Evaluation and optimisation of the soil carbon turnover routine in the MONICA model (version 3.3.1)

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Abstract. Simulation models are tools commonly used to predict changes in soil carbon stocks. Prior validation is essential, however, for determining the reliability and applicability of model results. In this study, the process-based biogeochemical model MONICA (Model of Nitrogen and Carbon dynamics on Agro-ecosystems) was evaluated with respect to soil organic
10 carbon (SOC) using long-term monitoring data from 46 German agricultural sites. A revision and parameterisation of equations, encompassing crop and fertiliser-specific C contents and the abiotic factors of soil temperature, soil water and clay content, were undertaken and included in the model. The modified version was also used for a Morris elementary effects screening method, which confirmed the importance of environmental and management factors to the model's performance. The model was then calibrated by means of Bayesian inference using the Metropolis-Hastings algorithm. The performance of
15 the MONICA model was compared with that of five established carbon turnover models (CCB, CENTURY, C-TOOL, ICBM and RothC). The original MONICA model systematically overestimated SOC decomposition rates and produced on average a ~17 % greater mean absolute error (MAE) than the other models. The modification and calibration significantly improved its performance, reducing the MAE by ~30 %. Consequently, MONICA outperformed CENTURY, CCB and C-TOOL, and produced results comparable with ICBM and RothC. Use of the modified model allowed mostly adequate reproduction of site-
20 specific SOC stocks, while the availability of a nitrogen, plant growth and water submodel enhanced its applicability compared with models that only describe carbon dynamics.

1 Introduction

The historic conversion from native to agriculturally managed soils released 116 Pg of soil organic carbon (SOC) as CO_2 to the atmosphere (Sanderman et al., 2018). If no measures are taken, current projections estimate a continuation in the decline
25 of SOC stocks, further intensifying the release of CO_2 from soils (Riggers et al., 2021; Zhao et al., 2021). However, the sequestration of atmospheric CO_2 in the soil is a climate change mitigation strategy that has substantial potential (Poeplau and Don, 2015; Fuss et al., 2018; Amelung et al., 2020). In particular, croplands are promising landscapes since they are depleted in SOC and generally accessible for cultivation practices (Poeplau and Don, 2015; Sanderman et al., 2017). Researchers have proposed various agronomic systems, including agroforestry, reduced tilling, and perennial and cover cropping, to optimise

30 agricultural opportunities for sequestering carbon and thus partly restore the historically lost SOC (Smith et al., 2005). Estimates of the global capacity of agricultural soils for sequestering carbon range from 2 to 5 Pg C a^{-1} (Fuss et al., 2018). However, identifying and promoting site-specific measures in a highly diverse agricultural landscape encompassing more than 570 million farms globally and affecting ~36 % of the earth's terrestrial surface is a challenge (FAO, 2014; Amundson and Biardeau, 2018; FAO, 2022). Despite the uncertainty in their practicability, international initiatives such as the '4 per mille'

35 and scientific working groups advise the rapid modification of agricultural systems so that they act as carbon sinks (Minasny et al., 2017; Amelung et al., 2020). In contrast, some scientists focus on the potential for underestimated detrimental effects and advocate a well-thought-out approach to sequestering SOC on a large scale (Sommer and Bossio, 2014; Lugato et al., 2014; Amundson and Biardeau, 2018; Murphy, 2020).

Process-based models are tools that can be used to assess mitigation capabilities of agricultural systems for different

40 environmental and management conditions and to analyse disagreements and accuracy of alternative estimation approaches. They are designed to represent the behaviour of natural systems based on a selection of mathematical formulations of generally accepted biophysical principles. Carbon turnover models predict the dynamics of SOC in response to management and climate. Agro-ecosystem models combine carbon and nitrogen turnover routines with algorithms that represent soil water dynamics, greenhouse gas emissions and crop growth and allow the simultaneous prediction of multiple variables from different domains

45 of the soil–plant-atmosphere continuum. Due to the explicit simulation of crop growth in response to weather, site conditions and crop management, agroecosystem models exhibit a clear advantage over simple soil carbon turnover models, as they can translate any changes in conditions that affect biomass input to soil, e.g., climate or crop management, directly to responses in SOC. Potentially, they could also simulate any feedback from those changing SOC dynamics to crop growth, e.g. via nitrogen release or immobilization. However, complex agro-ecosystem models have yet rarely been tested for their ability to do exactly

50 this, just as little as they have been compared to simple soil carbon turnover models to identify whether their advantage in applicability comes with a reduced predictive power (Smith et al., 1997; Farina et al., 2021). In any case, to ensure the adequate predictive capability of such models, complex or simple, a comprehensive validation process using measured data with different soil types, management practices and weather conditions is required.

MONICA (Model of Nitrogen and Carbon dynamics on Agro-ecosystems; Nendel et al., 2011; 2013) is a process-based agro-

55 ecosystem model that simulates the effects of spatiotemporal environmental variabilities and agricultural practices on crops yields and soil conditions (Nendel et al., 2014). It is a successor of the HERMES model, which was initially developed to estimate N dynamics in the soil-crop system (Kersebaum, 1995; 2007). A comprehensive SOC turnover submodule based on the Daisy model was added, replacing the simpler approach in the HERMES model (Hansen et al., 1991; Abrahamsen and Hansen, 2000). The holistic principle of the MONICA model supports research questions on the interlinkage between site

60 conditions, plant development and C and N dynamics in a changing climate. The performance of the MONICA model showed the capability to adequately reproduce crop yields for different crops (Asseng et al., 2013; Bassu et al., 2014; Kothari et al, 2022; Rötter et al., 2012; Salo et al., 2016) and more complex crop rotations (Kollas et al., 2015; Kostková et al., 2021). For SOC dynamics under bare fallow treatment, MONICA performed similar to simple C turnover models (Farina et al., 2021),

but furthermore demonstrated also good performance when simulating short-term high-resolution CO_2 exchange in a soil–
65 plant system (Specka et al., 2016). The Daisy model, another agro-ecosystem model, had already been tested against long-
term soil carbon experiments (Smith et al., 1997) and performed comparably well to soil carbon turnover models, such as
RothC (Coleman & Jenkinson, 1996) or CANDY (Franko, 1996).

The question now is whether the MONICA model with its carbon turnover submodule is able to adequately simulate changes
in SOC stocks considering different agronomic practices. It is also important to establish how the performance of a complex
70 model like MONICA compares with simple C turnover models such as RothC, ICBM, C-TOOL, CENTURY and CCB. To
answer these questions, we move along the following four objectives: i) to validate the SOC submodule of the MONICA model
with data from 46 agricultural long-term monitoring sites in Germany, ii) if necessary to enhance its applicability in view of
the diversification of the dataset used for validation, iii) to identify the main drivers of C sequestration and decomposition in
the model, and iv) if necessary to improve model performance.

2 Materials and Methods

2.1 Data composition and site characteristics
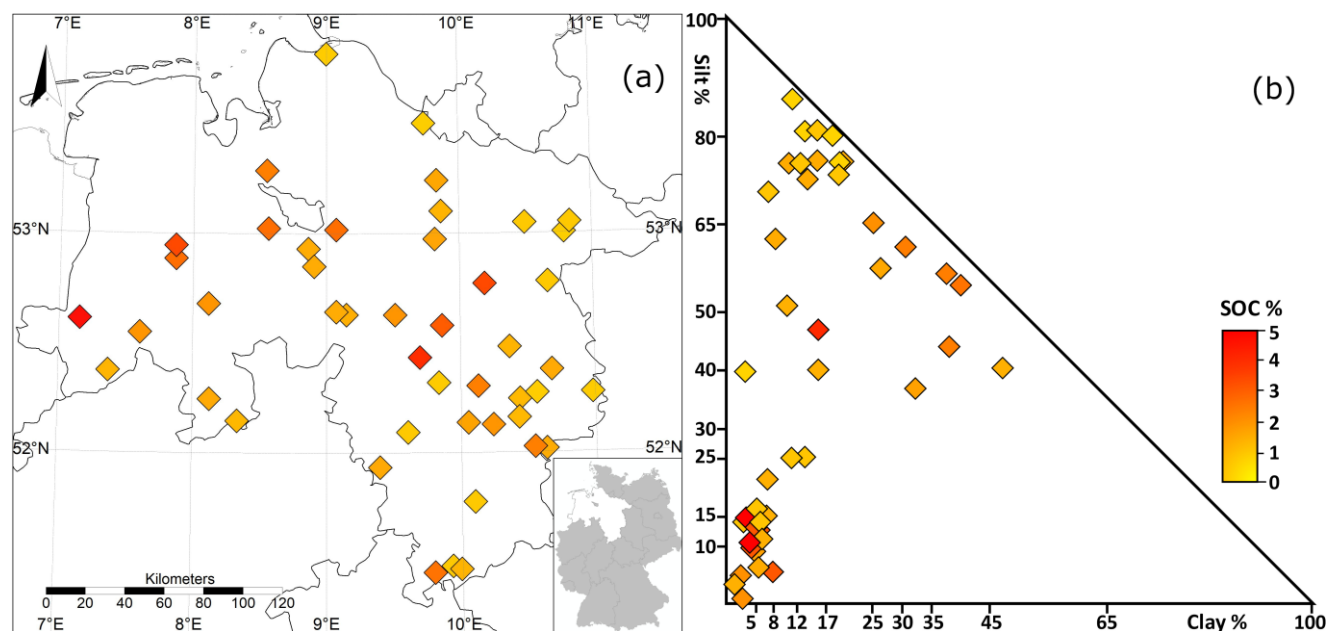
Figure 1: Location (a) and soil classification (b) of the 46 agricultural long-term monitoring sites (BDF) in Lower Saxony, Germany.

For calibration and validation of the SOC turnover subroutine of the MONICA model, long-term monitoring data collected by
the State Office for Mining, Energy and Geology (LBEG) were used. In 1990, the state government of Lower Saxony
80 (Germany) established a programme to permanently observe agricultural and forest soils in the federal territory. Its goal was

to monitor soil changes and identify adverse ecologic and economic implications, and thus predict the risk of imminent damage for the sustainable utilisation of soils.

The examined area is part of the North German Plain, with highlands up to 971 m in its southern region. It is located at latitudes between 51° and 54° North and longitudes between 7° and 11° East (Fig. 1a). It is characterised by a maritime climate, with mean annual temperatures ranging from 8.9 °C to 10.4 °C. Average annual precipitation ranges from 574 mm to 887 mm, with a positive climatic water balance in spring, autumn and winter, and a negative balance in summer.

The database consists of 48 agricultural sites, of which 42 are conventional agricultural systems and six are organic. Two sites were removed due to their observation periods being less than 10 years, leaving 46 entries for the analysis. The observation data consist of 33 sites that have been used for agriculture for more than a century. The remaining 13 sites were converted from forests, meadows or peatlands less than five decades before the start of the observation period. Soil sampling was generally conducted at regular intervals of between one and five years (Höper and Meesenburg, 2021). They were sampled at a depth of up to ~30 cm and consisted of four replicates, which were combined as mixed samples.

In addition, soil horizons were sampled every 10 years at a minimum depth of 1 m. During the soil monitoring programme, a total of 333 soil carbon measurements (replicates not included) were taken at the selected sites from 1992 until the end of 2015. Between the sites, the mineral composition was diverse, with soil textures predominantly sand (20 sites) or silt (21 sites), and less frequently clay soils (five sites; Fig. 1b). The key properties of the examined soils were highly variable between the sites, with SOC ranging from 0.84 % to 4.91 %, C/N from 5.25 to 23.66, bulk density from 1.14 g cm^3 to 1.76 g cm^3, and pH from 3.9 to 7.9. A special distinctiveness of the investigated area is the occurrence of so-called black sands, which are characterised by comparably high, stable C contents in sand-dominated locations (Vos et al., 2018). Site-specific details are summarised in the appendix (Table A1).

The soil and crop management at the sites was undertaken by the landowners or their associates. Therefore, there is a high degree of variability in the agronomic practices performed at the sites, such as tillage depth and type, fertiliser application, irrigation, crop rotation, harvest and crop residue management. The reported management reflects common agricultural practices for the site-specific regions (more details are given by Höper and Meesenburg (2021)). Control treatments were not available for the sites. Simulated soil temperatures and moisture contents were evaluated using data from 11 German long-term field experiments (Table A2; Kaiser and Ruser, 2000; Leidel et al., 2000; Heitkamp et al., 2009; Wolf et al., 2014; Flessa et al., 2017).

2.2 Model description

The MONICA model is comprised of subroutines that characterise plant growth, C and N dynamics, soil water and soil temperature (for a more detailed description see Nendel (2013)). Execution of the programme requires comprehensive input data, including weather, soil and management information, as well as parameter values describing model settings. Every model iteration simulates soil and plant conditions in a daily time step for soil depths up to two metres. Mathematically, the model is one-dimensional (point model) and the spatial representativeness of the model output depends primarily on the calibration

data. Generally, the spatial representation ranges between one m² and one km². The soil profile is partitioned into 20 layers,
115 each with 10 cm thickness of homogenous soil conditions. Soil input data are uniformly extrapolated to the relevant depth if
the entered soil horizons total less than two metres.
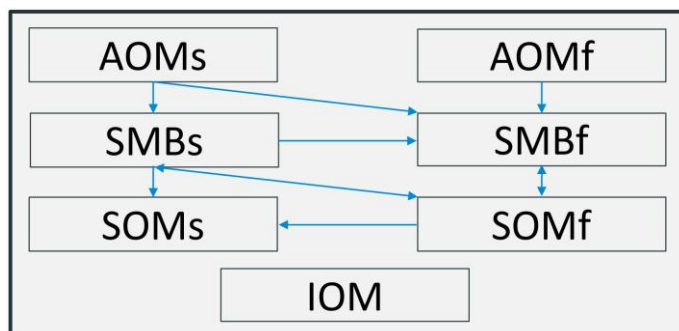
Figure 2. Conceptual SOC pool distribution in MONICA, subdivided into the groups added organic matter (AOM), soil microbial biomass (SMB), soil organic matter (SOM) and inert organic matter (IOM). Except for IOM, each group consists of a slow (s) and a fast (f) pool.

According to the Daisy model, the SOC content is partitioned into three groups composed of dead native soil organic matter (SOM), living soil microbial biomass (SMB) and added organic matter (AOM), each of which contains one rapidly (SOMf, SMBf, AOMf) and one slowly decomposing subpool (Fig. 2; SOMs, SMBs, AOMs; Hansen et al., 1991; Bruun et al., 2003). These compartments are calculated separately for each soil layer.

MONICA is usually initialised using a standard distribution of the SOC pools, which should correspond to an equilibrium at the beginning of the simulation (Bruun and Jensen, 2002). During the initialisation, inert SOC (IOM) is subtracted from the total SOC, based on equations adopted from Falloon et al. (1998). Then the SOM and SMB pools are determined by multiplying the manually entered SOC contents
130 by the corresponding distribution parameters. The AOM pools are initialised and altered by added C contents from crop residues and organic fertilisers. Organic fertiliser and crop types are parameterised with specific N contents (NH_4^+ and NO_3^-), decomposition rate coefficients and pool distributions. Furthermore, plant biomass accumulation is divided into five groups: root, leaf, shoot, fruit and sugar. For perennial plants, an additional permanent structure is considered. Each group has its own parameters, controlling the growth rate and the biomass that can potentially be harvested or added to the AOM pools. Following
135 a harvest command, accrued crop residues left on an agricultural field are immediately distributed to the AOM pools. Root remains are fragmented into the soil layers, depending on the root mass distribution in the soil profile. During tillage, the SOM content is distributed homogenously to the soil layers affected by tillage depth.

The initialised pools are managed by first-order degradation due to decomposition, each controlled by related parameters. This process follows a decay rate that is proportional to the pool size (Hansen et al., 1991). In addition, the decomposition rate of
140 organic matter is affected by soil temperature, soil water and clay contents (Bruun et al., 2003; Farina et al., 2021). These environmental conditions influence each pool as rate-modifying factors, with the exception of clay, which solely affects SMBs. Decomposed AOMf and AOMs are subsequently transferred to the SMB pools, from where fractions are respired to the atmosphere as CO_2 or alternatively converted to SOMs and SOMf, following a pool-specific carbon use efficiency coefficient. Biologically degraded SOMs and SOMf goes through a similar process and are either respired or returned to the SMB pools.
145 A detailed description of the SOM turnover subroutine is depicted in Hansen et al. (1991) and Nendel (2013).

2.3 Model modification

Sixteen crop plants (amaranth, asparagus, buckwheat, carrot, flaxseed, field bean, hemp, lupine, serradella, spelt, strawberry, summer onion, sunflower, turnip, vicia) and eight types of organic fertilisers (biogas slurry, calf slurry, dried poultry manure, hoof and horn meal, lime-stabilised sewage sludge, sheep manure, vinasse) were added to the model. This enabled the

150 reproduction of all occurring land management characteristics in the dataset. Parameter values for crop growth (Tesar, 1984; Jeuffroy and Ney, 1997; Allen et al., 1998), fertiliser nutrient contents (Möller and Schultheiß, 2015; Wiesler et al., 2016), crop residues and organic fertiliser decomposition and partitioning coefficients (Gilmour et al., 1998; Thuriès et al., 2001; Lashermes et al., 2009; Peltre et al., 2012; Semenov et al., 2019) were taken from the literature. Furthermore, a specific C content parameter (*CorgContent*) was defined for each crop and organic fertiliser, providing a more adequate replication of C

155 inputs in comparison with the original constant (Möller and Schultheiß, 2015; Wiesler et al., 2016; Ma et al., 2018). In addition, an overhaul of existing crop and fertiliser types was undertaken, ensuring distinct and more reasonable values for essential parameters. Crop growth was calibrated according to the measured yield data, while the residue-to-crop product ratio was adapted from the literature and an ensemble of allometric functions (Bolinder et al., 2007; Franko et al., 2011; Jacobs et al., 2020; Taghizadeh-Toosi, 2020). The differences between the allometric functions served as indicators of uncertainty in the

160 estimation of C inputs from crop residues.

The functions that describe the influence of soil temperature, water and clay content on decomposition were revised and parameterised. This enabled all relevant carbon turnover factors to be tested for uncertainty and subsequently calibrated. The optimal values of the parameters were used as the evidential basis for the prior probability distributions, which is necessary for model calibration by means of Bayesian inference. For further explanations, see Section 2.5 and Table 1.

165 Extensive research provided evidence of the influence of temperature on physical soil processes, microbial metabolisms and enzymatic activities (Pietkäinen et al., 2005; Haddix et al., 2006; Meyer et al., 2018). The novel temperature function is described by the following function Eq. (1):

$$f_{TOD} = \frac{1}{(1+e^{Td - (e+T_{opt})})Q_{10}/(T_{opt}/\pi)(-1+Q_{10}\frac{Td}{15.76})} , \tag{1}$$

where T is the daily soil temperature in °C, Q_{10} is the Q_{10} temperature coefficient specified by the parameter *QTenFactor*, and

170 T_{opt} is the optimal temperature for decomposition in °C defined by the parameter *TempDecOptimal* (Fig. B1). From literature analysis, a range of 38.2 ± 7.5 °C was identified for the optimal temperature (Pietikäinen et al., 2005; Richardson et al., 2012; Menichetti et al., 2015; Liu et al., 2018; Čapek et al., 2019) and 2.6 ± 0.6 for the Q_{10} temperature coefficient (Haddix et al., 2006; Vanhala et al., 2007; Conant et al., 2008; Gillabel et al., 2010; Meyer et al., 2018).

Microbial aerobic respiration is often driven by changing soil water contents up to a point where water becomes a restricting

175 property for the availability of oxygen (Curiel Yuste et al., 2007; Schimel et al., 2007). Based on analysis of incubation and field experiments, a Gaussian function was proposed, described by the following Eq (2):

$$f_{MOD} = e^{-18(WFPS_d - M_{opt})^2} , \tag{2}$$

where $WFPS_d$ is the daily water-filled pore space in % and M_{opt} is the optimal water-filled pore space (WFPS) for decomposition in % defined by the parameter *MoistureDecOptimal* (Fig. B2). Optimal WFPS for respiration was determined
180 to be around 59.5 ± 9.9 % (Linn and Doran, 1984; Doran et al., 1990; Liebig et al., 1995; Aon et al., 2001; Gabriel and Kellman 2014; Zhang et al., 2015).

Several studies have indicated that increasing soil clay contents have a negative effect on respiration due to the reinforced inaccessibility of SOM for microorganisms (Six and Paustian, 2014; Churchman et al., 2020; Fomina and Skorochod, 2020). In the MONICA model, the modulating effect of clay on the decomposition rates of SMBs is by default adjustable through the
185 parameter LimitClayEffect. An inverse logistic function was proposed, described by the following Eq (3):

$$f_{COD} = \frac{(1-LCE)}{(1+e^{-\pi + C * 16}) + LCE}, \tag{3}$$

where C is the site-specific clay content in kg kg^{-1} and LCE is the limit to which the clay content affects the mineralisation defined by the parameter *LimitClayEffect* (Fig. B3). Laboratory evidence suggests that soils amended with clay minerals reduce mineralisation rates by $\sim 1.5 \pm 1.1$ % for every percent of clay added to the soil, followed by a gradual reduction in effectiveness
190 after a certain clay content is exceeded (Nguye and Marschner, 2014; Schapel et al., 2018; Riaz and Marschner, 2020; Liddle et al., 2020).

2.4 Screening of variables relevant to soil carbon turnover

For the purpose of understanding the relationship between input variables, parameters and the predictive proficiency of the MONICA model, a comprehensive sensitivity analysis was conducted using an improved Morris elementary effects screening
195 method (Fig. C1; Morris, 1991; Campolongo et al., 2007; Pujol, 2009). The algorithm demonstrates the capability of determining the importance of parameters in relatively short computing times (Campolongo et al., 2007; Confalonieri et al, 2010). The effect of 51 variables was examined, of which 15 were environmental, six were related to the crop, four to the management and 26 to SOC turnover. Of the turnover parameters, six were unique for each fertiliser and crop residue type.

Depending on the tangible properties of the variables (factual or conceptual), the upper and lower bounds of each factor were
200 either derived from the literature or intuitively inferred from model characteristics. After building the parameter space, all the factors were scaled to a range from zero to one. During each iteration, one variable was changed to a random value in the defined space (levels = 5, grid.jump = 3) while the remaining variables were unaltered. Subsequently, the effect of the modified factor on SOC stocks was assessed by comparing the present iteration with the previous one. Based on the magnitude of the changes and the degree of the interaction with other factors, a parameter importance ranking was identified using the two
205 measures μ^* and σ (Campolongo et al., 2007).

Common recommendations for the number of model iterations are between 10 to 50 for each variable considered (Campolongo et al., 2007). However, several studies suggest a much higher number of iterations to achieve robust results (Sarrazin et al., 2016; Vanuytrecht et al., 2014; Pianosi et al., 2016). As a base value, 10^3 iterations and three runs were used in total to accomplish adequate coverage of possible input spaces and to estimate the uncertainty in the elementary effects.

The programming language R (version 4.1.0) with the graphical user interface RStudio (version 1.4.1717) and the sensitivity
package (Iooss et al., 2020; version 1.26.0) were used for screening of variables relevant to soil carbon turnover in the
MONICA model.

2.5 Parameter calibration

Bayesian inference techniques are an efficient approach for the calibration of process-based models (Vrugt, 2016; Van de
Schoot et al., 2021). Common methods for prior elicitation, likelihood specification and posterior distributions have been
extensively researched (Van de Schoot et al., 2021). For instance, Bayesian calibration has been applied successfully to forest
biomass growth (Van Oijen et al., 2005; Svenson et al., 2008) and agro-ecosystem models (Lehuger et al., 2009; Gurung et
al., 2020).

As a first step, the experimental data were divided into one subset for calibration and one for validation. In the process, the
data were clustered into three groups based on their initial SOC stocks in order to achieve a general representation of the
variability between the sites. From each of the groups, two thirds of the sites were randomly sampled for model calibration
(totalling 31 sites) and one third for model validation (totalling 15 sites). The parameter selection process for the calibration
was adapted from the results of the sensitivity analysis and the recommendations from Minunno et al. (2013), excluding local
parameters and 25 % of the least important parameters.

The calibration of the 14 selected parameters was undertaken by means of Bayesian statistical inference. In Bayesian
calibration, the determination of possible parameter values is based on evidential probabilities. Therefore, a comprehensive
literature analysis combined with expert knowledge provided the basis for the prior probability distribution, as shown in
Section 2.3 (Oakley and O'Hagan, 2007). A weakly informative prior probability distribution was constructed, with the mean
of all identified measurements as the most likely value and the observed lower and upper bounds as the variance for each
parameter (Table 1; Van de Schoot et al., 2021). At the beginning of the calibration, a random initial value (within the prior
probability distribution) was generated for each parameter, creating a parameter vector for each iteration. Random selection
for subsequent parameter values was limited by the boundaries and a squared range of 1-5 % of the previous value (Van Oijen
et al., 2005). During each iteration, MONICA was executed using the created parameter vector. Afterwards, the generated
output was compared with the measured data using a log-Laplace likelihood function suggested by Vrugt (2016), estimating
the association between simulated and observed values. This association was a representation of the posterior probability of
each parameter vector. The Laplace distribution showed robustness against outliers and random variations, resulting in less
biased and more consistent parameter estimates (Schoups and Vrugt, 2010).

A Markov Chain Monte Carlo (MCMC) method, namely the Metropolis-Hastings algorithm, was used for random sampling
from proposed values to estimate the posterior probability distribution (Fig. D1; Metropolis et al., 1954; Hastings, 1970;
Lehuger et al., 2009). In this process, the likelihood function serves as an indicator to generate a chain of accepted vectors. A
sample was accepted in the sequence if the current proposal divided by the previous proposal was greater than or equal to a
random number between 0 and 1.

Table 1. Description of the 14 parameters chosen for the calibration of the C submodel in MONICA. The prior distribution depicts the upper and lower bounds of possible parameter values inferred from the literature; the posterior distribution describes the median value of the posterior and the corresponding standard deviation.

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Parameter			Prior distribution				Posterior distribution	
Notation	Unit	Default value	Minimum value	Maximum value	Reference		Mean	Standard deviation
AOM_FastUtilizationEfficiency	%	0.1	0.1	0.65	Manzoni et al. (2012), Spohn et al. (2016), Qiao et al. (2019), Schroeder et al. (2020), Agumas et		0.27	1.2^{-3}
AOM_SlowUtilizationEfficiency	%	0.4	0.1	0.65			0.16	0.01
MoistureDecOptimal	%	NA	0.3	0.76	Linn and Doran (1984), Doran et al. (1990), Liebig et al. (1995), Aon et al. (2001), Gabriel and Kellman		0.45	3.1^{-3}
LimitClayEffect	%	0.25	0.01	0.5	Nguye and Marschner (2014), Schapel et al. (2018), Riaz and Marschner (2020), Liddle et al.		0.34	2.2^{-3}
QTenFactor	Q_{10}	NA	1.6	4.2	Haddix et al. (2006), Vanhala et al. (2007), Conant et al. (2008), Gillabel et al. (2010), Meyer et al.		3.15	0.04
PartSMB_Fast_to_SOM_Fast	%	0.6	0.4	0.8	Kindler et al. (2006), Liang and Balser (2011), Miltner et al. (2012), Kallenbach et al. (2016)		0.66	2.1^{-3}
PartSMB_Slow_to_SOM_Fast	%	0.6	0.4	0.8			0.61	1.3^{-3}
SMB_FastDeathRateStandard	d^{-1}	0.01	0.01	0.25	McGill et al. (1986), Joergensen et al. (1990), Throckmorton et al. (2012)		0.12	1.1^{-3}
SMB_SlowDeathRateStandard	d^{-1}	1.0^{-3}	3.0^{-4}	0.01			5.4^{-3}	5.6^{-5}
SMB_UtilizationEfficiency	d^{-1}	0.0	0.1	0.88	Throckmorton et al. (2012), Creamer et al. (2019), Buckeridge et al. (2020), Geyer et al. (2020), Paustian et al. (1992), Trumbore		0.43	6.6^{-3}
SOM_FastDecCoeffStandard	d^{-1}	1.4^{-4}	1.0^{-5}	2.8^{-4}	(2000), Gleixner (2013), Wang et al. (2016), Shi et al. (2020)		8.6^{-5}	1.9^{-5}
SOM_SlowDecCoeffStandard	d^{-1}	4.3^{-5}	6.4^{-7}	5.0^{-5}			1.3^{-5}	4.0^{-7}
SOM_FastUtilizationEfficiency	%	0.5	0.1	0.88	Manzoni et al. (2012), Saifuddin et al. (2019), Qiao et al. (2019), Schroeder et al. (2020), Agumas et		0.35	6.8^{-3}
SOM_SlowUtilizationEfficiency	%	0.4	0.1	0.88			.49	2.2^{-3}

Three chains with 10^5 iterations and different starting values were created to obtain a reliable inference on the posterior distribution (Van de Schoot et al., 2021). Following the recommendation of Van Oijen et al. (2005), the first ten percent of the accepted vectors were discarded as unrepresentative "burn-in" of the chains. Stationary distribution of the chains was verified using the empirical method developed by Gelman and Rubin (1992), which compares the variance between and within the

250 chains. An adequate convergence was obtained if the Gelman-Rubin shrink factor was < 1.2 (Brooks and Gelman, 1998). The R package coda (Plummer et al., 2006; version 0.19-04) was used for random sampling from the posterior distribution and verification of convergence.

2.6 Quantitative methods and model evaluation specifics

The statistical methods for evaluating the sampling distributions and model performances were conducted with the R packages
255 hydroGOF (Zambrano-Bigiarini, 2020; version 0.4-0) and car (Fox et al., 2019; version 3.0-12). Shapiro-Wilk (Shapiro and Wilk, 1965) and Levene's tests (Levene, 1960) were used on the measured fluxes to test the population for normality and variance equality. The Shapiro-Wilk test resulted in rejection of the null hypothesis for the variables ($p < 0.001$) inferring that the examined data were not normally distributed. The Levene's test resulted in acceptance of the null hypothesis, implying that the variance in the samples was homoscedastic ($p = 0.88$). Using the interquartile range criterion (Jargowsky and Yang,
260 2005), 19 outliers were identified in the data, which represents ~6 % of the cases. Q-Q plots confirmed the assumption that the data points were nonlinear distributed.

To improve representativeness, the years 1992 to 2000 were excluded from the analysis, reducing the observed number of cases from 333 to 223. The reason for this approach was justified by the inherent risk of statistical biases from overassessment due to the initialisation of simulated SOC stocks based on measured values. Furthermore, a spin-up period of 20 simulated
265 years was included to ensure a realistic reflection of available fresh matter in the soil, since the AOM pools were initialised equal to zero. The chosen period corresponds to the maximum duration of mineralisation of most AOM (Thuriès et al., 2001; Semenov et al., 2019).

Based on the reduced sample size and the abovementioned statistical assumptions, Kendall's rank correlation coefficient (T; Kendall, 1948) was selected to assess the statistical dependence between the observed and simulated values. Kendall's Tau is
270 the preferred method for measuring the ordinal association in data consisting of small sample sizes and outliers, and is more robust and efficient than the Pearson correlation coefficient and Spearman's rank correlation coefficient (Bonett and Wright, 2000). In addition, the model results were evaluated with the statistical methods mean absolute error (MAE; Willmott and Matsuura, 2005) for analysis of the average error between observed and simulated pairs, and with the Nash-Sutcliffe model efficiency coefficient (NSE; Nash and Sutcliffe, 1970) to assess the predictive skill of the simulation model (Fig. E1).
275 Furthermore, the slope-intercept was considered in order to better assess the behaviour of the models.

3 Results

3.1 Simulation of soil and plant growth variables

Soil temperatures were well predicted, with an average error of just under ~1.8 °C (Table 2). Low temperatures were slightly overestimated and high temperatures were underestimated. In comparison, soil water contents were predicted somewhat less
280 favourably, with an average error of ~5.4 %. Nonetheless, the predictive power was better than the mean of the observations.

The model tended to slightly overestimate moderate water contents and underestimate the measured variance. A comparison of the default with the calibrated plant growth parameters was not performed, since new crops were added to the model in order to simulate all possible systems. The calibrated plant growth model achieved a good representation of measured yields, with an average error of less than ~1.6 Mg DM ha^{-1} for all crops. Overall, there was a linear relationship between the calibrated and measured yields. However, there was a high variability in the estimation quality between different crop types. Some crops, such as triticale and potatoes, were modelled less adequately than silage maize and sugar beet, which was partly explained by the differences in the number of observations and the variance in yields. The C inputs were compared with the estimated mean of the allometric functions. The model produced an average error that was similar to the calculated difference between the allometric functions (~1 Mg DM ha^{-1}).

Table 2. Statistics describing the performance of the MONICA model regarding soil and plant growth variables. The statistics depicted: Kendall's rank correlation coefficient (T), Nash-Sutcliffe model efficiency coefficient (NSE), mean absolute error (MAE), slope and intercept for characterising the relationship between measured values (except values for C inputs, which were calculated with allometric functions), and simulated values.

Variables	T	NSE	MAE	Slope	Intercept
Soil temperature (n = 24576)	0.83***	0.86	1.84 [°C]	0.93	1.17
Soil water content (n = 9636)	0.63***	0.39	5.37 [%]	0.84	7.83
Yield (n = 846)	0.65***	0.81	1.62 [Mg DM ha^{-1}]	0.99	0.07
C-Input (n = 846)	0.4***	0.32	0.98 [Mg DM ha^{-1}]	0.88	0.38

*** p < .001, ** p < .01, * p < .05

3.2 Quantification of uncertainty and parameter calibration

A strong relationship was identified between SOC stock change rates and C inputs through organic fertilisation and crop residues (Table F1; *OrganicFertilisation, AOM_DryMatterContent, CorgContent*), given the upper and lower bounds of possible values. Additionally, these factors were less affected by other variables, noticeably by the lower σ on average in comparison with μ*. The climate variables temperature (*Temperature*) and precipitation (*Daily_Precipitation*) were also of greater importance to the SOC turnover rate, but had a higher standard deviation of elementary effects. Except for solar irradiance (*Global_Radiation*), temperature and precipitation, all other climatic factors had an insignificant effect on SOC stocks. Of all the variables that were initialised at the beginning and remained static during the model run, the initial SOC stock (*SoilOrganicMatter*) was the most influential, followed by the site-specific field capacity (*FieldCapacity*), soil bulk density (*SoilBulkDensity*) and pore volume (*PoreVolume*). Other site-specific variables (*Sand, Clay, PermanentWiltingPoint, pH, Sceleton*) were comparably less important. The management factors tillage (*Tillage*), mineral fertilisation (*MineralFertilisation*), irrigation (*Irrigation*) and N content in the added organic matter (*CN_Ratio_AOM_Slow, CN_Ratio_AOM_Fast, CN, AOM_NH4Content, AOM_NO3Content*) had a small effect on the SOC turnover rate. σ values

were generally higher than the overall importance of the input factors, indicating that there is probably an interaction and non-linear effect between the environmental and management variables.

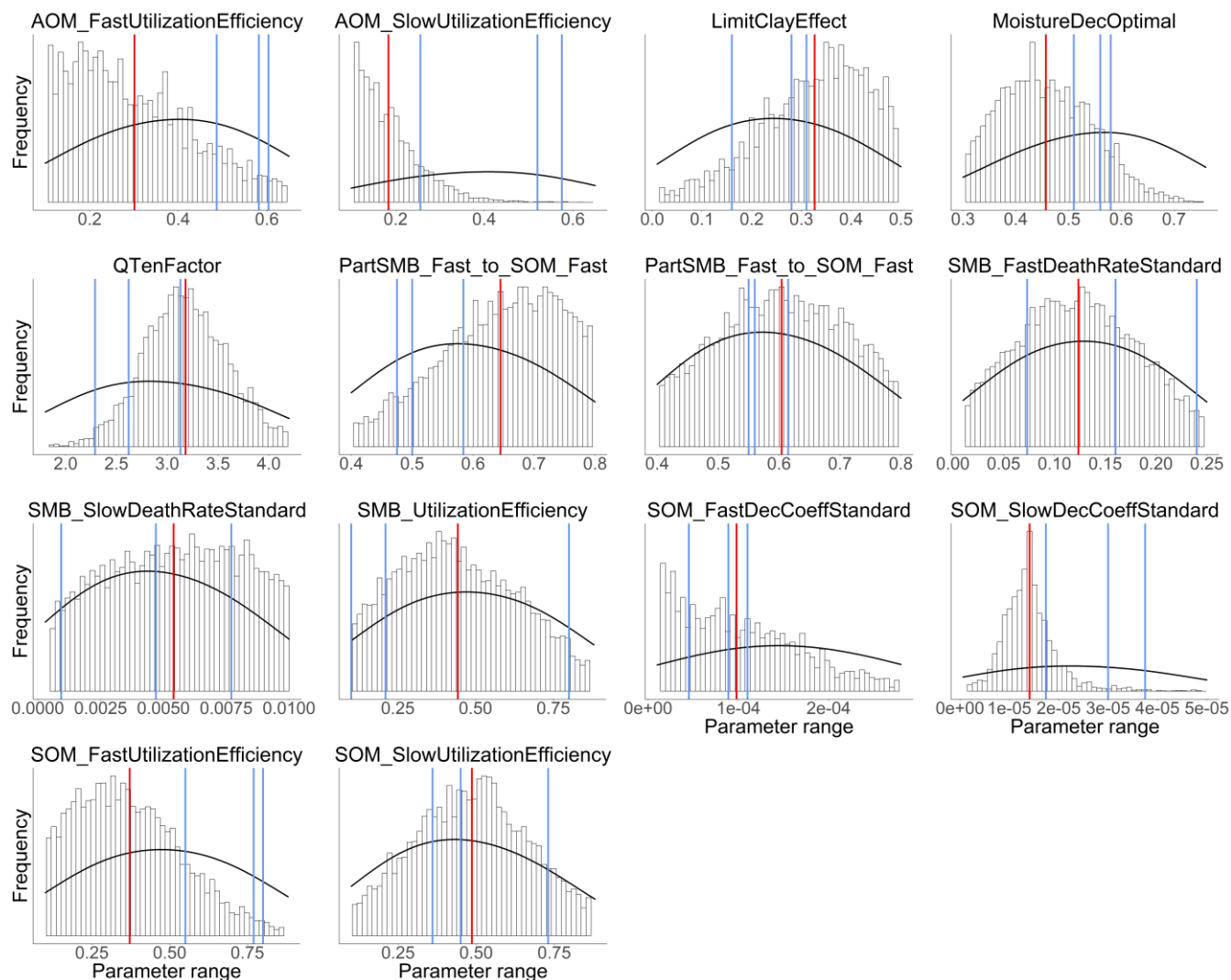



Figure 3. Posterior probability distribution (vertical bars) of 14 C submodel parameters calibrated over the range of the prior probability distribution (density curve). The posterior is depicted as the distribution of accepted values determined by random sampling from 10^5 proposed values. The vertical red lines represent the optimal values as the median of the posterior and the blue lines the initial values of each chain.

310   In comparison, the sensitivity of the turnover variables was less pronounced, and revealed three particularly important parameters. One C pool parameter (*SOM_SlowDecCoeffStandard)* and two weather-regulated decomposition factors (*QTenFactor, MoistureDecOptimal*). They were similar in significance to precipitation, initial SOC and field capacity. Changes in the parameters *QTenFactor* and *MoistureDecOptimal* were more important than in *LimitClayEffect*, which was in accordance with the influence of clay being limited to the SMBs pool. The effect of the factor describing the optimal

315   temperature for decomposition (*TempDecOptimal*) was negligible, but considerably gained in importance when a higher range of possible values for the variable temperature was included. Crop and fertiliser-specific turnover parameters were also sensitive and coincided with the importance of C inputs for SOC stock change rates. Less pronounced sensitivities involved the parameters describing the microbial dynamic (besides the parameter *SMB_UtilizationEfficiency*), distinguishable by the SMB in the name. Non-linear effects and interactions were more strongly pronounced in the SOC turnover parameters.

320   Figure 3 shows the results of the calibration process, with the posterior distribution curve representing the frequency of accepted values and the vertical red line indicating the optimal value as the median of the posterior. Bayesian inference reduced the uncertainty for all parameters, converging the posterior distribution to a normal or log-normal distribution. Throughout the random sampling, the entire plausible parameter space was covered and curve progressions showed that the optimal values were probably within the specified limits. During the calibration process, the posterior distribution of most parameters became

325   narrower in comparison with the prior, confirming the convergence to the optimal values. Similar to Lehuger et al. (2009), the acceptance rate for the proposed parameter values was ~25 %. Based on the results of the Gelman-Rubin diagnostics, adequate convergence was achieved for all parameters (Fig. G1), signifying the unbiased outcome of the MCMC. The correlations across parameters remained below a coefficient of 0.2, with the exception of *QTenFactor* and *SOM_SlowDecCoeffStandard* (Fig. H1).

3.3 Model performances

The statistical methods showed that the default MONICA model (MONICA$_{def}$) was inferior to the RothC, ICBM and C-TOOL models, and produced results similar to CENTURY and CCB (Table 3; Fig. 4; Fig. I1). The mean absolute error was the highest of all the models, but was within the margin of error compared with CENTURY and CCB for all sites and with CCB for the validation sites. Kendall's rank correlation coefficient and Nash-Sutcliffe model efficiency coefficient values were

335   within the margin of error for all models and sites, and therefore showed insignificant differences in the predictive skill of the models.

MONICA$_{def}$ estimations deviated on average by more than ~10 % from the measurements. Inclusion of a C content parameter for different crop residues and organic fertilisers in the MONICA model (MONICA$_{mod}$) improved the accuracy of predictions for C stock changes by ~16 % for all sites and by ~8 % for the validation sites, while the additional global parameter calibration

340   (MONICA$_{cal}$) improved MONICA's average estimation error by ~29 % for all sites and by ~30 % for the validation sites. Compared to MONICA$_{cal}$, only RothC achieved a lower MAE for all sites. The differences between RothC and MONICA$_{cal}$ were within the margin of error for the validation, but not for all sites. MONICA$_{cal}$, MONICA$_{mod}$, RothC, ICBM and C-TOOL produced errors within the average standard deviation between the replicates of the C measurements for the validation (7.76 Mg C ha^{-1}) and for all sites (5.01 Mg C ha^{-1}). In contrast, MONICA$_{def}$, CENTURY and CCB had errors that were not within

345   the average standard deviation for all sites. Each model was capable of producing results corresponding to the 95 % confidence interval for the validation sites (6.4 Mg C ha^{-1}). No model achieved the same for all measurements (3.3 Mg C ha^{-1}).

Table 3. Statistics describing the performance of the default (MONICA_def), the modified, uncalibrated (MONICA_mod), and the modified, calibrated (MONICA_cal) soil carbon submodel in MONICA in comparison with five C turnover models for all sites. The first ten years were excluded from the analysis (n = 223). The following statistics are depicted are Kendall's rank correlation coefficient (T), Nash-Sutcliffe model efficiency coefficient (NSE), mean absolute error (MAE) and the slope and intercept for characterising the relationship between measured and simulated C stocks.

Model	T	NSE	MAE [Mg C ha^{-1}]	Slope	Intercept
MONICA$_{def}$	0.83***	0.91	5.54	1.07	0.43
MONICA$_{mod}$	0.83***	0.93	4.68	1.11	-3.8
MONICA$_{cal}$	0.85***	0.95	3.91	1.02	-0.83
RothC	0.85***	0.96	3.69	1.03	-0.86
C-TOOL	0.83***	0.94	4.43	1.04	-0.23
ICBM	0.85***	0.94	3.99	0.93	3.43
CENTURY	0.79***	0.92	5.31	1.1	-4.19
CCB	0.8***	0.92	5.44	0.96	-1

*** p < 0.001, ** p < 0.01, * p < 0.05

Based on the outcomes of the NSE method, all models performed better than the mean of the measurements, with MONICA$_{def}$ achieving the lowest and RothC the highest scores. The differences between MONICA$_{cal}$ and MONICA$_{def}$ were within the margin of error. The Kendall rank coefficient displayed a significant statistical dependence between the simulated and measured data, with MONICA$_{cal}$, RothC and ICBM achieving the best results and CCB and CENTURY the weakest results. Overall the T coefficients were high, with MONICA$_{def}$ achieving average results and MONICA$_{cal}$ above average results in contrast with the other models. Of the six models, MONICA$_{cal}$ generally produced better results than CENTURY, C-TOOL and CCB. The slope and intercept of the statistical analysis indicated that MONICA$_{def}$ generally overestimated decomposition, while MONICA$_{cal}$ produced results closer to the measured data. Overall, MONICA$_{def}$ underestimated the SOC stocks at 24 sites and overestimated them at two sites, leaving 20 sites that were simulated within the standard deviation of the measurements. In comparison, MONICA$_{cal}$ underestimated the SOC stocks at six sites and overestimated them at nine sites, leaving 31 sites that were simulated within the standard deviation of the measurements. The calibration resulted in a model behaviour that was comparable with RothC.

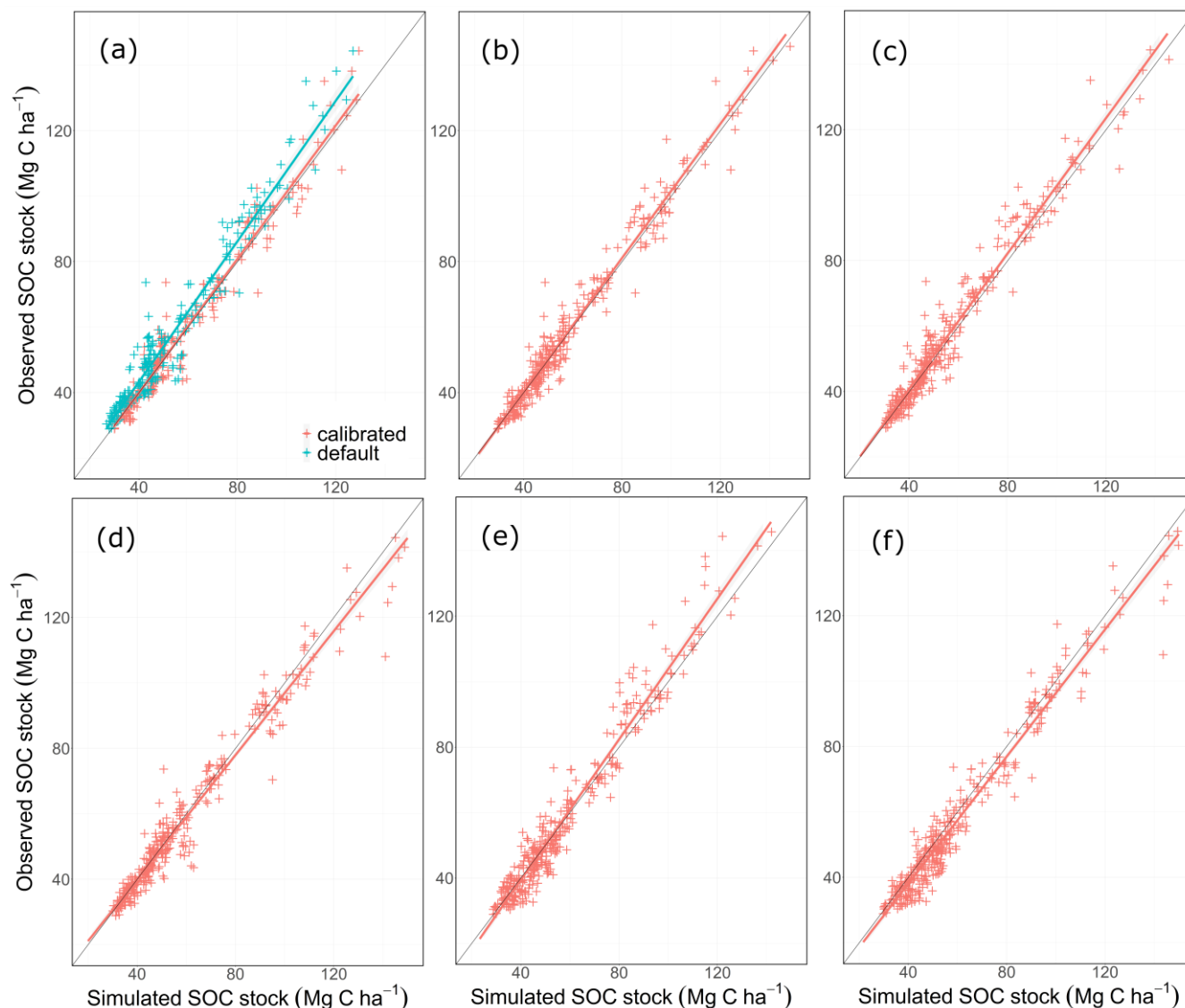



Figure 4. Linear regression between simulated and observed SOC stocks for 46 BDF sites, all years and 333 measurements for the MONICA (a), RothC (b), C-TOOL (c), ICBM (d), CENTURY (e) and CCB (f) models.

4 Discussion

4.1 Differences in models and performances

The MONICA model featured a plant growth, nitrogen and soil water submodel, while RothC, CCB, ICMB, CENTURY and C-TOOL did not. Smith et al. (1997) have argued that coupling C turnover with agro-ecosystem submodels will be necessary to simulate the impact of global change on agricultural production and SOC changes. However, they also assumed that simpler

370   models have an advantage in simulating various land uses due to the lack of adequate data and the necessary parameterisation
of complex models. Other literature shows that C turnover models coupled with sophisticated plant growth models do not
result in a better performance, highlighting that complex models are error prone (Mueller et al., 1996; Parton, 1996; Li et al.,
1997). This also applied to the data requirements of the MONICA model, which led to the necessary parameterisation of new
crop and fertiliser types in order to simulate all sites. Since there was no additional information in the database about the

375   organic fertiliser and crop residue contents, allometric functions and standard values had to be used to compare ingoing C
inputs. This could have led to an inaccurate reproduction of C inputs, since the specific properties of organic fertilisers and
crop residues are characterised by high standard deviations (Möller and Schultheiß, 2015; Wiesler et al., 2016; Ma et al., 2018).
However, the results of the present study did not confirm the assumption that model complexity is an indicator of predictive
ability in process-based models. Notable differences were driven by model characteristics that were partially controlled by

380   parameterisation. Each model showed weaknesses depending on the specific conditions at each site (Fig. J1 to J7). For instance,
MONICA_def overestimated decomposition at ~52 % at all sites in the dataset, but managed to predict the C losses at the
Fuhrberg and Vinnhorst sites better than MONICA_cal. These sites had undergone recent land-use change and had on average
higher initial SOC stocks and losses over the measured period. In temperate agro-ecosystems, conversion from grassland or
forest to arable land is considered to be one of the main reasons for C release (Vos et al., 2019; Poeplau et al., 2020).

385   Correspondingly, sites with recent land-use change showed the strongest decline in C stocks and were therefore a good
indicator of the downward trend in SOC stocks. Removing sites with a cultivation period < 50 years (n = 10) resulted in a
balanced C budget ($0.02 \, t \, C \, ha^{-1} \, a^{-1}$), which also produced an even better performance by the calibrated MONICA model. The
magnitude of changes in SOC stocks for all sites was comparable with other studies conducted in temperate climates,
representing a general trend for German arable land (Steinmann et al., 2016; Sanderman et al., 2017; Vos et al., 2019; Seitz et

390   al., 2022).

All the models underestimated the magnitude of variation in the natural SOC dynamics. The best simulation qualities were
generally achieved at sites with slowly changing SOC stocks. Some sites distinguished by dramatic changes in SOC were
especially problematic for the models to simulate, in particular the Dinklage site (BDF033_L), which was characterised by
high C losses, and in contrast the Reinhausen site (BDF051_L), which was characterised by increasing C stocks. CENTURY

395   and CCB, which underperformed in comparison with the other models, were the best at predicting these highly variable sites.
The other models overestimated the SOC stock for BDF033_L and underestimated it for BDF051_L. In MONICA_cal, these
two sites caused ~10 % of the mean absolute error.

At site BDF033_L, the farmer was cultivating winter barley and corn as energy crops, but experienced declining corn yields
over the measured period. This probably led to a negative C balance, even though the field was extensively fertilised with pig

400   and bull slurry. In MONICA, organic fertilisation was the best indicator of SOC changes, while there was no significant
relationship between the amount of organic amendments and SOC stock changes in the measured data. Accounting for the
quantitative variations in manure application supports the assumption that organic fertilisation is a weak predictor of SOC
turnover rates, and correspondingly was overestimated in the model. However, numerous studies clearly show a positive effect

of organic amendments on SOC stocks (Koishi et al., 2020; Gross and Glaser, 2021; Roß et al., 2022). Since the BDF data did
405   not feature control treatments, it was not possible to undertake a comparative analysis of the individual sites on the effect of
organic inputs. Another consideration concerned site-specific uncertainties in the necessary model input. The properties of
organic fertilisers and residues were unknown, and thus had to be supplemented with standard values from the literature. Each
fertiliser and crop type had fixed C ratios in the models, while in reality these inputs have a highly variable composition and
can vary greatly depending on farm management, animal diet, climate and soil conditions (Schnug et al., 1996; McCartney et
410   al., 2006; Cajamarca et al., 2019). There are three possible reasons for the disparity between the models and observations at
site BDF033_L: i) the fertiliser and crop-specific nutrient contents were inaccurate, leading to an overestimation of C inputs,
ii) the C losses from harvesting were underestimated, or iii) the site and environmental characteristics and their effect on SOC
decomposition were incorrectly simulated.

In comparison, at site BDF051_L the farmer had transitioned to a management with reduced tillage intensities and no
415   cultivation of root crops, and managed to increase the C content in the upper soil layer by ~30 Mg C ha^{-1} during the measured
time period. Although the effect of conservational tillage is still being debated, numerous studies support the effect measured
at site BFD051_L (Freibauer et al., 2004; Luo et al., 2010). In the MONICA model, tillage uniformly distributed SOC in the
affected soil layers. This was supported by the sensitivity analysis, which showed that tillage had an effect when there was
heterogeneity in the soil profile. However, based on this study's data, it is likely that the model underestimated the effect of
420   tillage on SOC decomposition. Possible reasons for the disparity were either an imbalance of C inputs into the topsoil or
missing functions that describe the processes involved, such as the effect of tillage on soil aggregation, hydraulic properties
and root growth (Mondal et al., 2020). The effect of tillage on SOC stocks is currently absent from other C turnover models,
such as RothC. Jordon and Smith (2022) suggest modifying the decomposition rate to achieve a realistic representation of the
effect of tillage on SOC, although they assume a small effect of tillage on SOC stocks.
425   Another probable reason for the differences between the model and measurements were the climate variables. In the data, only
precipitation correlated significantly with SOC stock change rates, while in MONICA this was important, but less so than
temperature. Sierra et al. (2015) argue that functions in process-based models tend to overestimate the temperature effect on
decomposition rates and this imbalance cannot be offset by the moisture effect. This was also likely in the MONICA model,
since the effect of temperature on decomposition was generally twice as strong as moisture. The importance of climate
430   variables is still being debated, with Luo et al. (2017) and Carvalhais et al. (2014) finding major effects on SOC changes, while
only minor effects have been established by Fujisaki et al. (2018) and Vos et al. (2019).

4.2 Implications for improving MONICA's soil carbon turnover model

Most biogeochemical models were developed many decades ago. This is also true of the C turnover submodel in MONICA,
which   is   based   on   the   Daisy   model   (Hansen   et   al.,   1991;   Abrahamsen   and
435   Hansen, 2000). Functions in C turnover models need to be improved in order to estimate SOC dynamics across a wide range
of spatiotemporal and agronomic conditions. Correspondingly, general insights into soil processes are needed to justify the

introduction of additional functions. There needs to be further examination of existing mechanisms and the interactive effects of site-specific characteristics, but also missing factors that affect microbial abundance, diversity and activity (Louis et al., 2016). Microorganisms in particular play a significant role in the soil carbon cycle by controlling the rate of mineralisation
440 and promoting plant growth by improving mineral nutrition or stimulating immune and stress reactions in crops (Tardy et al., 2015; Berg et al., 2017). Therefore, an understanding of the soil biosphere is key in order to simulate C and N dynamics adequately.

This study attempted to introduce two changes to the default MONICA model: i) a revision, parameterisation and calibration of the functions that describe the influence of soil temperature, water and clay content on decomposition, and ii) an integration
445 of a crop and fertiliser-specific C content parameter. Alteration of the temperature function involved mathematically reformulating the original exponential function into a surge function. The reason for this change was that the default equations did not consider the decreasing activity of soil microorganisms at temperatures above 30 °C (Liu et al., 2018; Fang and Moncrieff, 2001). This could lead to an overestimation of mineralisation, especially under globally increasing temperatures caused by climate change (DWD, 2020). Further improvements could be achieved by incorporating changes in the optimal
450 temperature for decomposition based on site-specific varieties of microbial communities and C qualities (Fierer et al., 2006).

The clay function was changed to an inverse logistic function on the basis of the original equation. Many studies suggest that increasing soil clay contents have an adverse effect on SOC decomposition due to the reinforced inaccessibility of organic substances for decomposers (Six and Paustian, 2014; Churchman et al., 2020; Liddle et al., 2020). Adversely, several studies have found no effect or even a positive effect (Wei et al., 2014; Fissore et al., 2016; Zhang et al., 2019). There are multiple
455 explanations for this disparity, e.g. different clay types, residue qualities and microbial diversities (Nguye and Marschner, 2014; Fissore et al., 2016). Without knowing the exact underlying processes to explain the variability and limitation of the available data, only minor changes were made when modifying the clay function. Removing the clay factor might reduce the complexity of the model, but could also negatively impact the predictive capability of the SOC turnover routine. Further evaluation is needed to assess the importance of the clay factor on decomposition and therefore change the function
460 comprehensively.

Furthermore, the soil moisture factor, which was originally based on the pressure potential derived from the van Genuchten model (Van Genuchten, 1980; Abrahamsen and Hansen, 2000), was simplified. The default moisture function was directly controlled by soil water content, soil texture, bulk density and SOC stocks, while the revised Gaussian function was only regulated by soil water content. It is argued that the effect of texture on the decomposing community could be overestimated,
465 since clay as a factor was already included in a function limiting mineralisation. In addition, the influence of soil texture on decomposition proved to be inconsistent, therefore requiring a better understanding of underlying processes (Doran et al., 1990; Scott et al., 1996), while the influence of bulk density was limited by model characteristics as it remained static after initialisation, hampering the reproducibility of management-related soil disturbances on the moisture function. Linn and Doran (1984) showed that the effect of ploughing on bulk density increased the microbial activity and decreased the optimum soil
470 moisture content, while other studies have also determined an impact of soil structure changes on the mineralisation rate

(Franzluebbers, 1999; Vilkiene et al., 2016). Another particularity of the default moisture function was that the optimal water-filled pore space for microbial activity was estimated to be between 10-25 % for sandy soils and 30-35 % for clay soils, while in the literature a general value of ~60 % is considered the most probable (Linn and Doran, 1984; Doran et al., 1990; Liebig et al., 1995; Aon et al., 2001; Gabriel and Kellman 2014; Zhang et al., 2015). Nevertheless, the simplification is probably less

475 error-prone and easier to understand, but at the same time not optimal for depicting the variance in microbial diversity and their environmental preferences. The inclusion of a more complex moisture function should be based on definite results.

As shown with the crop and fertiliser-specific C content parameter, increasing the complexity of a model can be important for improving performance. According to the sensitivity analysis and the results, the addition of a C content parameter was necessary for the representation of realistic C inputs. However, the limited availability of data hinders the determination of

480 precise C contents in plant residues and organic fertilisers. As mentioned in Section 4.1, values adopted from the literature could have a detrimental effect on the model's predictive capability. Therefore, measurements of C contents in crop residues and organic fertilisers should be part of the experimental design when investigating changes in SOC stocks. To further optimise the model, it is recommended that the C inputs from crop residues be divided into aboveground and belowground biomass. Currently there is no distinction between roots, shoot and leaves in the AOM pool, even though there is good evidence that

485 roots in particular have a much longer mean residence time in the soil (Poeplau et al., 2021; Sokol and Bradford, 2019).

5 Conclusions

The evaluation of the original MONICA model revealed weaknesses in estimating SOC trends. Using the default parameterisation, decomposition was overestimated and the statistical performance was inferior compared with the C turnover models RothC, ICBM and C-TOOL. MONICA's capabilities of simulating SOC trends increased when more realistic functions

490 were introduced describing the effects of soil temperature, moisture and clay on decomposition, coupled with a better specification of C inputs from organic fertilisation and crop residues. A sensitivity analysis displayed the importance of C inputs and the environmental factors of temperature and precipitation, as well as the corresponding parameters on SOC variation in the model. A successful calibration by means of Bayesian inference improved the parameterisation of the C turnover subroutine. The modified and calibrated MONICA model outperformed CENTURY, C-TOOL and CCB and

495 produced similar results in comparison to RothC and ICBM. With the exception of certain sites, adequate reproduction of SOC stock change rates was achieved. However, none of the investigated models was capable of simulating each site satisfactorily. All the models underestimated the variation in measured SOC dynamics that occurred at sites with particular properties and management practices. This paper reveals that even complex, biogeochemical models are capable of performing similarly or even better than the more simplistic C turnover models. It is likely that with better data availability and continuing optimisation

500 of functions representing natural processes, MONICA's performance can be further enhanced.

Appendix A. General properties of each monitoring and experimental site

Table A1. Description of the long-term monitoring sites (n = 46) used for calibration and validation of the carbon turnover routine in the MONICA model. SOC = soil organic carbon, MAT = mean annual temperature, MAP = mean annual precipitation.

Site	Location	Start date	End date	Sand [%]	Clay [%]	Bulk density [g cm⁻³]	SOC [%]	C/N ratio	MAT [°C]	MAP [mm]	Average fertiliser amendments [kg N yr⁻¹]**	Land-use change
BDF001_L	Timmerlah	1992	2014	6	16	1.34	1.2	8.6	9.8	632	151 UAN / 27 POM	1760
BDF002_L	Druette	1994	2014	3	15	1.35	1.2	11.0	9.8	631	181 UAN / 20 POM	1760
BDF003_L	Ehmen	1993	2014	70	7	1.37	1.4	10.6	9.9	689	151 CAN / 17 POM	1945
BDF004_L	Hemmendorf	1995	2014	2	16	1.38	1.1	10.6	9.8	769	139 UAN, CAN / 40 PIS	1780
BDF005_L	Reinshof	1991	2014	11	16	1.24	1.0	10.0	9.3	632	183 UAN / 4 LFR	1775
BDF006_L*	Mariental	1991	2014	60	16	1.48	0.9	10.4	9.6	601	132 CAN	1760
BDF007_L	Barum	1996	2014	21	7	1.4	1.0	10.0	9.4	674	155 UAN, U / 54 POM	1776
BDF008_L*	Hofschwicheldt	1991	2014	17	38	1.33	2.3	9.3	9.8	617	169 UAN, U / 33 PIS	1833 / 1992 drainage
BDF009_L*	Hornburg	1991	2014	2	23	1.29	1.5	9.4	9.4	666	158 UAN / 26 POM	1760
BDF010_L*	Uesen	1996	2013	86	4	1.36	2.7	17.7	9.7	680	113 CAN, AS / 22 LFR	1897 / 1997 drainage
BDF012_L	Buehren	1991	2014	14	26	1.35	1.6	10.7	10.0	740	180 UAN / 46 PIS	1775
BDF013_L	Ottenstein	1995	2014	4	21	1.15	1.6	7.9	8.9	759	160 CAN, AS / 22 PIM	1760
BDF014_L*	Neuhauserfelde	1991	2011	38	11	1.34	0.9	9.3	9.3	887	177 CAN, AS, U / 44 LFR	1878
BDF016_L	Tetendorf	1992	2014	81	4	1.43	1.4	17.0	9.3	796	119 CAN, U / 148 LFR, SS	1897
BDF017_L*	Lueder	1993	2014	82	6	1.46	1.1	12.0	9.2	735	102 CAN, AS / 81 LFR	1777

BDF019_L	Ganderkesee	1992	2014	75	6	1.37	2.7	15.7	9.7	693	66 UAN	1780
											184 CAS, LFR	
BDF021_L	Groenheimer Moor	1997	2014	84	5	1.45	2.8	18.3	10.0	770	79.8 POM	1945
												< 1990 = pasture
BDF022_L*	Voxtrup	1993	2014	44	16	1.33	1.4	9.9	10.0	830	82 CAN, AS	1920
											101 CAS, PIS	
BDF024_L*	Dalumer Moor	1992	2014	82	4	1.18	4.7	26.3	10.3	783	129 CAN, AS	1978
											171 CAS	
BDF026_L	Vechtel	1995	2014	95	4	1.34	2.0	13.1	10.1	766	112 CAN, AS	1935
											151 CAS, PIS	
BDF027_L	Barrien	1996	2014	32	5	1.26	1.5	11.5	9.7	715	134 UAN, U	1897
											166 SS	
BDF031_L	Vinnhorst	1992	2011	37	16	1.14	4.0	13.5	9.9	645	120 CAN	1980
BDF032_L	Markhausen	1992	2015	81	4	1.39	3.3	54.8	9.9	777	47 UAN, CAN	1957
											172 CAS, PIS	
BDF033_L	Dinklage	1992	2014	92	3	1.32	1.8	13.1	9.8	765	19 CAN	1973
											199 PIS, CAS	
BDF036_L*	Stuetensen	1993	2014	78	7	1.56	0.9	12.9	9.4	628	175 PIM, CAM	1780
BDF037_L	Schladen	1994	2014	6	31	1.42	2.3	7.9	9.6	708	165 UAN, U	1901
											17 CAS	
BDF039_L	Handeloh	1994	2014	86	5	1.63	1.7	17.3	9.5	824	142 CAN, AS, UAN	1901
											96 PIS, LFR	
BDF042_L	Fuhrberg	1995	2014	87	8	1.33	3.1	10.3	10.0	692	35 UAN, CAN, AS	1950
											54 LFR	< 1988 = pasture
BDF043_L	Oldershausen	1996	2014	2	18	1.38	1.0	9.1	9.3	841	178 U, UAN, AS	1784
											35 SS	
BDF045_L*	Riddagshausen	1994	2014	58	3	1.66	0.8	8.7	9.8	631	12 CAS	1759
BDF046_L*	Rodewald	1997	2014	31	31	1.42	2.0	9.9	10.3	687	114 CAN	1965
											48 PIS, LFR	
BDF047_L	Hiddestorf	1994	2014	2	12	1.37	0.9	8.7	10.0	705	141 UAN, U	1781
											28 CAM	
BDF049_L	Glissen	1994	2014	87	6	1.45	1.5	12.8	10.0	730	65 CAN	1955
											43 PIS	
BDF050_L	Bockheber	1994	2014	78	5	1.43	1.3	12.1	9.3	790	27 CAN	1776

										58 SHM	
BDF051_L	Reinhausen	1995	2014	12	49	1.43	1.3	6.1	9.3	628	169 UAN, AS	1784
											28 CAM	
BDF052_L	Suestedt	1996	2014	14	11	1.29	1.5	10.9	9.7	715	109 UAN	1899
											101 PIS	
BDF056_L*	Meinersen	1996	2014	96	2	1.42	1.3	18.9	9.8	625	62 CAN, UAN	1945
											6 LFR	
BDF057_L	Starkshorn	1997	2014	81	5	1.55	3.3	15.8	9.2	807	118 CAN, AS	1777
											53 POM	
BDF058_L	Kuingdorf	1996	2014	12	12	1.42	1.2	9.0	9.8	832	130 UAN, AS	1846
											62 CAS, PIS	
BDF059_L	Wuelferode	1997	2014	64	12	1.76	0.9	9.0	10.1	684	42 CP	1781
BDF063_L*	Meyenburg	1998	2014	4	36	1.3	2.4	10.3	10.0	818	213 U, AS	1980
											27 POM	1982 drainage
BDF064_L	Hohenzethen	1998	2014	79	7	1.44	1.0	12.9	9.7	574	145 UAN, CAN	1998
											401 CAS, LFR	
BDF065_L*	Juehnde	1998	2014	3	41	1.39	2.6	9.6	9.5	642	130 CAN, U	1775
											34 PIS, LFR	
BDF067_L*	Listrup	1998	2014	83	6	1.48	1.2	15.5	10.4	770	114 CAN, AN, UAN	1985
											116 PIS	
BDF069_L	Wendhausen	1998	2014	6	22	1.44	1.7	9.7	10.0	698	1.3 HOM	1845
BDF070_L	Sehlde	2001	2014	9	25	1.48	1.9	8.8	9.6	753	164 U, CAN	1845
											5 CAS	

* **sites used for model validation (n = 15).**

505   ** **predominantly applied fertilisers, N values for organic amendments were estimated; UAN = urea ammonium nitrate; U = urea; CAN = calcium ammonium nitrate; AS = ammonium sulfate; AN = ammonium nitrate; POM = poultry manure; PIS = pig slurry; PIM = pig manure; LFR = liquid fermentation residues; SS = sewage sludge; CAS = cattle slurry; CAM = cattle manure; SHM = sheep manure; CP = compost; HOM = horse manure.**

510 **Table A2. Description of the long-term experimental sites (n = 11) used for validation of simulated soil temperatures and water contents. SOC = soil organic carbon, MAT = mean annual temperature, MAP = mean annual precipitation.**

Site	Start date	End date	Sand [%]	Clay [%]	Bulk density [g cm⁻³]	SOC [%]	pH	C/N ratio	MAT [°C]	MAP [mm]
Bornim	2010	2016	75	6	1.65	1.0	6.6	9.8	8.7	503
Braunschweig	2008	2011	62	8	1.55	0.9	6.1	12.2	9.1	617
Dedelow	2010	2016	59	10	1.45	0.8	6.8	7.5	8.4	485
Göttingen	1992	1996	5	30	1.25	1.6	7.5	8.2	8.7	634
Hennef	1997	2000	8	22	1.31	0.9	6.5	7.5	10.3	837
Hohenheim	2010	2016	3	19	1.3	1.7	7.1	8.4	8.3	688
Hohenschulen	2010	2016	60	11	1.43	1.8	6.7	15.2	8.9	732
Merbitz	2010	2016	16	16	1.42	1.2	7.4	10.9	9.0	520
Rostock	1993	1998	70	7	1.5	0.9	6.5	1.1	9.3	593
Scheyern	1989	1996	22	23	1.25	1.9	6.1	9.4	7.4	833
Viehhausen	2011	2022	14	22	1.4	1.06	6.1	5.1	7.5	797

Appendix B. Revised environmental decomposition factors

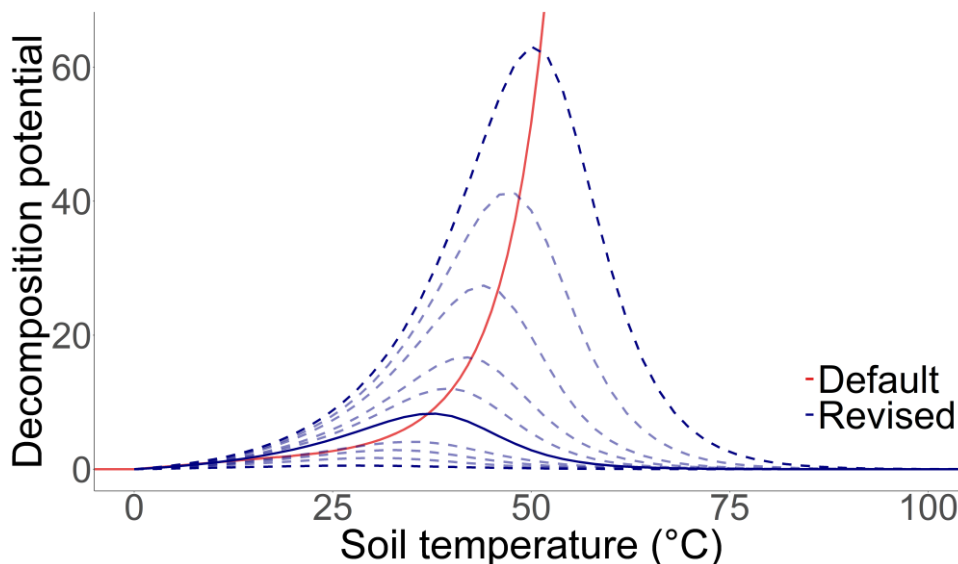
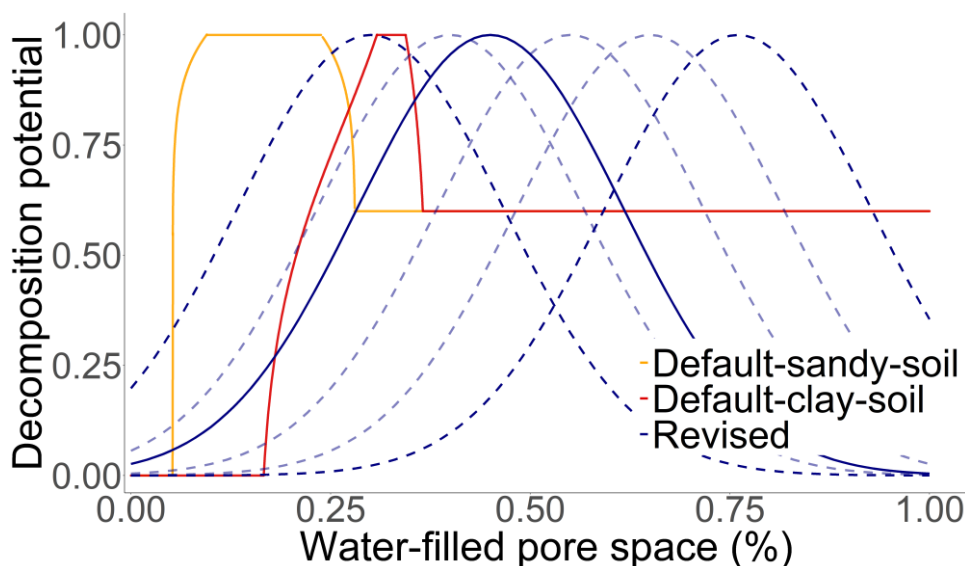
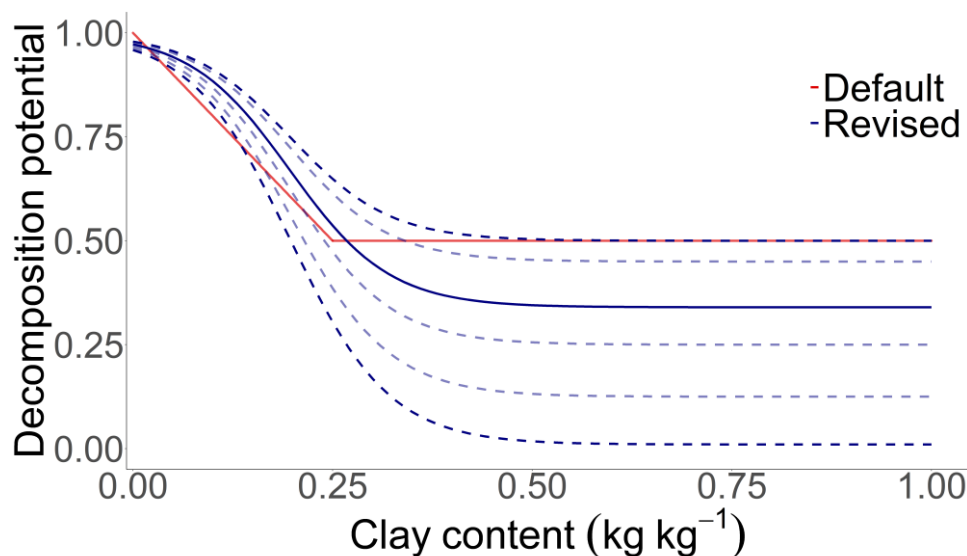
515 **Figure B1. The effect of soil temperature on the soil organic carbon decomposition rate. The solid red line describes the standard function that is integrated in the MONICA model. The solid blue line describes the revised function as the best fit curve to the observed data, while the dashed blue lines describe the possible properties of the equation according to the prior probability distribution of the parameters QTenFactor and TempDecOptimal.**

520 **Figure B2. The effect of soil moisture on the soil organic carbon decomposition rate. The solid red line describes the standard function for a clay soil (clay = 60 %, sand = 10 %, bulk density = 1.4, soil organic carbon = 2 %), while the solid yellow line describes the standard function for a sandy soil (clay = 4 %, sand = 80 %, bulk density = 1.4, soil organic carbon = 1 %). The solid blue line describes the revised function as the best fit curve to the observed data, while the dashed blue lines describe the possible properties of the equation according to the prior probability distribution of the parameter MoistureDecOptimal.**

525

Figure B3. The effect of clay on the soil organic carbon decomposition rate. The solid red line describes the standard function that is integrated in the MONICA model. The solid blue line describes the revised function as the best fit curve to the observed data, while the dashed blue lines describe the possible properties of the equation according to the prior probability distribution of the parameter LimitClayEffect.

530

Appendix C. Morris elementary effects screening method

Table C1. Mathematical formulation of the sensitivity measures of the Morris elementary effects screening method, where r is the number of input factors and EE is the elementary effect of each input.

Description	Equation	Properties
Ranking of the input factor (Campolongo et al., 2007)	$\mu_j^* = \dfrac{1}{r}\sum_{i=1}^{r}|EE_{ij}|$	Mean estimates of the distribution of the absolute values of each input
Standard deviation of the input factor (Morris, 1991)	$\sigma_j = \sqrt{\dfrac{1}{r}\sum_{i=1}^{r}\left(EE_{ij} - \dfrac{1}{r}\sum_{i=1}^{r}(EE_{ij})\right)^2}$	Standard deviation estimates of the absolute values of each input

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Appendix D. Metropolis-Hastings algorithm

Table D1. Mathematical formulation of the Metropolis-Hastings algorithm and the log-Laplace likelihood function, where θ is the parameter vector, Y the dataset used for calibration, t the number of iterations, σ the standard deviation inferred jointly from the error between observation and simulation, ω the model input, ƒ(ωi;θi) the model run with the parameter vector θ, α the acceptance ratio and δ a random vector from a multivariate distribution.

Description	Equation	Properties
Log-Laplace likelihood function (Vrugt, 2016)	$\mathcal{L}(\theta\|Y) = -\sum_{i=1}^{t} log\left(\frac{1}{2\sigma_i}\right) - \sum_{i=1}^{t}\left(\frac{e^{Y-f(\omega_i;\theta_i)}}{\sigma_i}\right)$	Continuous probability distribution, assumes all error residuals equally
Metropolis-Hastings algorithm (Metropolis et al., 1953; Hastings, 1970)	$\theta^* = \theta_{i-1} + \delta$ $\alpha = \sum_{i=1}^{t}\frac{\mathcal{L}(\theta^*\|Y)}{\mathcal{L}(\theta_{i-1}\|Y)} = \sum_{i=1}^{t}\frac{\mathcal{L}(Y\|\theta^*)\mathcal{L}(\theta^*)}{\mathcal{L}(Y\|\theta_{i-1})\mathcal{L}(\theta_{i-1})}$	Accept θ^* if $\alpha \geq u$ where u is a uniform random value between 0 and 1, otherwise if $\alpha < u$ reject θ^*

Appendix E. Model evaluation methods

Table E1. Quantitative methods used for data and model evaluation, where S is the predicted value, O is the observed value, n is the number of cases and k is the number of divided subgroups from n.

Statistical method	Equation	Value range	Properties
Kendall rank correlation coefficient (Kendall, 1938; 1948)	$T = \dfrac{\sum_{i<j}(sgn(S_i - S_j) \cdot sgn(O_i - O_j))}{\frac{n(n-1)}{2}}$	−1 - 1	If $T > 0$ positive association between two variables, otherwise if $T = 0$ no association, or if $T < 0$ negative association
Mean absolute error (Willmott and Matsuura, 2005)	$MAE = \sqrt{\dfrac{1}{n}\sum_{i=1}^{n}(S_i - O_i)}$	> 0	If $MAE = 0$ perfect agreement between two variables, otherwise if $MAE > 0$ decreasing agreement
Nash-Sutcliffe model efficiency coefficient (Nash and Sutcliffe, 1970)	$NSE = 1 - \dfrac{\sum_{i=1}^{n}(S_i - O_i)^2}{\sum_{i=1}^{n}(O_i - \bar{O})^2}$	< 1	If $NSE = 1$ perfect agreement between two variables, otherwise if $NSE < 1$ decreasing agreement

540

545

Appendix F. Results of the Morris elementary effects screening method

550 **Table F1. Results of the Morris method depicting the effects of 51 variables on soil carbon changes. The variables are subdivided into site, management, crop and fertiliser-specific and carbon turnover factors. The parameters *CorgContent* (C content of crop residues and organic fertilisers), *MoistureDecOptimal* (optimal water-filled pore space for decomposition), *TempDecOptimal* (optimal temperature for decomposition) and *QTenFactor* (Q$_{10}$ temperature coefficient) were included as modifications to the model. The overall importance of an input factor is specified by two sensitivity measures: the mean (μ*) and the standard deviation of the elementary effect (σ) of each parameter. Higher values in the two sensitivity measures imply a greater importance of the variable on 555 changes in soil carbon stocks. The uncertainty is presented as the standard deviation between the results of three separate Morris elementary effects screening runs.**

Variables	Absolute mean of elementary effect μ* [g C kg^{-1}]	Standard deviation of elementary effect σ [g C kg^{-1}]
Site factor		
Temperature	11.43 ± 0.46	15.9 ± 0.58
SoilOrganicCarbon	4.72 ± 0.14	8.08 ± 0.26
SoilBulkDensity	4.44 ± 0.14	4.96 ± 0.04
Daily_Precipitation	4.16 ± 0.28	8.55 ± 0.71
FieldCapacity	3.7 ± 0.07	7.15 ± 0.49
PoreVolume	2.46 ± 0.13	4.85 ± 0.49
Global_Radiation	1.48 ± 0.05	3.37 ± 1.28
Clay	0.62 ± 0.04	1.4 ± 0.11
CN	0.53 ± 0.05	2.77 ± 0.77
PermanentWiltingPoint	0.24 ± 0.01	0.87 ± 0.12
Sand	0.19 ± 0.02	0.54 ± 0.18
pH	0.0 ± 0.0	0.02 ± 0.02
Wind_Speed	0.0 ± 0.0	0.0 ± 0.0
Relative_Humidity	0.0 ± 0.0	0.0 ± 0.0
Sceleton	0.0 ± 0.0	0.0 ± 0.0
Management factor		
OrganicFertilization	18.36 ± 0.68	16.43 ± 0.44
MineralFertilization	1.03 ± 0.05	3.39 ± 0.26
Irrigation	0.37 ± 0.03	1.01 ± 0.12
Tillage	0.22 ± 0.03	1.42 ± 0.42
Crop and fertiliser factor		
AOM_DryMatterContent	17.75 ± 0.32	16.25 ± 0.39
CorgContent	7.74 ± 0.41	8.67 ± 0.3
PartAOM_to_AOM_Slow	2.82 ± 0.11	4.78 ± 0.11
PartAOM_to_AOM_Fast	2.41 ± 0.12	4.23 ± 0.15
AOM_SlowDecCoeffStandard	1.62 ± 0.14	4.84 ± 0.5

AOM_FastDecCoeffStandard	1.51 ± 0.07	4.94 ± 0.79
PartAOM_Slow_to_SMB_Slow	0.57 ± 0.06	2.05 ± 0.44
CN_Ratio_AOM_Fast	0.44 ± 0.06	1.89 ± 0.57
PartAOM_Slow_to_SMB_Fast	0.44 ± 0.04	1.18 ± 0.39
CN_Ratio_AOM_Slow	0.43 ± 0.06	2.53 ± 0.29
AOM_NH4Content	0.18 ± 0.03	0.95 ± 0.14
AOM_NO3Content	0.18 ± 0.03	0.72 ± 0.15
Carbon turnover factor		
QTenFactor	4.77 ± 0.08	8.56 ± 0.1
MoistureDecOptimal	4.02 ± 0.15	6.71 ± 0.15
SOM_SlowDecCoeffStandard	3.73 ± 0.18	7.11 ± 0.14
SOM_FastUtilizationEfficiency	2.5 ± 0.1	4.44 ± 0.33
AOM_FastUtilizationEfficiency	2.06 ± 0.03	4.87 ± 0.63
PartSOM_Fast_to_SOM_Slow	1.96 ± 0.04	3.5 ± 0.04
SOM_FastDecCoeffStandard	1.75 ± 0.09	3.05 ± 0.16
SMB_UtilizationEfficiency	1.63 ± 0.11	3.59 ± 0.24
AOM_SlowUtilizationEfficiency	1.34 ± 0.15	3.35 ± 0.16
PartSMB_Fast_to_SOM_Fast	1.25 ± 0.09	2.37 ± 0.36
PartSMB_Slow_to_SOM_Fast	0.94 ± 0.03	2.24 ± 0.25
SOM_SlowUtilizationEfficiency	0.81 ± 0.01	1.65 ± 0.07
SMB_FastDeathRateStandard	0.65 ± 0.02	2.0 ± 0.09
SMB_SlowDeathRateStandard	0.53 ± 0.08	1.69 ± 0.49
SMB_FastMaintRateStandard	0.38 ± 0.04	1.29 ± 0.48
SMB_SlowMaintRateStandard	0.37 ± 0.02	0.92 ± 0.07
LimitClayEffect	0.25 ± 0.02	0.63 ± 0.13
PartSOM_to_SMB_Slow	0.1 ± 0.0	0.13 ± 0.0
PartSOM_to_SMB_Fast	0.04 ± 0.01	0.13 ± 0.13
TempDecOptimal	0.0 ± 0.01	0.1 ± 0.17

Appendix G. Gelman and Rubin shrink factor

Figure G1. Evolution of the Gelman and Rubin shrink factor. The 50 % (solid line) and 97.5 % (dashed red line) quantiles show the convergence of the calibration.

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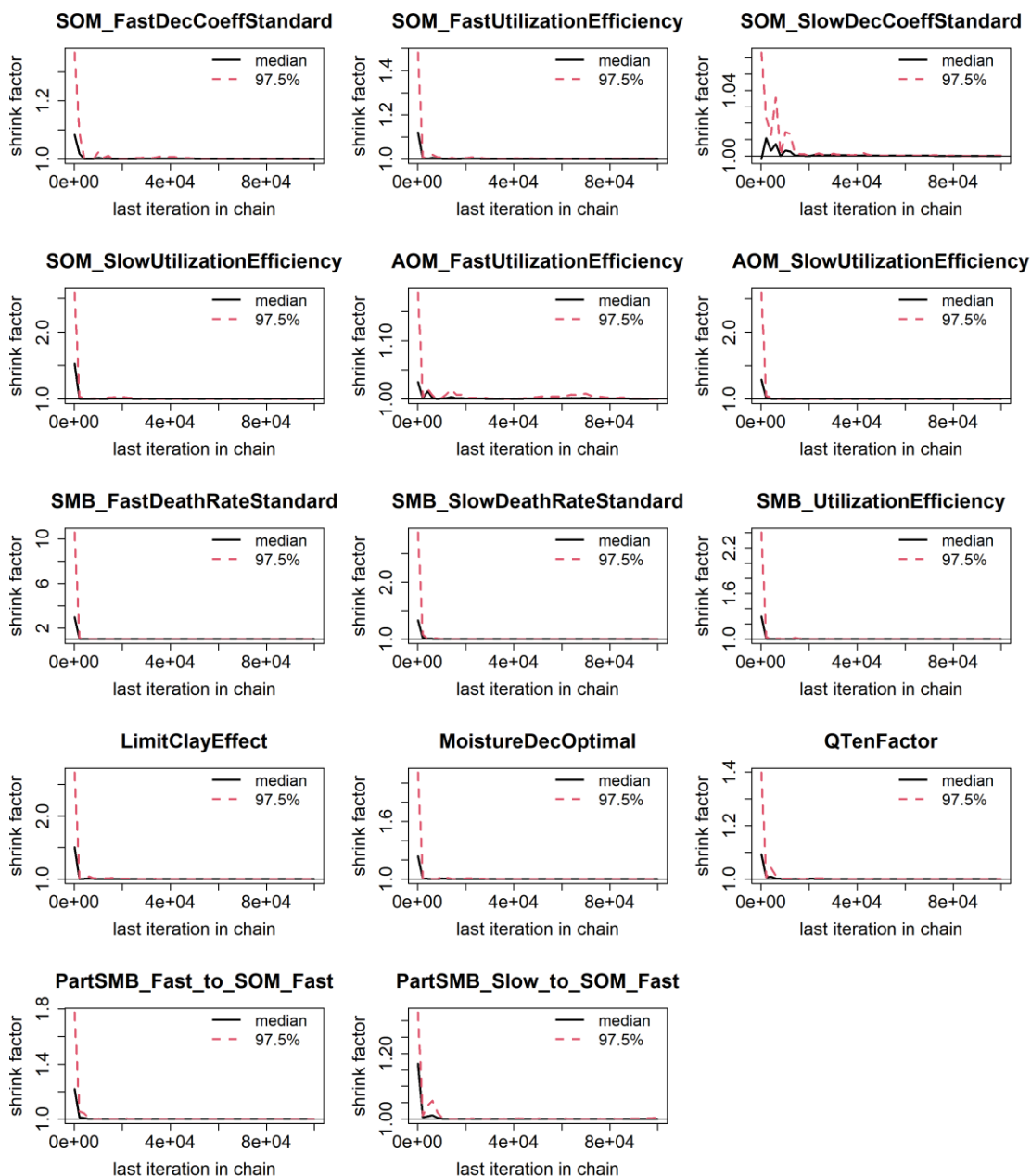
Appendix H. Correlations between C turnover parameters

Figure H1. Correlation matrix quantifying the statistical relationship between the prior probability distributions of the SOM submodel parameters.

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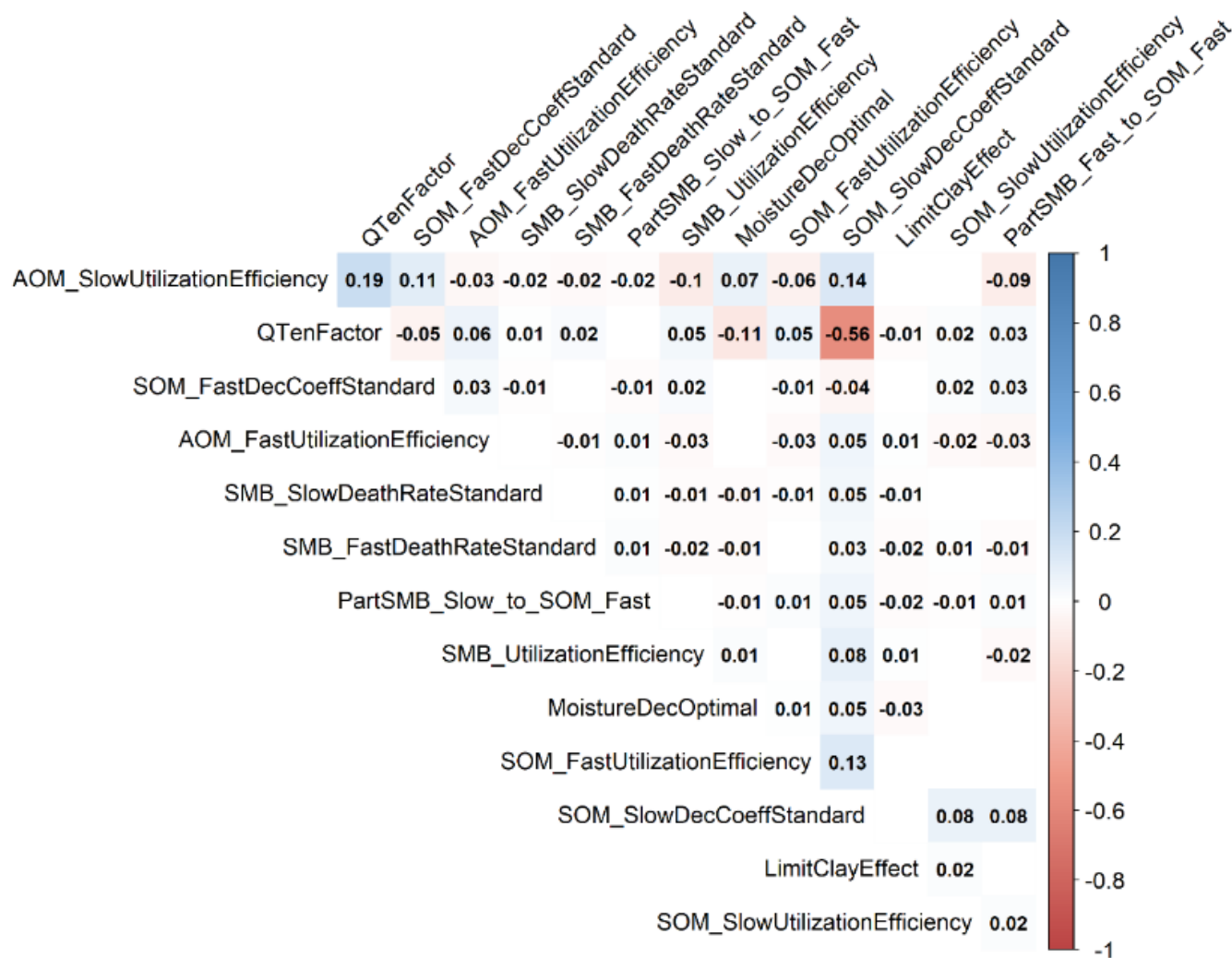

Appendix I. Model performance

Table I1. Statistics describing the performance of the default (MONICA$_{def}$), the modified, uncalibrated (MONICA$_{mod}$) and the modified, calibrated (MONICA$_{cal}$) soil carbon submodel in MONICA in comparison with five C turnover models for 15 validation sites. The first ten years were excluded from the analysis (n = 70). The statistics depicted are Kendall's rank correlation coefficient (T), Nash-Sutcliffe model efficiency coefficient (NSE), mean absolute error (MAE) and slope and intercept for characterising the relationship between measured and simulated C stocks.

Description	T	NSE	MAE (Mg C ha^{-1})	Slope	Intercept
MONICA$_{def}$	0.85***	0.93	5.41	1.07	0.54
MONICA$_{mod}$	0.81***	0.94	4.97	1.08	-2.14
MONICA$_{cal}$	0.83***	0.97	3.81	1.02	-0.41
RothC	0.83***	0.97	3.68	1.04	-0.57
C-TOOL	.85***	0.95	4.63	1	2.68
ICBM	.87***	0.95	3.6	0.91	4.86
CENTURY	.84***	0.95	4.32	0.88	4.37
CCB	.8***	0.93	4.95	0.89	3.29

*** p < 0.001, ** p < 0.01, * p < 0.05

Appendix J. Model performances

Figure J1. Performance of the default MONICA model for each individual site. The first 10 measured years were excluded.

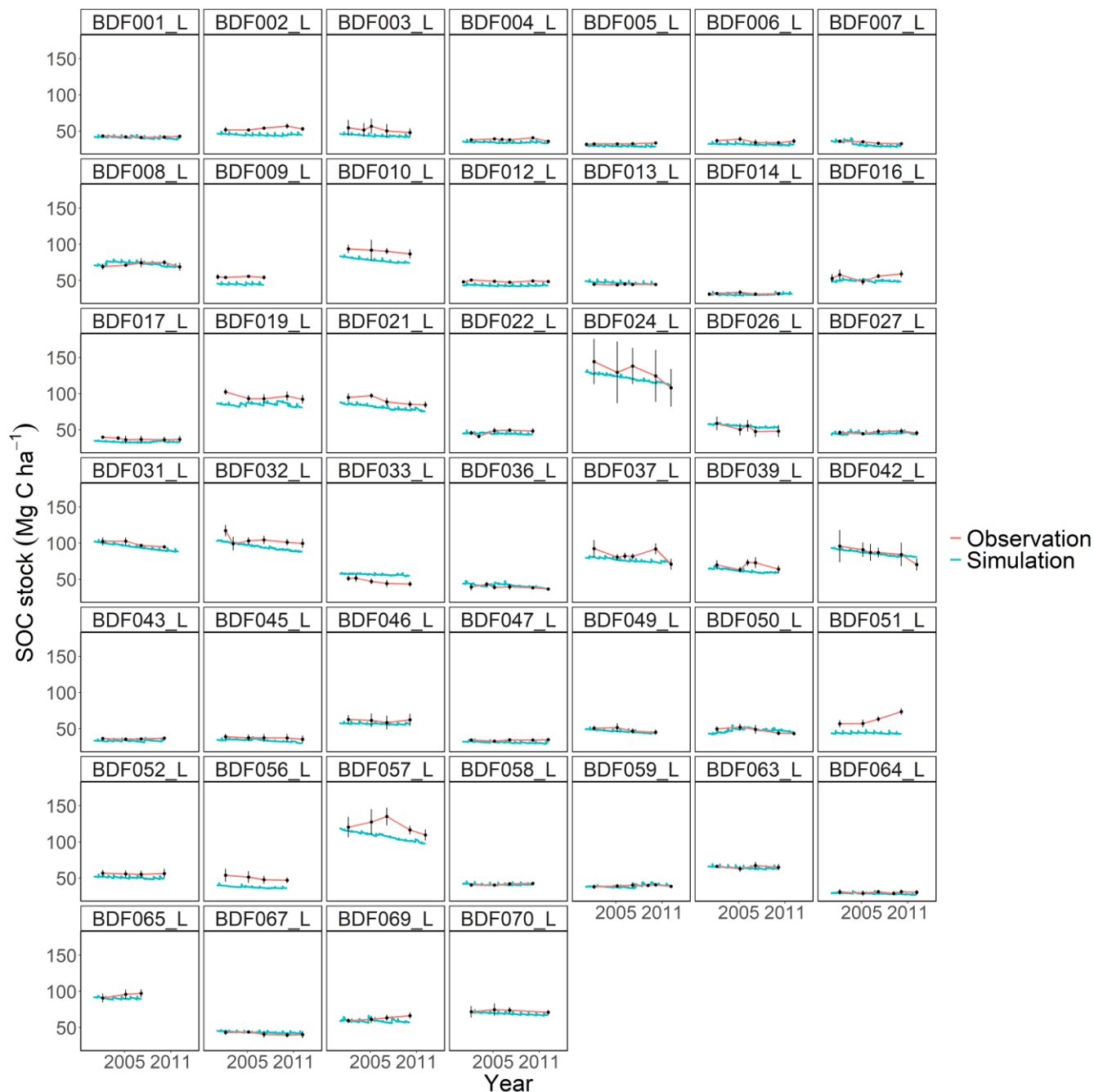



Figure I2. Performance of the modified and calibrated MONICA model for each individual site. The first 10 measured years were excluded.

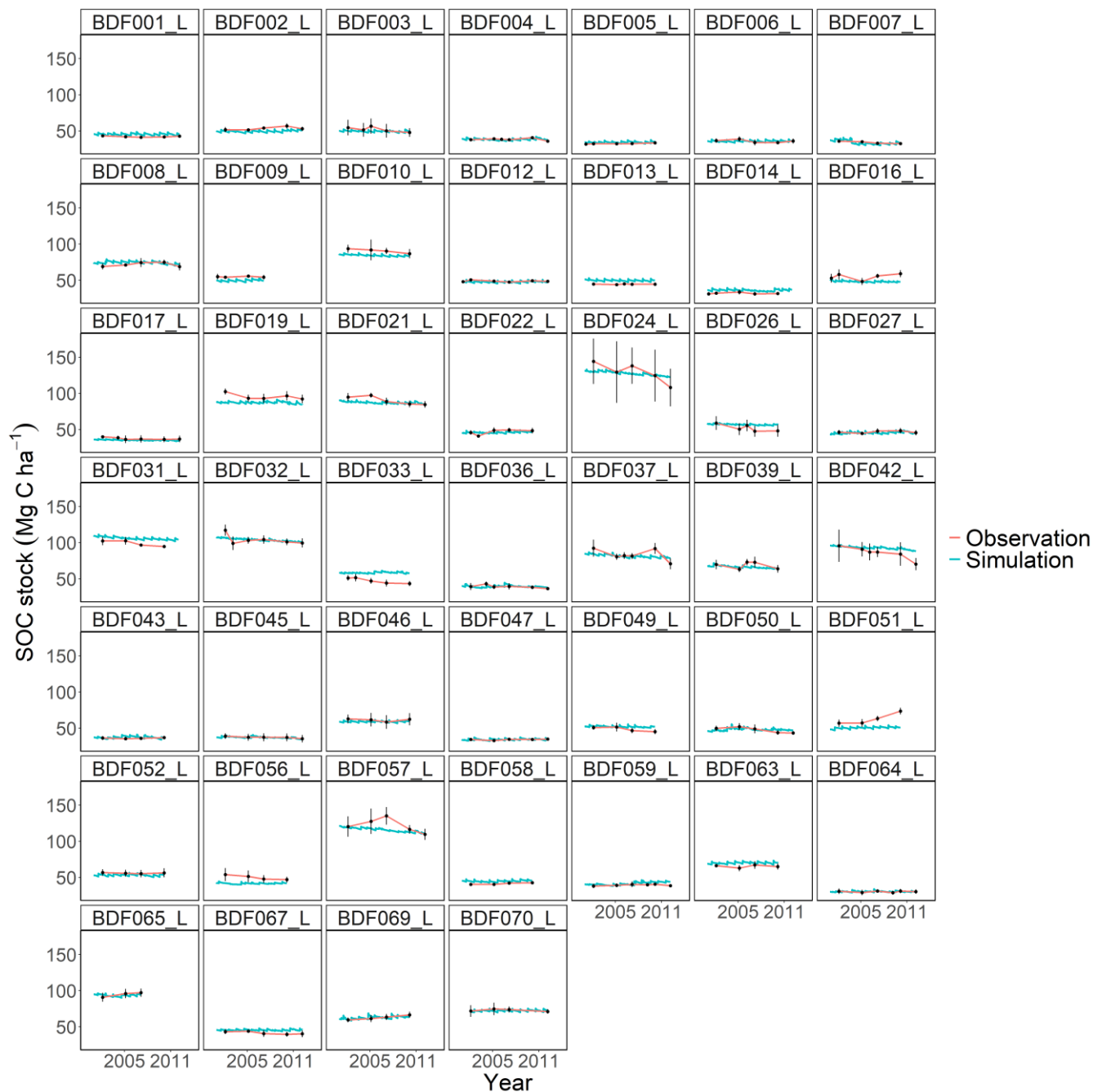
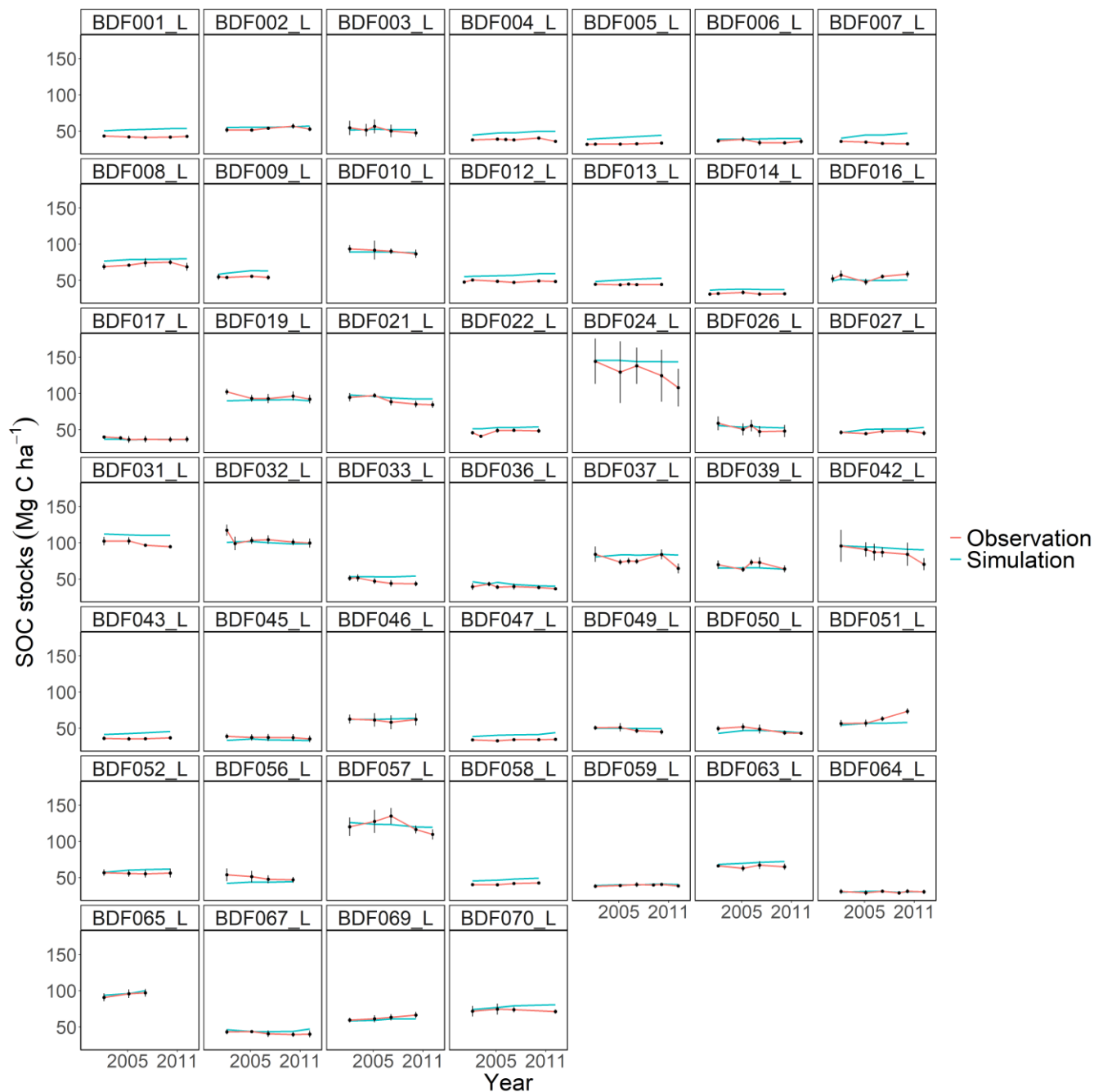



Figure I3. Performance of the CCB model for each individual site. The first 10 measured years were excluded.

585 **Figure I4. Performance of the CENTURY model for each individual site. The first 10 measured years were excluded.**

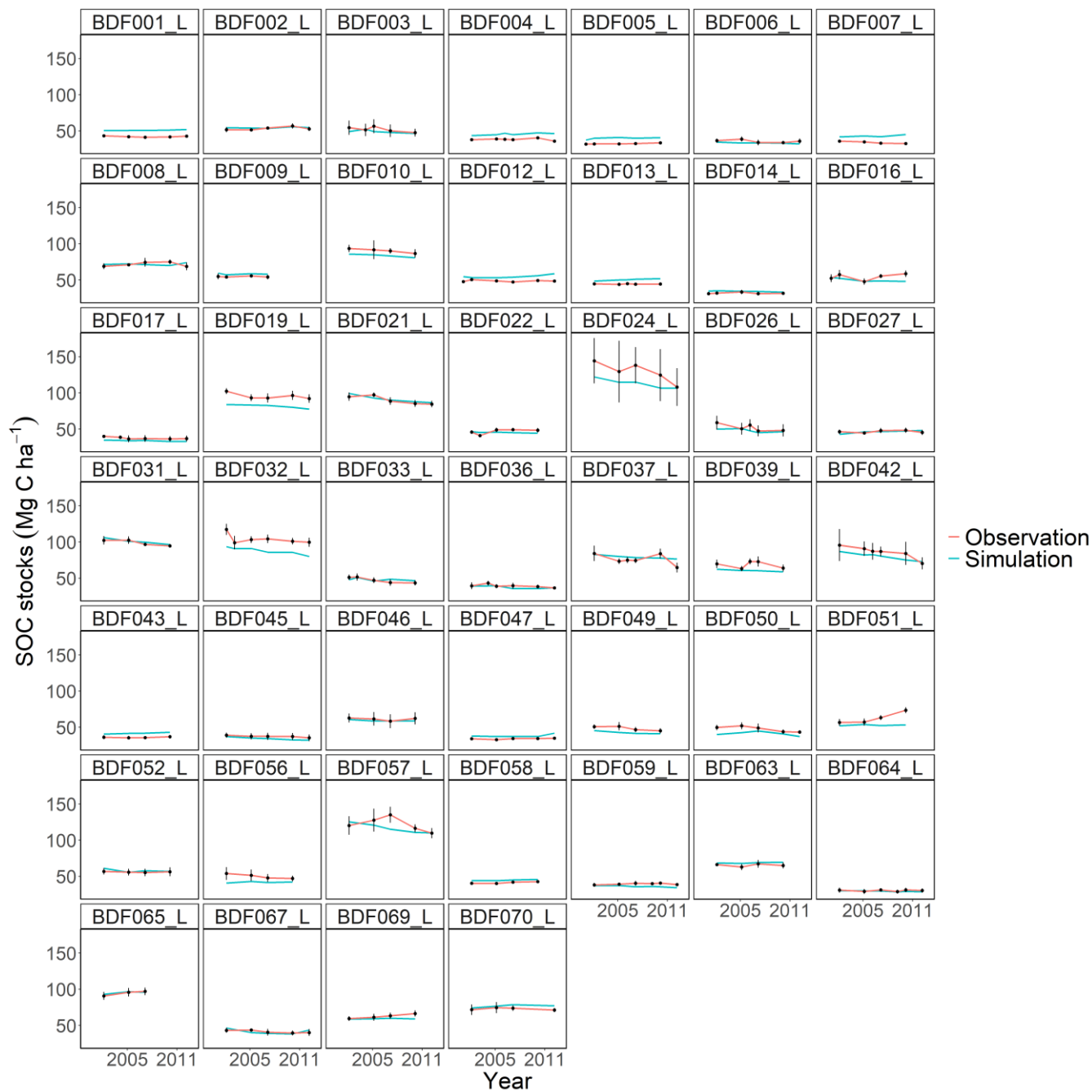
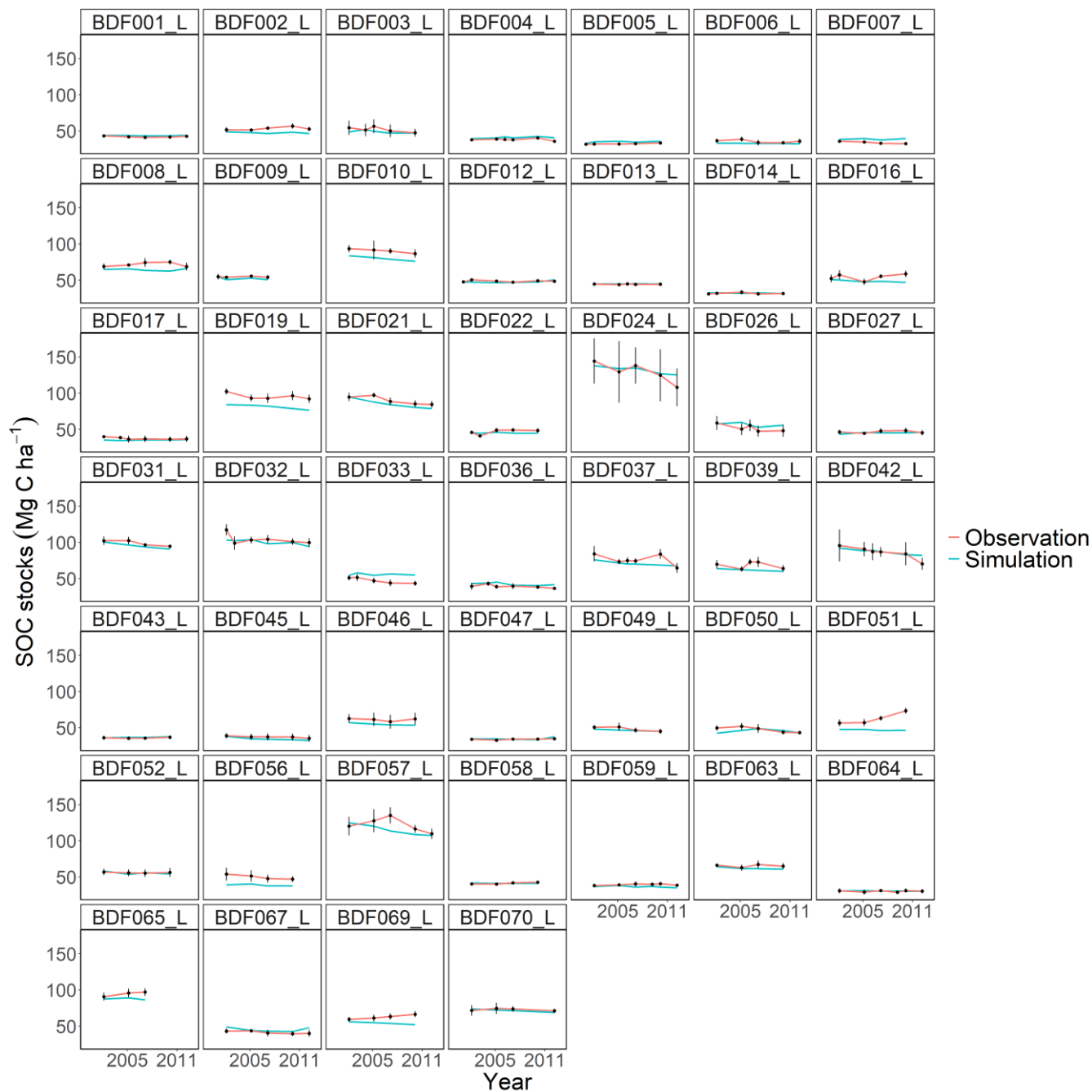



Figure I5. Performance of the C-TOOL model for each individual site. The first 10 measured years were excluded.

595 **Figure I6. Performance of the ICBM model for each individual site. The first 10 measured years were excluded.**

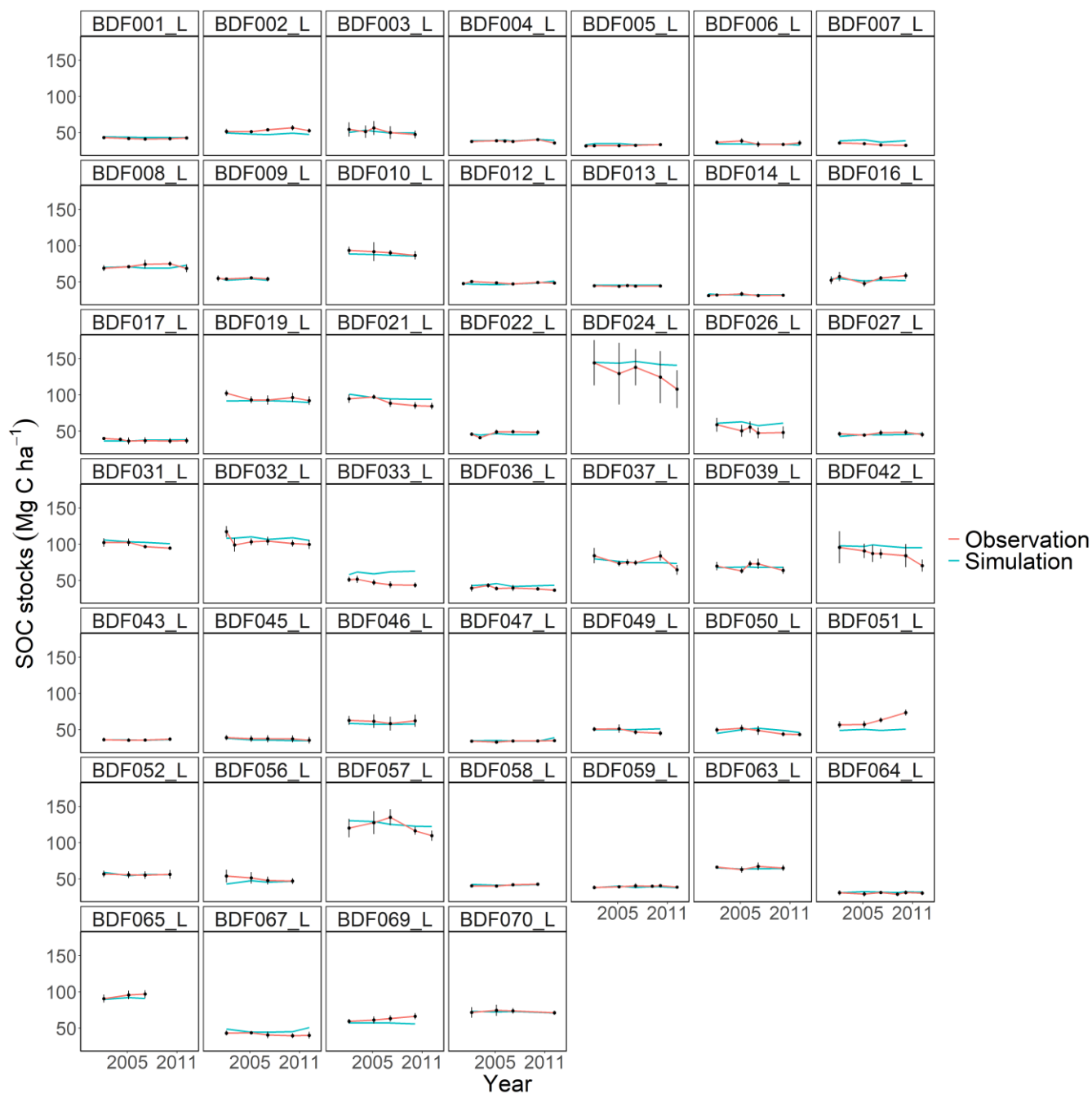
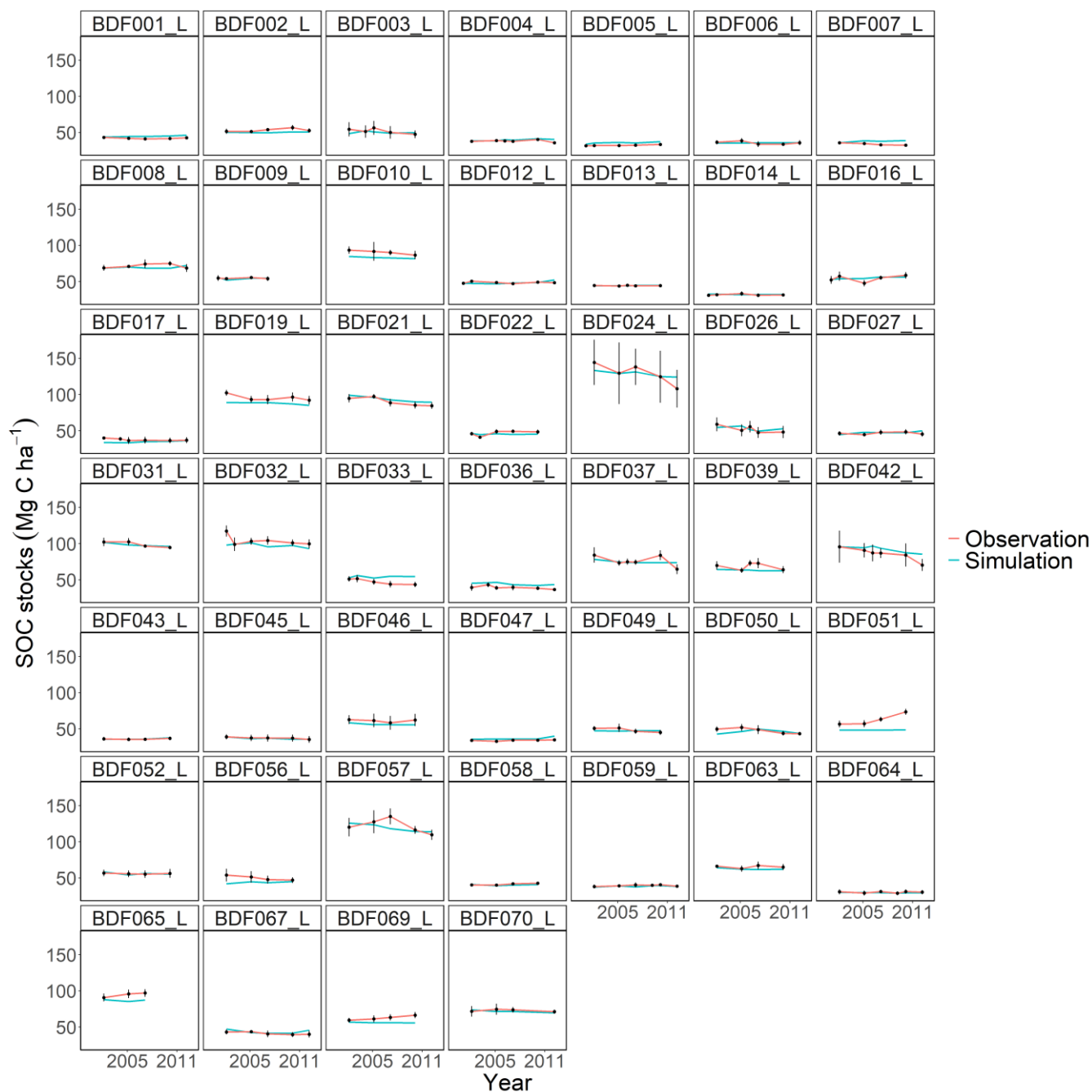



Figure I7. Performance of the RothC model for each individual site. The first 10 measured years were excluded.

Data availability

600 The described long-term soil monitoring data is available from the State Office for Mining, Energy and Geology of the Lower Saxony state government. Restrictions apply to the availability of these data, which were used under license for this study. Data are available from the State Office for Mining, Energy and Geology of the Lower Saxony state government. Additional data is freely available under the CREATIVE COMMONS ATTRIBUTION 4.0 INTERNATIONAL LICENSE.

Code availability

605 Code for model analysis is available at https://doi.org/10.6084/m9.figshare.22652047.v1.

Software availability

R 4.1.0: A language and environment for statistical computing; Copyright © 1990 W.N. Venables Copyright© 1992 W.N. Venables & D.M. Smith Copyright© 1997 R. Gentleman & R. Ihaka Copyright© 1997, 1998 M. Maechler Copyright© 1999-
610 2021 R Core Team-Copyright© 2021 The R Foundation for Statistical Computing; Hardware: Any PC; Software requirement: It compiles and runs on a wide variety of UNIX platforms and similar systems (including FreeBSD and Linux), Windows and MacOS. Free Software under the terms of the Free Software Foundation's; Language: C, C++ and Fortran; Size: installation package for Windows 64 bit: 85.4 MB.

MONICA 3.3.1; Claas Nendel, Xenia Specka & Michael Berg-Mohnicke; https://github.com/zalf-
615 rpm/monica/releases/tag/3.3.1.208 Contact information on the website; Hardware: Any PC; Software requirement: Windows, MacOS or Linux; Cost: free; Size: installation for Windows 64 bit: 6.49 MB.F

Author's contributions

Konstantin Aiteew prepared the data, developed and implemented the modifications and executed the MONICA model simulations. He developed the scripts for model performance evaluation, sensitivity analysis and calibration and also wrote
620 the manuscript.

Jarno Rouhiainen gave advice in developing the sensitivity analysis algorithm, helped interpret MONICA's source code and supported the drafting of the manuscript.

Claas Nendel advised on the execution of the MONICA model and supported in drafting of the manuscript.

René Dechow prepared the simulation of the C turnover models CCB, CENTURY, C-TOOL, ICBM and RothC. He also gave
625 advice, material and support in drafting of the manuscript.

Competing interests

The authors declare that they have no conflict of interest.

Acknowledgements

We are very grateful to the state government of Lower Saxony and the State Office for Mining, Energy and Geology for
630   collecting and providing the data used in this model evaluation and optimisation. We are also grateful to Prof. Dr. Claas Nendel
and Michael Berg-Mohnicke for their support regarding questions about the MONICA model. We also thank our colleagues
PD. Dr. Axel Don, Dr. Christopher Poeplau, Dr. Felix Seidel, Dr. Florian Schneider, Dr. Julia Schröder and Dr. Balazs Grosz
for their support and comments on the manuscript. We would specifically like to thank the German Federal Ministry of Food
and Agriculture and the Agency for Renewable Resources for financially supporting our work.

635

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