibility and customization extendability. This module provides a user-friendly platform for creating and managing GP kernels, streamlining the process of model construction and enhancing the adaptability of the analysis to diverse problems. For paleo-environmental applications, multiple choices of building kernels

have been adopted in various studies (e.g., Walker et al., 2021; Hay et al., 2015; Kopp et al., 2016, 2014, 2009), and some examples ~~will be~~ are shown in section 4.2.

3.3.6 Physical models

Physical models. A physics-based model simulates real-world changes with predictive capabilities anchored in the causal mechanisms delineated by the laws of physics (~~see Table 1~~) (Saltzman, 2001; Farrell and Clark, 1976). Comparatively, statistical models mostly depend on data-driven correlations, often overlooking ~~foundational~~ fundamental physical principles (e.g., mass or energy conservation). Examples in paleo-environment research include using global circulation models to understand the response of the climate system to different climate forcings (Kageyama et al., 2018) and employing ice sheet dynamic models to quantify past ice sheet response to climate change (DeConto and Pollard, 2016; Tarasov et al., 2012). In the realm of paleo sea-level change modeling, the ~~glacial isostatic adjustment (GIA)~~ GIA model is a widely adopted tool to characterize sea-level changes driven by the gravitational, rotational ~~and~~ and deformational (GRD) effects resulting from the redistribution of ice and water mass (e.g., Lin et al., 2023a; Whitehouse, 2018). The predictive power of such a model is contingent upon underlying formulation and core physical parameters (~~Kendall et al., 2005~~) (Peltier et al., 2015; Kendall et al., 2005; Peltier, 2004), such as the history of ice sheet fluctuations and the rheological properties of the Earth’s interior for a GIA model (Lin and Yousefi, 2025; Austermann et al., 2025). Validating the physical model against observational data ~~will~~ should allow a more accurate representation of ~~spatial teleconnections~~ spatially dependent patterns of sea-level change, including those linked to sea-level fingerprints (~~Lin et al., 2021~~) (Lin et al., 2021; Mitrovica et al., 2021), in stark contrast to statistical models that might merely presume correlation diminishes with distance (Walker et al., 2021).

Although PaleoSTeHM does not include a specific type of physics-based model (Figure 2), it offers multiple options to incorporate physical model outputs into final estimates (see examples in section 4.2). Users can use PaleoSTeHM to probabilistically calibrate physical model ensembles conditioned upon observational data. For instance, latent paleoenvironmental processes can be modeled as a combination of physical model ensembles conditioned on different physical parameter combinations, using a Dirichlet distribution prior. PaleoSTeHM also supports using a physical model as a mean function in a GP model. In this context, the GP covariance function essentially models the ~~residuals~~ residuals—those processes not captured by the physical model—between observations and the predicted mean function. Additionally, PaleoSTeHM facilitates the construction of ~~sampling covariance functions~~ sampling covariance functions derived from a physical model ensemble, further enhancing its utility in model integration and assessment (Hay et al., 2015). All of these capabilities are demonstrated in section 4.2, with accompanying source code provided in the PaleoSTeHM GitHub page (see code and data availability).

3.4 Gaussian Process Kernel module

The GP Kernel module in PaleoSTeHM is a cornerstone for modeling spatial and temporal variations in paleo-environmental data based on GP priors (~~Figure 1~~ Figures 1 and 2). It ~~encompasses~~ includes a variety of ~~commonly used kernels in paleo-environmental studies, including~~ widely used kernels as described in Rasmussen and Williams (2006). In paleoenvironmental studies, examples of kernel applications include the linear (or dot-product) kernel (Khan et al., 2017), radial basis function kernel (Cahill et al., 2015), rational quadratic kernel (Turner et al., 2023; Hay et al., 2015), Matérn kernel (Walker et al., 2021; Kopp et al., 2016),

Table 2. Summary of Gaussian Process kernels in PaleoSTeHM, based on Stein (2012) and Rasmussen and Williams (2006).

Kernel Name	Supports Spatial Data [*]	Differentiability	Equation [†]
Radial Basis Function	Yes	Infinitely differentiable	$k(X, X') = \sigma^2 \exp\left(-\frac{1}{2} \frac{ X - X' ^2}{\ell^2}\right)$
Rational Quadratic	Yes	Infinitely differentiable	$k(X, X') = \sigma^2 \left(1 + \frac{ X - X' ^2}{2\alpha\ell^2}\right)^{-\alpha}$
Periodic	No	Infinitely differentiable	$k(X, X') = \sigma^2 \exp\left(-2 \frac{\sin^2(\pi(X - X')/p)}{\ell^2}\right)$
2/1 Matérn	Yes	Non-differentiable	$k(X, X') = \sigma^2 \exp\left(-\frac{ X - X' }{\ell}\right)$
3/2 Matérn	Yes	Once differentiable	$k(X, X') = \sigma^2 \left(1 + \sqrt{3} \frac{ X - X' }{\ell}\right) \exp\left(-\sqrt{3} \frac{ X - X' }{\ell}\right)$
5/2 Matérn	Yes	Twice differentiable	$k(X, X') = \sigma^2 \left(1 + \sqrt{5} \frac{ X - X' }{\ell} + \frac{5}{3} \frac{ X - X' ^2}{\ell^2}\right) \exp\left(-\sqrt{5} \frac{ X - X' }{\ell}\right)$
Linear (or dot product)	No	Once differentiable	$k(t, t') = \sigma^2(t - \gamma) \cdot (t' - \gamma)$
Sampling covariance kernel	Yes	Not applicable	$k(X, X') = Cov(m(X), m(X'))^\#$
Polynomial	No	λ times differentiable	$k(t, t') = \sigma^2(\gamma + t \cdot t')^\lambda$
Constant	No	Not applicable	$k(X, X') = \sigma^2$
Temporal white noise	No	Not applicable	$k(t, t') = \sigma^2 \delta(t, t')$
Spatial white noise	Yes	Not applicable	$k(x, x') = \sigma^2 \delta(x, x')$

^{*} All GP kernels can calculate temporal covariance, except the spatial white noise kernel.

[†] X represents spatio-temporal location, incorporating both the age and spatial coordinates of the data; t denotes the age of the sample; and x indicates the spatial coordinates. σ^2 = variance; ℓ = a positive characteristic length-scale parameter; α = a scale mixture parameter, when $\alpha \rightarrow \infty$, the rational quadratic kernel is equivalent to the radial basis function kernel; γ = offset or shift parameter, adjusting the baseline level of the kernel's output; p = periodicity parameter for the Periodic kernel, defining the cycle length of repeating patterns; λ represents the degree of the polynomial, an integer determining the complexity of the model for the polynomial kernel.

[#] $Cov(m(X), m(X'))$ indicates the sampling covariance between outputs at different spatio-temporal points, derived from deterministic models under varying physical parameter assumptions. Here, $m(X)$ denotes the output at a specific location and time from a suite of physical models assuming different parameters.

and periodic kernel (Meltzner et al., 2017). These kernels characterize features such as ~~stationarity, isotropy, smoothness,~~
290 stationarity, isotropy, smoothness and periodicity in Gaussian processes (~~see definitions in Table 1; Ashe et al., 2019~~)(Ashe et al., 2019)
. Detailed kernel information is given in Table 2.

Each kernel possesses unique characteristics and necessitates specific parameters (Table 2). For instance, the linear kernel produces linear trends identical to a temporally linear model, suitable for modeling signals with long temporal length scales (e.g., tectonic and GIA in Common Era and future sea level modeling; Kopp et al., 2016, 2014). The radial basis kernel and the
295 Matérn family of kernels are highly generalizable and allow specification of the degree of differentiability (Table 2), making

them suitable for representing physical processes with different levels of smoothness. For example, the GRD effects related to GIA are spatio-temporally smooth, while sediment compaction-induced sea-level rise can be much more localized and rough (i.e., less differentiable, [Kopp et al., 2016](#); [Mitrovica et al., 2011](#)).

In the GP Kernel module of PaleoSTeHM v1.0, all kernels are designed for process-level modeling [to capture temporal and/or spatial correlations](#), except for the temporal and spatial ~~white-noise-white noise~~ kernels, which ~~add-additional-account~~ [for additional measurement errors or unstructured variability by introducing](#) serially uncorrelated uncertainty at the data level ([equation 5](#)). Apart from the linear ~~and-white-noise-kernels~~kernel, all included kernels are stationary and isotropic ([see-Table 1](#)). To enhance kernel construction flexibility, PaleoSTeHM supports combining different kernels, either additively, multiplicatively, or both. [Additive combinations capture independent contributions from distinct processes, such as long-term trends or periodic variations, treating them as separate effects. In contrast, multiplicative combinations create interactions between processes, resulting in more structured patterns. For example, multiplying a periodic kernel with a linear kernel produces a periodic variation with an amplitude that increases or decreases linearly over time, effectively modeling phenomena where seasonal patterns intensify or diminish progressively \(Görtler et al., 2019\).](#)

Designed for spatio-temporal data analysis, all GP kernels in PaleoSTeHM support temporal data (represented as a 1-dimensional vector) ~~-~~and most of kernels support spatial data (represented as 2-dimensional matrix including latitude and longitude; see Table 2). ~~The-spatial-correlation-is-computed-for-spatial-kernels-based-on-~~[Temporal kernel correlations are calculated using the 1-dimensional Euclidean distance between time points, while spatial kernel correlations are derived from the 1-dimensional geographical radial distance](#)~~between-data-points,~~ [calculated based on the spherical distance between pairs of longitude and latitude under the assumption of a purely spherical Earth geometry.](#) Users can choose to build a temporal or spatial kernel by switching a parameter in each kernel function.

3.5 Analysis Choices module

To accommodate diverse computational resources and varying requirements for the trade-off between modeling robustness and computational demands, the Analysis Choices module offers multiple methods for Bayesian inference of model parameters as defined in the Modeling Choices module (Figure 2). This flexibility ensures users can optimize their analyses based on available technology and specific modeling needs. Unlike deterministic methods (e.g., least-squares), which have been extensively implemented in other ~~packages~~[studies \(e.g., Crichton et al., 2023; Lin et al., 2021\)](#), PaleoSTeHM focuses on developing Bayesian probabilistic approaches that more effectively manage the inherent uncertainties associated with paleo data.

3.5.1 Fully Bayesian analysis

A fully Bayesian analysis requires assigning prior probability distributions to all model parameters, allowing them to take on a range of ~~probable-values~~[values, potentially with different probabilities](#). These priors can either incorporate informative prior knowledge or remain uninformative and vague. Since the posterior distribution is shaped by both the priors and the likelihood of the observed data, it often becomes complex and analytically intractable. ~~Markov-Chain-Monte-Carlo (MCMC)-~~[Markov Chain Monte Carlo \(MCMC\)](#) methods are crucial in this case as they enable the efficient exploration and approximation of the

posterior distribution. PaleoSTeHM supports two advanced MCMC samplers: Hamiltonian Monte Carlo (HMC; Neal et al., 2011) and the No-U-Turn sampler (NUTS; Hoffman et al., 2014), which provide more efficient sampling performance than traditional Metropolis-Hastings MCMC (Hastings, 1970).

HMC significantly ~~enhances the improves~~ sampling efficiency over traditional Metropolis-Hastings MCMC by ~~utilizing~~ leveraging gradients of the probability distribution to ~~inform guide~~ the sampling process, which involves generating random samples from the underlying latent probability distribution. This method reduces autocorrelation (the correlation between successive samples, in a Markov Chain, indicating how dependent the current sample is on previous ones), thereby increasing the effective sample size ~~per iteration~~ (the number of independent samples, accounting for autocorrelation) per iteration (a single step in the sampling process where the algorithm generates a new sample) and enabling faster convergence. ~~Based~~ Building on HMC, NUTS further ~~improves upon this by automatically adjusting~~ enhances efficiency by automatically adapting the path length ~~and effectively~~ (the distance traversed in parameter space during a single Hamiltonian trajectory) and managing the step size (the distance traveled in parameter space at each leapfrog step during Hamiltonian dynamics, Bingham et al., 2019). NUTS eliminates the need for manual tuning of these parameters, facilitating more ~~efficient effective~~ exploration of complex, high-dimensional ~~distributions typical~~ posterior distributions commonly encountered in Bayesian analysis.

Compared to other analysis choices such as empirical Bayesian models or variational Bayesian models (~~details provided below~~ Table 1), a fully Bayesian model offers a more comprehensive estimation of the ~~relative~~ uncertainties associated with model parameters (Piecuch et al., 2017). It also offers a direct framework for sample age measurement uncertainty in an ~~EIV manner~~ (Table 1 errors-in-variable (EIV) manner, Dey et al., 2000). However, the nature of MCMC-based samplers means they are computationally more demanding. Particularly within the EIV framework, where the number of sampling parameters increases linearly with data size, this leads to a polynomial increase in the computational power required (Belloni and Chernozhukov, 2009), which can be significant and unaffordable when dealing with large datasets or complex models.

3.5.2 Empirical Bayesian analysis

Unlike Fully Bayesian analysis, which requires full probability distributions for prior and posterior, Empirical Bayesian analysis offers a practical alternative. This approach approximates a fully Bayesian treatment where parameters at the highest level of the hierarchy are fixed at their most likely values rather than being integrated out. This optimization is typically achieved using the maximum likelihood estimate, leading to a posterior distribution that is conditional on the data and these optimized parameters:

$$p(f|y, \hat{\theta}_s, \hat{\theta}_d) \propto p(y|f, \hat{\theta}_d) p(f|\hat{\theta}_s) \quad (12)$$

here, the posterior probability of the latent processes f is inferred, assuming that the hyperparameters at the data and process levels ($\hat{\theta}_d$ and $\hat{\theta}_s$) are known and fixed. While the existing code base allows for explicit bounds to be set on hyperparameters for the maximum likelihood estimate (e.g., Ashe et al., 2019; Kopp et al., 2016), it does not provide for an explicit prior distribution for the parameters. In other words, it only support a uniformly distributed prior information, limiting the ability to incorporate informative prior knowledge. By leveraging Pyro’s variational inference capabilities (details ~~below~~ in section

3.5.3), PaleoSTeHM enables users not only to optimize hyperparameters using ~~the~~their maximum likelihood estimate but also to define ~~prior~~many commonly-used distributions for each prior model parameter explicitly. This allows optimization to be conducted in a maximum *a posteriori* probability estimation manner, assuming the variational distribution is a Dirac
 365 delta function. In PaleoSTeHM, by default, the optimization is achieved using Adam, a stochastic optimizer (Kingma and Ba, 2014). While empirical Bayesian analysis generally requires fewer computational resources than fully Bayesian methods, it is important to note that, assuming hyperparameters at the data and process levels are known and fixed may lead to substantial underestimation in the inference uncertainty (Piecuch et al., 2017).

3.5.3 Variational Bayesian analysis

370 Considering the computational expense required to perform MCMC in fully Bayesian analysis and the limitations of Empirical Bayesian methods that fail to account for the uncertainty of hyperparameters, PaleoSTeHM also supports variational Bayesian analysis, which emerges as an efficient intermediary. Rather than directly sampling from the posterior distribution through MCMC, variational Bayesian methods aim to approximate the true posterior probability distribution ($p(f, \theta_s, \theta_d | y)$) with a simpler, parametric probability distribution ($q(f | \phi)$). Thus, Bayesian inference is transformed from a sampling chal-
 375 lenge into an optimization problem—known as variational inference—requiring significantly fewer computational resources while facilitating uncertainty estimation (Wingate and Weber, 2013).

In PaleoSTeHM, variational Bayesian analysis is achieved by optimizing the variational parameters ϕ to minimize the Kullback-Leibler (KL) divergence, a metric to effectively measure the difference between two distributions:

$$\phi = \arg \min_{\phi} \text{KL}[q(f, \theta_s, \theta_d | \phi) || p(f, \theta_s, \theta_d | y)] \quad (13)$$

380 For more details above KL divergence, readers can refer to Blei et al. (2017). Adam facilitates this minimization ~~and~~ the variational distribution for PaleoSTeHM is a normal distribution by default. In contrast to MCMC-based fully Bayesian analysis, which often requires computational power that increases polynomially with the number of data points, the optimization-driven approach of variational Bayesian analysis generally scales linearly (Ko et al., 2024; Hoffman and Blei, 2015). Consequently, variational methods can handle larger datasets more effectively, making them suitable for large-scale problems prohibitively
 385 for full Bayesian analysis.

3.5.4 Incorporation of temporal uncertainty

PaleoSTeHM provides two methods to incorporate temporal uncertainty into final estimations. The first method uses EIV framework (~~Cahill et al., 2015~~)(Cahill et al., 2015; Dey et al., 2000), which directly incorporates temporal uncertainty through MCMC sampling of the distribution. The second approach adopts the ~~noisy-input~~noisy-input framework (McHutchon and
 390 Rasmussen, 2011), which applies a first-order Taylor series approximation—a linear expansion around each input point—to account for errors in the independent variable, time, thereby translating these into equivalent errors in the dependent variable:

$$f(x_i, t_i) \approx f(x_i, \hat{t}_i) + \epsilon_i^t \frac{\partial f(x_i, \hat{t}_i)}{\partial t} \quad (14)$$

here \hat{t}_i and ϵ_i^t are the same as in equation 6, standing for mean observational age and age uncertainty, respectively. The integration of temporal uncertainty within PaleoSTeHM is executed alongside each process level model (Figure 2). All process-
395 level models are implemented using an EIV framework, while for the GP models, both EIV and noisy-input frameworks are available (Figure 2).

3.6 Model Validation

After implementing and optimizing a hierarchical model in PaleoSTeHM, it is essential to perform a model validation step to further ensure the robustness and reliability of the trained model. This process involves evaluating how well the model fits the
400 observed data, assessing its predictive accuracy, and diagnosing potential issues such as convergence problems. PaleoSTeHM includes a range of techniques for model validation, such as residual analysis, posterior predictive checks, MCMC convergence diagnostics (e.g., effective sample size and Gelman-Rubin statistic, Gelman and Rubin, 1992), visual inspections (e.g., optimization trace plots, true versus predicted plot), simulation validation and cross-validation methods. These tools allow users to critically
405 examine the model's assumptions, quantify uncertainties, and compare competing models to select the most appropriate one for their specific paleo-environmental application.

To complement these validation techniques, we demonstrate their application in various case studies presented in section 4. Each case study incorporates specific model validation methods tailored to the modeling and analysis choices used. For example, prior and posterior predictive checks are employed to evaluate the performance of optimized models (section 4.1.1); residual plots, weighted mean square error (wMSE) and cross-validation are used to assess the performance of different
410 process-level models (sections 4.1.2 and 4.2); and effective sample size and the Gelman-Rubin statistic ensure good model convergence for MCMC-based analyses (section 4.1.4). Detailed implementations and usage of these validation methods for various PaleoSTeHM experiments, including those mentioned above but not covered in detail in the following sections, are available in the PaleoSTeHM tutorials (see code and data availability).

4 ResultsCase studies

415 This section presents illustrative ~~results-case studies~~ using a tutorial format to ~~enhance-demonstrate~~ PaleoSTeHM's usability. All codes and data are accessible and actively managed on the PaleoSTeHM GitHub page (~~see-code-and-data-availability~~).

~~Firstly, we demonstrate various data levels, process levels, and analysis choice modeling techniques for time series analysis~~
~~https://github.com/radical-collaboration/PaleoSTeHM~~). The case studies include:

1. ~~Reconstruction of temporal sea-level changes~~ using coral reef data from the Great Barrier Reef ~~and-with different data~~
420 ~~level models (section 4.1.1).~~
2. ~~Reconstruction of temporal sea-level changes using~~ salt marsh data from New Jersey ~~and North Carolina~~. ~~Subsequently,~~
~~we provide examples of reconstructing with different process-level models (section 4.1.2).~~

3. Reconstruction of temporal sea-level changes using salt marsh data from North Carolina with different Bayesian inference methods (section 4.1.4).
- 425 4. Reconstruction of spatio-temporal sea-level changes in using various geological proxies from the US Atlantic coast using
with different process-level models .(section 4.2).

Although these examples focus on modeling paleo sea-level, additional tutorials are available for analyzing other paleoenvironmental data, such as ocean temperature anomalies and concentration of carbon dioxide.

The prior and posterior distributions and analysis choice for each model are provided in Table A1. It should be noted this
 430 section only briefly describes the modeling results; for a more systematic analysis of paleo-environmental modeling results
 based on different statistical techniques, the user can refer to Ashe et al. (2019), PAGES2k Consortium (2019) ,and Tingley
 et al. (2012).

4.1 Time series analysis

4.1.1 Data level modeling

435 In this section, we examine the impact of the data-level model on inference results. Although numerous paleo-environmental
 applications commonly assume that proxy reconstruction uncertainties are normally distributed (Ashe et al., 2019; Khan et al.,
 2019; Tingley et al., 2012), certain types of proxies may exhibit different forms of uncertainty. ~~A typical example is the~~
~~For instance,~~ coral reef sea-level ~~indicator, where reconstruction uncertainty can also be represented by a~~ proxies indicate
~~past sea-level changes through a quantifiable relationship between the coral's living-habitat depth and the concurrent sea~~
 440 ~~level (Hibbert et al., 2016). Representations of coral living-habitat depth uncertainties are often modeled using either a normal~~
~~distribution (e.g., Khan et al., 2019) or a~~ uniform distribution (Lin et al., 2021) ~~according to species-specific living-habitat range~~
~~(Hibbert et al., 2016).~~

~~To illustrate the impact of the data-level model.~~ To illustrate such data-level impact on inference results, we apply a tempo-
 rally linear model within an EIV framework to coral reef data from the Great Barrier Reef (Yokoyama et al., 2018) .~~We use~~
 445 using two alternative data-level models. The first can be expressed as:

$$\epsilon_1^y \sim U(\tau_l - \omega_1, \tau_u + \omega_1) \quad (15)$$

where U indicates a uniform distribution between lower and upper ranges defined by specific coral species (τ_l and τ_u) and an
 additional ~~white~~-noise, defined by hyperparameter ω_1 , ~~the~~ which follows a prior distribution of:

$$\omega_1 \sim U(10^{-4}, 10) \quad (16)$$

450 The second data level model can be represented as:

$$\epsilon_2^y \sim N(\mu_2, \sqrt{\sigma_2^2 + \omega_2^2} \underbrace{\sigma_2^2 + \omega_2^2}_{\text{prior}}) \quad (17)$$

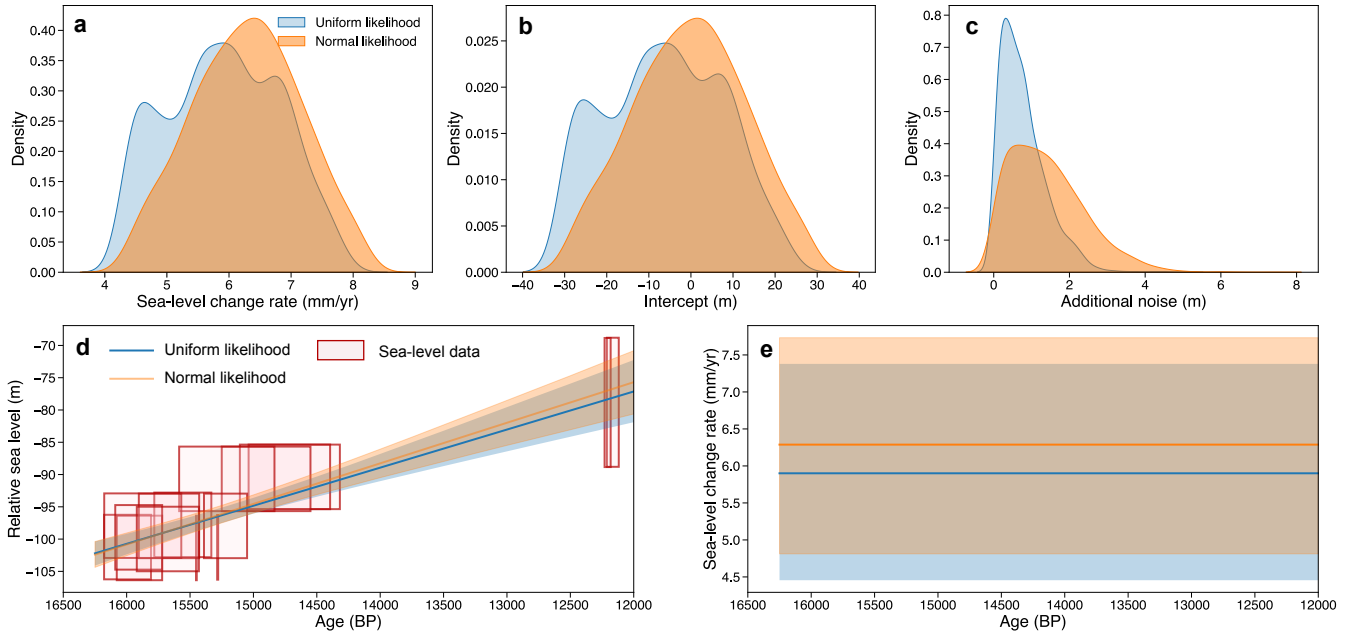


Figure 3. Data level models impact on temporal sea-level change inference at the Great Barrier Reef. The upper panel displays the posterior Posterior probability density functions of model parameters sea-level change rate (a), intercept (b) and standard deviation of additional noise (c), assuming either a uniform likelihood (blue) or a normal likelihood (orange). The bottom panel presents the inferred-Inferred mean sea-level trends (d) and rates (e) along with a 90% credible interval, where sea-level data are represented by red boxes with horizontal range representing indicating $\pm 2\sigma$ age uncertainty and vertical range indicating the reconstructed maximum and minimum sea-level range determined by coral species (i.e., τ_l and τ_u in equation 15). A negative RSL value indicates that the local RSL in the Great Barrier Reef was lower than present-day levels, reflecting the significant amount of water stored in continental ice sheets. CI = credible interval, BP = before present.

where N indicates a normal distribution with mean μ_2 and a standard deviation σ_2 , both of which are determined by specific coral species τ and ω_2 is an additional white-noise-hyperparameter noise hyperparameter, following the same prior distribution as ω_1 . The same prior distributions for each parameter are used for both data-level models, which are represented as non-informative uniform distributions. The characteristics of these non-informative priors are evident in the Prior Predictive Check (Figure A1), which reveals a wide and flat spread of predictions, reflecting the absence of observational influence at this stage.

For both models, the posterior distribution is determined by 11,000 posterior samples drawn from a NUTS sampler, with the first 1,000 samples discarded as burn-in steps. The Posterior Predictive Checks (Figure A1) illustrate that the posterior predictions for both models align closely with the observed data, suggesting successful model convergence. It can be seen in Figure 3 that, although the inference results from different data level models are overall similar, there are still some noticeable differences in the inferred sea-level change trend and rate. Uniform and normal likelihoods yield average sea-level rates of 5.91 mm/yr (4.45-7.38 mm/yr; 90% credible interval, CI) and 6.29 mm/yr (4.81-7.73 mm/yr), respectively. These likelihood

assumptions also produce considerably different additional noise parameter distributions. Therefore, users should select an appropriate data-level model to better represent the specific characteristics of different paleoenvironmental data.

465 4.1.2 Process level modeling

To demonstrate the impact of different process level models on inferring paleo sea-level time series, we will use the same data level model together with process level models introduced in section 3.3.2, employing non-informative priors (Table A1). The sea-level data used here is a ~~near-continuous~~ near continuous core record from single cores of salt-marsh sediment from Leeds Point (New Jersey) covering the Common Era (Kemp et al., 2013). For this database, a normal likelihood data level
470 model is adopted with sea-level reconstruction uncertainties provided by the original study. Here we will test three process level models: (a) temporally linear model; (b) change-point model (assuming 3 change points; following Ashe et al., 2019); (c) Gaussian Process model with an ~~RBF~~ radial basis function (RBF) kernel. Posterior distributions for ~~model~~ models *a* and *b* were ~~sampled-through~~ obtained using a variational Bayesian ~~manner~~ approach, while model *c* ~~was-sampled-using~~ employed an empirical Bayesian ~~approach~~ method.

475 Figure 4 shows estimated RSL trends and rates of RSL change for each process model. The ~~results~~ resulting trends and fit to the data (quantified using wMSE) differ significantly due to the fundamentally different model formulations. The temporally linear model can only estimate an averaged trend and rate of sea-level change and will never predict an accelerated RSL change. Consequently, it exhibits the highest wMSE (2.53) with a systematic errors with strong temporal correlations displayed in residual plots (Figure A2), both of which reflect a poor fit to the observations.

480 Comparatively, the change-point model ~~presents-is able to capture~~ a noticeable change in RSL rate from 1.49 mm/yr (1.26-1.70 mm/yr) between -500 CE and 1839 CE (1824-1852 CE) to 3.91 mm/yr (3.72-4.10 mm/yr) after 1839 CE. ~~Such-an-estimate is-suitable-for-finding~~ This added flexibility substantially improves the model's fit to the data, achieving a wMSE of 0.38 and producing a less structured error distribution (Figure A2). Such flexibility makes the change-point model particularly suitable for identifying the time of emergence ~~for-in~~ various environmental change ~~problems (e.g., Walker et al., 2022; Caesar et al., 2021; Lyu et al., 2014).~~
485 ~~-contexts (e.g., Walker et al., 2022; Caesar et al., 2021; Lyu et al., 2014).~~

As a non-parametric approach, the GP model ~~can-produce~~ produces continuous distributions of RSL change rates over time, allowing for the estimation of multiple inflection points (Walker et al., 2022). ~~However, using the RBF kernel, which is-infinitely-differentiable,~~ This flexibility results in the lowest wMSE (0.31), alongside minimal temporal structure in the residuals (Figure A2), indicating the best overall fit to the observations. However, the infinite differentiability of the RBF kernel
490 can lead to overly smooth changes in RSL rate—an assumption inappropriate in many environmental statistics (Stein, 2012). predictions in time series analysis, potentially oversmoothing sharp changes that are critical in many environmental contexts, such as abrupt sea-level rise (Lin et al., 2021), ocean circulation slowdowns (Caesar et al., 2018), and extreme events like heavy rainfall (Stein, 2012). Alternative kernels (e.g., Matérn kernels) can provide alternative levels of differentiability.

4.1.3 Analysis choices

495 ~~Using similar near-continuous sediment core~~

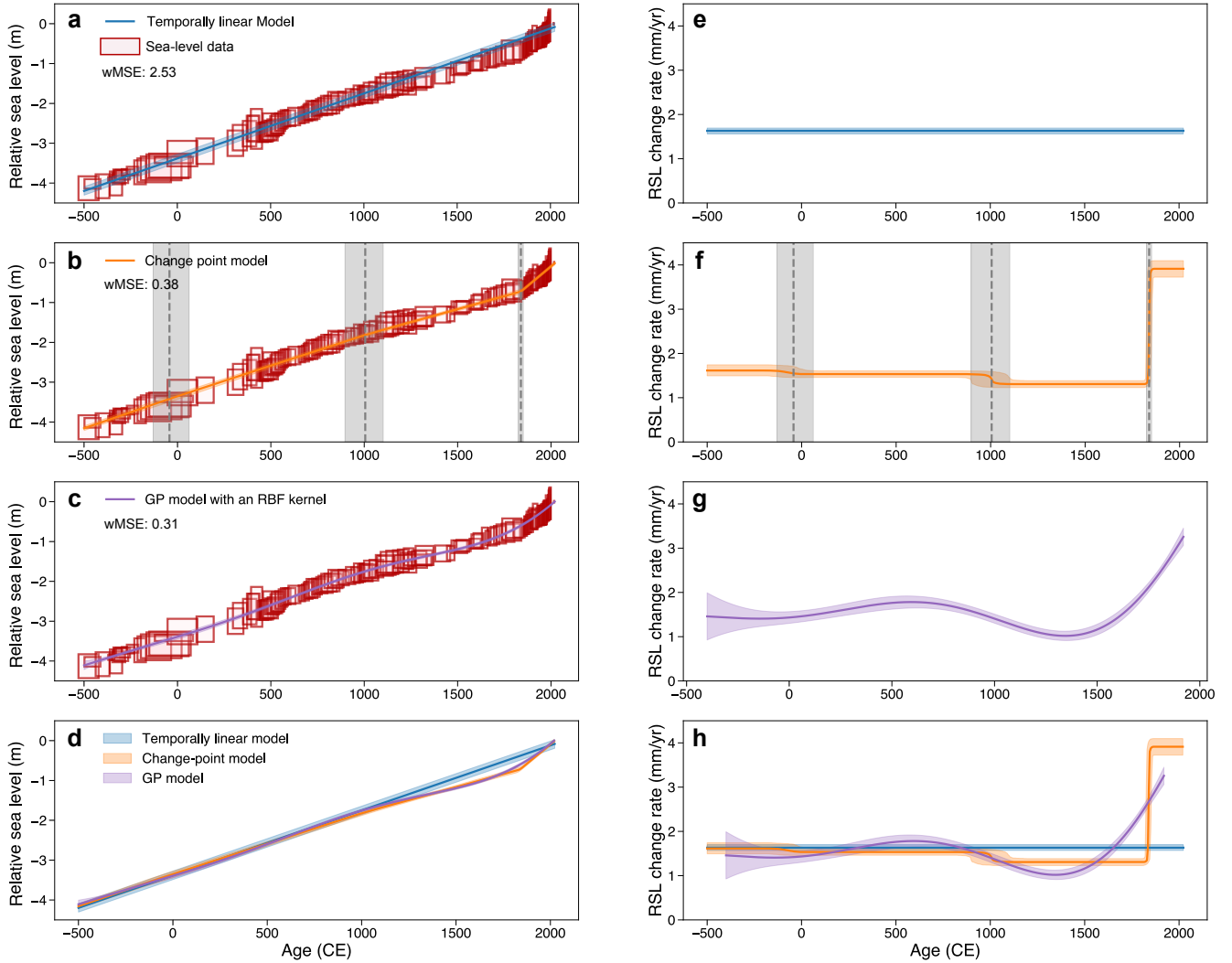


Figure 4. Process level models impact on temporal sea-level change inference at New Jersey. Common Era sea-level comparison of linear model (row-1a, e), change-point model (assuming 3 change points; row-2b, f) and Gaussian Process model with an RBF kernel (c, g), where input data are continuous cores. Output includes estimates of RSL (column-1a-d) and rates of RSL change (column-2e-h), which are each shown with mean and 90% credible intervals. Paleo sea-level data here are modeled using a normal likelihood. The horizontal and vertical range of red boxes indicates $\pm 2\sigma$ age and relative sea-level reconstruction uncertainties, respectively. Sea-level data here were reconstructed using near-continuous record from single cores of salt-marsh sediment from Leeds Point (New Jersey, Kemp et al., 2013). wMSE = weighted mean square error; CE = Common Era.

4.1.4 Analysis choices

Using similar near *continuous core* data from Sand Point, North Carolina (Kemp et al., 2011), we illustrate the effects of analysis choices on RSL inference. Here, we only use a subset of the original data to better demonstrate the difference between various analysis choices. The adopted data and process level model employ a normal likelihood with a GP model using an RBF kernel (Table A1). The hyperparameters will be sampled using empirical, fully Bayesian ~~and~~ and variational Bayesian methods. For the fully Bayesian method, the posterior distribution is determined by ~~2,200~~ 5,500 posterior samples drawn from a NUTS sampler, with the first 200 samples discarded as burn-in steps. For the empirical and variational Bayesian methods, the hyperparameters were optimized using the Adam optimizer over ~~500~~ 1000 iterations (Kingma and Ba, 2014). The run times of each implementation are reported on a 2023 MacBook Pro with an Apple M2 Pro chip.

For MCMC-based fully Bayesian analysis, PaleoSTeHM employs the Gelman-Rubin statistic (Gelman and Rubin, 1992) to verify that the Markov chains have converged to a stationary phase, indicating good convergence. Additionally, the effective sample size is used to assess the amount of information retained, accounting for the correlation in the sequence (Bürkner, 2017). Typically, a Gelman-Rubin statistic of less than 1.1 and an effective sample size greater than 1000 suggest reliable sampling of the posterior distribution. In this case, the analysis meets these criteria with a Gelman-Rubin statistic of 1.0 and an effective sample size exceeding 3000. For empirical and variational Bayesian methods, validation is typically conducted through the inspection of optimization trace plots (plots showing how optimization target function improve with each iteration), where successful optimization is characterized by a steadily decreasing loss function and parameters convergence over iterations. These conditions are also satisfied in this analysis, as illustrated in the corresponding optimization trace plots (Figure A3 and A4).

Figure 5 compares posterior distributions of RSL trend and rate of change and the computational time for each analysis choice. The empirical Bayesian method requires the least computational power, only providing a point estimate of hyperparameters without accounting for their underlying uncertainty. Although more computationally demanding, the fully Bayesian method captures the hyperparameter uncertainties effectively. As an intermediary, variational Bayesian method requires slightly more computational time compared to empirical method but can derive a variational posterior distribution that is largely similar to that obtained by the fully Bayesian method through MCMC sampling. In contrast, the point estimate ~~by~~ obtained through the empirical Bayesian method ~~lies-falls~~ at the third percentile of the posterior hyperparameter distributions ~~by~~ derived from the fully Bayesian method, ~~indicating a strong bias highlighting a significant bias introduced by the overly simplistic approach.~~

Because of the near continuous sea-level data with smoothly rising sea-level trend in North Carolina, the inference results from these three methods are similar. However, given that geological sea-level data ~~is~~ are often sparsely distributed across both spatial and temporal domains and may subject to abrupt change in rate, neglecting the underlying uncertainty of hyperparameters by empirical Bayesian method can result in a significant underestimation of the final inference uncertainty compared with fully Bayesian method.

4.2 Spatio-temporal analysis

~~The spatio-temporal analysis is~~ Spatio-temporal analysis presents a common challenge in paleo-environmental studies, ~~for~~ example, how to reconstruct continuous spatiotemporal such as reconstructing continuous spatio-temporal signals from sparse

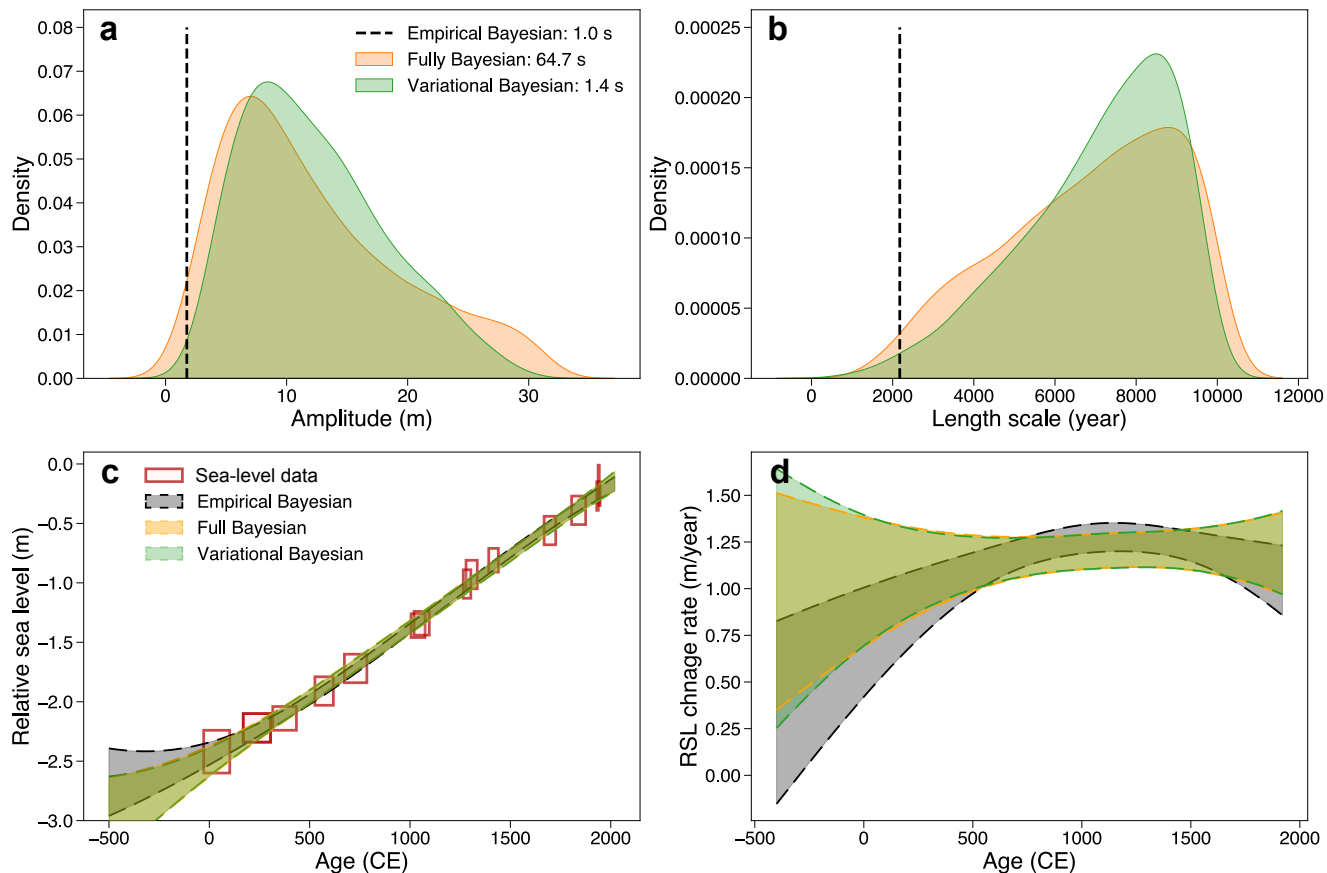


Figure 5. Analysis choices impact on temporal sea-level change inference at North Carolina. **Row 1, (a-b)** GP model hyperparameters optimization results along with the required computational time in second (based on a 2023 MacBook Pro with an Apple M2 Pro chip). **Row 2, common (c-d)** Common Era sea-level comparison between three analysis choices, the results indicate 90% credible interval of RSL change trend (**left**) and rate (**right**). **Paleo sea-level data here are modeled using a normal likelihood.** The horizontal and vertical range of red boxes indicates $\pm 2\sigma$ age and relative sea-level reconstruction uncertainties, respectively. Sea-level data here were reconstructed using near-continuous record from single cores of salt-marsh sediment from Sand Point (North Carolina, Kemp et al., 2011)

and noisy data. ~~For this perspective, PaleoSTeHM offers~~ To address this, PaleoSTeHM provides a range of options from approaches, spanning purely statistical to purely physical approaches. Utilizing methods. Here, we present an illustrative example to recover the spatio-temporal RSL pattern and its associated uncertainty. This analysis utilizes a sea-level database containing 1,043 proxy records spanning from 11 ka to the present, compiled by Ashe et al. (2019) from previous studies (Kemp et al., 2017a, b, 2015, 2014, 2013; Khan et al., 2017; Engelhart and Horton, 2012), ~~here we attempt to recover the spatio-temporal RSL pattern along with its associated uncertainty.~~ The database includes sea-level proxies such as salt marsh, mangrove, beach rock, and coral. All records were used to train the model except for 51 sea-level data points from New York

(Engelhart and Horton, 2012), which were reserved for cross-validation—a technique used to evaluate model performance on unseen data (shown as a gray dot in Figure 6m-p).

540 We demonstrate four process level models that have been used in previous studies: (i) a GP model with a zero mean function and multiple isotropic kernels (Ashe et al., 2019); (ii) a GP model with the mean function determined by a GIA model and multiple isotropic kernels (Walker et al., 2021; Kopp et al., 2016); (iii) a GP model with a zero mean function and a sampling covariance kernel determined by a GIA model ensemble (Kopp et al., 2009); and (iv) a purely GIA model ~~ensemble~~ ensemble (Lin et al., 2023a). All models assume a data-level model with a normal likelihood determined by RSL reconstruction un-
545 certainty and an additional white noise term. For this analysis, we implement a noisy-input framework to address temporal uncertainty and use the empirical Bayesian method to optimize hyperparameters for models *i*, *ii*, and *iii*, while model *iv* is optimized through the variational Bayesian method (Table A1).

For model *i*, we follow the kernel structure as in Ashe et al. (2019), which can be expressed as:

$$f(X) \sim GP(0, K_1(X, X')) \quad (18)$$

550

$$K_1(X, X') = g(t) + r(x, t) + l(x, t) \quad (19)$$

where $g(t)$ represents a spatially-uniform covariance function, while $r(x, t)$ and $l(x, t)$ are regional and local varying isotropic covariance functions, respectively. These are characterized by a 3/2 Matérn temporal kernel (Table 2) for $g(t)$, and a product of a 3/2 Matérn temporal kernel and a 1/2 Matérn spatial kernel 2) for $r(x, t)$ and $l(x, t)$, which are ~~differentiated by~~ distinguished
555 by their prior distributions of hyperparameters.

Similarly, model *ii* can be written as:

$$f(X) \sim GP(GIA(X), K_2(X, X')) \quad (20)$$

$$K_2(X, X') = r(x, t) + l(x, t) \quad (21)$$

560 here, the mean RSL expectation is determined by RSL prediction from ICE_7G ice model with VM5a Earth model (Roy and Peltier, 2018; Peltier et al., 2015), and $r(x, t)$ and $l(x, t)$ are the same as in equation 19. We do not include $g(t)$ ~~as we assume the mean function already captures it~~ kernel here as the RSL prediction derived from the ICE_7G model—embedded within the GP mean function and rigorously calibrated against comprehensive RSL and geodetic datasets across North America Roy and Peltier (2018); Peltier et al. (2015)—is assumed to adequately capture all spatially uniform signals.

565 Model *iii* can be denoted as:

$$f(X) \sim GP(0, K_3(X, X')) \quad (22)$$

$$K_3(X, X') = Cov(m(X), m(X')) \cdot exp(-|t - t'|/\tau^2) \quad (23)$$

here, $Cov(\cdot)$ here indicates a sampling covariance function through a physical model ensemble m . In this context, the covariance between data points is not directly determined by their spatio-temporal proximity but instead depends on the variance within the physical model ensemble, conditioned on various combinations of physical parameters. For this iv , where m includes an ensemble of forward GIA models: (a) ICE_7G ice model with VM5a Earth model; (b) PaleoMIST ice model (Gowan et al., 2021) with 71 km lithosphere and 0.3 and 70×10^{21} Pa s upper and lower mantle viscosity; and (c) ANU ice model (Lambeck et al., 2014) with 71 km lithosphere and 1 and 10×10^{21} Pa s upper and lower mantle viscosity. To expand the variability of physical model predictions, we create six synthetic GIA model outputs by enlarging or shrinking these three GIA model outputs by 1.5. Therefore, this physical model ensemble consists of predictions from nine models. More details about the physics-based GIA model used here can be found in Lin et al. (2023b). To stabilize the estimate and reduce variability related to finite sample size, we applied a temporal Gaussian taper function to this kernel, controlled by a parameter τ . Following Hay et al. (2015); Kopp et al. (2009), we set τ to 3000 years.

And lastly, model iv can be written as a weighted mean of different physical models:

$$f(X) = \sum_{n=1}^N \nu_i \text{GIA}_i(X) \quad (24)$$

$$\nu \sim \text{Dirichlet}(\alpha_d) \quad (25)$$

In this model, ν represents the relative weights associated with each GIA model. These probabilities follow a Dirichlet distribution (or multivariate beta distribution) characterized by a concentration parameter α_d . A value greater than 1 for α_d indicates a preference for a more evenly distributed probability across all models. In contrast, a value less than 1 indicates a preference for more concentrated probabilities on fewer models (Lin et al., 2023b). For this experiment, we set α_d according to each GIA model prediction fit to RSL observation (using weighted root mean square as a metric, see Table A1).

A comparison of RSL inference results between different spatio-temporal process level models is provided in Figure 6. At the purely statistical end of the process model spectrum, model i correlates RSL from various locations and times based solely on their spatio-temporal proximity, a property derived from the adopted isotropic kernels. According to model i , the RSL change along the US Atlantic coast during the Holocene was dominated by a spatially uniform signal (produced by the ~~regional-common~~ spatially-uniform kernel, $g(t)$, in equation 19; Table A1), which contributed to more than 25 m of RSL rise. In contrast, $r(x, t)$ and $l(x, t)$ only produce up to 5 m of spatially variable RSL signal, resulting in virtually no spatial pattern in the mean RSL prediction of this model. In the temporal domain, multiple studies have demonstrated that GP models like model i can accurately recover multi-millennial sea-level variation trends at locations with abundant sea-level observations, such as Florida, as shown in Figure 6a (Tang et al., 2023; Ashe et al., 2019; Cahill et al., 2015). This is further supported by the residual plot (Figure 6e), which exhibits a low wMSE (1.1) and minimal temporal structure in the residuals, indicating a good agreement with the observed data.

However, spatial inferences based on isotropic and stationary kernels of model i are often considered overly simplistic (Stein, 2005a), partly due to the sparse nature of geological data and the complexity of environmental change mechanisms. As seen in

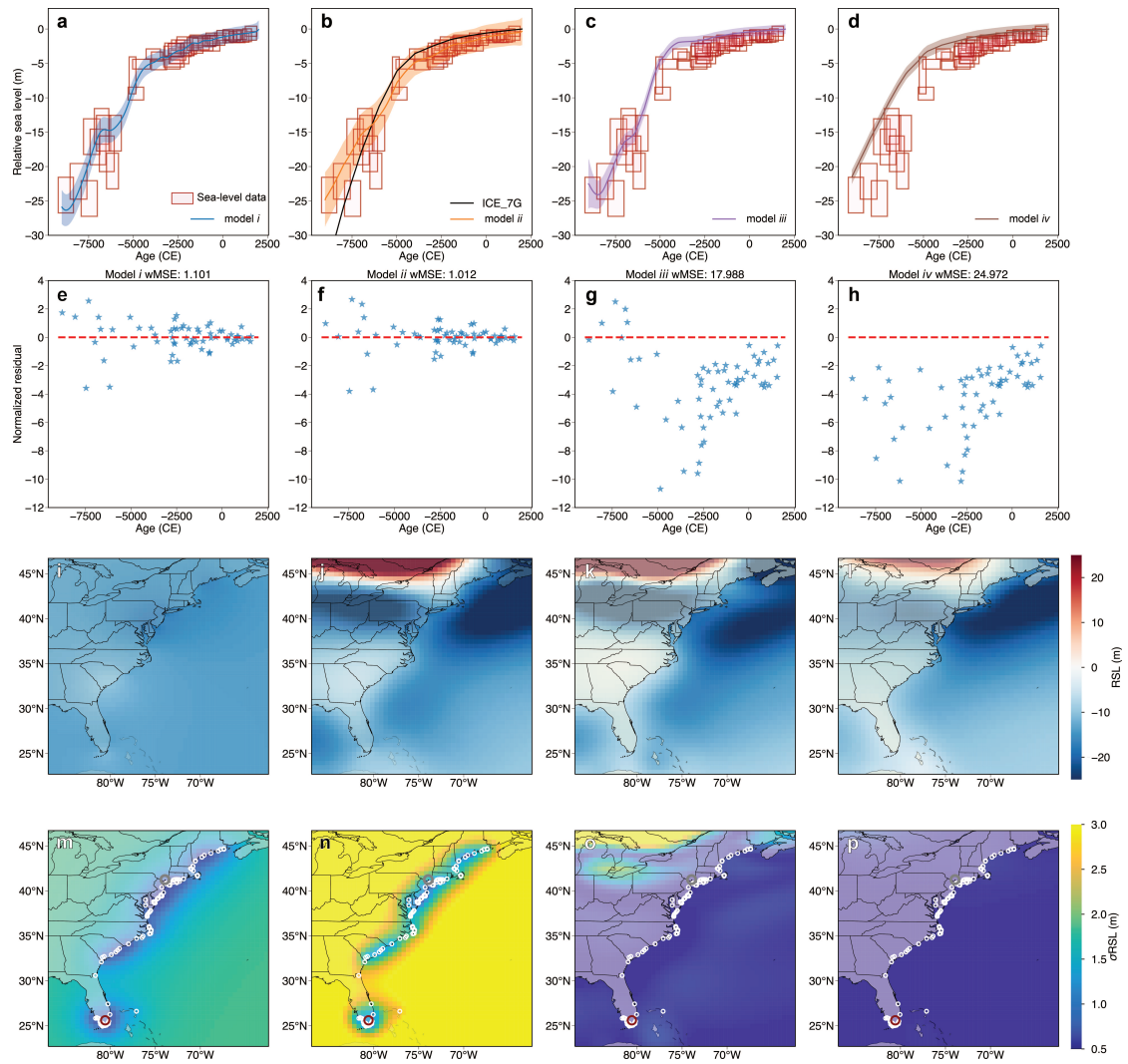


Figure 6. Process-level models—The impact on of process level modeling choices for spatio-temporal sea-level change inference along the US Atlantic coast. Each column represents a different process level model. (a-d) Row 1 displays the mean Mean and 90% credible or confidence interval for RSL predictions at Florida (indicated by a red dot in Row 3 m-p). The horizontal and vertical range of red boxes indicates $\pm 2\sigma$ age and relative sea-level reconstruction uncertainties, data from Khan et al. (2017). (b) For Model ~~iii~~ ii, the GP mean function, determined by the ~~ICE7G~~ ICE_7G model (Roy and Peltier, 2018; Peltier et al., 2015), is depicted with a black line(b). (e-h) Row 2 shows Normalized residuals, which represent the difference between observed and predicted values normalized by observational uncertainty, for Florida RSL predictions generated by each process-level modeling choice. Red dashed lines represent 0 error. The weighted mean square error for each model is given in title. (i-l) Mean RSL prediction for the year -5,500 CE. (i-l m-p) Row 3 illustrates the standard Standard deviation of the RSL prediction for the year -5,500 CE, where white dots indicate locations where each sea-level data collected and gray dot represent location of New York where 51 data points were held for testing purpose.

Figure 6im, geological sea-level data are mostly collected across paleo-coastal areas. Therefore, RSL inferences from model *i* are only representative of coastal areas (as opposed to terrestrial or marine areas) and cannot adequately reflect the physical knowledge of paleo sea-level change (e.g., the RSL uncertainty caused by the existence of the Laurentide Ice Sheet).

605 Model *ii* uses a deterministic GIA model (ICE_7G ice model with VM5a Earth model, Roy and Peltier, 2018; Peltier et al., 2015) as the GP mean function. By harnessing a physics-based model, model *ii* captures intricate spatial sea-level variation patterns due to the GIA-induced GRD effects (Figure 4). In this setup, the covariance functions describe residuals between the GIA model and RSL observations (mostly captured by $r(x, t)$ in equation 21). ~~At~~ Similar to model *i*, RSL predictions for Florida by model *ii* closely align with observations, demonstrating a low wMSE (1.0) and an unstructured residual distribution. In the spatial domain, at -5500 CE, model *ii* suggests the GIA model underestimates ~ 10 m RSL at New Jersey (Figure A4A5), which may reflect oversimplified physics (e.g., 3D rheology; Austermann et al., 2013) (e.g., neglecting 3D solid Earth rheology; Austermann et al., 2013), biased sampling of physical parameters (such as poorly-constrained ice history), or missing physical processes in the GIA model (e.g., sediment isostatic adjustment; Lin et al., 2023a). Because model *ii* assumes no uncertainty in GIA modeling, the uncertainty quantification here also relies solely on the radial distance from RSL data points (Figure 6jn).

615 Model *iii* employs utilizes a kernel constructed by sampling ~~covariance between different~~ covariances between various forward GIA models based upon on alternative ice and Earth models. ~~The incorporation of relevant physics in~~ By incorporating relevant physical processes into GP kernel construction enables, model *iii* to capture effectively captures anisotropic behaviors, non-stationarities, heterogeneities and teleconnections (Table 1) inherent in the physical processes, and teleconnections that are intrinsic to the physical dynamics of RSL change that cannot be easily described by normal but are challenging to represent with standard classes of covariance function functions (Table 2). For example instance, the size of the Laurentide Ice Sheet is positively correlated exhibits a positive correlation with RSL around the northern Great Lakes while negatively correlated displaying a negative correlation with RSL in peripheral bulge regions like such as New Jersey (Figure 6g). Whereas potential problems for this method include the computational burden required for k).

Model *iii* also has certain limitations, including the computational cost of thoroughly sampling physical model parameters, and structural errors in the physical model and the presence of structural errors within the physical models, such as oversimplifications or missing processes. Also, this method's omitted processes. As shown in the residual plot of Figure 6g, there is a significant mismatch between RSL observations and model *iii* predictions, reflected in a high wMSE value (17.99) and pronounced temporal structure in the residuals. This poor fit may stem from biased sampling of ice and Earth models, as well as the reliance on an oversimplified 1-D rheology. Furthermore, the model prioritizes fitting regions with denser data distribution, such as the mid-northern US Atlantic coast. For instance, model *iii* provides a good fit to unseen RSL observations in New York (Figure 7c,g), but this prioritization comes at the expense of accuracy in regions with sparser data, such as Florida, where substantial misfits are observed. Additionally, the posterior mean and standard deviation are not directly interpretable as generated by this method are less directly interpretable compared to those produced by model *iv*.

635 Lastly, model Model *iv* represents the purely physical end of the process level spectrum. It is equivalent to a process-level spectrum and is formulated as a weighted linear combination of physical models according to, with weights determined by data-model misfits (e.g., chi-square; Lin et al., 2021; Li et al., 2020; Lambeck et al., 2014) (e.g., wMSE and chi-square misfit; Lin et al., 2021; Li et al., 2020).

The mean and uncertainty estimated by this method reflect the parametric uncertainty of a certain inherent in a given physical model, allowing for direct physical interpretation of prediction results interpretation of physical parameters, such as calculating deriving posterior distributions of global ice history (Creel et al., 2024). However, this method is susceptible to model structural errors approach is also susceptible to structural errors within the model, similar to those observed in model *iii*. For instance, the uncertainty here is underestimated due to the The limited sample size of physical parameters (only nine models were used here). Additionally, due to the limited in this analysis—and the model’s tendency to prioritize fitting denser sea-level data in mid-northern locations result in uncertainty estimates that appear underestimated and biased, as illustrated by substantial misfits to observations (Figures 6d and h). Furthermore, the difficulty in directly quantifying certain physical parameters often leads to oversimplified model predictions. For example, the scarcity of direct constraints on ice history, the millennial-resolution ice models used in forward GIA modeling cannot capture centennial (Dalton et al., 2020) reduces the ability of forward GIA models to resolve centennial-scale sea-level variation as effectively as models *i-iii* variations. This limitation makes these models less effective in capturing centennial-scale variability compared to models *i-iii* (Figure 6).

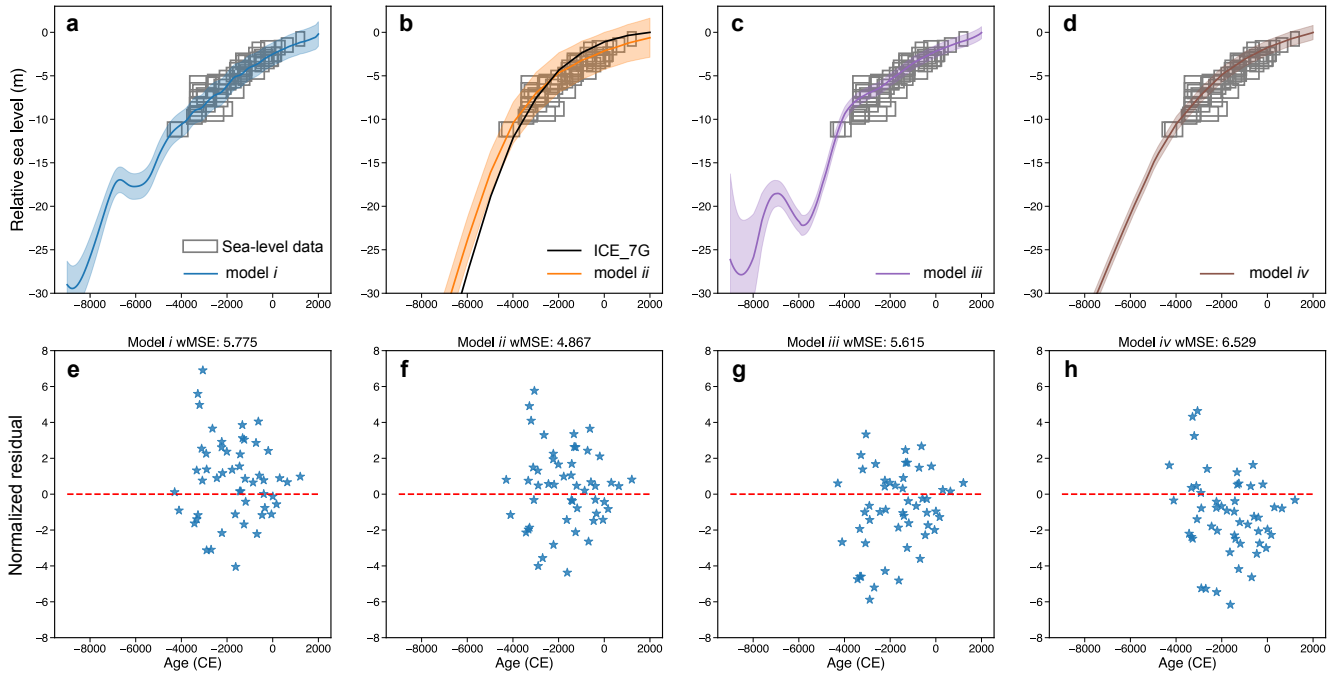


Figure 7. Performance of different process-level modeling choices on unseen sea-level data from New York. (a-d) Mean and 90% credible or confidence intervals for RSL predictions at New York (gray dot location shown in Figure 6m-p). The gray boxes represent $\pm 2\sigma$ uncertainties in both age and relative sea-level reconstructions, with data from Engelhart et al. (2009). For Model *ii*, the GP mean function, derived from the ICE_7G model (Roy and Peltier, 2018; Peltier et al., 2015), is shown as a black line (b). (e-h) Normalized residuals, which represent the difference between observed and predicted values normalized by observational uncertainty, for New York RSL predictions generated by each process-level modeling choice. Red dashed lines represent 0 error. The weighted mean square error for each model is given in title.

Due to the dense distribution of sea-level data along the mid-northern US Atlantic coast, all models effectively capture the general trend of RSL variation observed in the withheld data from New York (Figure 7), as evidenced by low wMSE values and minimal temporal structure in the residual plots. While model *ii* achieves the lowest wMSE and model *iv* exhibits slight bias in its residuals, all models demonstrate comparable overall performance in reconstructing sea-level changes. In contrast, significant variation in model performance is observed for RSL predictions at Florida (Figure 6)–a–h), where models *iii* and *iv* show substantial misfit to RSL observations. This misfit stems from the requirement to preserve the overall consistency of the physical systems constrained by the ensemble of physical models.

It is important to note that high-quality and standardized datasets, such as those available for the mid-northern US Atlantic coast, are rare in many paleoenvironmental fields, such as deep-sea isotopes or ice core records (Shackleton et al., 2021; Lemieux-Dudon et al., 2021). Consequently, users must carefully evaluate factors such as data availability, computational resources, the need for interpretability, and the level of understanding of underlying physical processes when selecting a process level model. Generally, physics-based models offer superior interpretability and better extrapolation capabilities to spatio-temporal locations with minimal data, as they are rooted in well-established theoretical frameworks. However, discovering and validating new physical laws can be time-intensive and often computationally demanding. In contrast, machine learning or statistical approaches provide flexibility and computational efficiency but often face challenges in extrapolating non-linear functions (Xu et al., 2020; Goodfellow et al., 2016). Whereas, they require large volumes of training data and rigorous validation to ensure consistency with physical principles. For a more detailed discussion on the integration of physics-based and machine learning models, readers are referred to Lai et al. (2024).

5 Discussion

5.1 Generalization for paleo-environmental problems

~~Although all functionality demonstrations in this paper focus on~~ This paper focuses on demonstrating the functionality of PaleoSTeHM in paleo-sea-level applications, ~~the great flexibility of the hierarchical model.~~ However, the flexibility of hierarchical models, where any statistical model can be interpreted as a hierarchical model enables the PaleoSTeHM modeling hierarchically, allows the PaleoSTeHM framework to be readily applicable to most a wide range of paleo-environmental problems. Employing a hierarchical model to address various paleo-environmental problems enhances transparency by clearly By using hierarchical models, transparency is enhanced by distinguishing between modeling assumptions and analysis-analytical methods, as well as between process variability and separating process variability from observational noise.

~~Because of the commonality between~~ The common characteristics of paleo sea-level and datasets, such as sparsity and discreteness, are shared by many other paleo-environmental datasets that are characterized by sparse and discrete, such as, including paleo temperature (PAGES2k Consortium, 2017), past ice sheet thickness (Small et al., 2019), and sediment deposition depth (Wang et al., 2018). As a result, the data and process-level models introduced process-level models introduced in this paper can be readily generalized to those these paleo-environmental fields. For example, Tingley et al. (2012) and Stein (2005b) suggested it is reasonable to use proposed that using a GP model is a reasonable approach to describe latent some-space-time

climate processes ~~like-, such as~~ annual mean surface temperature anomalies and daily wind speed; Lin et al. (2023a) applied a spatial GP model to recover the spatial pattern of Holocene coral reef depth based on a Holocene coral reef deposition depth database across the Great Barrier Reef (Hinestrosa et al., 2022); and Caesar et al. (2018) implemented a change-point model on multiple proxy datasets to detect significant reductions in the strength of the Atlantic Meridional Overturning Circulation.

Beyond the process level models featured in PaleoSTeHM v1.0, various approaches have been employed for paleo-environmental analyses. Common techniques for addressing problems in this field include principal component analysis, equivalent to the empirical orthogonal function method when temporal aspects are considered, autoregressive models ~~;~~ and generalized additive models. For instance, Shakun and Carlson (2010) used an empirical orthogonal function approach to detect modes of deglacial temperature variability ~~;~~ and Piecuch et al. (2017) adopted a degree-1 autoregressive model to reconstruct sea-level evolution using tide gauge data, ~~and Simpson (2018)~~ Simpson (2018) and Upton et al. (2023) developed a series of generalized additive models to model ~~paleo-ecological time-series~~ paleo-ecology and paleo sea-level, respectively. While the reimplementations of these models in PaleoSTeHM is beyond the scope of this paper, doing so would benefit from the framework's multiple analysis options and its capacity for smooth integration with flexible data and parameter-level models.

695 5.2 Future developments

From a scientific perspective, numerous promising directions exist for further development of PaleoSTeHM.

Existing data-level models only support a common class of likelihood. However, in paleo-environmental studies, it is typical for proxy data to be subject to complex likelihoods (Ashe et al., 2022; Hibbert et al., 2016). For instance, organic matter that has been radiocarbon-dated undergoes a calibration procedure to account for the time-evolving atmospheric carbon concentration, which can yield a data chronology characterized by multi-modal distributions that significantly differ from each other. Similarly, it is common for paleo-environmental studies to use multiple types of proxy data with different likelihoods to infer a common signal. Recently, new approaches have been developed to account for ~~nonparametric~~ non-parametric proxy distributions within a hierarchical modeling framework (e.g., Ashe et al., 2022), which could better characterize the underlying uncertainty but can be computationally expansive.

705 While PaleoSTeHM allows users to specify any number of change points (m in equation 10) in the model, determining the optimal number of change points can be challenging and may require additional modeling strategies. Recent advancements, such as Bayesian transdimensional models (e.g., Sambridge, 2016; Bodin et al., 2012; Gallagher et al., 2011), provide a flexible framework by treating the number of change points as an unknown parameter, allowing it to be inferred alongside other model parameters. Incorporating such approaches into PaleoSTeHM is a potential avenue for future development to address this complexity in abrupt paleoenvironmental change problems.

The current GP Kernel module incorporates commonly used kernel options that are stationary, isotropic ~~;~~ and ~~space-time separable (definitions in Table 1)~~ space-time separable. While these assumptions simplify calculations significantly, they may not be suitable for some environmental applications. For example, temperature and dew point variations often exhibit strong non-stationary behavior influenced by diverse geographic and atmospheric conditions (Poppick and Stein, 2014).
715 Additionally, the assumption of stationarity may cause rate uncertainty estimates to fail in properly reflecting the reduced

uncertainty expected during periods with abundant data, as shown in Figure 5d, largely because the model's variance is uniformly applied across the entire temporal domain (Heinonen et al., 2016). Furthermore, temperature anomalies over the last two millennia (Mann et al., 2008) demonstrate strong space-time interactions, which cannot be captured by a space-time separable kernel (Tingley et al., 2012). Developing a scientifically richer class of kernel structures could be an important future advancement for PaleoSTeHM. However, given the fundamental differences across various paleo-environmental problems, generalizing sophisticated kernel structures to multiple fields remains challenging.

Another outstanding issue for GP based process level models is scalability, the standard GP models included in PaleoSTeHM v1.0 cannot scale well to large data sets (>10 thousands data points) due to the computational cost, which increases at a rate of $\mathcal{O}(n^3)$, where n is the number of data points (Hensman et al., 2013). Thus, implementing alternative classes of GP models within PaleoSTeHM to model large data sets, especially when incorporating modern environmental observations, which often consist of millions of data points, is an important next step for PaleoSTeHM to develop in the future. Some potentially efficient methods include sparse GP (Quinonero-Candela and Rasmussen, 2005), stochastic variational GP (Hensman et al., 2013), and exact GP with black-box matrix-matrix inference (Wang et al., 2019).

Building upon machine learning infrastructure, another promising direction for the future development of PaleoSTeHM is integrating spatio-temporal hierarchical modeling with machine learning-based emulators as a process-level model. An emulator indicates a statistical model that mimics the behavior of the physics-based simulator but is computationally cheap to run (Reichstein et al., 2019), which is particularly useful for fast sensitivity analysis, model parameter calibration, and derivation of confidence intervals for the estimate. The use of statistical emulators trained by physical models will enable hierarchical models to capture the non-stationary physical systems better and enable better interpretation of the modeling results. For paleo-environment, Holden et al. (2019) presents a GP-based emulator for an atmosphere-ocean general circulation model with intermediate-complexity, and (Lin et al., 2023b) and Lin et al. (2023b) developed a neural network-based emulator for GIA-induced global sea-level change.

More broadly, PaleoSTeHM has been developed by a small team specialized in modeling paleo sea-level changes over multi-millennial time scales. Moving forward, a critical objective is to expand PaleoSTeHM into a larger-scale paleo-environmental community project, where modules are developed autonomously by diverse research teams. The design of PaleoSTeHM, which allows modules to act as wrappers for independently developed code, is specifically intended to facilitate this collaborative effort.

6 Conclusion



Paleo-environmental records provide critical out-of-sample information essential for contextualizing current global changes and testing models used to simulate future environmental scenarios. However, our understanding of past environmental changes is often complicated by the sparse nature of geological records, geochronological uncertainties, and the indirect relationships between proxies and ecological variables. Hierarchical modeling offers a conceptually straightforward framework to address these challenges, though the limited availability of user-friendly software often hinders it. PaleoSTeHM offers a flexible and

open-source platform that facilitates the rapid and easy implementation of hierarchical models for paleo-environmental applications. The inclusion of multiple process-level models in PaleoSTeHM allows it to be readily applicable across a broad spectrum of paleo-environmental studies. ~~In contrast~~ Additionally, its flexibility allows for customization to meet the specific needs of diverse paleo-environmental problems, such as using different Gaussian Process kernels or substituting alternative process-level models.

Code and data availability. The development version of PaleoSTeHM is available under an MIT license in a Git version-controlled repository at <https://github.com/radical-collaboration/PaleoSTeHM> (last access: 16 January 2025). The latest release is archived on Zenodo with the identifier <https://doi.org/10.5281/zenodo.12730141> (Lin et al., 2024). Documentation of PaleoSTeHM is available at <https://paleostehm.org/>. All codes required to generate results and figures shown in section 4 are available in the repository. Video tutorials are available at https://youtube.com/playlist?list=PLR4-1Y89NM_x3zwnxc5nI2mU3pplGzIa3&si=5VoDvpZAWwLE2by4.

Author contributions. YL developed the PaleoSTeHM architecture and modules. REK and SJ conceived the project and REK supervised and administered it. AR guide software engineering of PaleoSTeHM, ELA provided YL guidance in statistical modeling. All authors contributed to code development and the writing and editing of the paper.

Competing interests. The authors declare no competing interests.

Acknowledgements. We thank Chris Piecuch, Roger Creel, Fiona Hibbert, Jeremy Ely, Sönke Dangendorf  and the participants of PaleoSTeHM workshops for their valuable feedback in the development of PaleoSTeHM. Yucheng Lin, Robert E. Kopp, Alexander Reedy  and Shantenu Jha are supported by the U.S. National Science Foundation under awards 2002437 and 2148265. The authors acknowledge PALSEA, a working group of the International Union for Quaternary Sciences (INQUA) and Past Global Changes (PAGES), which received support from the Swiss Academy of Sciences and the Chinese Academy of Sciences.

References

- Aitken, A. C.: IV.—On least squares and linear combination of observations, *Proceedings of the Royal Society of Edinburgh*, 55, 42–48, 1936.
- Ashe, E. L., Cahill, N., Hay, C., Khan, N. S., Kemp, A., Engelhart, S. E., Horton, B. P., Parnell, A. C., and Kopp, R. E.: Statistical modeling of rates and trends in Holocene relative sea level, *Quaternary Science Reviews*, 204, 58–77, 2019.
- Ashe, E. L., Khan, N. S., Toth, L. T., Dutton, A., and Kopp, R. E.: A statistical framework for integrating nonparametric proxy distributions into geological reconstructions of relative sea level, *Advances in Statistical Climatology, Meteorology and Oceanography*, 8, 1–29, 2022.
- Austermann, J., Mitrovica, J. X., Latychev, K., and Milne, G. A.: Barbados-based estimate of ice volume at Last Glacial Maximum affected by subducted plate, *Nature Geoscience*, 6, 553–557, 2013.
- Belloni, A. and Chernozhukov, V.: On the computational complexity of MCMC-based estimators in large samples, 2009.
- Bingham, E., Chen, J. P., Jankowiak, M., Obermeyer, F., Pradhan, N., Karaletsos, T., Singh, R., Szerlip, P., Horsfall, P., and Goodman, N. D.: Pyro: Deep universal probabilistic programming, *Journal of machine learning research*, 20, 1–6, 2019.
- Blaauw, M.: Methods and code for ‘classical’ age-modelling of radiocarbon sequences, *Quaternary geochronology*, 5, 512–518, 2010.
- Blei, D. M., Kucukelbir, A., and McAuliffe, J. D.: Variational inference: A review for statisticians, *Journal of the American statistical Association*, 112, 859–877, 2017.
- Bodin, T., Sambridge, M., Tkalčić, H., Arroucau, P., Gallagher, K., and Rawlinson, N.: Transdimensional inversion of receiver functions and surface wave dispersion, *Journal of geophysical research: solid earth*, 117, 2012.
- Bürkner, P.-C.: brms: An R package for Bayesian multilevel models using Stan, *Journal of statistical software*, 80, 1–28, 2017.
- Caesar, L., Rahmstorf, S., Robinson, A., Feulner, G., and Saba, V.: Observed fingerprint of a weakening Atlantic Ocean overturning circulation, *Nature*, 556, 191–196, 2018.
- Caesar, L., McCarthy, G., Thornalley, D., Cahill, N., and Rahmstorf, S.: Current Atlantic meridional overturning circulation weakest in last millennium, *Nature Geoscience*, 14, 118–120, 2021.
- Cahill, N., Kemp, A. C., Horton, B. P., and Parnell, A. C.: Modeling sea-level change using errors-in-variables integrated Gaussian processes, 2015.
- Carlin, B. P., Gelfand, A. E., and Smith, A. F.: Hierarchical Bayesian analysis of changepoint problems, *Journal of the royal statistical society: series C (applied statistics)*, 41, 389–405, 1992.
- Creel, R. C., Austermann, J., Kopp, R. E., Khan, N. S., Albrecht, T., and Kingslake, J.: Global mean sea level likely higher than present during the holocene, *Nature Communications*, 15, 10 731, 2024.
- Cressie, N. and Wike, C. K.: *Statistics for spatio-temporal data*, John Wiley & Sons, 2015.
- Crichton, K. A., Wilson, J. D., Ridgwell, A., Boscolo-Galazzo, F., John, E. H., Wade, B. S., and Pearson, P. N.: What the geological past can tell us about the future of the ocean’s twilight zone, *Nature Communications*, 14, 2376, 2023.
- Dalton, A. S., Margold, M., Stokes, C. R., Tarasov, L., Dyke, A. S., Adams, R. S., Allard, S., Arends, H. E., Atkinson, N., Attig, J. W., et al.: An updated radiocarbon-based ice margin chronology for the last deglaciation of the North American Ice Sheet Complex, *Quaternary Science Reviews*, 234, 106 223, 2020.
- DeConto, R. M. and Pollard, D.: Contribution of Antarctica to past and future sea-level rise, *Nature*, 531, 591–597, 2016.
- Dey, D. K., Ghosh, S. K., and Mallick, B. K.: *Generalized linear models: A Bayesian perspective*, CRC Press, 2000.

- Engelhart, S. E. and Horton, B. P.: Holocene sea level database for the Atlantic coast of the United States, *Quaternary Science Reviews*, 54, 12–25, 2012.
- Engelhart, S. E., Horton, B. P., Douglas, B. C., Peltier, W. R., and Törnqvist, T. E.: Spatial variability of late Holocene and 20th century sea-level rise along the Atlantic coast of the United States, *Geology*, 37, 1115–1118, 2009.
- Farrell, W. and Clark, J. A.: On postglacial sea level, *Geophysical Journal International*, 46, 647–667, 1976.
- Gallagher, K., Bodin, T., Sambridge, M., Weiss, D., Kylander, M., and Large, D.: Inference of abrupt changes in noisy geochemical records using transdimensional changepoint models, *Earth and Planetary Science Letters*, 311, 182–194, 2011.
- Gelman, A. and Rubin, D. B.: Inference from iterative simulation using multiple sequences, *Statistical science*, 7, 457–472, 1992.
- Goodfellow, I., Bengio, Y., and Courville, A.: *Deep learning*, MIT press, 2016.
- Gowan, E. J., Zhang, X., Khosravi, S., Rovere, A., Stocchi, P., Hughes, A. L., Gyllencreutz, R., Mangerud, J., Svendsen, J.-I., and Lohmann, G.: A new global ice sheet reconstruction for the past 80 000 years, *Nature communications*, 12, 1–9, 2021.
- Görtler, J., Kehlbeck, R., and Deussen, O.: A Visual Exploration of Gaussian Processes, *Distill*, <https://doi.org/10.23915/distill.00017>, <https://distill.pub/2019/visual-exploration-gaussian-processes>, 2019.
- Hastings, W. K.: *Monte Carlo sampling methods using Markov chains and their applications*, 1970.
- Hay, C. C., Morrow, E., Kopp, R. E., and Mitrovica, J. X.: Probabilistic reanalysis of twentieth-century sea-level rise, *Nature*, 517, 481–484, 2015.
- Heaton, T. J., Köhler, P., Butzin, M., Bard, E., Reimer, R. W., Austin, W. E., Ramsey, C. B., Grootes, P. M., Hughen, K. A., Kromer, B., et al.: Marine20—the marine radiocarbon age calibration curve (0–55,000 cal BP), *Radiocarbon*, 62, 779–820, 2020.
- Heinonen, M., Mannerström, H., Rousu, J., Kaski, S., and Lähdesmäki, H.: Non-stationary gaussian process regression with hamiltonian monte carlo, in: *Artificial Intelligence and Statistics*, pp. 732–740, PMLR, 2016.
- Hensman, J., Fusi, N., and Lawrence, N. D.: Gaussian processes for big data, *arXiv preprint arXiv:1309.6835*, 2013.
- Hibbert, F. D., Rohling, E. J., Dutton, A., Williams, F. H., Chutcharavan, P. M., Zhao, C., and Tamisiea, M. E.: Coral indicators of past sea-level change: A global repository of U-series dated benchmarks, *Quaternary Science Reviews*, 145, 1–56, 2016.
- Hinestrosa, G., Webster, J. M., and Beaman, R. J.: New constraints on the postglacial shallow-water carbonate accumulation in the Great Barrier Reef, *Scientific Reports*, 12, 1–18, 2022.
- Hoffman, M. and Blei, D.: Stochastic structured variational inference, in: *Artificial Intelligence and Statistics*, pp. 361–369, PMLR, 2015.
- Hoffman, M. D., Gelman, A., et al.: The No-U-Turn sampler: adaptively setting path lengths in Hamiltonian Monte Carlo., *J. Mach. Learn. Res.*, 15, 1593–1623, 2014.
- Holden, P. B., Edwards, N. R., Rangel, T. F., Pereira, E. B., Tran, G. T., and Wilkinson, R. D.: PALEO-PGEM v1. 0: a statistical emulator of Pliocene–Pleistocene climate, *Geoscientific Model Development*, 12, 5137–5155, 2019.
- Hunter, J. D.: Matplotlib: A 2D graphics environment, *Computing in Science & Engineering*, 9, 90–95, <https://doi.org/10.1109/MCSE.2007.55>, 2007.
- Islam, A. R. M. T., Karim, M. R., and Mondol, M. A. H.: Appraising trends and forecasting of hydroclimatic variables in the north and northeast regions of Bangladesh, *Theoretical and Applied climatology*, 143, 33–50, 2021.
- Kageyama, M., Braconnot, P., Harrison, S. P., Haywood, A. M., Jungclaus, J. H., Otto-Bliesner, B. L., Peterschmitt, J.-Y., Abe-Ouchi, A., Albani, S., Bartlein, P. J., et al.: The PMIP4 contribution to CMIP6—Part 1: Overview and over-arching analysis plan, *Geoscientific Model Development*, 11, 1033–1057, 2018.

- Kay, R. F., Vizcaíno, S. F., Bargo, M. S., Spradley, J. P., and Cuitiño, J. I.: Paleoenvironments and paleoecology of the Santa Cruz Formation (early-middle Miocene) along the Río Santa Cruz, Patagonia (Argentina), *Journal of South American Earth Sciences*, 109, 103 296, 2021.
- Kemp, A. C., Horton, B. P., Donnelly, J. P., Mann, M. E., Vermeer, M., and Rahmstorf, S.: Climate related sea-level variations over the past two millennia, *Proceedings of the National Academy of Sciences*, 108, 11 017–11 022, 2011.
- 845 Kemp, A. C., Horton, B. P., Vane, C. H., Bernhardt, C. E., Corbett, D. R., Engelhart, S. E., Anisfeld, S. C., Parnell, A. C., and Cahill, N.: Sea-level change during the last 2500 years in New Jersey, USA, *Quaternary Science Reviews*, 81, 90–104, 2013.
- Kemp, A. C., Bernhardt, C. E., Horton, B. P., Kopp, R. E., Vane, C. H., Peltier, W. R., Hawkes, A. D., Donnelly, J. P., Parnell, A. C., and Cahill, N.: Late Holocene sea-and land-level change on the US southeastern Atlantic coast, *Marine Geology*, 357, 90–100, 2014.
- Kemp, A. C., Hawkes, A. D., Donnelly, J. P., Vane, C. H., Horton, B. P., Hill, T. D., Anisfeld, S. C., Parnell, A. C., and Cahill, N.: Relative
850 sea-level change in Connecticut (USA) during the last 2200 yrs, *Earth and Planetary Science Letters*, 428, 217–229, 2015.
- Kemp, A. C., Hill, T. D., Vane, C. H., Cahill, N., Orton, P. M., Talke, S. A., Parnell, A. C., Sanborn, K., and Hartig, E. K.: Relative sea-level trends in New York City during the past 1500 years, *The Holocene*, 27, 1169–1186, 2017a.
- Kemp, A. C., Kegel, J. J., Culver, S. J., Barber, D. C., Mallinson, D. J., Leorri, E., Bernhardt, C. E., Cahill, N., Riggs, S. R., Woodson, A. L., et al.: Extended late Holocene relative sea-level histories for North Carolina, USA, *Quaternary Science Reviews*, 160, 13–30, 2017b.
- 855 Kemp, A. C., Wright, A. J., Edwards, R. J., Barnett, R. L., Brain, M. J., Kopp, R. E., Cahill, N., Horton, B. P., Charman, D. J., Hawkes, A. D., et al.: Relative sea-level change in Newfoundland, Canada during the past 3000 years, *Quaternary Science Reviews*, 201, 89–110, 2018.
- Kendall, R. A., Mitrovica, J. X., and Milne, G. A.: On post-glacial sea level–II. Numerical formulation and comparative results on spherically symmetric models, *Geophysical Journal International*, 161, 679–706, 2005.
- Khan, N. S., Ashe, E., Horton, B. P., Dutton, A., Kopp, R. E., Brocard, G., Engelhart, S. E., Hill, D. F., Peltier, W., Vane, C. H., et al.: Drivers
860 of Holocene sea-level change in the Caribbean, *Quaternary Science Reviews*, 155, 13–36, 2017.
- Khan, N. S., Horton, B. P., Engelhart, S., Rovere, A., Vacchi, M., Ashe, E. L., Törnqvist, T. E., Dutton, A., Hijma, M. P., and Shennan, I.: Inception of a global atlas of sea levels since the Last Glacial Maximum, *Quaternary Science Reviews*, 220, 359–371, 2019.
- Khan, N. S., Ashe, E., Moyer, R. P., Kemp, A. C., Engelhart, S. E., Brain, M. J., Toth, L. T., Chappel, A., Christie, M., Kopp, R. E., et al.: Relative sea-level change in South Florida during the past ~ 5000 years, *Global and Planetary Change*, 216, 103 902, 2022.
- 865 Kingma, D. P. and Ba, J.: Adam: A method for stochastic optimization, *arXiv preprint arXiv:1412.6980*, 2014.
- Ko, J., Kim, K., Kim, W. C., and Gardner, J. R.: Provably Scalable Black-Box Variational Inference with Structured Variational Families, *arXiv preprint arXiv:2401.10989*, 2024.
- Kopp, R. E., Simons, F. J., Mitrovica, J. X., Maloof, A. C., and Oppenheimer, M.: Probabilistic assessment of sea level during the last interglacial stage, *Nature*, 462, 863–867, 2009.
- 870 Kopp, R. E., Horton, R. M., Little, C. M., Mitrovica, J. X., Oppenheimer, M., Rasmussen, D., Strauss, B. H., and Tebaldi, C.: Probabilistic 21st and 22nd century sea-level projections at a global network of tide-gauge sites, *Earth’s future*, 2, 383–406, 2014.
- Kopp, R. E., Kemp, A. C., Bittermann, K., Horton, B. P., Donnelly, J. P., Gehrels, W. R., Hay, C. C., Mitrovica, J. X., Morrow, E. D., and Rahmstorf, S.: Temperature-driven global sea-level variability in the Common Era, *Proceedings of the National Academy of Sciences*, 113, E1434–E1441, 2016.
- 875 Kopp, R. E., Garner, G. G., Hermans, T. H., Jha, S., Kumar, P., Reedy, A., Slangen, A. B., Turilli, M., Edwards, T. L., Gregory, J. M., et al.: The Framework for Assessing Changes To Sea-level (FACTS) v1. 0: a platform for characterizing parametric and structural uncertainty in future global, relative, and extreme sea-level change, *Geoscientific Model Development*, 16, 7461–7489, 2023.

- Lai, C.-Y., Hassanzadeh, P., Sheshadri, A., Sonnewald, M., Ferrari, R., and Balaji, V.: Machine learning for climate physics and simulations, *Annual Review of Condensed Matter Physics*, 16, 2024.
- 880 Lambeck, K., Rouby, H., Purcell, A., Sun, Y., and Sambridge, M.: Sea level and global ice volumes from the Last Glacial Maximum to the Holocene, *Proceedings of the National Academy of Sciences*, 111, 15 296–15 303, 2014.
- Laplace, P.-S.: Sur les approximations des formules qui sont fonctions de tres grands nombres et sur leur application aux probabilites, *Œuvres complètes*, 12, 301–345, 1810.
- Lemieux-Dudon, B., Blayo, E., Petit, J.-R., Waelbroeck, C., Svensson, A., Ritz, C., Barnola, J.-M., Narcisi, B. M., and Parrenin, F.: Consistent
885 dating for Antarctic and Greenland ice cores, *Quaternary Science Reviews*, 29, 8–20, 2010.
- Li, T., Wu, P., Wang, H., Steffen, H., Khan, N. S., Engelhart, S. E., Vacchi, M., Shaw, T. A., Peltier, W. R., and Horton, B. P.: Uncertainties of glacial isostatic adjustment model predictions in North America associated with 3D structure, *Geophysical Research Letters*, 47, e2020GL087 944, 2020.
- Lin, Y. and Yousefi, M.: *Barystatic Sea-Level Changes Since the Last Glacial Maximum*, vol. 6, pp. 26–38, Elsevier, 2025.
- 890 Lin, Y., Hibbert, F. D., Whitehouse, P. L., Woodroffe, S. A., Purcell, A., Shennan, I., and Bradley, S. L.: A reconciled solution of Meltwater Pulse 1A sources using sea-level fingerprinting, *Nature communications*, 12, 1–11, 2021.
- Lin, Y., Whitehouse, P. L., Hibbert, F. D., Woodroffe, S. A., Hinestrosa, G., and Webster, J. M.: Relative sea level response to mixed carbonate-siliciclastic sediment loading along the Great Barrier Reef margin, *Earth and Planetary Science Letters*, 607, 2023a.
- Lin, Y., Whitehouse, P. L., Valentine, A. P., and Woodroffe, S. A.: GEORGIA: A graph neural network based EmulatOR for glacial isostatic
895 adjustment, *Geophysical Research Letters*, 50, e2023GL103 672, 2023b.
- Lin, Y., Kopp, R., Reedy, A., Jha, S., and Turilli, M.: PaleoSTeHM v1.0: a modern, scalable spatio-temporal hierarchical modeling framework for paleo-environmental data (v1.0), <https://doi.org/10.5281/zenodo.12730140>, 2024.
- Lyu, K., Zhang, X., Church, J. A., Slangen, A. B., and Hu, J.: Time of emergence for regional sea-level change, *Nature Climate Change*, 4, 1006–1010, 2014.
- 900 Mann, M. E., Zhang, Z., Hughes, M. K., Bradley, R. S., Miller, S. K., Rutherford, S., and Ni, F.: Proxy-based reconstructions of hemispheric and global surface temperature variations over the past two millennia, *Proceedings of the National Academy of Sciences*, 105, 13 252–13 257, 2008.
- McHutchon, A. and Rasmussen, C.: Gaussian process training with input noise, *Advances in neural information processing systems*, 24, 2011.
- 905 Meltzner, A. J., Switzer, A. D., Horton, B. P., Ashe, E., Qiu, Q., Hill, D. F., Bradley, S. L., Kopp, R. E., Hill, E. M., Majewski, J. M., et al.: Half-metre sea-level fluctuations on centennial timescales from mid-Holocene corals of Southeast Asia, *Nature Communications*, 8, 14 387, 2017.
- Mitrovica, J., Gomez, N., Morrow, E., Hay, C., Latychev, K., and Tamisiea, M.: On the robustness of predictions of sea level fingerprints, *Geophysical Journal International*, 187, 729–742, 2011.
- 910 Mitrovica, J. X., Tamisiea, M. E., Davis, J. L., and Milne, G. A.: Recent mass balance of polar ice sheets inferred from patterns of global sea-level change, *Nature*, 409, 1026–1029, 2001.
- Neal, R. M. et al.: MCMC using Hamiltonian dynamics, *Handbook of markov chain monte carlo*, 2, 2, 2011.
- PAGES2k Consortium: A global multiproxy database for temperature reconstructions of the Common Era, *Scientific data*, 4, 2017.
- PAGES2k Consortium: Consistent multidecadal variability in global temperature reconstructions and simulations over the Common Era,
915 *Nature Geoscience*, 12, 643–649, <https://doi.org/10.1038/s41561-019-0400-0>, 2019.

- Paszke, A., Gross, S., Chintala, S., Chanan, G., Yang, E., DeVito, Z., Lin, Z., Desmaison, A., Antiga, L., and Lerer, A.: Automatic differentiation in PyTorch, 2017.
- Paszke, A., Gross, S., Massa, F., Lerer, A., Bradbury, J., Chanan, G., Killeen, T., Lin, Z., Gimelshein, N., Antiga, L., Desmaison, A., Kopf, A., Yang, E., DeVito, Z., Raison, M., Tejani, A., Chilamkurthy, S., Steiner, B., Fang, L., Bai, J., and Chintala, S.: PyTorch: An Imperative Style, High-Performance Deep Learning Library, in: *Advances in Neural Information Processing Systems 32*, pp. 8024–8035, Curran Associates, Inc., <http://papers.neurips.cc/paper/9015-pytorch-an-imperative-style-high-performance-deep-learning-library.pdf>, 2019.
- Peltier, W., Argus, D., and Drummond, R.: Space geodesy constrains ice age terminal deglaciation: The global ICE-6G_C (VM5a) model, *Journal of Geophysical Research: Solid Earth*, 120, 450–487, 2015.
- Peltier, W. R.: Global glacial isostasy and the surface of the ice-age Earth: the ICE-5G (VM2) model and GRACE, *Annu. Rev. Earth Planet. Sci.*, 32, 111–149, 2004.
- Piecuch, C. G., Huybers, P., and Tingley, M. P.: Comparison of full and empirical Bayes approaches for inferring sea-level changes from tide-gauge data, *Journal of Geophysical Research: Oceans*, 122, 2243–2258, 2017.
- Pollack, A., Campbell, J. E., Condon, M., Cooper, C., Coronese, M., Doss-Gollin, J., Hegde, P., Helgeson, C., Kwakkel, J., Lesk, C., et al.: Peer-reviewed climate change research has a transparency problem. The scientific community needs to do better, 2024.
- Poppick, A. and Stein, M. L.: Using covariates to model dependence in nonstationary, high-frequency meteorological processes, *Environmetrics*, 25, 293–305, 2014.
- Quinonero-Candela, J. and Rasmussen, C. E.: A unifying view of sparse approximate Gaussian process regression, *The Journal of Machine Learning Research*, 6, 1939–1959, 2005.
- Rasmussen, C. E. and Williams, C. K.: *Gaussian processes for machine learning*, vol. 2, MIT press Cambridge, MA, 2006.
- Reichstein, M., Camps-Valls, G., Stevens, B., Jung, M., Denzler, J., Carvalhais, N., et al.: Deep learning and process understanding for data-driven Earth system science, *Nature*, 566, 195–204, 2019.
- Reimer, P. J., Austin, W. E., Bard, E., Bayliss, A., Blackwell, P. G., Ramsey, C. B., Butzin, M., Cheng, H., Edwards, R. L., Friedrich, M., et al.: The IntCal20 Northern Hemisphere radiocarbon age calibration curve (0–55 cal kBP), *Radiocarbon*, 62, 725–757, 2020.
- Roy, K. and Peltier, W.: Relative sea level in the Western Mediterranean basin: A regional test of the ICE-7G_NA (VM7) model and a constraint on late Holocene Antarctic deglaciation, *Quaternary Science Reviews*, 183, 76–87, 2018.
- Saltzman, B.: *Dynamical paleoclimatology: generalized theory of global climate change*, vol. 80, Elsevier, 2001.
- Sambridge, M.: Reconstructing time series and their uncertainty from observations with universal noise, *Journal of Geophysical Research: Solid Earth*, 121, 4990–5012, 2016.
- Shackleton, S., Menking, J. A., Brook, E., Buizert, C., Dyonisius, M. N., Petrenko, V. V., Baggenstos, D., and Severinghaus, J. P.: Evolution of mean ocean temperature in Marine Isotope Stage 4, *Climate of the Past*, 17, 2273–2289, 2021.
- Shakun, J. D. and Carlson, A. E.: A global perspective on Last Glacial Maximum to Holocene climate change, *Quaternary Science Reviews*, 29, 1801–1816, 2010.
- Shennan, I.: Handbook of sea-level research: framing research questions, *Handbook of sea-level research*, pp. 3–25, 2015.
- Simpson, G. L.: Modelling palaeoecological time series using generalised additive models, *Frontiers in Ecology and Evolution*, 6, 149, 2018.
- Small, D., Bentley, M. J., Jones, R. S., Pittard, M. L., and Whitehouse, P. L.: Antarctic ice sheet palaeo-thinning rates from vertical transects of cosmogenic exposure ages, *Quaternary Science Reviews*, 206, 65–80, 2019.
- Stein, M. L.: Nonstationary spatial covariance functions, Unpublished technical report, 2005a.
- Stein, M. L.: Space-time covariance functions, *Journal of the American Statistical Association*, 100, 310–321, 2005b.

- Stein, M. L.: Interpolation of spatial data: some theory for kriging, Springer Science & Business Media, 2012.
- 955 Tan, F., Khan, N. S., Li, T., Meltzner, A. J., Majewski, J., Chan, N., Chutcharavan, P. M., Cahill, N., Vacchi, M., Peng, D., et al.: Holocene relative sea-level histories of far-field islands in the mid-Pacific, *Quaternary Science Reviews*, 310, 107995, 2023.
- Tang, Y., Zheng, Z., Huang, K., Chen, C., Chen, Z., Lu, H., Wu, W., Lin, X., Zhang, X., and Li, H.: Holocene Evolution of the Pearl River Delta: Mapping Integral Isobaths and Delta Progradation, *Journal of Marine Science and Engineering*, 11, 1986, 2023.
- Tarasov, L., Dyke, A. S., Neal, R. M., and Peltier, W. R.: A data-calibrated distribution of deglacial chronologies for the North American ice complex from glaciological modeling, *Earth and Planetary Science Letters*, 315, 30–40, 2012.
- 960 Tingley, M. P. and Huybers, P.: Recent temperature extremes at high northern latitudes unprecedented in the past 600 years, *Nature*, 496, 201–205, 2013.
- Tingley, M. P., Craigmile, P. F., Haran, M., Li, B., Mannshardt, E., and Rajaratnam, B.: Piecing together the past: statistical insights into paleoclimatic reconstructions, *Quaternary Science Reviews*, 35, 1–22, 2012.
- 965 Turner, F. E., Buck, C. E., Jones, J. M., Sime, L. C., Vallet, I. M., and Wilkinson, R. D.: Reconstructing the Antarctic ice-sheet shape at the Last Glacial Maximum using ice-core data, *Journal of the Royal Statistical Society Series C: Applied Statistics*, 72, 1493–1511, 2023.
- Upton, M., Parnell, A., and Cahill, N.: reslr: An R package for relative sea level modelling, *arXiv preprint arXiv:2306.10847*, 2023.
- Vacchi, M., Joyse, K. M., Kopp, R. E., Marriner, N., Kaniewski, D., and Rovere, A.: Climate pacing of millennial sea-level change variability in the central and western Mediterranean, *Nature communications*, 12, 4013, 2021.
- 970 Virtanen, P., Gommers, R., Oliphant, T. E., Haberland, M., Reddy, T., Cournapeau, D., Burovski, E., Peterson, P., Weckesser, W., Bright, J., van der Walt, S. J., Brett, M., Wilson, J., Millman, K. J., Mayorov, N., Nelson, A. R. J., Jones, E., Kern, R., Larson, E., Carey, C. J., Polat, İ., Feng, Y., Moore, E. W., VanderPlas, J., Laxalde, D., Perktold, J., Cimrman, R., Henriksen, I., Quintero, E. A., Harris, C. R., Archibald, A. M., Ribeiro, A. H., Pedregosa, F., van Mulbregt, P., and SciPy 1.0 Contributors: SciPy 1.0: Fundamental Algorithms for Scientific Computing in Python, *Nature Methods*, 17, 261–272, <https://doi.org/10.1038/s41592-019-0686-2>, 2020.
- 975 Walker, J. S., Kopp, R. E., Shaw, T. A., Cahill, N., Khan, N. S., Barber, D. C., Ashe, E. L., Brain, M. J., Clear, J. L., Corbett, D. R., et al.: Common Era sea-level budgets along the US Atlantic coast, *Nature Communications*, 12, 1841, 2021.
- Walker, J. S., Kopp, R. E., Little, C. M., and Horton, B. P.: Timing of emergence of modern rates of sea-level rise by 1863, *Nature Communications*, 13, 966, 2022.
- Walter, R. M., Sayani, H. R., Felis, T., Cobb, K. M., Abram, N. J., Arzey, A. K., Atwood, A. R., Brenner, L. D., Dassié, É. P., DeLong, K. L., et al.: The CoralHydro2k Database: a global, actively curated compilation of coral $\delta^{18}\text{O}$ and Sr/Ca proxy records of tropical ocean hydrology and temperature for the Common Era, *Earth System Science Data Discussions*, 2022, 1–56, 2022.
- 980 Wang, K., Pleiss, G., Gardner, J., Tyree, S., Weinberger, K. Q., and Wilson, A. G.: Exact Gaussian processes on a million data points, *Advances in neural information processing systems*, 32, 2019.
- Wang, Z., Saito, Y., Zhan, Q., Nian, X., Pan, D., Wang, L., Chen, T., Xie, J., Li, X., and Jiang, X.: Three-dimensional evolution of the Yangtze River mouth, China during the Holocene: impacts of sea level, climate and human activity, *Earth-Science Reviews*, 185, 938–955, 2018.
- 985 Whitehouse, P. L.: Glacial isostatic adjustment modelling: historical perspectives, recent advances, and future directions, *Earth surface dynamics*, 6, 401–429, 2018.
- Wilks, S. S.: The large-sample distribution of the likelihood ratio for testing composite hypotheses, *The annals of mathematical statistics*, 9, 60–62, 1938.
- 990 Wingate, D. and Weber, T.: Automated variational inference in probabilistic programming, *arXiv preprint arXiv:1301.1299*, 2013.

- Wright, A. J., Edwards, R. J., van de Plassche, O., Blaauw, M., Parnell, A. C., van der Borg, K., de Jong, A. F., Roe, H. M., Selby, K., and Black, S.: Reconstructing the accumulation history of a saltmarsh sediment core: Which age-depth model is best?, *Quaternary Geochronology*, 39, 35–67, 2017.
- 995 Xu, K., Zhang, M., Li, J., Du, S. S., Kawarabayashi, K.-i., and Jegelka, S.: How neural networks extrapolate: From feedforward to graph neural networks, *arXiv preprint arXiv:2009.11848*, 2020.
- Yokoyama, Y., Esat, T. M., Thompson, W. G., Thomas, A. L., Webster, J. M., Miyairi, Y., Sawada, C., Aze, T., Matsuzaki, H., Okuno, J., et al.: Rapid glaciation and a two-step sea level plunge into the Last Glacial Maximum, *Nature*, 559, 603–607, 2018.

Appendix A: Additional model information

Table A1. Summary of Model Characteristics. The posterior is reported with a mean value with 90% credible interval. [GBR = Great Barrier Reef](#), NJ = New Jersey, NC = North Carolina, EIV = errors in variable, NI = noisy input.

Task	Analysis Choice	Data Level	Process Level	Parameter Level <u>Prior for Parameters</u>	Posterior
GBR coral time series	Fully Bayesian; EIV	Uniform likelihood with additional white noise	Temporally linear	$\alpha \sim U(-30, 30)$ m	-6.3 (-28.3, 16.2)
				$\beta \sim U(-10, 10)$ mm/yr	5.9 (4.5, 7.4)
				$\omega \sim U(0.0001, 10)$ m	0.8 (0.1, 2.0)
GBR coral time series	Fully Bayesian; EIV	Normal likelihood with additional white noise	Temporally linear	$\alpha \sim U(-30, 30)$ m	-0.2 (-22.8, 21.9)
				$\beta \sim U(-10, 10)$ mm/yr	6.3 (4.8, 7.7)
				$\omega \sim U(0.0001, 10)$ m	1.4 (0.01, 1.97)
NJ salt marsh time series	Variational Bayesian; EIV	Normal likelihood	Temporally linear	$\alpha \sim U(-5, 5)$ m	-3.38 (-3.47, -3.30)
				$\beta \sim U(-10, 10)$ mm/yr	1.63 (1.57, 1.69)
NJ salt marsh time series	Variational Bayesian; EIV	Normal likelihood	Change-point model	$\alpha_1 \sim U(-15, 0)$ m	-4.92 (-5.01, -4.84)
				$\beta_1 \sim U(-10, 10)$ mm/yr	1.6 (1.5, 1.8)
				$\beta_2 \sim U(-10, 10)$ mm/yr	1.53 (1.46, 1.60)
				$\beta_3 \sim U(-10, 10)$ mm/yr	1.31 (1.22, 1.39)
				$\beta_4 \sim U(-10, 10)$ mm/yr	3.92 (3.73, 4.10)
				$\gamma_2 \sim U(-476, 1020)$ CE	-42.0 (-139.6, 58.2)
				$\gamma_3 \sim U(23, 1518)$ CE	1004.3 (893.1, 1106.0)
NJ salt marsh time series	Empirical Bayesian; NI	Normal likelihood	Gaussian Process with one RBF kernel	$\ell \sim U(1, 5000)$ yr	1038
				$\sigma \sim U(1, 22.4)$ m	20.79
NC salt marsh time series	Empirical Bayesian; NI	Normal likelihood	Gaussian Process with one RBF kernel	$\ell \sim U(1, 10000)$ yr	2175
				$\sigma \sim U(1, 100)$ m	1.75
NC salt marsh time series	Fully Bayesian; EIV	Normal likelihood	Gaussian Process with one RBF kernel	$\ell \sim U(1, 10000)$ yr	6889 (2845, 9766)
				$\sigma \sim U(1, 100)$ m	12.5 (3.4, 28.1)
NC salt marsh time series	Variational Bayesian; NI	Normal likelihood	Gaussian Process with one RBF kernel	$\ell \sim U(1, 10000)$ yr	7305 (3559, 9509)
				$\sigma \sim U(1, 100)$ m	10.4 (3.8, 18.4)

Task	Analysis Choice	Data Level	Process Level	Parameter Level	Posterior
US Atlantic spatio-temporal analysis	Empirical Bayesian; NI	Normal likelihood with additional white noise	Gaussian Process with a zero mean function and multiple isotropic kernels	$\omega \sim U(0.01, 10)$ m	0.02
				$\ell_g \sim U(100, 20000)$ yr	11567
				$\sigma_g \sim U(0.01, 33.3)$ m	30.7
				$\sigma_r \sim U(0.2, 10)$ m	1.7
				$\ell_{r,x} \sim U(319, 1593)$ km	345
				$\ell_{r,t} \sim U(500, 5000)$ yr	3254
				$\sigma_l \sim U(0.1, 3.3)$ m	0.14
				$\ell_{l,x} \sim U(64, 319)$ km	317.4
US Atlantic spatio-temporal analysis	Empirical Bayesian; NI	Normal likelihood with additional white noise	Gaussian Process with ICE_7G as mean function and multiple isotropic kernels	$\ell_{l,t} \sim U(100, 2000)$ yr	1978
				$\omega \sim U(0.01, 10)$ m	0.02
				$\sigma_r \sim U(0.2, 10)$ m	6.0 <u>6.1</u>
				$\ell_{r,x} \sim U(319, 1593)$ km	1586
				$\ell_{r,t} \sim U(500, 5000)$ yr	4684 <u>4683</u>
				$\sigma_l \sim U(0.1, 3.3)$ m	0.13 <u>0.12</u>
				$\ell_{l,x} \sim U(64, 319)$ km	312
				$\ell_{l,t} \sim U(100, 2000)$ yr	1990 <u>1970</u>
US Atlantic spatio-temporal analysis	Empirical Bayesian; NI	Normal likelihood with additional white noise	Gaussian Process with zero mean and a sampling kernel determined by a GIA model ensemble	$\omega \sim U(0.01, 10)$ m	0.3 <u>0.4</u>
US Atlantic spatio-temporal analysis	Variational Bayesian; NI	Normal likelihood with additional white noise	A GIA model ensemble (consist of 9 individual models)	$\omega \sim U(0.01, 0.5)$	0.498 (0.497, -0.498) <u>0.042</u> <u>(0.028, 0.056)</u>
				$\nu_1 \sim Beta(0.22, 0.78)$	0.04 (0.03, 0.05)
				$\nu_2 \sim Beta(0.02, 0.98)$	0. (0., 0.)
				$\nu_3 \sim Beta(0.02, 0.98)$	0. (0., 0.)
				$\nu_4 \sim Beta(0.09, 0.91)$	0. (0., 0.)
				$\nu_5 \sim Beta(0.01, 0.99)$	0. (0., 0.)
				$\nu_6 \sim Beta(0.02, 0.98)$	0. (0., 0.)
				$\nu_7 \sim Beta(0.16, 0.84)$	0.38 (0.35, 0.39) <u>0.36</u> (0.42)
				$\nu_8 \sim Beta(0.29, 0.71)$	0.36 (0.33, 0.38) <u>0.35</u> <u>(0.32, 0.37)</u>
				$\nu_9 \sim Beta(0.17, 0.83)$	0.22 (0.20, 0.24)

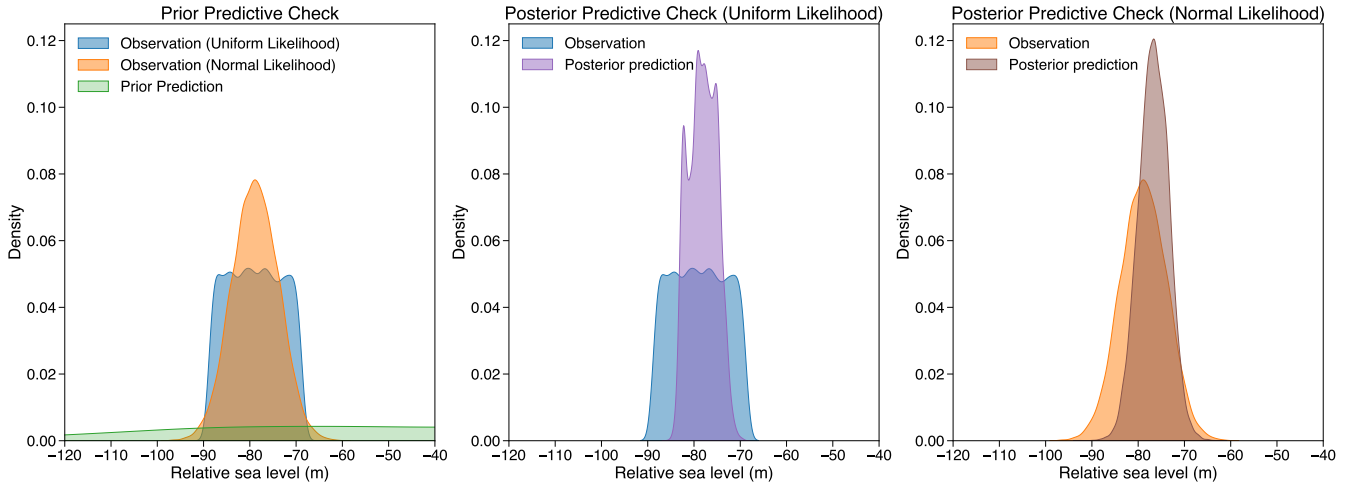


Figure A1. Model $\hat{\mu}$ prediction on -5500-CE RSL at Prior and Posterior Predictive Checks for modeling results presented in the US Atlantic coast main text, section 4.1.1. (a) Mean relative sea level prediction. A random data point was selected for illustrative purposes. Left, which is a combination of mean function Prior predictions compared with observational data, assuming uniform (blue) and covariance function normal (orange) likelihood functions. (b) Relative sea level prediction by Gaussian Process mean function (described by a glacial isostatic adjustment model). Middle, Posterior predictions compared with ICE_7G ice model and VM5a Earth model). (c) Covariance observational data, assuming a uniform likelihood function induced mean relative sea level change. (d) Covariance Right, Posterior predictions compared with observational data, assuming a normal likelihood function induced relative sea level standard deviation.

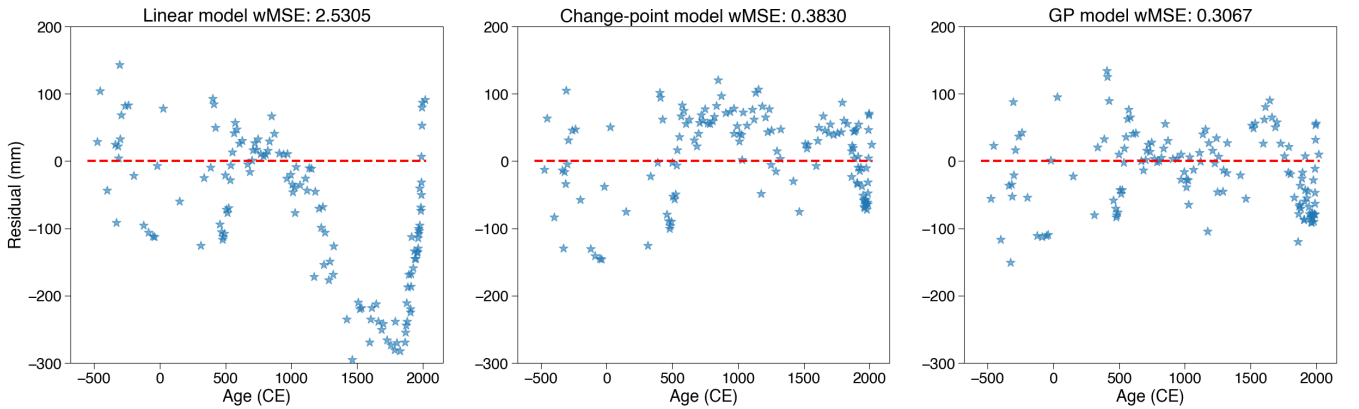


Figure A2. Residual plots for three process level model introduced in main text, section 4.1.2. The weighted mean square error (wMSE) for each model is given as figure titles. Red dashed line represents 0 error.

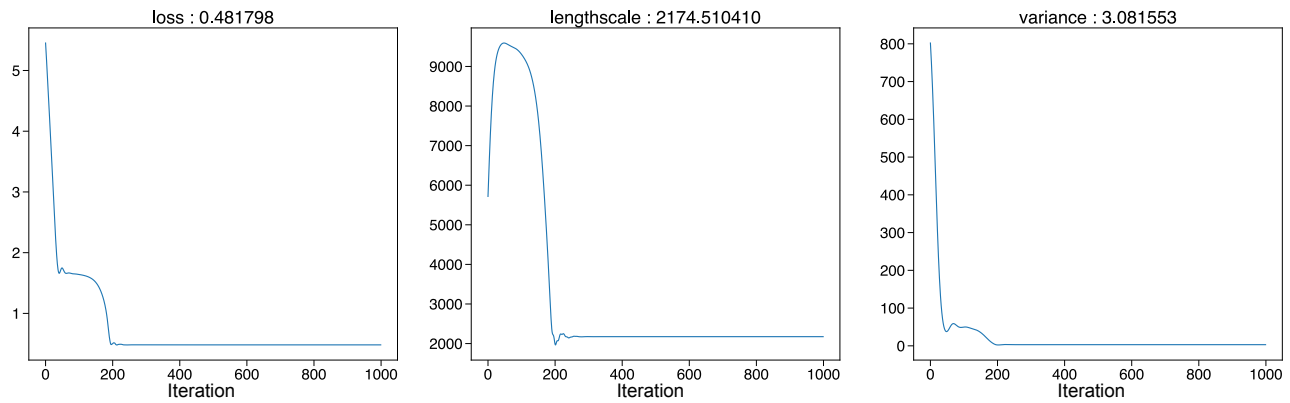


Figure A3. Optimization trace plot of empirical Bayesian analysis introduced in main text, section 4.1.3.

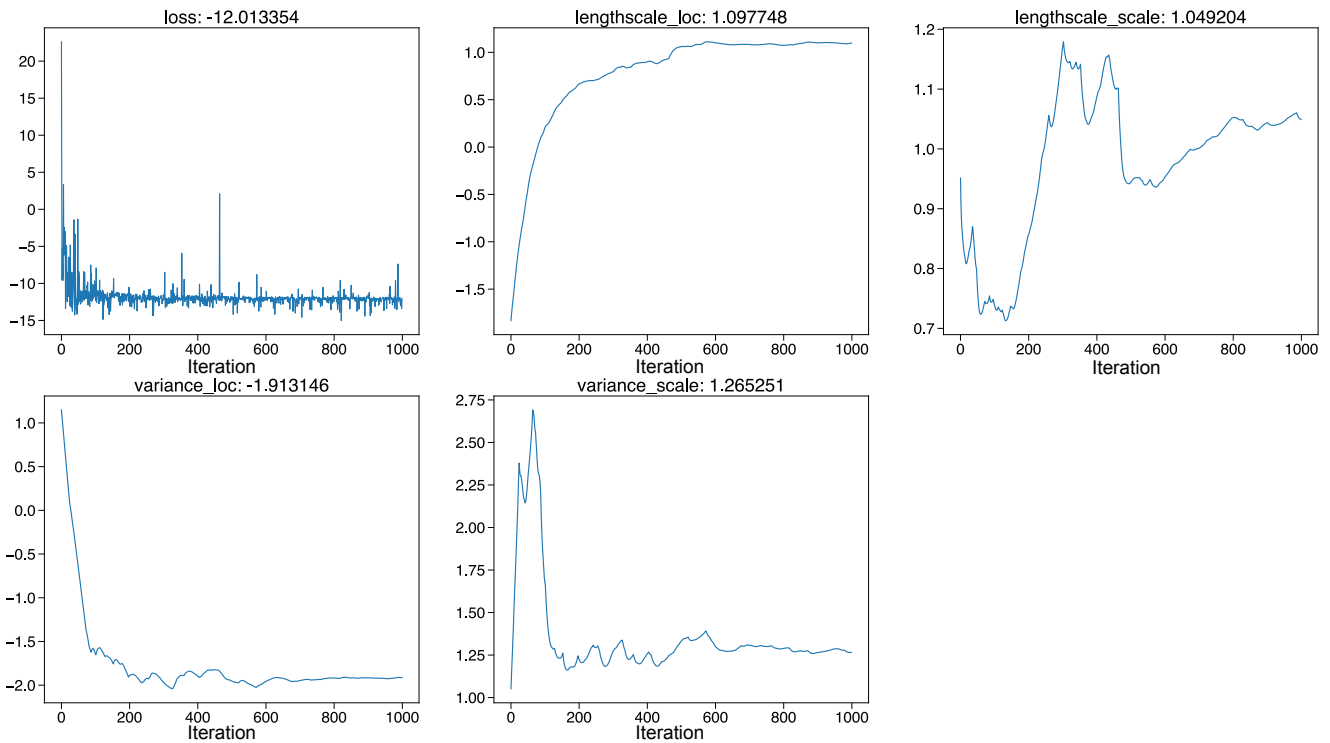


Figure A4. Optimization trace plot of Variational Bayesian analysis introduced in main text, section 4.1.3. Note that the parameters shown here are variational inference parameters in Pyro, which are optimized to approximate the posterior distribution but do not directly correspond to the actual parameters of probability distribution.

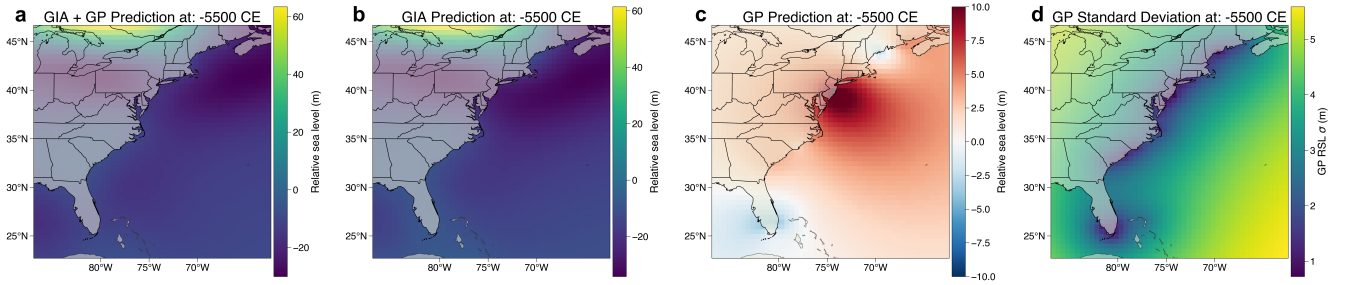


Figure A5. Model prediction from model *ii* for relative sea level at time point -5500 CE along the US Atlantic coast. (a) Mean relative sea-level prediction, representing the sum of components (b) and (c). (b) Relative sea-level prediction derived from the Gaussian Process mean function, based on a glacial isostatic adjustment model incorporating the ICE 7G ice model and VM5a Earth model (Roy and Peltier, 2018; Peltier et al., 2015). (c) Relative sea-level change induced by the GP covariance function. (d) Standard deviation of relative sea level, as estimated by the GP covariance function.