

# Supplementary Information

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47 **Figure S16.** Average diurnal profiles for SO<sub>2</sub>, Sh-IndOA and the sum of industrial and shipping factors  
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49 **Figure S17.** Residuals values of WLS models for **(a)** OP<sub>AA</sub> and **(b)** OP<sub>DTT</sub>. An outlier point (07/19/2018  
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53 distribution.

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55 *List of tables :*

56 **Table S1.** List of elements measured with the Xact, their respective MDLs and the percent of  
57 measurements above the MDL.

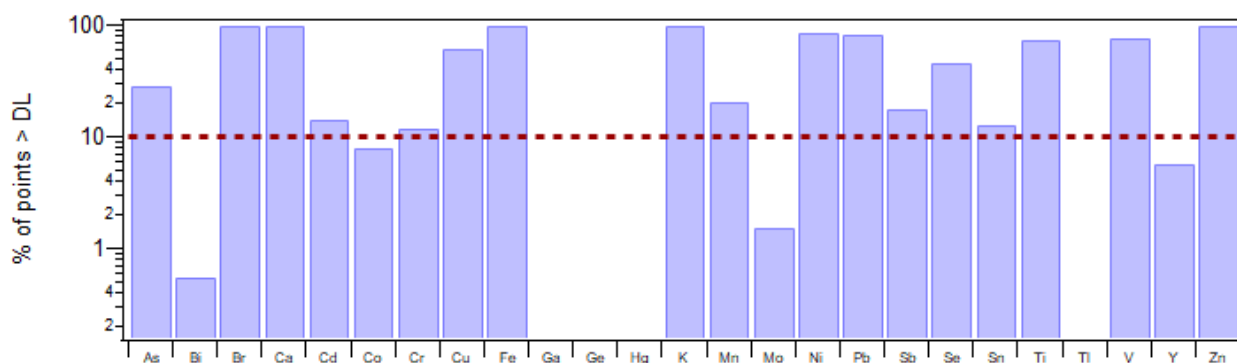
58 **Table S2.** Summary of statistics for the different PMF tests carried out on the WFP datasets of metals.  
59 18 downweight conditions were tested for the PMF inputs. The matrix including a S2N downweight  
60 and errors below MDLs downweighted with  $\alpha_i = 2 \times r_{p95}$  (test n°12) was selected as final inputs.

61 **Table S3.** Intrinsic  $OP_{AA}$  and  $OP_{DTT}$  ( $OP_m$ ) provided by weighted robust linear regression with an M-  
62 estimator expressed in  $nmol \cdot min^{-1} \cdot \mu g^{-1}$  of sources provided by **(a)**  $PMF_{organics}$  (scenario 1) and **(b)**  
63  $PMF_{metals}$  (scenario 2) over the OP sampling campaign ( $n = 90$ ). Values are the mean  $\pm$  standard deviation  
64 from bootstraps runs for both OP assays. The model parameters  $R^2_{adjusted}$  and Pearson's correlation  
65 between model OP and observed OP are mentioned on the right.

66 **Table S4.** Factors identification for the  $PMF_{PM1}$  analyses between 5 and 12 factors. The green cells  
67 represent the base case identification for the related factors. The remaining undefined factors for each  
68 solution corresponded to mixed profiles not attributed to a specific source. The red squares are the base  
69 cases used as reference profile constraints.

70 **Table S5.** Pearson's correlation coefficients between  $OP_{vAA}$  and  $OP_{vDTT}$  to the PM sources identified by  
71  $PMF_{PM1}$  model.

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73

74 **Figure S1.** Percent of values above the MDL for each element measured with the Xact.

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Elements	MDL (ng.m <sup>-3</sup> )	Values > MDL (%)
As	0.11	29.3
Bi	0.23	0.6
Br	0.18	99.8
Ca	0.52	99.5
Cd	4.4	14.4
Co	0.24	8.1
Cr	0.2	12.0
Cu	0.14	63.3
Fe	0.3	100
Ga	0.1	0
Ge	0.1	0
Hg	0.21	0.2
K	2	100
Mn	0.25	20.6
Mo	0.84	1.5
Ni	0.17	86.6
Pb	0.22	84.8
Sb	9	18.0
Se	0.14	46.4
Sn	7.1	12.8
Ti	0.28	76.0
Tl	0.2	0
V	0.21	76.7
Y	0.48	5.8
Zn	0.12	100

78 **Table S1.** List of elements measured with the Xact, their respective MDLs and the percent of  
 79 measurements above the MDL.

80

## 81 Error matrix downweights for PMF<sub>metals</sub>

82 Polissar et al. (1998) (Polissar et al., 1998a) first introduced an uncertainty of  $5/6 \times MDL$  for data below  
83 MDL (set to  $MDL/2$ ). The purpose was to provide relative errors for these values 2 to 5 times greater  
84 than the maximum relative errors of the data exceeding the MDLs. Here, several uncertainties values  
85 were tested for data below MDL by conducted a panel of PMF runs with 2 to 8 factors. The errors were  
86 calculated by applying a downweight coefficient ( $\alpha$ ) to the previous formula from Polissar et al. (1998)  
87 (Polissar et al., 1998a):

$$88 \sigma_{i,j} = \alpha \times \frac{5}{6} MDL_i \text{ if } x_{i,j} < MDL_i \quad (S1)$$

89 For all the elements  $i$ ,  $\alpha$  was set to 6, 10 and 14 in order to obtain a ratio of 2, 3.5 and 5, respectively,  
90 with the maximum relative error found in the dataset, i.e. 476% for Sn (the value corresponds to the 95<sup>th</sup>  
91 percentile instead of the max value to avoid outlier effects). Another test consisted in applying a  
92 dependant  $\alpha$  based on the maximum relative error (95<sup>th</sup> percentile) for each element  $i$  ( $r_{p95}$ ):

$$93 \alpha_i = 2 \times r_{p95} \quad (S2)$$

94 Where 2 was used to determine the same ratio between the relative error of data below the MDLs from  
95 Polissar et al. (1998) equation (167 %, considering the  $\frac{5}{6} MDL / \frac{1}{2} MDL$  calculation) and the maximum  
96 relative error for the data greater than the MDLs (50%) found in Polissar et al. (1998) (Polissar et al.,  
97 1998a) dataset. A last test was performed with  $\alpha=1$  (i.e. no downweight) for the comparison. Each PMF  
98 analysis was also conducted with and without 1/S2N downweight (Visser et al., 2015). The tests were  
99 performed on the WFP dataset and the results were synthetized in Table S2. Here we focus on the 5F-  
100 solutions results as they resolved unmixed factors and represented a statistically relevant number of  
101 factor (see section 2.4.2 in the main text).

102 For all PMF solutions, applying the 1/S2N downweight provided lower scaled residuals as shown by  
103 the narrower width of fits. The solutions with  $\alpha=1$  (i.e. no errors downweight for data  $<MDLs$ ) were  
104 discarded due to less satisfactory mass reconstructions and residuals and higher average unexplained  
105 variations. The unexplained variation is a dimensionless quantity which indicates how much variation  
106 (in time or in each variable) is not explained by the factors (Canonaco et al., 2013). Thus, the unexplained  
107 variation of the  $i^{\text{th}}$  point for the factor  $k^{\text{th}}$  is:

$$108 UEV_{ik} = \frac{\sum_{j=1}^m (|e_{ij}| / \sigma_{ij})}{\sum_{j=1}^m ((\sum_{k=1}^p |g_{ik} \cdot f_{kj}| + e_{ij}) / \sigma_{ij})} \quad (S3)$$

109  $UEV$  is further calculated for data with  $S2N > 2$  ( $UEV_{\text{real}}$ ) or for noisy data ( $UEV_{\text{noisy}}$ ).

110 The remaining tests gave comparable explained variations, mass reconstitutions and residuals. The  
 111 uncertainties calculated with  $\alpha_i = 2 \times r_{p95}$  (test n°12 in Table S2) were finally selected as error inputs  
 112 for the data below the MDLs since this solutions resolved 5 unmixed factors with the best mean and  
 113 median diurnal patterns for each identified source.

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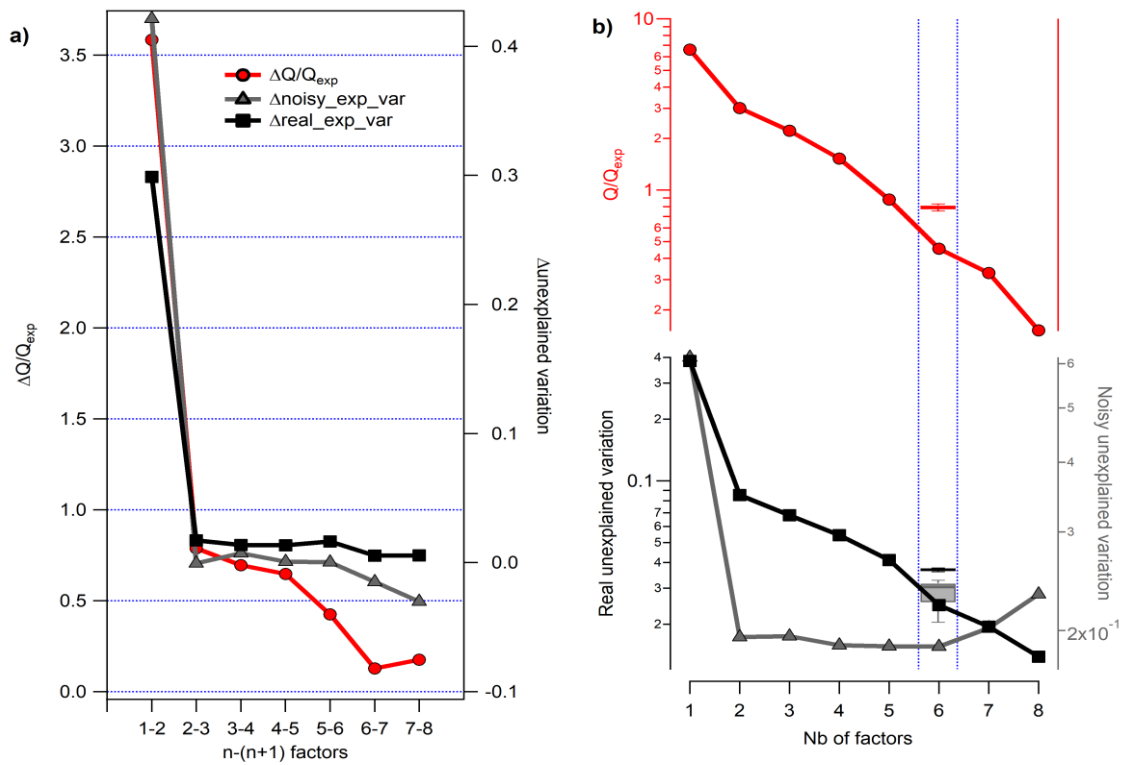
Tests	N°	Unexplained Variations			$\Sigma_{\text{factors vs } \Sigma_{\text{metals}}}$		Sc residuals	
		F5 EV_Noise	F5 EV_Real	F5 EV_Sum	F5 Slope	F5 R <sup>2</sup>	F5 center	F5 width
noDW_noS2N	1	0.215	0.049	0.264	0.987	0.958	0.094	0.337
noDW_S2N	2	0.221	0.042	0.263	0.987	0.959	0.059	0.201
DW6_ALL_noS2N	3	0.187	0.048	0.235	0.989	0.973	-0.023	0.137
DW6_ALL_S2N	4	0.203	0.033	0.236	1.004	0.985	-0.002	0.014
DW6_SPEC_noS2N	5	0.187	0.041	0.228	1.000	0.979	-0.010	0.089
DW6_SPEC_S2N	6	0.203	0.033	0.236	1.003	0.986	-0.002	0.014
DW10_ALL_noS2N	7	0.187	0.040	0.227	1.010	0.993	-0.020	0.074
DW10_ALL_S2N	8	0.201	0.033	0.234	1.015	1.000	-0.002	0.005
DW10_SPEC_noS2N	9	0.187	0.040	0.227	1.010	0.993	-0.020	0.073
DW10_SPEC_S2N	10	0.201	0.033	0.234	1.014	1.000	-0.002	0.005
Roll_DW_ALL_noS2N	11	0.188	0.041	0.229	1.007	0.988	-0.013	0.091
Roll_DW_ALL_S2N	12	0.203	0.034	0.237	1.010	0.995	-0.001	0.005
Roll_DW_SPEC_noS2N	13	0.187	0.041	0.228	1.006	0.988	-0.013	0.091
Roll_DW_SPEC_S2N	14	0.203	0.034	0.237	1.011	0.995	-0.002	0.011
DW14_ALL_noS2N	15	0.189	0.040	0.229	1.016	1.000	-0.023	0.063
DW14_ALL_S2N	16	0.204	0.032	0.236	1.017	0.999	-0.002	0.005
DW14_SPEC_noS2N	17	0.189	0.040	0.229	1.016	1.000	-0.023	0.063
DW14_SPEC_S2N	18	0.203	0.033	0.236	1.023	1.000	-0.001	0.003

115 **Table S2.** Summary of statistics for the different PMF tests carried out on the WFP datasets of metals.  
 116 18 downweight conditions were tested for the PMF inputs. The matrix including a S2N downweight  
 117 and errors below MDLs downweighted with  $\alpha_i = 2 \times r_{p95}$  (test n°12) was selected as final inputs.

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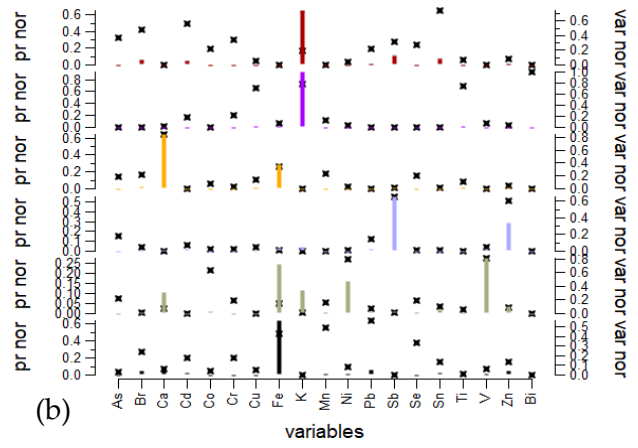
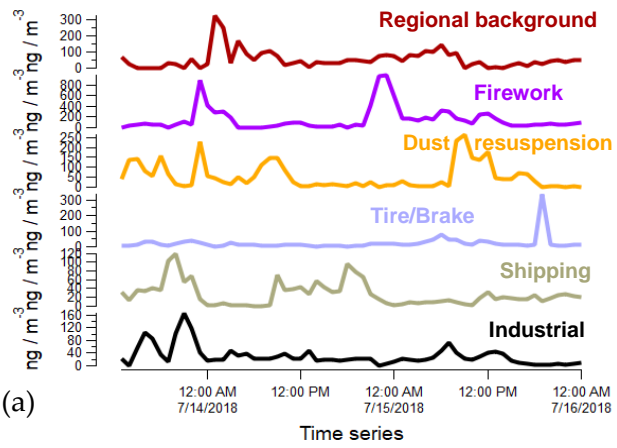


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122 **Figure S2. (a)** Changes in  $\Delta Q/Q_{exp}$ ,  $\Delta UEV_{real}$  and  $\Delta UEV_{noisy}$  for n-(n+1)-factor PMF<sub>metals</sub> runs and **(b)**  
 123  $Q/Q_{exp}$ ,  $UEV_{real}$  and  $UEV_{noisy}$  for PMF<sub>metals</sub> runs from 1 to 8 factors. These PMF runs are performed for the  
 124 WFP dataset. The box plots located in the blue dashed-line area represent the values for the finalized 6-  
 125 factors bootstrap solution using the total metals dataset.

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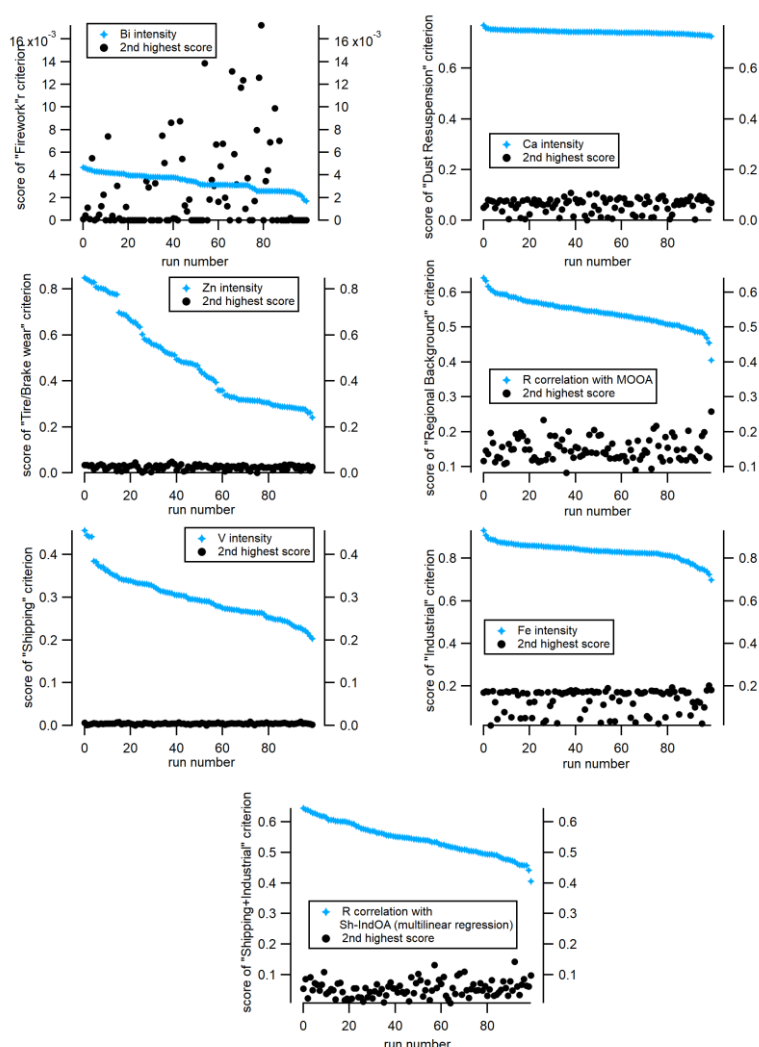
129 **Figure S3. (a)** Factors time series and **(b)** profiles from the PMF solution using the FDP dataset. The  
 130 regional background factor profile was constrained with an a-value of 0.1.

131



132 **Criteria selection for PMF<sub>metals</sub>**

133 A first type of criterion was the use of the dominant element in the related factor. Thus, the Bi, Ca, Zn,  
134 V and Fe intensity in profiles were monitored for the Firework, Dust resuspension, Tire/brake wear,  
135 Shipping and Industrial factors, respectively. Then we inspected the r Pearson correlation with MOOA  
136 for the regional background. A last criterion was the r Pearson correlation based on the multilinear  
137 regression analysis of both shipping and industrial vs SO<sub>2</sub> concentrations. The statistical acceptance of  
138 a run was based on the comparison between the criterion scores of a factor and the second highest scores  
139 from the remaining factors (Fig. S4). For all criteria the second highest scores were much lower in every  
140 run, with some rejected scores for the firework criterion. In total 25% of the runs were discarded based  
141 on this criterion, and the remaining runs were averaged into a unique solution.



142  
143 **Figure S4.** Criteria scores for the 100 bootstrapped runs from the PMF<sub>metals</sub>. Each graph represents one  
144 criterion for the different factors. The blue markers are for the factor criterion scores and the black  
145 markers represent the second highest scores attributed to one of the remaining factors.

146

147 **Scenario and regression model selection for OP apportionment**

148 Three scenarii in the construction of the matrix of the source factors contribution to PM mass identified  
149 by the three PMF have been considered to make the best use of the results from the different PMF:

- 150 • Scenario 1: OP apportionment from independent variables with the OA factors from PMF<sub>organics</sub> (83  
151 observations), following:

$$152 \quad OP = G \times \beta_g + \varepsilon \quad (S4)$$

- 153 • Scenario 2: OP apportionment from independent variables considering only the metals factors from  
154 PMF<sub>metals</sub> (90 observations), following:

$$155 \quad OP = H \times \beta_h + \varepsilon \quad (S5)$$

- 156 • Scenario 3: OP apportionment taking as independent variables PM<sub>1</sub> factors from PMF<sub>PM1</sub> (78  
157 observations), following (Eq. S6). In this configuration, the firework episode has been removed from  
158 the data as the sources from the PMF<sub>PM1</sub> analysis have been determined without including the  
159 firework metal factor.

$$160 \quad OP = I \times \beta_i + \varepsilon \quad (S6)$$

161 In (Eq. S4, S5, S6), OP vector (p×1) is the observed OP expressed in volume unit, G matrix (g × (p+1)) of  
162 g sources (plus the intercept) is determined by PMF<sub>organics</sub>, H matrix (h × (p+1)) of h sources (plus the  
163 intercept) is determined by PMF<sub>metal</sub>, I matrix (i × (p+1)) of i sources (plus the intercept) is determined  
164 by PMF<sub>PM1</sub>, and ε vector (p×1) is the discrepancy between the model and the observations.

165 Three models were tested for the three scenarii (e.g. 9 solutions): weighted least squares linear  
166 regression (WLS), weighted robust multiple linear regression with an iterative M-estimator, and partial  
167 least square regression (PLS):

- 168 • WLS regression considers the uncertainties σ of the OP measurements by minimizing the weighted  
169 sum of squares function (WSS):

$$170 \quad WSS(\beta) = \sum_{i=1}^p w_i (y_i - \sum_{j=1}^n x_{ij} * \beta_j)^2, \quad w_i = \frac{1}{\sigma_i} \quad (S7)$$

171 where y<sub>i</sub> is the measured OP (p observations), x<sub>ij</sub> is the values of n sources determined by PMF and σ<sub>i</sub>  
172 is the OP uncertainties. This method already used in this purpose in previous studies (Borlaza et al.,  
173 2021; Weber et al., 2018, 2021) well suited to extracting maximum information from small data sets.  
174 Ordinary Least Squares (OLS) is a simple special case of WLS where σ = 1.

- 175 • Linear weighted robust regression methods by M-estimator minimizes the function ρ:

$$176 \quad M(\beta) = \sum_{i=1}^p \rho(w_i (y_i - \sum_{j=1}^n x_{ij} * \beta_j)) \quad (S8)$$

$$\rho_k(x) = \begin{cases} \frac{x^2}{2} & \text{if } |x| < k = 1.5 \\ k \left( |x| - \frac{k}{2} \right) & \text{if } |x| \geq k = 1.5 \end{cases} \quad (\text{S9})$$

178 Based on similar work in Grange et al. (2022), Huber's function  $\rho$  and  $k=1.5$  were used in this study.  
 179 This technique is adapted to data sets presenting particular events (de Menezes et al., 2021), as fireworks  
 180 on 13<sup>th</sup> and 14<sup>th</sup> of July -National day of France- in our data set. Indeed, the regression by successive  
 181 iterations implies lower weights on outliers, which tends to underestimate these points. We can note  
 182 WLS regression is a simple special case where  $\rho(x) = x^2$ .

- 183 • PLS regression is a method that reduces the predictors to a smaller set of uncorrelated components  
 184 and performs least squares regression on these components. It is especially useful when dependent  
 185 variables are highly correlated. Moreover, unlike multiple regression, PLS does not imply that the  
 186 predictors are fixed but can be measured with error, making PLS more robust to measurement  
 187 uncertainties.

188

#### 189 **OP apportionment from PMF<sub>organics</sub> (scenario 1) and PMF<sub>metals</sub> (scenario 2)**

190 M-estimator inversion model's results issued from PMF<sub>organics</sub> (scenario 1) and PMF<sub>metals</sub> (scenario 2)  
 191 alone are respectively presented in Table S 3a. and Table S 3b.  $\beta$  coefficients (i.e intrinsic OP, see 2.5)  
 192 obtained by M-estimator model from PMF<sub>metals</sub> display values an order of magnitude higher than those  
 193 issued from PMF<sub>organics</sub> inversion. This stress the importance of metals in OP apportionment, for both  
 194 assays. Among the organic factors, only the Sh-IndOA factor seems to be slightly more sensitive to  
 195 OP<sub>VDTT</sub>. The Firework factor constrains a significant part of the data, implying a fairly high Pearson's  
 196 correlation coefficient between OP<sub>model</sub> and OP<sub>observed</sub>. Nevertheless,  $R^2_{\text{adjusted}}$  of both M-estimator  
 197 inversion models in scenario 1 (only organic fraction of PM is considered) indicated that the percentage  
 198 of OP<sub>AA</sub> and OP<sub>DDT</sub> variance explained by the models is weak. On the other hand, several studies  
 199 highlighted the role of Secondary Organic Aerosol (SOA) in the oxidative potential indicating that  
 200 apportion OP from the metallic data alone is an incomplete step. Finally, the bootstrap method (see 2.5)  
 201 applied to the four M-estimator models in these two scenarii did not achieve their convergence and are  
 202 therefore not robust. Overall, this confirms that OP reflects the overall redox-activity of wide spectra of  
 203 multispecies of organics, inorganics, metals and synergistic/antagonistic reactions between these  
 204 compounds, and assess the importance to consider all these chemical compounds in the OP  
 205 apportionment process.

(a)

	Intercept	COA	HOA	LOOA	MOOA	Sh-IndOA	$R^2$ adjusted	$r$ (OP <sub>observed</sub> /OP <sub>model</sub> )
OP <sub>AA</sub>	0.19 ± 0.04	0.00 ± 0.02	0.02 ± 0.01	0.10 ± 0.04	0.04 ± 0.02	0.24 ± 0.09	0.27	0.48***
OP <sub>DTT</sub>	0.38 ± 0.10	-0.04 ± 0.10	-0.12 ± 0.05	0.23 ± 0.04	0.1 ± 0.07	1.41 ± 0.15	0.41	0.51***

(b)

	Intercept	Firework	Industrial	Regional background	Shipping	Tire brake	$R^2$ adjusted	$r$ (OP <sub>observed</sub> /OP <sub>model</sub> )
OP <sub>AA</sub>	0.36 ± 0.02	1.57 ± 0.16	3.21 ± 0.42	2.30 ± 0.33	-0.74 ± 0.68	n.c.	0.66	0.73***
OP <sub>DTT</sub>	0.61 ± 0.12	4.17 ± 0.83	-1.24 ± 1.64	7.11 ± 2.30	17.0 ± 5.3	6.0 ± 4.60	0.38	0.61***

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207

**Table S3.** Intrinsic OP<sub>AA</sub> and OP<sub>DTT</sub> (OP<sub>m</sub>) provided by weighted robust linear regression with an M-estimator expressed in nmol.min<sup>-1</sup>.μg<sup>-1</sup> of sources provided by (a) PMF<sub>organics</sub> (scenario 1) and (b) PMF<sub>metals</sub> (scenario 2) over the OP sampling campaign (n = 90). Values are the mean ± standard deviation from bootstraps runs for both OP assays. The model parameters R<sup>2</sup><sub>adjusted</sub> and Pearson's correlation between model OP and observed OP are mentioned on the right.

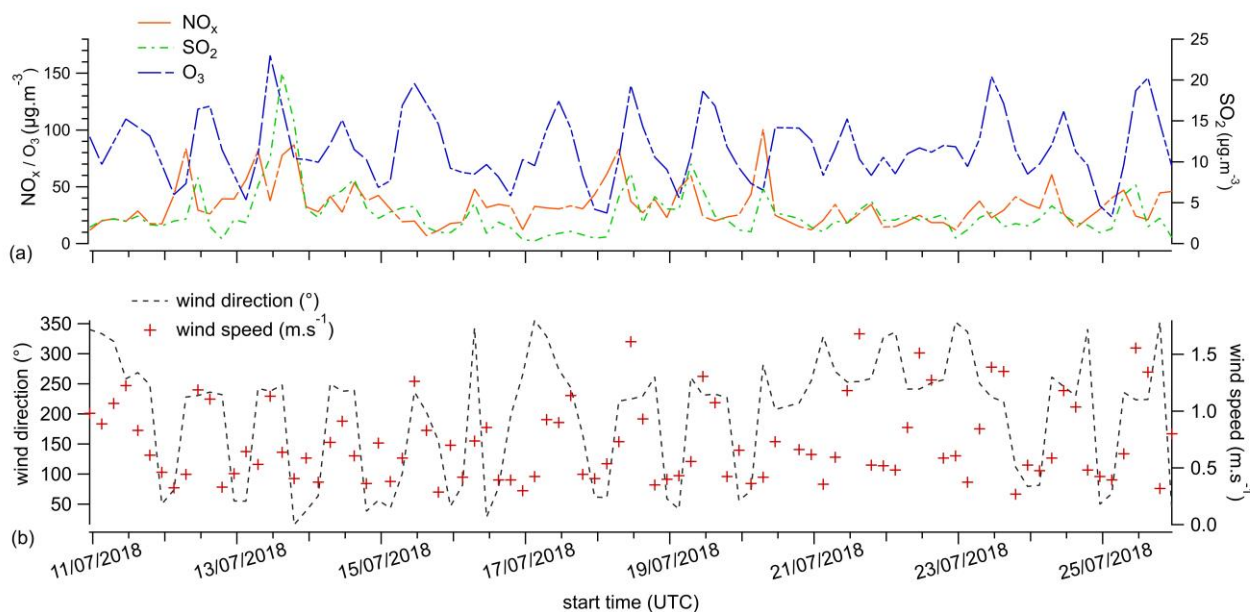
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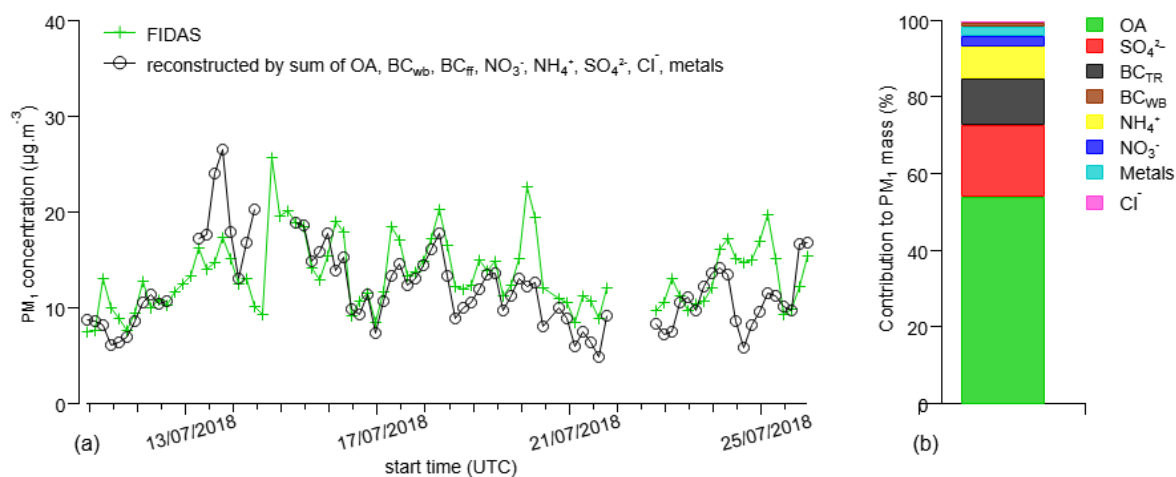


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**Figure S5.** (a) NO<sub>x</sub>, SO<sub>2</sub> and O<sub>3</sub> concentration and (b) wind speed and direction during OP measurement period.

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218 **Figure S6. (a)** Comparison between time series of PM<sub>1</sub> measured by FIDAS and time series of particulate  
 219 fraction reconstituted by the sum of chemical components ( $r_s = 0.47$ ,  $p < 0.001$ ); **(b)** Contribution to PM<sub>1</sub>  
 220 of chemical components (%) measured from 11<sup>th</sup> July 2018 to 25<sup>th</sup> July 2018 (included firework episode,  
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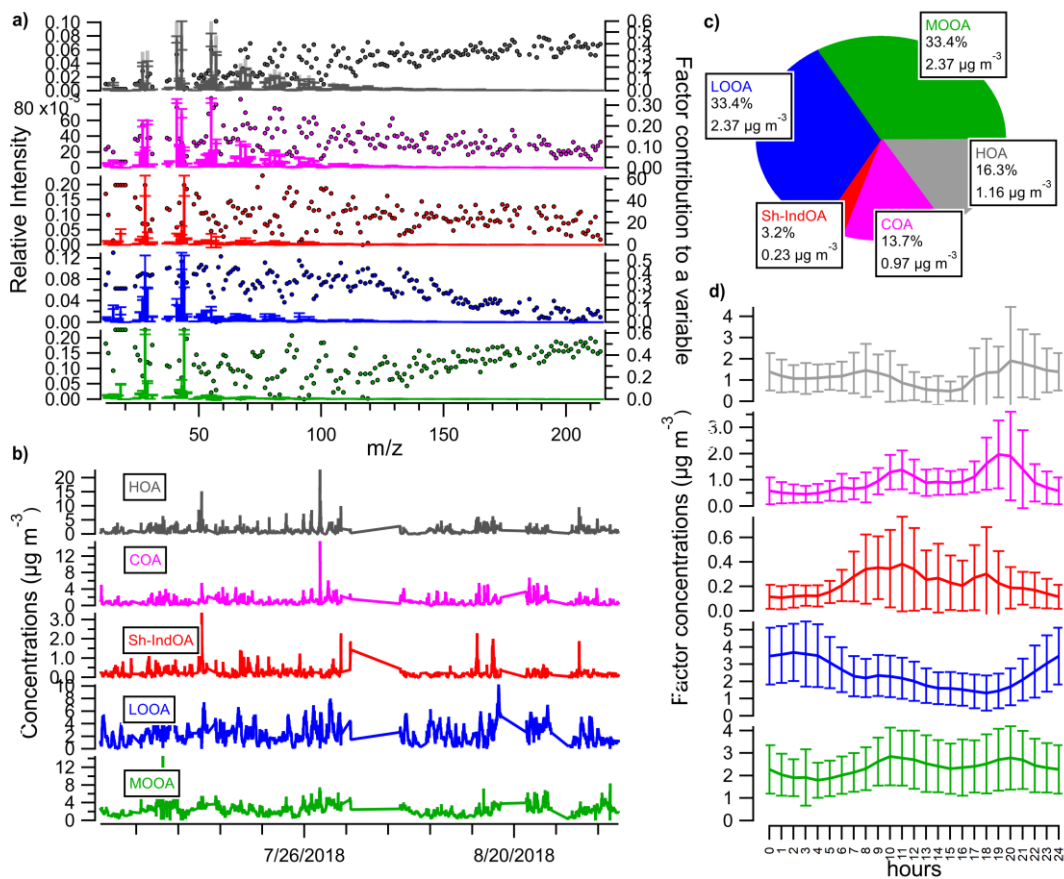
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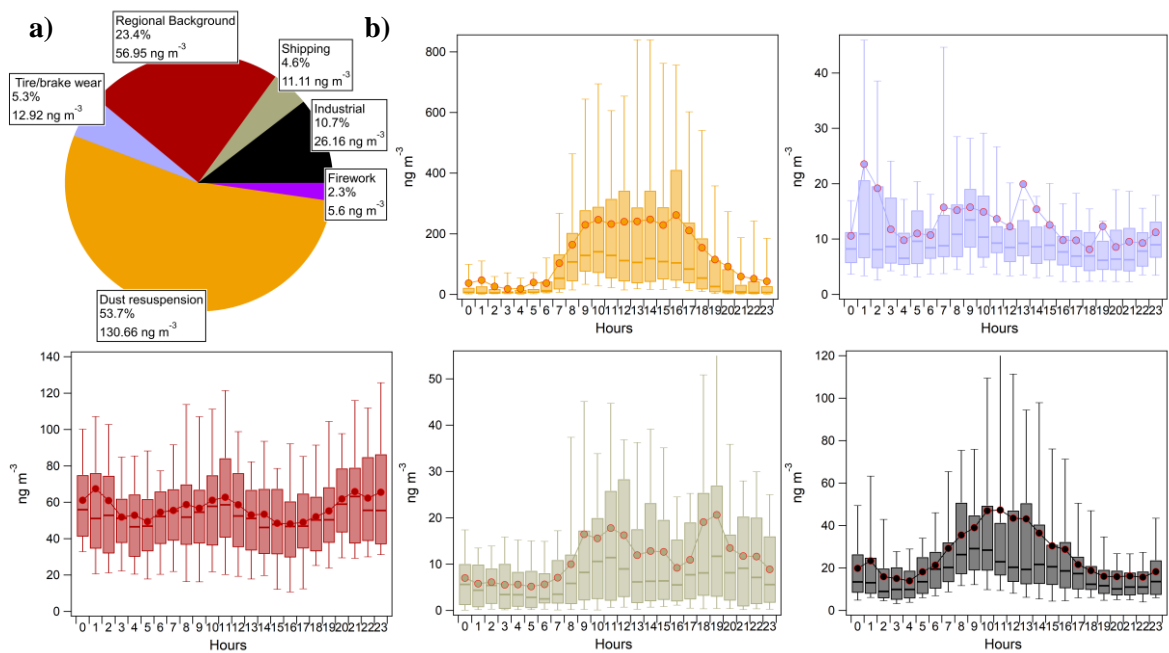
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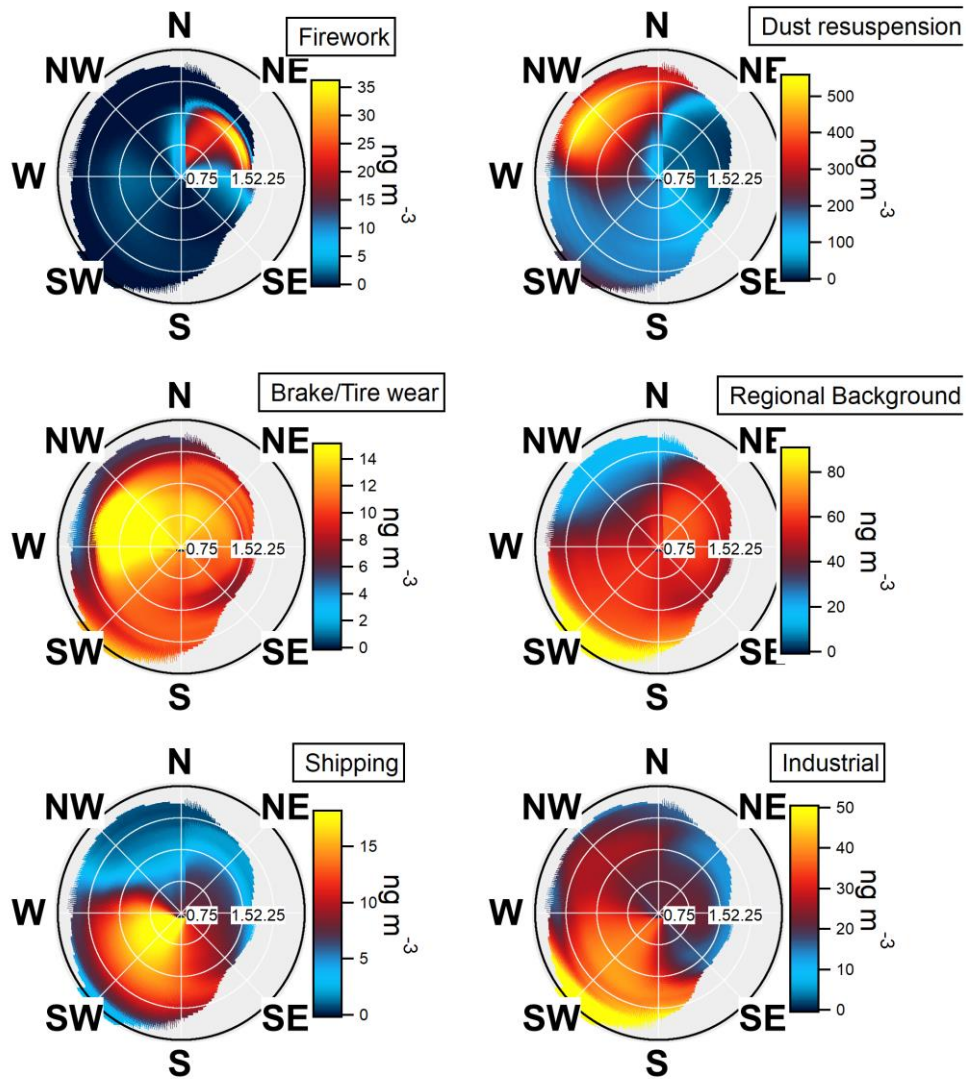
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238 **Figure S8. (a)** Pie chart contributions and **(b)** average diurnal profiles of factors from the  $\text{PMF}_{\text{metals}}$   
 239 analysis. For the diurnal plots the red dots correspond to the mean, the bands are the median, the bottom  
 240 and top of the boxes represent the 25<sup>th</sup> and 75<sup>th</sup> percentile respectively, and the ends of the whiskers are  
 241 for the 10<sup>th</sup> and 90<sup>th</sup> percentiles.

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244 **Figure S9.** NWR plots for the different factors from the PMF<sub>metals</sub> analysis.

245

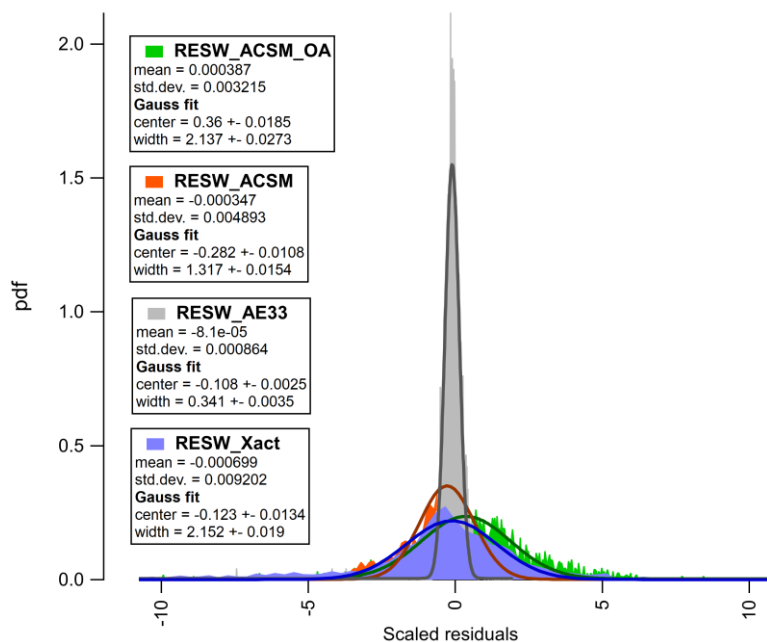


246 **C-value weighting**

247 The instrument weight was controlled by applying a scaling factor (i.e. C-value) to the uncertainties of  
 248 each group of components (Slowik et al., 2010):

$$249 (\sigma'_{i,j})_s = \frac{(\sigma_{i,j})_s}{C_s} \tag{S10}$$

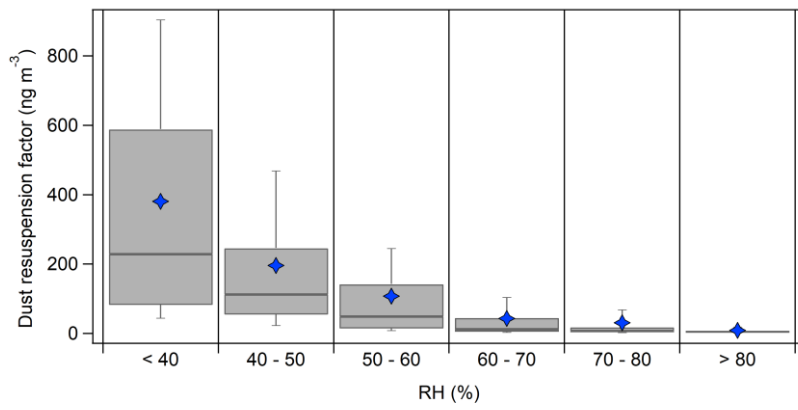
250  $\sigma$  represents the uncertainties,  $C$  the scaling value applied to the  $s$  datasets. Here we distinguished the  
 251 PMF<sub>organics</sub> (ACSM\_OA), PMF<sub>metals</sub> (Xact), ACSM inorganics (ACSM) and BC (AE33) datasets. A well  
 252 balanced solution should show magnitude of scaled residuals independent from the instrument. Since  
 253 their scaled residuals were rather in the same range, a C-value of 1 was chosen for ACSM\_OA, Xact and  
 254 ACSM datasets and resulted in unweighted results. However, we applied a C-value of 5 to the AE33  
 255 dataset, meaning that dataset of BC concentrations were upweighted. The overlapping of scaled  
 256 residuals from the different instrument datasets is shown in Figure .



257  
 258 **Figure S10.** Probability density function of scaled residuals for the standalone ACSM\_OA, ACSM, AE33  
 259 and Xact datasets.

260

261



262

263 **Figure S11.** Box plots of dust resuspension factor concentrations for different relative humidity (RH)  
 264 bins in %. The concentrations are enhanced under low RH conditions. The blue diamonds are the mean,  
 265 the bars inside the boxes the median, the bottom and top of the boxes are the 25<sup>th</sup> and 75<sup>th</sup> percentile,  
 266 respectively, and the ends of the whiskers are the 10<sup>th</sup> and 90<sup>th</sup> percentiles.

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270 **Factors identification and rotational ambiguity exploration for PMF<sub>PM1</sub>**

271 Seed runs between 1 and 12 factors were examined. The factors interpretability was based on profiles  
 272 consistency and our expectations from the factors composition. The summarizes the occurrence of 8  
 273 well-identified factors in all runs between 5 and 12 factors. The choice of a 8-factors solution is supported  
 274 also by mathematical diagnostics ( $\Delta Q/Q_{exp}$ , mass reconstruction,  $\Delta UEV$  – not presented here) which  
 275 showed that realistic solutions can be found up to 5 factors. While some factors are easily resolved in  
 276 most of the solutions (e.g. dust resuspension) some others are retrieved from an elevated number of  
 277 factor (e.g. shipping and cooking are found in up to 9 factors-solution).

278 Therefore, the solution was constrained using base case profiles (Table ). The biomass burning, cooking  
 279 and industrial factors were constrained as they presented unstable profiles across the different runs.  
 280 Constraining the industrial factor allow an improved separation of the shipping factor (see the  
 281 discussion below).

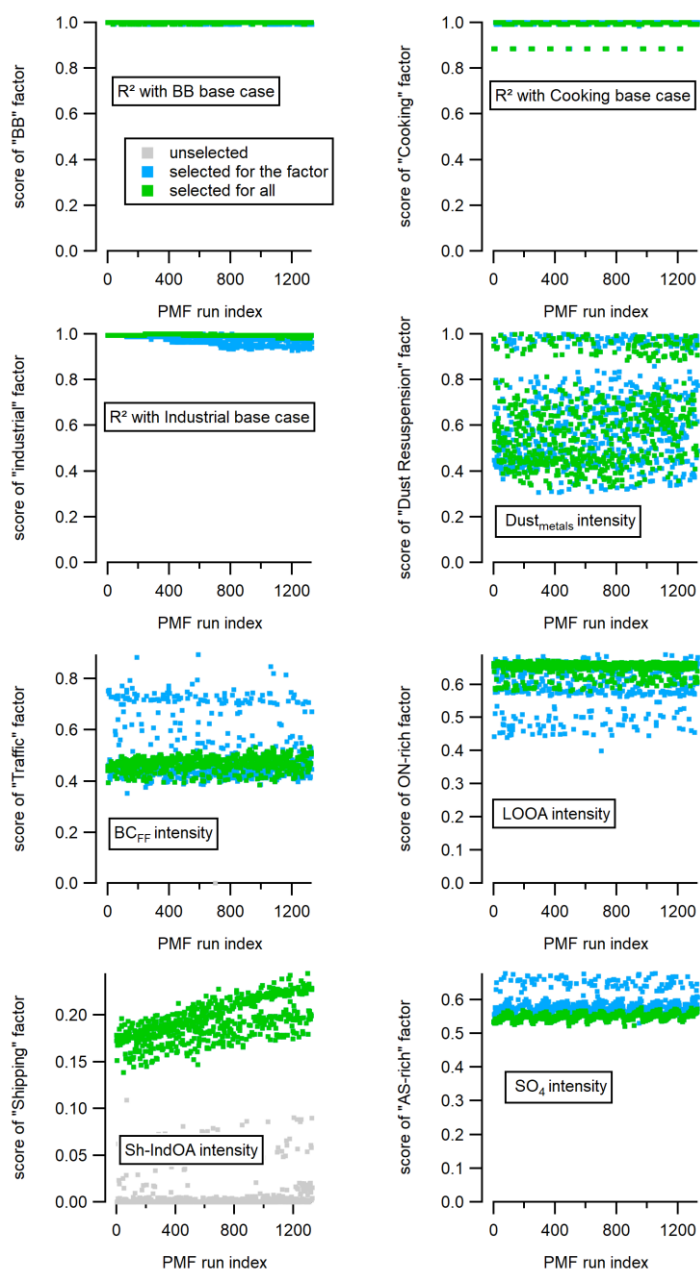
282

	5F	6F	7F	8F	9F	10F	11F	12F
<b>Traffic</b>								
<b>Dust</b>								
<b>AS-rich</b>								
<b>ON-rich</b>								
<b>Industrial</b>								
<b>Shipping</b>								
<b>BB</b>								
<b>Cooking</b>								

283 **Table S4.** Factors identification for the PMF<sub>PM1</sub> analyses between 5 and 12 factors. The green cells  
 284 represent the base case identification for the related factors. The remaining undefined factors for each  
 285 solution corresponded to mixed profiles not attributed to a specific source. The red squares are the base  
 286 cases used as reference profile constraints.

287 To inspect the best combination of a-values for the profile constraints, we performed a-values sensitivity  
 288 analyses by scanning a-values from 0 to 0.5 with increment of 0.05, leading to 1330 outcomes. The  
 289 goodness of the solutions was examined with a criteria selection list and the scores are presented in the  
 290 Figure . First, the R<sup>2</sup> correlations between biomass burning, cooking and industrial factors with their  
 291 corresponding constraint were monitored. Then, we monitored the intensity of the dominant variable  
 292 in the related factor profiles: Dust<sub>metals</sub> for dust resuspension, BC<sub>FF</sub> for traffic, LOOA for ON-rich, SO<sub>4</sub><sup>2-</sup>  
 293 for AS-rich and Sh-IndOA for shipping. Sh-IndOA was inspected instead of shipping<sub>metals</sub> to ensure a  
 294 clear separation between shipping and industrial factors since Sh-IndOA is assumed to only be  
 295 attributed to these two factors. For the first seven criteria, the scores were much higher than the second  
 296 highest scores (not displayed in the graph). Therefore, some runs were only discarded based on the

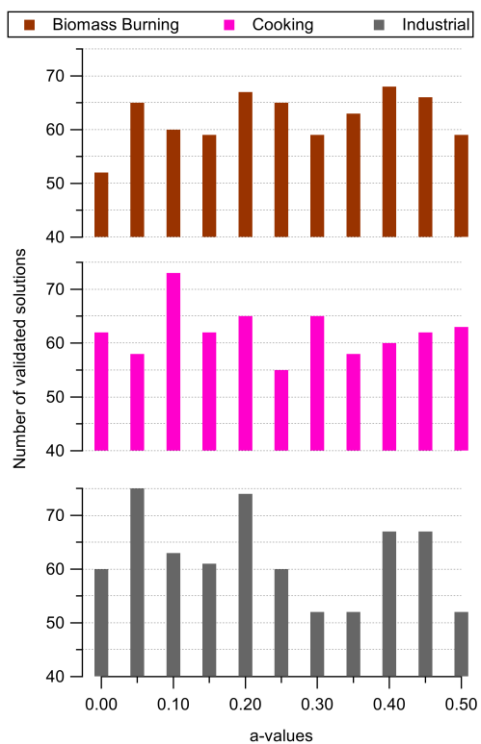
297 shipping criterion as we only selected the runs whose Sh-IndOA intensity was in the same range than  
 298 the base case profile from the preliminary analyses. Moreover, the selected runs (green markers in  
 299 Figure ) showed similar scores intensity for traffic, ON-rich, AS-rich and dust resuspension than those  
 300 found in their respective base case profile. In the end, the same criteria list was used for the bootstrap  
 301 runs selection.



302  
 303 **Figure S12.** Criteria scores for the a-values sensitivity test runs from the PMF<sub>PM1</sub>. Each graph represents  
 304 one criterion per factor. The grey markers are the unselected runs, the blue markers are the selected  
 305 runs for the related factor and the green markers are the effectively chosen runs.

306

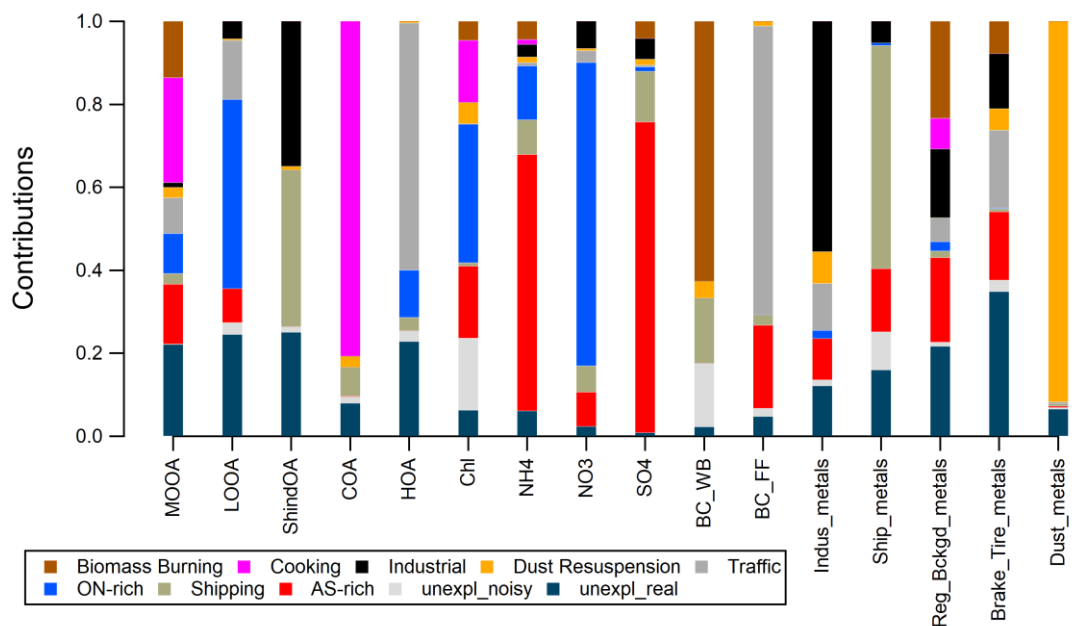
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308

309 **Figure S13.** Number of accepted solutions based on the  $PMF_{PM1}$  criteria list for the different a-values  
 310 explored in the sensitivity test. A-values associated to the greatest number of validated solutions  
 311 were chosen for the bootstrap PMF runs (i.e. 0.4 for biomass burning, 0.1 for cooking and 0.05 for industrial  
 312 constrained profiles).

313



314

315 **Figure S14.** Relative contributions of PM<sub>1</sub> factors profiles and unexplained variations from the PMF<sub>PM1</sub>  
 316 analysis.

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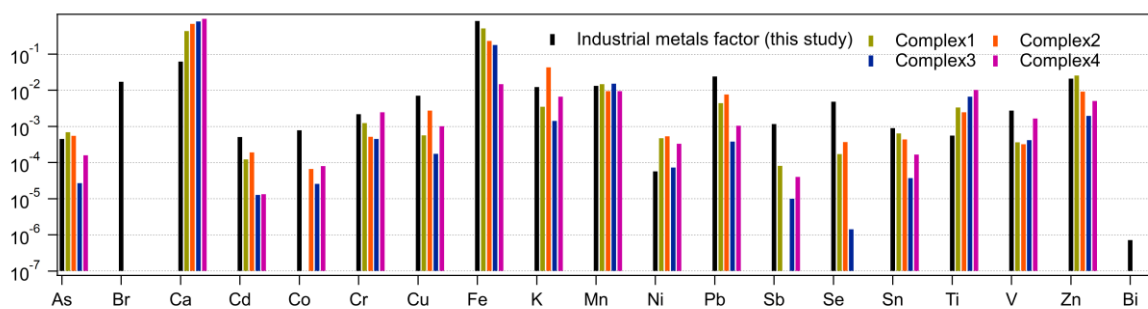
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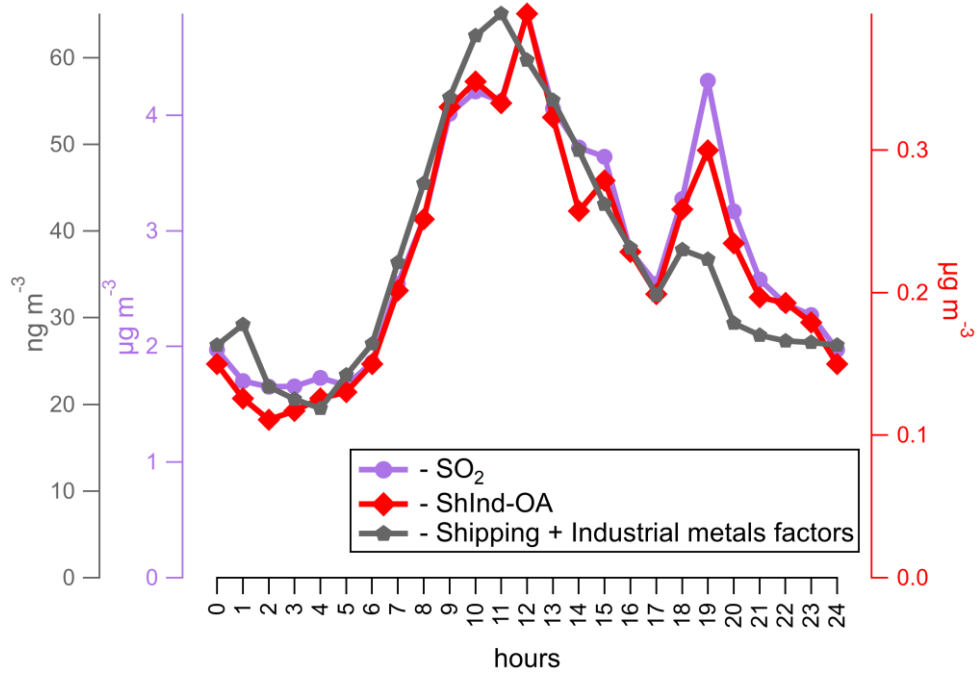


323

324 **Figure S15.** Comparison between the industrial metals profile from our study and ICP-MS profiles for  
 325 the PM<sub>2.5</sub> fraction in the industrial area of Fos-sur-mer (Sylvestre et al., 2017). Complex n°1 is a cast iron  
 326 converter complex, complex n°2 is a ore iron converter complex, complex n°3 is a blast furnace slag  
 327 storage and complex n°4 is an ore terminal.

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331 **Figure S16.** Average diurnal profiles for  $\text{SO}_2$ , Sh-IndOA and the sum of industrial and shipping factors  
 332 from the  $\text{PMF}_{\text{metals}}$ .

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338 **Associations between both OP and sources of PM**

339 Pearson's correlation coefficients (r) between the source factor contributions identified by the PMF<sub>PMI</sub>  
 340 and both OP assays are presented in Table with the idea to provide a first estimate of the associated  
 341 sources with OP. We note that no source strongly correlates alone to both OP assays, but moderate  
 342 correlations (0.3 < r < 0.5) can be noted for both OP vs. Traffic source (OP<sub>VAA</sub>: r=0.40, p<0.001 - OP<sub>VDIT</sub>:  
 343 r=0.34, p<0.01) and Shipping source (OP<sub>VAA</sub>: r=0.32 - OP<sub>VDIT</sub>: r=0.30, p<0.01). OP<sub>VAA</sub> also correlates  
 344 moderately with Industrial source (r=0.41, p<0.001) and ON-rich source (r=0.32, p<0.01). Finally, OP<sub>VDIT</sub>  
 345 displays a mild correlation with AS-rich source (r=0.36, p<0.01), but this correlation might be attributed  
 346 to a collinearity with PM mass (r OP<sub>VDIT</sub> vs SO<sub>4</sub><sup>2-</sup>=0.46, r OP<sub>VDIT</sub> vs NH<sub>4</sub><sup>+</sup> = 0.47 - p<0.001).

347

	<b>Biomass Burning</b>	<b>Cooking</b>	<b>Industrial</b>	<b>Dust resuspension</b>	<b>Traffic</b>	<b>ON- rich</b>	<b>Shipping</b>	<b>AS- rich</b>
<b>OP<sub>VAA</sub></b>	0.15	0.18	0.41***	0.13	0.40***	0.32***	0.32**	0.17
<b>OP<sub>VDIT</sub></b>	0.12	-0.02	0.14	0.14	0.34**	0.19	0.30	0.36**

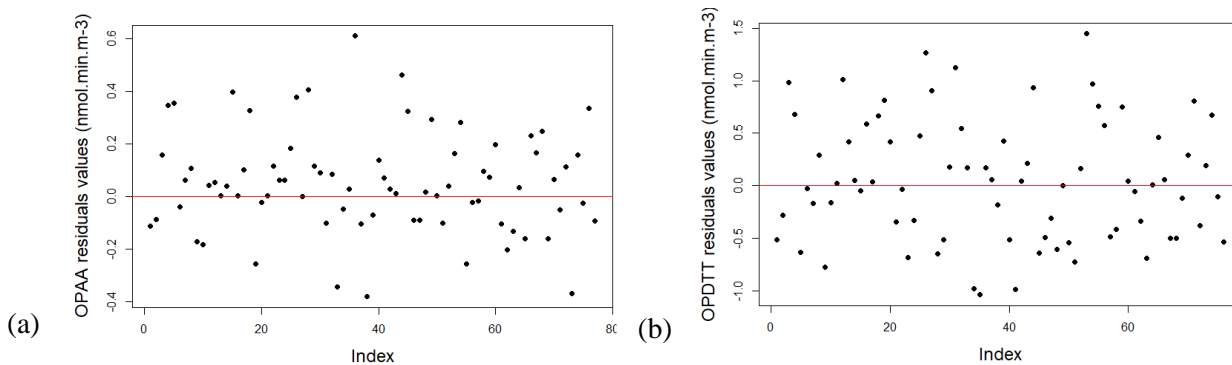
\*\*\*p < 0.001, \*\*p < 0.01

348 **Table S5.** Pearson's correlation coefficients between OP<sub>VAA</sub> and OP<sub>VDIT</sub> to the PM sources identified by  
 349 PMF<sub>PMI</sub> model.

350

351



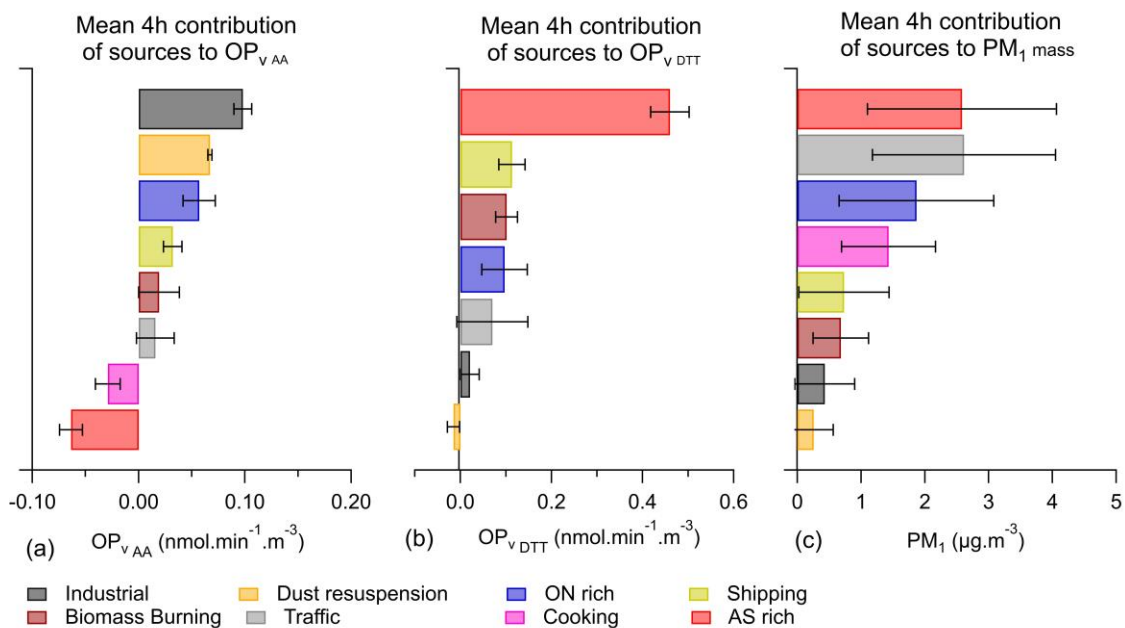


**Figure S17.** Residuals values of WLS models for **(a)**  $OP_{AA}$  and **(b)**  $OP_{DTT}$ . An outlier point (19 July 2018 03:00) was withdrawn to ensure homoscedasticity of residuals values.

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**Figure S18.** Mean contribution of the sources identified by  $PMF_{PM1}$  over the OP sampling campaign ( $n = 86$ ) to **(a)**  $OP_{AA}$ , **(b)**  $OP_{DTT}$ , **(c)**  $PM_1$ . Error bars represents the standard deviation of the data distribution.

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