

Answer to review #2 of manuscript:
“Could old tide gauges help estimate past atmospheric variability?”
for Climate of the Past

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Reviews in black, answers in blue, quotes from revised manuscript in bold.

This work uses historical tide gauge records in the French coast as proxy data to improve the atmospheric 20th century reanalysis. Specifically, the sea level observations are used to constrain the number of ensemble members in the reanalysis, selecting and weighting them according to their correspondence to observed storm surges. This is an interesting idea. The approach discards model realisations that, although physically consistent with the scarce available mean sea level pressure observations, are inconsistent with observations of storminess from coastal tide gauges. I am not a statistician and therefore I am not qualified to comment on the details in section 4. My comments are focused on the results and, mostly, on sections 2 and 3 on data description and processing. I have found major issues in these sections, including a basic conceptual misunderstanding of the generation of storm surges. Also, there are missing details in data processing and I have concerns on the simple linear regression model that has been used to relate storm surges and sea level pressure anomalies.

We thank Reviewer #2 for the fruitful comments, corrections, remarks and suggestions. We believe that the review greatly improved the clarity and precision of the paper. More details are given below.

In the revised manuscript, we have chosen to do some major modifications, driven by the comments of both reviewers.

- i Comments of reviewer # 2 on the physical mechanisms driving storm surge generation have been taken into account and important modifications were made both in the manuscript’s text and in the data pre-processing. The effect of mean-sea-level is now removed using yearly medians. A reference pressure averaged over the ocean is used to estimate the statistical relationship between sea-level and local pressure.
- ii The linear model was modified for a local-linear model, details of which are given in section 3.1 of the revised manuscript. However, we want to stress here that the change in the model only mildly affected the result of our algorithm. However, we believe that the revised manuscript makes a clearer interpretation of the results and of the model’s advantages and limitations, thanks to the added value of the reviewer’s comments suggestions.

Another major concern is that the authors state in the conclusions that they use the tide gauge time series as a barometer record not assimilated in the reanalysis. However, they model the storm surges from sea level accounting for both pressure and geostrophic winds. I do not understand the interest of including the winds here (which by the way are not modelled correctly, in my opinion). If the purpose is to incorporate only pressure-like measurements, then pressure gradients should not be included at all. Or at least, the pressure-like record should be reconstructed without accounting for wind effects. All these issues are detailed in the following. My recommendation is that the manuscript should undergo major revisions before being considered for publication.

Accounting for this comment, we have chosen to remove the part of the statistical model related to pressure gradients. In our modified statistical model, we estimate directly the probability distribution of pressure based on the tide gauge record. More details are given in the article and copied below.

Other major modifications to the manuscript have been made, which are listed here for clarity purposes:

- i We have removed the use of the Saint-Nazaire tide gauge, and kept only the Brest tide gauge record. In the way our statistical model was built, adding the Saint-Nazaire tide gauge observation was not making a significant change to the results, as in most cases the response of the Saint-Nazaire tide gauge to variations of pressure was nearly identical to the one of Brest. Therefore, adding it only complicated the paper’s statement. Only a more sophisticated model would have been able to make smart use of the slight differences between both sea-level responses, and this is out of the scope of this article.

- ii We have used independent pressure records for the city of Brest, coming from the EMULATE project and from recently released Météo France archive data. This motivated us further both to keep only the Brest tide-gauge and not to incorporate pressure-gradients in the statistical model. This new set of observations is described in section 2.5 of the revised manuscript. It greatly helps to identify the potential and the limitations of our method, as explained below and in section 5 of the revised manuscript.

Section 2: This section needs further details in the description of the data set. Preprocessing of data needs improvement.

- 1. L.62: define the spatial and temporal resolution of the reanalysis.

- 1. We have added the following sentences [line 71](#):

It has a temporal resolution of 3 hours, and uses a spectral triangular model in space with truncation of T254 (approximately 75km at the equator). There are 64 vertical levels, up to .3mb.

- 2. L. 70: I assume the variable used is mean sea level pressure (MSLP), correct?

2. Correct. In section 2.2 of the revised article called “Preprocessing of mean-sea-level pressure”, at the beginning we state [90](#):

In this work, we are using only the mean sea-level pressure (MSLP) variable from 20CRv3.

This whole new section is copied below in answer to comment 4.

- 3. L. 81-82: please, remove.

3. We have removed the following sentence: “This double constraint of data scarcity and high variations in sampling frequency in the 19th century legitimizes the search for other sources of observations to study centuries-old atmospheric variability.”

4. L. 83: anomalies of what? I assume this means that mean SLP anomalies are computed with respect to climatologies for 2 periods. Please, specify.

4. This has changed slightly (anomalies are not considered for the satellite-era period anymore). This is specified more precisely in the new subsection 2.2 called “Preprocessing of mean-sea-level pressure”, which states [line 90](#):

In this work, we are using only the mean sea-level pressure (MSLP) variable from 20CRv3. We make two different preprocessings of this variable.

A first preprocessing is used for the statistical relationship between the local pressure and the surge. As the latter is driven by a physical phenomenon called the “inverse barometer effect” which will be introduced in the next section, we consider the difference between the MSLP interpolated at the city of Brest (4.49504°W, 48.3829°N), and the MSLP averaged over all members of 20CRv3 and over the North-Atlantic ocean (using the reanalysis’ land mask and averaging from 98°W to 12°E and from 0°N to 69°N), similarly to Ponte (1994). This spatial-averaged pressure is noted $\overline{\text{MSLP}}^{\text{ocean}}(t)$ and depends only on time.

A second preprocessing of MSLP is used to compute the probability of transition from one member of the reanalysis to another in the Hidden Markov Model (HMM) presented in section 3.1. For this purpose, we consider seasonal anomalies of MSLP with respect to a climatology computed from the period 1847-1890, because the HMM is run only for those years. The reference MSLP climatology for calendar day d and hour h is given by the average over days between $d - 30$ and $d + 30$, hours between $h - 3$ and $h + 3$, and all years 1847-1890. This reference MSLP is noted $\overline{\text{MSLP}}^{\text{clim}}$ and depends on latitude and longitude.

5. Description of tide gauges (section 2.2): please, define the temporal sampling of tide gauges, and their overlapping periods. The statement that one time series can be used to fill in gaps of the another is incorrect. They are highly correlated at low frequencies but not necessarily at high frequencies (i.e. storms). Actually, they are not even in the same location, as stated later in the text, so differences in storm surges may be expected.

5. The overlapping periods need not be specified anymore since the Saint-Nazaire tide gauge is not used anymore. The temporal sampling was specified in the following sentence [line 106](#):

The Brest sea-level record from this database starts in 1846 and has a hourly sampling.

Indeed, it is incorrect to state that one series can be used to fill the gaps of the other. This is only partially true for 12h-averages of sea-level, of which the tide was removed as well as the yearly median. Anyway, we do not use the Saint-Nazaire tide gauge anymore, so the statement was removed in the revised article.

- 6. L. 97: actual → modern, current, present-day?

- 6. Indeed. The corrected sentence is now [line 108](#):

This combination of historical and modern records is at the foundation of the methodology exposed in the next section.

7. The removal of MSL effects in the tide gauge processing is incorrect. The impact of MSL should not be removed in this way, because there are other MSL variations at interannual and decadal time scales that would

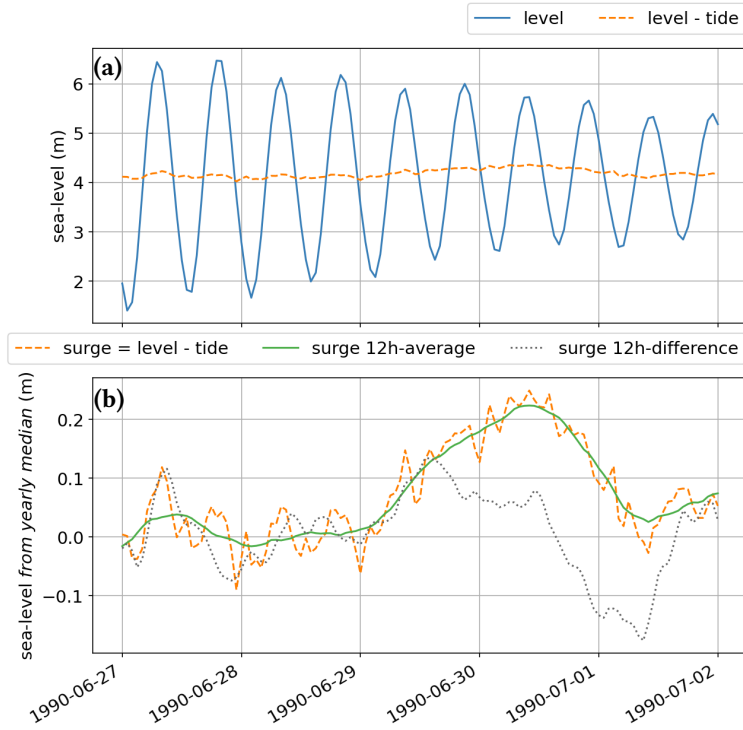


Figure 3 of revised manuscript, reproduced here. An example output of the different stages of preprocessing of the sea-level signal used in this work. (a) Raw level before (full, blue line) and after (dashed, orange line) removing the tidal part of the signal. (b) Sea-level after removing the yearly median value: the surge $h(t)$ (1h sampling, orange dashed curve), the centered 12h-average of the surge $\bar{h}^{12h}(t)$ (green full curve), and the 12h difference between 3h-averages of the surge $\Delta\bar{h}^{3h}(t)$ (gray dotted curve).

remain. A better approach is to remove mean (or median) yearly averages. It also avoids the problem of selecting an arbitrary change point to calculate different linear trends.

7. This was changed in the revised manuscript. We are now removing the yearly median after having removed the tidal part of the signal, as detailed in section 2.4 of the revised manuscript. We copy-paste this section below (starting [line 111](#)):

As mentioned earlier, the part of the sea-level which responds to atmospheric processes is the surge (also called “storm surge” or “skew surge”). To access the surge, one first has to remove the tidal part of the signal, and then to remove yearly variations of the mean-sea-level (at interannual and decadal scale), such as sea-level rise (Cazenave and Llovel, 2010). In this work, we are also interested in moving averages and differences of the surge. All these steps are exemplified in Fig. 3.

We first compute the tidal constituents of the raw sea-level (blue curve, Fig. 3.a) using U-Tide (Codiga, 2011), which performs harmonic (Fourier) decomposition with prescribed frequencies corresponding to planetary movements. The tidal constituents are computed over two different periods, one is 1847-1890, and the second is 1981-2015. Removing the tidal part of the signal gives the orange dashed line of Fig. 3.a, which has a temporal average value of $\sim 4\text{m}$ for the Brest tide gauge.

Then, we remove the yearly median value of the sea-level, which allows to access the surge (orange dashed line of Fig. 3.b). We choose to remove the median and not the mean because the mean can in principle be influenced by the number and magnitude of extremes in a given year, which can be linked to the number and magnitude of storms passing in a given year. This second step allows to access the surge which is noted $h(t)$ in the following:

$$h(t) = H(t) - \text{Tide}_H(t) - \text{median}[H(t'), t' \in \text{year}(t)] ,$$

where $H(t)$ denotes the raw sea-level, $\text{Tide}_H(t)$ is the tidal part of the signal computed from H , and $\text{year}(t)$ is the year in which time t is found.

Note from Fig. 3.b that the surge fluctuates at hourly scale, part of which are oscillations which are not due to variations in atmospheric pressure. These oscillations are either due to tide-surge interactions (Horsburgh and Wilson, 2007) or to measurement errors in the 19th century leading

to phase shifts. Such oscillations can dominate the surge signal in Brest where the tidal amplitude is large. Furthermore, tide-surge interactions lead to stronger surges in low-tide and weaker surges in high-tide (Horsburgh and Wilson, 2007). As these phenomena are not linked to atmospheric processes, we chose to filter them out with a simple 12h-average (green full curve in Fig. 3). This also implies that these 12-hours-averaged surges will only respond to atmospheric events persisting for more than 12 hours. Given the spatial resolution of 20CRv3, smaller-scale events are likely not to be represented in the MSLP fields used in this study. In the following, we note $\bar{h}^{12h}(t)$ the 12h-average of the surge:

$$\bar{h}^{12h}(t) = \frac{1}{12} \sum_{t'=-6}^{t'=+6} h(t+t').$$

Furthermore, as we are using sea-level observations to estimate atmospheric pressure, we also want to measure the amplitude of local time-variations of the surge. Indeed, as will be further explained in section 3.1, the sea-level response to variations of pressure depends on the time-scale of these variations. More precisely, the “inverse barometer” is an approximation that is only valid for slow variations of pressure. Accordingly, when observing fast variations of the surge, one expects deviations from the inverse barometer approximation. We therefore compute the difference between the surge at time t and at time $t - 12h$, choosing the 12h-interval again to filter out oscillations at a period close to 12h. Furthermore, since the reanalysis is run at 3h-resolution, we perform a 3h-moving average of the surge before computing the difference. This difference is noted $\Delta\bar{h}^{3h}(t)$ and defined by the following equation:

$$\Delta\bar{h}^{3h}(t) := \frac{1}{3} \sum_{t'=-2}^{t'=+1} [h(t+t') - h(t-12+t')].$$

8. L. 108-109 on detiding: which is the period used for detiding?

8. We specify this in the revised manuscript, see above and section 2.4 of the revised manuscript. We recall here [line 116](#):

The tidal constituents are computed over two different periods, one is 1847-1890, and the second is 1981-2015.

The first period is used to remove the tidal signal in 1847-1890 while the second is used to remove the tidal signal in 1981-2015.

9. L. 113: phenomenons→ phenomena

9. This was corrected in the revised manuscript.

10. L. 110: A 12-h filtering removes part of the storm surge variability, especially that related to storms.

10. Indeed. We have not found an other simple method that allows to suppress the oscillations unrelated to atmospheric variations from the tide gauge signal.

11. L. 114-115: smaller scale events are unlikely to be recorded in 20cR at 2deg resolution

11. This is true and we should have noted it from the start. This argument is now included in the revised manuscript as ([line 132](#)):

Given the spatial resolution of 20CRv3, smaller-scale events are likely not to be represented in the MSLP fields used in this study.

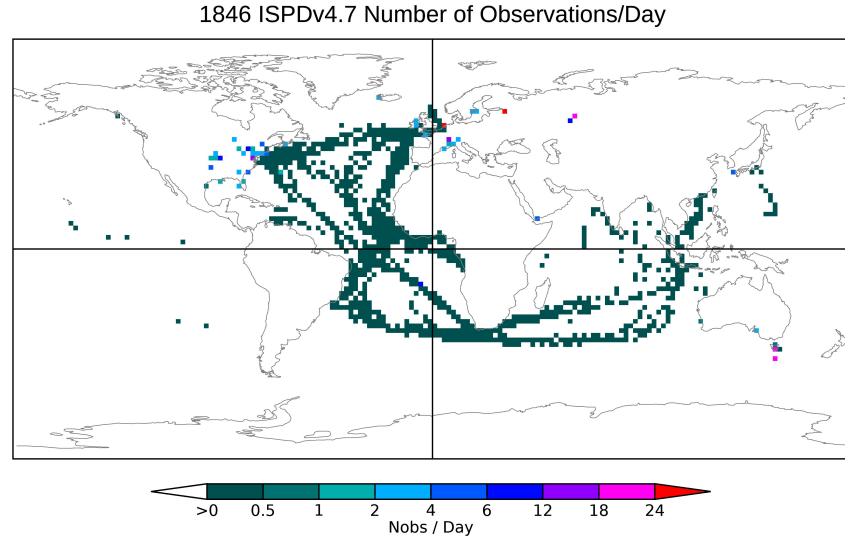
Section 3: This section contains fundamental errors in the understanding of the processes that generate storm surges. Also, the linear regression model is questionable as applied here, and should be better justified. More details:

As stated above, we have made several modifications to the MSLP pretreatment and to the statistical model, including an ocean-averaged reference pressure, discarding pressure gradients, and using a local-linear model. More details are given below.

12. L. 123-124: hydrostratic equilibrium is a general law not only affecting this region. Inverted barometer has been widely applied everywhere.

12. We believe that the new formulation does not imply that the inverse barometer would be an effect only valid in Brest. See [line 162](#):

The filtered surges described in section 2.4 respond to sub-seasonal variations in atmospheric pressure. First, the sea-level is sensitive to pressure variations. An approximation called the “inverse barometer effect” (Roden and Rossby, 1999) states that an increase (respectively, decrease) of 1hPa in pressure at the mean sea-level leads to a decrease (respectively, increase) in sea-level of approximately 1cm. This approximation is valid for slow variations of atmospheric pressure compared to the typical time of dynamic adjustment of the sea-level (Bertin, 2016).



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Figure A. Not part of revised manuscript. Number of observations per day assimilated in 20CRv3 for year 1846.

13. L. 127-131: this is incorrect. The main mechanism of wind as storm surge generation is not the Ekman transport, but the piling up of waters due to wind blowing perpendicular to the coast. Check your own results, where you can actually see it (e.g. line 163). Please, change this in the abstract too.

13. This was modified in the text, see below. The abstract does not mention directly the effects of wind anymore as we do not use it in our statistical relationship. See [line 167](#):

Moreover, the piling up of waters due to wind blowing perpendicular to the coast is responsible for positive (respectively, negative) surge when the wind stress is directed towards (respectively, away from) the coast. This effect depends non-linearly on the wind amplitude (Bryant and Akbar, 2016; Pineau-Guillou et al., 2018). Statistical correlation between pressure variations and wind intensity and direction are responsible for deviations from the inverse barometer approximation (Ponte, 1994).

14. L. 1334-134: I assume that SLP here (and used later in equation 1) is the SLP anomaly with respect to time at the given grid point. If this intend to represent the inverted barometer effect, this is incorrect. The anomaly should be computed with respect to the average mean sea level pressure over the world oceans in the reanalysis. See for example Ponte 1994 (<https://doi.org/10.1029/94JC00217>).

14. Our use of MSLP anomaly with respect to seasonality was incorrect. We now use the average over the North-Atlantic ocean -as in Ponte (1994)- from the reanalysis as a reference pressure. We have also tried to use the average over the world oceans (not only the North-Atlantic), and it only slightly changes the performance of the statistical method, so we decided to keep it as is. Using the North-Atlantic ocean average may be more meaningful in the 19th century when the distribution of observations is not uniform, see Fig. A of this document (this figure does not appear in the revised manuscript and is only here for discussion purposes).

Note that we still use the climatological anomaly to compute the probability of transition from one member to another in the HMM algorithm. This is all described in section “Preprocessing of mean-sea-level pressure”, [line 92](#):

A first preprocessing is used for the statistical relationship between the local pressure and the surge. As the latter is driven by a physical phenomenon called the “inverse barometer effect” which will be introduced in the next section, we consider the difference between the MSLP interpolated at the city of Brest (4.49504°W, 48.3829°N), and the MSLP averaged over all members of 20CRv3 and over the North-Atlantic ocean (using the reanalysis’ land mask and averaging from 98°W to 12°E and from 0°N to 69°N), similarly to Ponte (1994). This spatial-averaged pressure is noted $\overline{\text{MSLP}}^{\text{ocean}}(t)$ and depends only on time.

A second preprocessing of MSLP is used to compute the probability of transition from one member of the reanalysis to another in the Hidden Markov Model (HMM) presented in section 3.2. For this purpose, we consider seasonal anomalies of MSLP with respect to a climatology computed from the period 1847-1890, because the HMM is run only for those years. The reference MSLP climatology

for calendar day d and hour h is given by the average over days between $d - 30$ and $d + 30$, hours between $h - 3$ and $h + 3$, and all years 1847-1890. This reference MSLP is noted $\overline{\text{MSLP}}^{\text{clim}}$ and depends on latitude and longitude.

15. L. 134: interpolated how? Linearly?

15. Yes, we use linearly interpolated MSLP at the city of Brest. This was made clear in modified section 3.1 “Local Linear Regression (LLR) between surges and mean-sea-level pressure” [line 188](#):

The predicted variable is $\text{MSLP}(t) - \overline{\text{MSLP}}^{\text{ocean}}(t)$ where $\text{MSLP}(t)$ is the value of the MSLP linearly interpolated at the city of Brest from the reanalysis.

16. Equation 1: please, define all the terms. I do not believe this is a good way to separate the forcings. I think it would be better to use perpendicular and parallel to the coast MSLP gradients (geostrophic winds). In this case, the parallel will be insignificant and the model would be simplified. This separation has a physical meaning and would also allow to interpret the results easily.

16. This equation was suppressed as we do not use the MSLP gradients anymore in the statistical relationship, as suggested by reviewer # 2.

17. Also, why is the term Cov introduced? How does it change the results?

17. This term was introduced because there are high correlations between the observed surges in Brest and Saint-Nazaire, especially after 12h-filtering. This term is removed now since we use only Brest’s tide gauge.

18. Still in the same equation, are all coefficients significant? what happens if a step-wise regression is used instead? I would expect gamma coefficients to be discarded.

18. We are now using a local-linear relationship where all coefficients are re-estimated for each new value of the regressors, which are now the surge’s 12h-average and 12h-difference. More details on the model are given in section 3.1 “Local Linear Regression (LLR) between surges and mean-sea-level pressure”.

19. Tables 1 and 2: units are missing.

19. These tables were removed in the revised version of the manuscript.

20. L. 170: “1.02 on physical grounds” → actually not, the inverted barometer effect is a simplification and does not work at the coast.

20. This line is not present in the revised version of the manuscript.

21. L. 171-172: “This justifies...” → I do not think this statement is correct. I believe a step-wise regression would tell you if the use of these parameters is justified or not. Here it has not been proved, not even with uncertainties of the parameters.

21. This line is not present in the revised version of the manuscript. The new method is not a global linear-regression, and therefore step-wise regression cannot be used to justify the choice of parameters. We have chosen to justify our model by other means, see below.

22. L. 174: wave setup would be removed with the 12-h averaging, so it does not play any role.

22. Indeed our understanding of the physical mechanism called “wave setup” was not correct. This line is not present in the revised version of the manuscript.

23. Section 3.2: To demonstrate that the same coefficients can be used during another period, this is not the correct approach. I think a better approach would be to use the recent period to generate 19th-century-like observations, i.e., downgrading the number and location of observations, and compute the coefficients of the LR. These downgraded, coarse-resolution coefficients would be directly comparable to the ones in the section above with the full resolution. It would be important to calculate uncertainties in all cases in order to compare the coefficients.

23. The methodology proposed by reviewer #2 to demonstrate that the coefficients can be used in both periods is interesting, and would definitely be the best way to prove that the statistical relationship derived in recent years also holds in previous years. However, if our understanding of what you suggest is correct, it is not feasible for us in practice. Indeed, we neither have access to the numerical weather prediction model, nor to the data assimilation scheme used to generate 20CRv3, and we are not in capacity to re-run the reanalysis with downgraded number of observations. This is beyond our capabilities and out of the scope of our study.

This is the reason why we chose to include independent pressure observations, as stated earlier. This provides an independent data to compare to both 20CRv3 and our statistical estimate of pressure based on surges. Aside from the fact that it is a more robust and convincing justification, it also allows to better understand the limitations of our methodology. More details are given in the manuscript and below.

For these reasons, and because it would not be feasible with our new, local-linear regression, we have chosen not to replicate this section.

24. Fig 5: it is unclear what is shown here. Are coefficients of the modern period used here? Please, specify. The description of the results in fig 5 is confusing and does not reflect, in my opinion, the results. For example, in l. 188 the bias referred to one of the cases is visible in all (5a and 5b); l. 193: which interpretation is referred to here?; l. 196 states that there is no clear sign of bias in individual members but I think the bias is similar to that in the ensemble average. The bias is not only due to ensemble mean but also to limitation of coarse resolution data.

24. This Figure was removed from the final manuscript because the method was changed as stated earlier.

25. L. 206-220: this discussion on the differences in coefficients should consider uncertainties to ensure that the use of two different periods lead to different values. Values slightly larger/smaller do not provide confidence in the results. Therefore, the interpretations are not reliable (e.g. differences attributed to ensemble averaging).

25. Again, this discussion was removed.

26. L. 225-226: I think the consistency of the coefficients and the explained variance in modern and old periods is a consequence of the dominance of the inverted barometer effect. This is probably not true for extreme values generated by strong winds associated to storms, but it holds for mean storm surge variability driven by pressure changes.

26. This comment also motivated us to remove the pressure-gradient part of our algorithm. We believe that our new algorithm, although still mostly imperfect, is able to capture part of these two regimes: the moderate values where the inverse barometer is more valid, and the extreme values where the effect of wind causes the statistical relationship to differ from the inverse barometer. However, our algorithm is still biased towards already observed values of pressure, and therefore has a hard time modelling the largest absolute values of pressure, as explained further in section “Local-linear regression (LLR) between surges and mean-sea-level pressure”. However, we believe that it is still a better model than the previous linear one.

27. Summarising, I believe the LR model should be modified to consider winds parallel and normal to the coast in each case, and use step-wise (or similar) approach to remove the terms that do not explain more variance but introduce noise. Also, uncertainties of the parameters should be calculated. The use of the coefficients calculated in the modern period should be tested in downgraded modern data to prove that they are usable in older periods. The entire comparison and discussion of the two periods should be modified accordingly and simplified in case this is proved.

27. In conclusion, we have chosen to follow the alternative path suggested by reviewer #2, that is to discard the use of MSLP gradients in our model. We do not calculate the uncertainties in the parameter estimation of our new model because the parameters change at each time t . Rather, we chose to evaluate the consistency of the *result* of our algorithm, that is, the value of the predicted average $m(t)$ and variance $\text{var}(t)$ of the distribution of $\text{MSLP}(t) - \overline{\text{MSLP}}^{\text{ocean}}(t)$, as defined in the text. To check the consistency of the result of our algorithm with the ground truth in the period 1980-2015, we compared $m(t)$ estimated from our algorithm with the true value $\text{MSLP}(t) - \overline{\text{MSLP}}^{\text{ocean}}(t)$ in a density-scatter plot (Fig. 4.a of the revised manuscript), and to assess the consistency of our estimate of variance we checked that the variable $\text{MSLP}(t) - \overline{\text{MSLP}}^{\text{ocean}}(t)$ shifted by $m(t)$ and rescaled by $\text{var}(t)^{1/2}$ follows a standard Gaussian distribution using a histogram compared to a standard normal probability distribution function (Fig. 4.b of the revised manuscript). These tests demonstrate the relevance of the model when applied in the period 1980-2015.

To assess the relevance of our algorithm in the period 1848-1890, we can use the independent observations of pressure, after making a yearly-shift to avoid constant differences in pressure as an observational artifact (see section 2.5 of the revised manuscript). These pressure observations are compared to our estimate using surge observations, and to the estimate based on the reanalysis. Several examples are shown in section 5. (figures 11, 12, 13, 14), with one example (figure 13) where we also show the time series of 10m-wind as estimated from 20CRv3, because it helps interpreting the erroneous estimation of pressure from our LLR algorithm for a specific storm. The main conclusion of section 5. is that wind is the key driver of uncertainties in our LLR-based estimation of pressure (one value of surge cannot be unambiguously attributed to either pressure or wind, at least in the way we modelled it). We give some more details here that answer part of reviewer #2's questions.

In Fig. B of the present document, we show the statistical relationship between the daily 10m-wind direction as estimated from the reanalysis versus the daily value of surge. Projecting the wind component along the direction $\pi/4$ shows a non-trivial relationship between wind and surges, as shown in Fig C of the present document. These figures are here for discussion purposes and are not part of the revised manuscript.

In the revised manuscript, Fig. 13 shows that the same low value of MSLP (as estimated both by the 20CRv3 reanalysis and the independent pressure observations for the city of Brest) can lead to different values of surge, and we interpret this as the result of a decrease in the amplitude of wind from an exceptionally high value ($\approx 15\text{m/s}$) to a moderate value. When the wind intensity is exceptionally high, the estimate of pressure from the surge-based LLR is biased, while it is not when the wind intensity decreases to more moderate, typical values. We believe that this is the main limitation of our study. However, the study can still be useful if improved, and our combination of HMM and statistical relationship could be used to assess other reanalysis products from other independent observations.

Section 4: I am not an expert in HMM and I am not qualified to comment on the details of the method explained here. It would be good that an expert statistician reviews this part. I nevertheless would recommend defining all variables in the equations, as well as the acronyms (e.g. NWP), as it makes it very difficult to follow as it is now. Other comments follow:

28. L. 270: this is a rather large area. Are the results sensitive to this choice?

28. Yes, the results are, in principle, sensitive to the choice of this area, as it is used to compute the probability of transition from one member to another. However, as MSLP fields are highly space-correlated, the choice of the

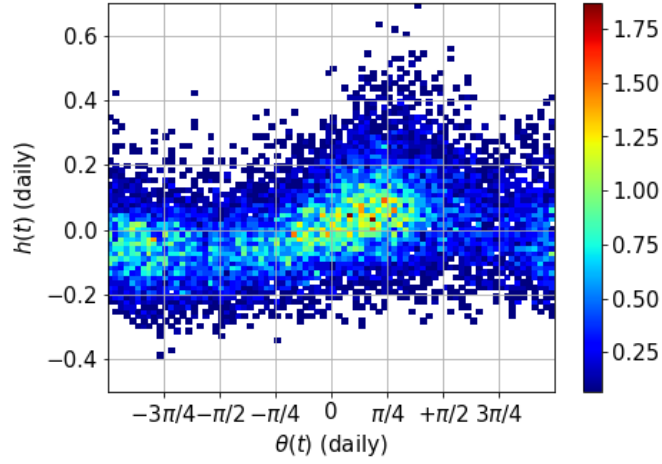


Figure B. *Not part of revised manuscript.* The effect of wind direction on the daily-average of surge, using daily 10m-wind data linearly interpolated at the city of Brest from the ensemble mean of 20CRv3, for the period 1981-2015. This shows a preferred direction at $\sim \pi/4$ for positive surges and opposite direction $\sim -3\pi/4$ for negative surges.

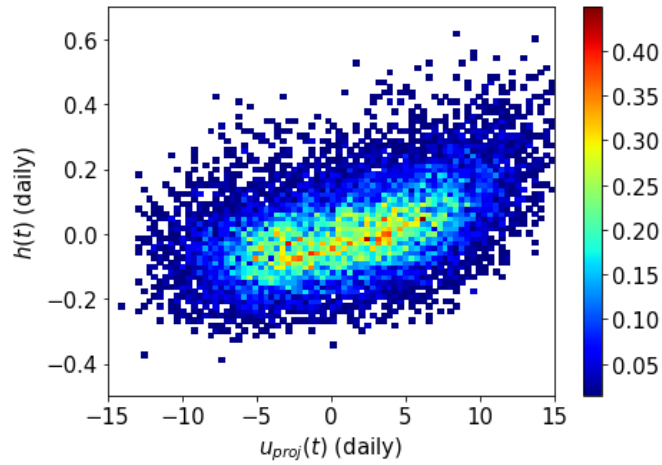


Figure C. *Not part of revised manuscript.* Same as Fig. B but for projection of wind on the direction $\pi/4$.

extent of the area is of lesser influence than the choice of the parameter θ , which we set to a value which allows to maximize the likelihood of the reanalysis based on the surge observations. In particular, high values of θ correspond to high probabilities of transition from one member to another, and low values of θ correspond to low probabilities of transition. Whatever the choice of area, in the limit of $\theta \rightarrow 0$ transition from one member to another are impossible, and in the limit $\theta \rightarrow +\infty$ all transitions between members are allowed. Therefore, once the area is chosen, the critical parameter is θ .

29. L. 336-337: does this mean that the storm surges constrain the data more in winter when they have a stronger signal?

Yes this is correct.

30. L. 361: then \rightarrow than

The sentence in the revised manuscript is unchanged “Note, as well, that the area of influence is greater for $\delta\mu_{\text{HMM}}$ then for $\delta\mu_{\text{HMM}}$, because of the time-propagation of corrections thanks to the smoothing HMM algorithm.” We do not understand why “then” should be replaced by “that” here.

31. Section 4.3 is very illustrative of the potential of the approach. However, more information would be useful to understand why including a single (or two) new records essentially rules out 79 out of 80 model ensemble members. In particular, how many sea level pressure records were included in this period (or shortly before the storms) in the area? How are they distributed? It would be also useful to see a case with more sea level pressure observations (early or mid- 20th century for example).

Fig. 10 of the revised manuscript shows the distribution in space of the number of observations assimilated in the reanalysis during the months of the studied events.

We have included a case in year 1888 where the reanalysis is more constrained by pressure observations. The four time series in Fig. 9 clearly show how the reanalysis is more and more constrained over time and therefore has reduced uncertainties. The case of Fig. 14 (in year 1888) also shows higher values of the effective number of members compared to the other cases.

Conclusions section:

32. L. 405-408: I understood from the LR model that winds were incorporated into the model as SLP gradients. Then, why if wind is an important driver of storm surges may limit the use of the tide gauge? In fact, if wind is not accounted for and wants to be removed, then the LR model could include only the inverted barometer effect with the adjusted parameter at each tide gauge location.

32.The proposed approach is the one we have chosen in the revised manuscript. The revised conclusion now reads [line 405](#):

This study is a proof of concept for the use of century-old tide gauge data as a means of understanding past atmospheric subseasonal variability. Surges of Brest allow to assess part of the atmospheric variability that was uncaught in global 20CR reanalyses based on pressure observations. Weighing 20CR members according to surge observations reduces the effective ensemble size, and implies significant deviations in members-averaged sea-level pressure in the Bay of Biscay. Through the second-half of the 19th century, these deviations diminish and the effective ensemble size rises, however they remain non-negligible. Independent pressure observations in the city of Brest are coherent with pressure estimations from the reanalysis and the surge-based local-linear relationship. Such comparisons also show that the reconstruction of pressure based on surges is ambiguous due to the influence of winds, so that biases between the surge-base and the reanalysis-based pressure estimates can last for several days.

This work has several potential applications. First, replicating this work with other tide gauges could help to validate reanalyses like 20CRv3 against independent data, and to potentially identify anomalous trends or wrong estimation of specific events. Combining our statistical approach with the physics-based approach of Hawkins et al. (2023) could allow to have both a precise estimate from a high-fidelity coastal model and a good quantification of uncertainties. Second, tide gauges could be used to constrain regional scale atmospheric simulations in order to better estimate the magnitude and spatial extent of known past severe storms. Third, tide gauge records could be combined with direct observations of atmospheric pressure to give statistical estimates of atmospheric fluctuations in the 19th century without the use of a Numerical Weather Prediction model, such as the optimal interpolation of Ansell et al. (2006) based on direct pressure observations only, or the analogue upscaling of Yiou et al. (2014) for the short period 1781-1785 of dense observations in western Europe. Finally, this work could be replicated in a more general context, using other types of variables and observations, learning the relationship between observations and large-scale features using recent observations and precise reanalyses, and applying these statistical relationship in the past to uncover past large-scale events. In particular, the hidden-Markov model algorithm outlined here could be replicated to weigh ensemble members according to independent observations.

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