

1 Exploring the joint probability of precipitation and soil moisture 2 over Europe using copulas

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7 **Abstract.** The joint probability of precipitation and soil moisture is here investigated over Europe
8 with the goal to extrapolate meaningful insights on the potential joint use of these variables for
9 the detection of agricultural droughts within a [multivariate](#) probabilistic modeling framework. The
10 use of copulas is explored, ~~being the framework as a parametric approach~~ often used in
11 hydrological studies for the analysis of bivariate distributions. The analysis is performed for the
12 period 1996-2020 on the [empirical frequencies derived from](#) ERA5 precipitation and
13 LISFLOOD soil moisture datasets, both available as part of the Copernicus European Drought
14 Observatory. The results show an overall good correlation between the [empirical frequency series](#)
15 ~~derived from the~~ two [standardized datasets-series](#) (Kendall's $\tau = 0.42 \pm 0.1$), but also clear spatial
16 patterns in the tail-dependence derived with both non-parametric and parametric approaches.
17 About half of the domain shows symmetric tail-dependences, well reproduced by the Student-t
18 copula; whereas the rest of the domain is almost equally split between low and high tail-
19 dependences ([both](#) modeled with the Gumbel family of copulas). These spatial patterns are
20 reasonably reproduced by a random forest classifier, suggesting that this outcome is not driven by
21 chance. This study stresses how a joint use of [standardized](#) precipitation and soil moisture for
22 agriculture drought characterization may be ~~more~~ beneficial in areas with strong low tail-
23 dependence, and how this behavior should be carefully considered in [multivariate](#) drought studies.

24 1. Introduction

25 Agricultural drought, defined as a condition of unusually high precipitation shortages and/or soil
26 water deficits causing adverse effects on crop yields and production (Panu and Sharma, 2002), is
27 probably the most recognized of the four main drought types or phases (Wilhite and Glantz, 1985).
28 This is mainly due to the more direct and easier to understand impacts compared to the other types
29 of droughts (Mishra and Singh, 2010). The scientific literature on agricultural drought has
30 produced/provides a very-large number-variety of indices (WMO and GWP, 2016), with the aim of
31 reproducing the temporal dynamics and the effects of crop water deficit through a combination of
32 climatic observations, hydrological modeling, and remote sensing data (Zargar et al., 2011).

33 The difficulty in capturing the multi-facet nature of agricultural drought events across the
34 world with a single index-approach (Sivakumar et al., 2011) is confirmed by the absence of
35 consensus in the scientific literature on the most reliable agricultural drought index. However,
36 Despite the large range of available indices, some common characteristics can be identified, such
37 as the focus on some proxy variables of plant water availability – through soil moisture (Dutra et
38 al., 2008), actual evapotranspiration (Anderson et al., 2011) or basic meteorological information
39 (Vicente-Serrano et al., 2010) – and the need to account for deviations from long-term conditions
40 (i.e., use of standardized anomalies/scores).

41 Meteorological drought indicators computed on appropriate aggregation time scales
42 (McKee et al., 1993; Vicente-Serrano et al., 2010) have demonstrated a good capability to
43 represent agricultural drought conditions in several case studies (e.g., Bachmair et al., 2018;
44 Mohammed et al., 2022; Tian et al., 2018). They have been successfully integrated in a number of
45 operational drought monitoring systems, thanks to their minimal input data requirements and ease
46 of use. Among those indices, the Standardized Precipitation Index (SPI, McKee et al., 1993)
47 computed on short-to-medium-aggregation periods (i.e., SPI-3 and -6) is often adopted as a
48 suitable proxy variable for agricultural droughts (WMO, 2012).

49 As highlighted by Sheffield and Wood (2007), simplified indices for drought monitoring,
50 such as the Palmer Drought Severity index (PDSI; Palmer, 1965) or the previously mentioned
51 meteorological indicators, have been slowly integrated with indices directly based on modeled soil
52 moisture data. This transition is fostered by thanks to the increasing availability worldwide of
53 process-based hydrological models. Soil moisture percentile, or similarly standardized quantities,
54 are often used for this scope (Mo and Lettenmeier, 2013; Xia et al., 2014). The ever-growing

55 records of remote sensing-based estimates of soil moisture are becoming an additional data source
56 to support the development of dedicated soil moisture-~~based~~ drought indices (Cammalleri et al.,
57 2017; Carrão et al., 2016).

58 In the context of agricultural drought, an overall good agreement between SPI and soil
59 moisture indices has been demonstrated over a large range of agricultural practices, crop types and
60 climatic conditions. Halwatura et al. (2017) showed how SPI-3 represents a good approximation
61 of modeled soil moisture over three different climatic regions in eastern Australia. Sims et al.
62 (2002) found high correlation between short-term precipitation deficit and soil moisture variations
63 in North Carolina, while Ji and Peters (2003) highlighted the high correlation between SPI-3 and
64 vegetation growth over croplands and grasslands in the U.S. Great Plains. Wang et al. (2015)
65 observed a good matching between soil moisture dynamics and SPI at the scale of 1-3 months
66 when testing various indices over China. In Europe, Manning et al. (2018) highlighted how
67 precipitation is the main driver of soil moisture droughts for a set of both dry and wet sites.

68 In spite of the above-mentioned consistencies, ~~the outcome of any of a drought analysis is~~
69 ~~will be inevitably affected by~~ the index selected to characterize drought conditions over a certain
70 study region, ~~will inevitably affect the outcome of the drought analysis,~~ as ~~also~~ highlighted by
71 Quiring and Papakryiakou (2003) in testing different indices over the Canadian prairies. These
72 ~~a~~Authors suggest that a variety of drought indices should always be tested to determine the most
73 appropriate one for ~~each specific a given~~ application. It follows that the synergy between multiple
74 indices can be exploited by the use of multivariate indicators (Hao and Singh, 2015), a family of
75 approaches that encompasses a variety of merging strategies, including combined cascading
76 indices (Cammalleri et al., 2021a; Rembold et al., 2019), composite and integrated approaches
77 (Brown et al., 2008; Svoboda et al., 2002), and joint probability functions (Bateni et al., 2018; Hao
78 and AghaKouchak, 2013; Kanthavel et al., 2022).

79 The latter ~~category class of approaches~~, in particular, aims at capturing the complex
80 statistical ~~inter~~dependence among different drought-related variables (Hao and Singh, 2015), and
81 it has seen a growing relevance in many hydrological applications thanks to the introduction of
82 copula functions and their ability to model a wide range of dependence structures (Nelsen, 2006;
83 Salvadori et al., 2007; Joe, 2015). In the field of drought indices, the approach proposed by Kao
84 and Govindaraju (2010) for the computation of the Joint Deficit Index (JDI) has been applied to a

85 variety of drought-related quantities over different regions, often including precipitation and soil
86 moisture (i.e., Dash et al., 2019; Kwon et al., 2019).

87 ~~A One~~ key feature in using ~~of~~ joint probability is the possibility to characterize the so-called
88 tail-dependence (TD), namely the asymptotical dependence ~~of~~ at the extremes (Frahm et al., 2005).
89 While TD has received large attention in the scientific literature of hydrological extremes (e.g.,
90 Aghakouchak et al., 2010; Poulin et al., 2007; Serinaldi, 2008), its use is largely unexploited in
91 ~~limited attention to this quantity is still given in the~~ studies focusing on combined drought indices.

92 Studies on the marginal distribution of either precipitation or soil moisture usually adopt
93 ~~have somewhat converged on adopting~~ the Gamma distribution for precipitation and the Beta
94 distribution for soil moisture. The use of the Gamma family for the implementation of the SPI at
95 different accumulation periods has become a standard practice in many applications (e.g., Mo
96 and Lyon, 2015; Yuan and Wood, 2013). While other distributions have also proven to be reliable,
97 such as the exponentiated Weibull (Pieper et al., 2020) and the Person Type III (Ribeiro and Pires,
98 2016), fitting the Gamma is still the most adopted ~~one~~ approach. Over Europe, Stagge et al. (2015)
99 demonstrated how the Gamma outperformed the other tested distributions across all accumulation
100 periods and regions.

101 A more limited number of applications based on soil moisture data are available in the
102 scientific literature compared to SPI. The use of the Beta distribution for soil moisture data has
103 been introduced as early as the late '70s, with the pioneer study of Ravelo and Decker (1979),
104 following the consideration that soil moisture is a double-bounded quantity, ranging between
105 residual and saturation. Sheffield et al. (2004) successfully applied this standardization for drought
106 analyses over the US, while the same distribution has been adopted by Cammalleri et al. (2016)
107 on modeled data over Europe. Most recently, the Beta distribution was also used to characterize
108 the frequency of global satellite soil moisture data (Sadri et al., 2020).

109 Conversely, no standard approaches have been identified for the application of copulas to
110 model the bivariate joint distribution of precipitation and soil moisture, mainly due to the large
111 variety of probabilistic structures than may be observed between these two quantities. Common
112 fitting strategies rely on the application of various copula families to identify the optimal best
113 ~~fitting~~ for each specific site (e.g., Hao and AghaKouchak, 2013), or are based on an a-priori
114 selection of a copula family following empirical evidence (e.g., Dixit and Jayakumar, 2021).
115 Independently from the selection strategy, the adopted copula implicitly ~~assumes~~ ~~carries with it an~~

116 underling TD behavior, which influenceimpacts on extreme detection should be properly
117 accounted.

118 A comprehensive study on the joint probabilistic dynamics of ~~these two~~
119 quantitiesprecipitation and soil moisture, ~~and on their bivariate distribution~~, is currently lacking
120 in the scientific literature of multivariate drought modeling. Hence, the main goal of this study is
121 to fill this gap, by investigating the mutual relationship between the empirical frequencies of
122 precipitation (cumulated over 3 months, as for SPI-3) and soil moisture datasets as available over
123 Europe as part of the European Drought Observatory of the Copernicus Emergency Management
124 Service (EDO, <https://edo.jrc.ec.europa.eu>).

125 A large set of copulas is tested for this purpose across the entire European domain, to
126 identify an optimal modeling of the dependence especially in proximity of the tails (given its major
127 role in extreme detection). The spatial distribution of the results is analyzed to infer evidence of
128 common patterns and behavior, which may support future operational applications based on
129 similar parametric approaches.

131 2. Materials and Methods

132 2.1 Precipitation and soil moisture datasets

133 The study focuses on Europe and makes use of the dataset of indicators available over the region
134 as part of EDO. Precipitation data accumulated over consecutive 3-month periods are used here,
135 as the quantity at the base of the SPI-3 index. Hourly total precipitation maps from the ECMWF
136 ERA5 global atmospheric reanalysis model ([https://www.ecmwf.int/en/forecasts/dataset/ecmwf-](https://www.ecmwf.int/en/forecasts/dataset/ecmwf-reanalysis-v5)
137 [reanalysis-v5](https://www.ecmwf.int/en/forecasts/dataset/ecmwf-reanalysis-v5)) are collected through the Copernicus Climate Change Service (C3S,
138 <https://climate.copernicus.eu/>) and cumulated at monthly updates (no missing values are present
139 in the reanalysis dataset). This dataset has proven to be quite reliable over Europe for drought
140 analyses (e.g., Cammalleri et al., 2021b; van der Wiel et al., 2022), as it is currently employed in
141 near-real time as part of the operational tools of EDO. Empirical frequencies of 3-month
142 precipitation are derived from the rainfall records, in order to obtain a non-parametric calculation
143 of the standardized anomaly, dataset conceptually analogous to the commonly used SPI-3, but
144 without the possible artifact introduced by the fitting of a theoretical distribution (i.e., Gamma
145 distribution) (see Sořáková et al., 2014). From here on, we will refer to this dataset as standardized
146 precipitation.

147 Soil moisture records over the entire European domain are derived from the simulations of
148 the LISFLOOD distributed hydrological rainfall–runoff model (de Roo et al., 2000). LISFLOOD
149 runs in near-real time as part of the European Flood Awareness System (Thielen et al., 2009), and
150 it provides daily soil moisture maps for the root zone at a spatial resolution of 5-km. Daily modeled
151 data are averaged at monthly scale and converted into a Soil Moisture Index (SMI) as in
152 Seneviratne et al. (2010). The model is calibrated and validated over an extensive network of river
153 discharge stations following the procedure described in Arnal et al. (2019), and it has been
154 successfully tested for drought analyses over Europe as part of EDO for the computation of the
155 Soil Moisture Anomaly (SMA) index (Cammalleri et al., 2015). Similar to As for the precipitation,
156 empirical frequencies are computed from the monthly soil moisture data in order to obtain a non-
157 parametric calculation of the standardized anomaly, SMA, thus but independent from a theoretical
158 fitting (i.e., Beta distribution). We will refer to this dataset as standardized soil moisture from
159 hereafter.

160 In this study, data collected ~~foren~~ the most recent 25 years (1996-2020) are used as a
161 common period. This period is chosen to minimize the effects of non-stationarity in precipitation
162 records and to avoid the inclusion of early LISFLOOD records that are affected by a lower number
163 of ground meteorological stations in the forcing (Thieming et al., 2022). The time series of both
164 standardized precipitation and soil moisture at grid-cell scale are preliminarily tested for auto-
165 correlation using the partial auto-correlation function (PACF, Box and Jenkins, 1976). This
166 analysis returned positive and statistically significant (95% confidence interval) values only at for
167 lag = 1, suggesting a substantial absence of auto-correlation beyond what is expected for time
168 series with smooth temporal dynamics such as derived on a 3-month cumulative data precipitation
169 and soil moisture.

170 The 300 maps (12 months × 25 years) for the two standardized datasets are then spatially
171 interpolated on a common Lambert azimuthal equal-area (LAEA) projection on a regular grid of
172 5-km using the nearest neighbor algorithm. This is done to preserve the high-resolution
173 information of the soil moisture and by considering the smooth spatial dynamics of precipitation
174 accumulated over 3 months.

175 **2.2 Copula families**

176 The introduction of copulas in multivariate probability modeling has provided to hydrologists a
177 flexible tool to reproduce the joint probability of multiple dependent variables characterized by a

178 variety of marginal distributions (De Michele and Salvadori, 2003; Salvadori and De Michele,
179 2004).

180 Limiting the focus on bivariate variables, the joint probability distribution, F , of two
181 random variables (X_1 and X_2) can be expressed, thanks to the Sklar's theorem, as:

$$182 \quad F(x_1, x_2) = C(F_1(x_1), F_2(x_2)) \quad (1)$$

183 where F_1 and F_2 are the marginal distribution of X_1 and X_2 , respectively, and C is the copula
184 function (Salvadori et al., 2007).

185 A large variety of parametric formulations has been introduced in the literature to explicitly
186 link the marginal [distributions](#) to the joint [probability distributions](#), with some of the most common
187 copula families used in hydrology belonging to the Elliptical and Archimedean copulas (Chen and
188 Guo, 2019). Two [non-parametric](#) measures of dependence play a major role in parametric copula
189 inference. The Kendall rank correlation coefficient (τ) is commonly used as a [non-parametric](#)
190 measure of overall ordinal association, while the so-called Tail-Dependence (TD) [coefficients](#); (
191 Salvadori et al., 2007) [are](#) used to estimate the asymptotical degree of dependence in the upper
192 and lower extremes (upper tail-dependence, λ_U , and lower tail-dependence, λ_L , respectively). [The](#)
193 [estimation/assessment of the non-parametric TD non-parametrically is not an easy task, as](#)
194 [highlighted by Serinaldi et al. \(2015\), as it aims at assessing an reconstructing an asymptotic](#)
195 [behavior/value starting from a finite sample. Several formulations are proposed in the scientific](#)
196 [literature \(see Frahm et al., 2005\), and The non-parametric values of both TDs can be evaluated](#)
197 [following the method proposed by Schmidt and Stadtmueller \(2006\) is here used to evaluate/obtain](#)
198 [the non-parametric values/estimates of both TD coefficients.](#)

199 In this study, the parametric bivariate probability of [standardized](#) precipitation and soil
200 moisture is assessed [by](#) using the R package "VineCopula" (Aas et al., 2009; Dissman et al., 2013).
201 The Akaike Information Criterion (AIC, Stoica and Selen, 2004) is used to select, for each spatial
202 [grid](#) cell, the best fitting copula among the wide range of families available [in the package](#). The
203 main properties of some relevant copulas are reported in Table 1, as they will be useful to interpret
204 the successive results.

205 In particular, from the data in Table 1 it is important to highlight how the BB7 copula is a
206 combination of [the](#) Joe and Clayton [copulas](#), of which it inherits the tail-dependences, and how the
207 TD behavior of a copula can be inverted (i.e., the upper tail-dependence can become the lower and
208 *vice versa*) by simply considering the reciprocal marginals (commonly known as rotated forms,

identified by the suffix 180). [Information from both non-parametric and parametric approaches are here jointly used to discriminate between different TD behaviors.](#)

Table 1. Main copulas analyzed in this study and their ~~coefficients for the~~ upper and lower tail-dependence ~~coefficients~~ (λ_L and λ_U , respectively).

Copula	λ_L	λ_U
Gaussian	0	0
Student-t	$2t_{v+1}\left(-\sqrt{v+1}\sqrt{\frac{1-\rho}{1+\rho}}\right)$	$2t_{v+1}\left(-\sqrt{v+1}\sqrt{\frac{1-\rho}{1+\rho}}\right)$
Gumbel	0	$2 - 2^{\frac{1}{\theta}}$
Clayton	$2^{-\frac{1}{\theta}}$	0
Joe	0	$2 - 2^{\frac{1}{\delta}}$
BB7	$2^{-\frac{1}{\theta}}$	$2 - 2^{\frac{1}{\delta}}$

Even if a copula is selected as the optimal based on the AIC, this does not necessarily exclude that other copulas may perform similarly. For this reason, we introduced a further test

based on the relative likelihood criterion (Burnham and Anderson, 2002), $\exp\left(\frac{AIC_{\min} - AIC_i}{2}\right)$, to establish the likelihood that an AIC value of a given copula (AIC_i) is ~~statistically~~-significantly different than the minimum value (AIC_{\min}) obtained for the optimal solution.

2.3 Random forest classification of selected copulas

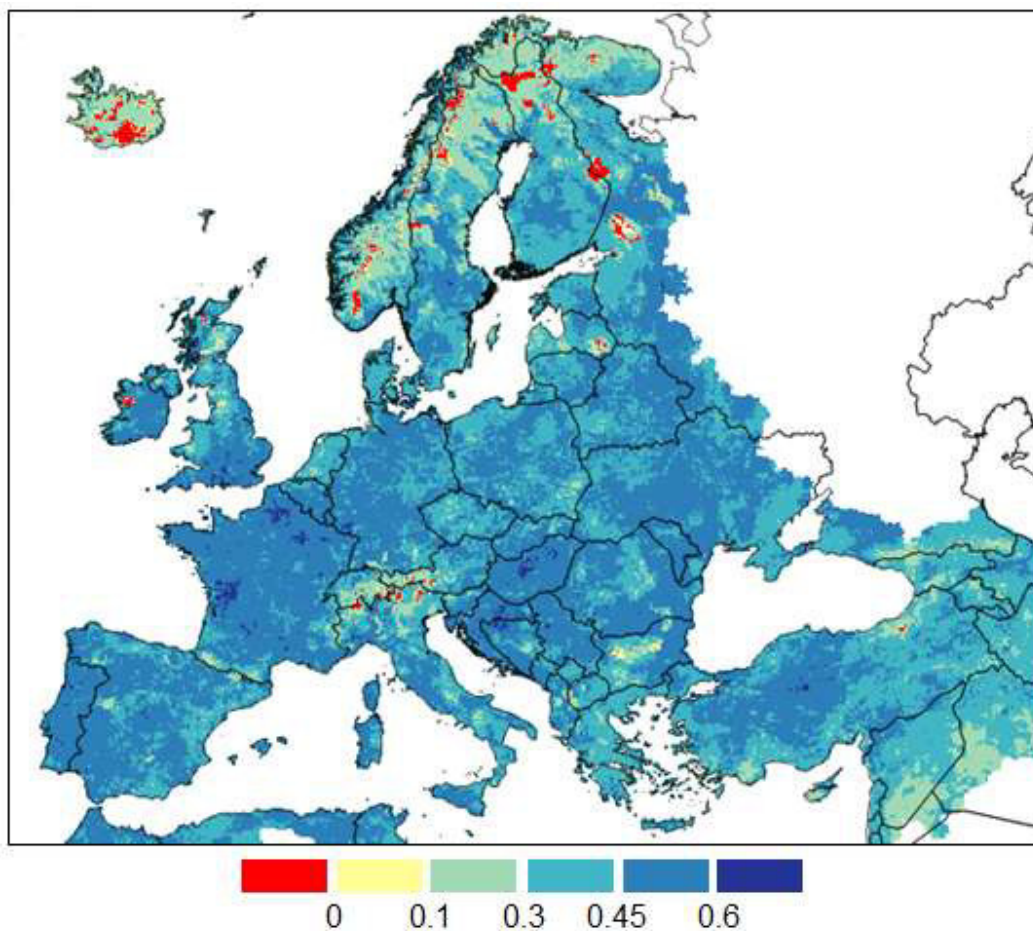
The interpretation of the selected copula functions may help highlighting the transferability of the observed results over different contexts. For this reason, the observed spatial distribution [of the selected copulas](#) is analyzed through a random forest classifier (Breiman, 2001), in order to find evidences of reproducible patterns beyond simple chance.

As input features we consider a set of commonly available variables, such as: ground elevation, annual average temperature, annual total precipitation, precipitation seasonality (ratio

228 between total precipitation in warm and cold months), annual average Normalized Difference
229 Vegetation Index (NDVI), annual average soil moisture, and soil type. As hyperparameters for the
230 random forest, we tuned the number of trees (ntree) and the number of features randomly sampled
231 at each split (mtry) using the “randomForest” R package (Breiman, 2001).

232 3. Results

233 A preliminary analysis of the degree of correlation between the monthly ~~empirical frequencies~~
234 ~~of standardized~~ 3-month precipitation and soil moisture (~~analogous to non-parametric SPI-3 and~~
235 ~~SMA~~) is tested on the full timeseries of each grid cell using the ~~non-parametric Kendall rank~~
236 ~~correlation coefficient (also known as~~ Kendall’s τ), as depicted in Fig. 1 for the entire European
237 domain.



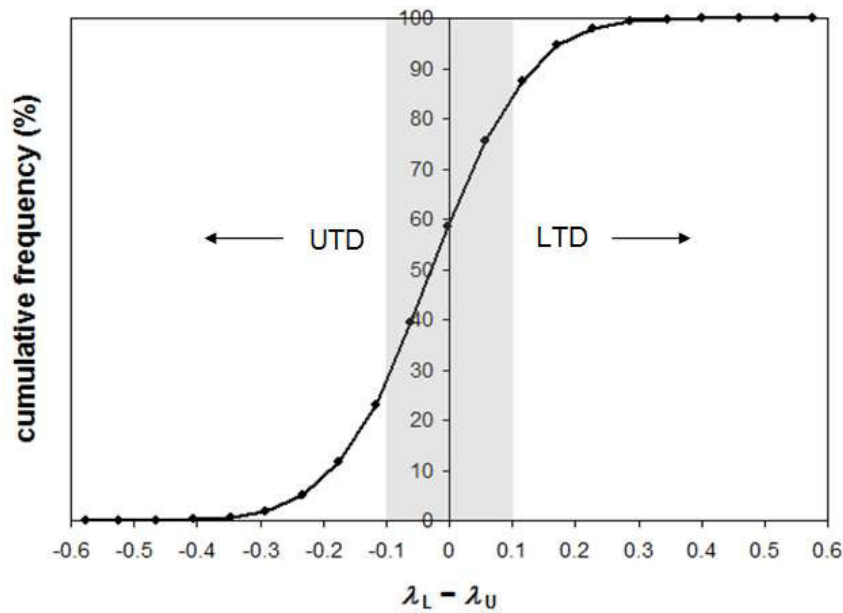
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239 **Fig. 1.** Spatial distribution of the Kendall’s τ between monthly ~~empirical frequencies~~
240 ~~of standardized~~ 3-month precipitation and soil moisture. Roughly, values lower than 0.1 are not
241 statistically significant at $p = 0.05$ (two-tails).
242

243 The results reported in Fig. 1 confirms the expected direct relation between the two
244 ~~variables~~quantities, with a relatively homogeneous distribution of medium/high ~~(between 0.3 and~~
245 ~~0.5)~~ correlation τ values ~~between 0.3 and 0.5~~ ($\tau = 0.42 \pm 0.1$). Limited regions with low (and
246 sometimes even slightly negative) τ values are sporadically observed, mostly ~~concentrated~~ over
247 the Alps, Iceland and the coldest regions of the Scandinavia peninsula. ~~Low~~ ~~C~~correlations over
248 these regions are likely ~~affected~~ ~~related to~~by the presence of snow coverage during extended
249 periods of the year. Overall, the observed τ values cannot be considered statistically significant (at
250 $p = 0.05$) only for less than 2% of the domain.

251 The analysis of the non-parametric tail-dependence values is summarized in the plot
252 depicted in Fig. 2, where the cumulative frequency of the difference between the empirical λ_L and
253 λ_U values is reported. ~~The range of TD values in Fig. 2 for which it is possible to exclude significant~~
254 ~~asymmetry in the tail dependence coefficients~~ ~~Symmetric behaviors in Fig. 2 can be~~ identified
255 by setting a maximum value for $|\lambda_L - \lambda_U|$. To ~~identify~~ ~~define~~ this threshold, ~~the~~ non-parametric TD
256 ~~coefficients~~ were re-computed on shuffled time series (to artificially reconstruct conditions of null
257 ~~tail~~ ~~dependencies~~), and the $|\lambda_L - \lambda_U|$ value corresponding to a cumulative frequency of 90% of the
258 grid cells after the shuffling was detected as threshold, corresponding to ~~a value of~~ 0.1. ~~This value~~
259 ~~can be seen as a lower limit to identify symmetric dependence.~~

260 The plot in Fig. 2 highlights how the majority (about 50%) of the grid cells can be
261 considered characterized by a symmetric behavior in the tail-dependence ~~coefficients according to~~
262 ~~the above mentioned criterion~~ ($|\lambda_L - \lambda_U| < 0.1$), whereas the rest of the grid cells are almost equally
263 split between a predominance of the Upper Tail-Dependence (UTD, corresponding to negative
264 differences) or a predominance of Lower Tail-Dependence (LTD, positive differences).

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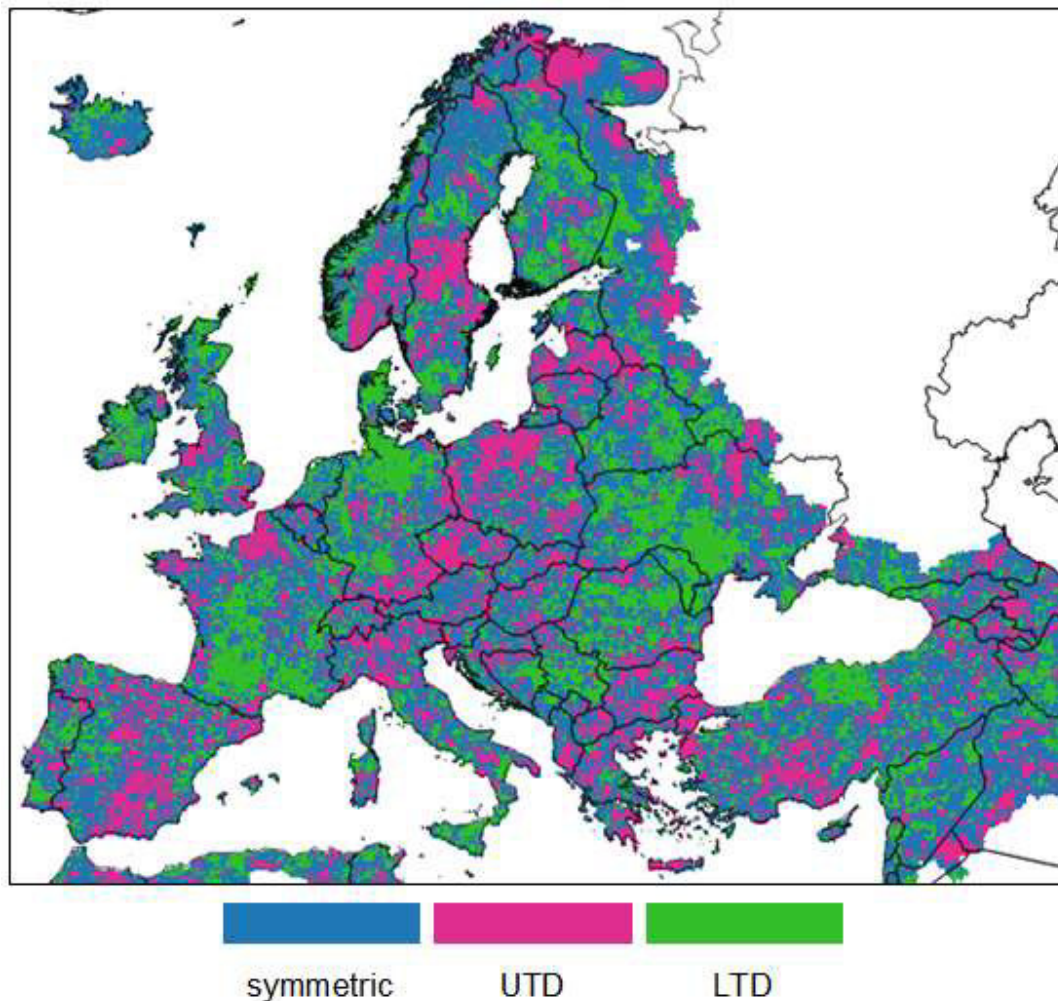


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267 **Fig. 2.** Analysis of the frequency of the empirical tail-dependence [coefficients](#). The plot shows the
 268 cumulative frequency distribution of the differences between the empirical λ_L and λ_U values
 269 computed according to Schmidt and Stadtmueller (2006). The domain with a roughly symmetric
 270 behavior ($|\lambda_L - \lambda_U| < 0.1$) is highlighted by the grey box area.
 271

272

The results reported in Fig. 2 were used to divide the entire domain in three categories
 273 (symmetric, LTD and UTD) as depicted in Fig. 3. This map shows evidence of some coherent
 274 spatial patterns, such as the predominance of LTD in southern France, southern Italy, northern
 275 Germany and Denmark, and western Ukraine (among others), and a clustering of UTD in Poland,
 276 Czechia-Republic, southern Scandinavia, and Greece. The symmetric condition seems overall
 277 more spread across the entire domain, also thanks to the higher frequency, with a slightly
 278 predominance over northern Europe (i.e., northern Scandinavian peninsula and Iceland).
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Fig. 3. Spatial distribution of the three categories derived from the differences in the empirical tail-dependence coefficients.

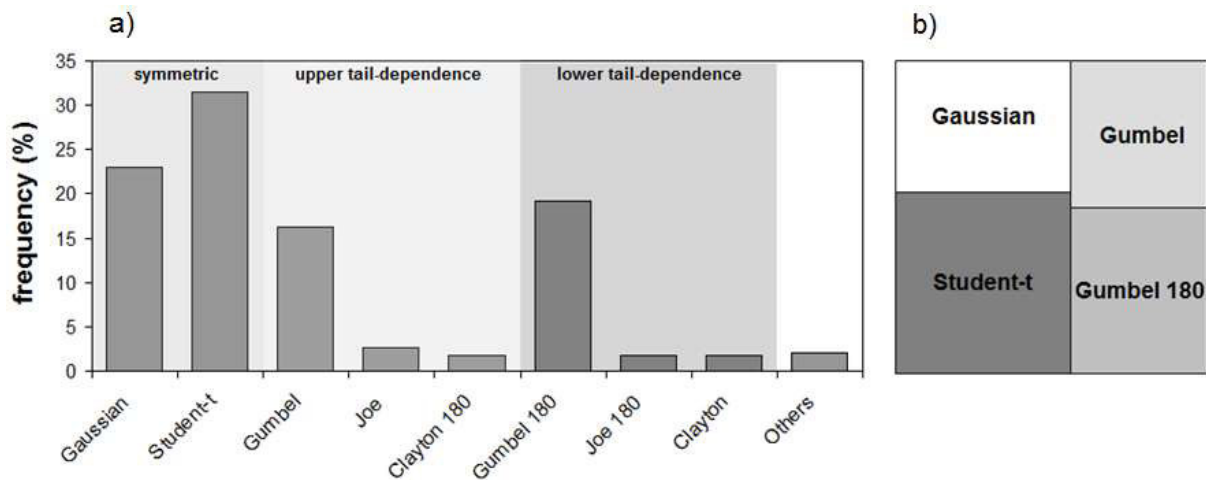
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Given the results ~~observed in terms of~~ the tail-dependence assessment, ~~is it~~ is useful to focus the copula parametric analysis on the capability to reproduce such patterns instead of finding the single copula that can perform reasonably well over the entire domain. Indeed, the search for the optimal copula based on the minimum AIC returns the BB7 as the optimal onesolution in about 80% of the domain (not shown). This result is ~~thanks a consequence of~~ the BB7 flexibility of its formulation (being derived from a combination of two purely asymmetric functions), which allows reproducing both symmetric and asymmetric tail-dependence coefficients according to ~~depending on~~ the values assumed by the two parameters. However, the fact that a single flexible copula works well over a large range of conditions may hide the key spatial patterns observed in the TD analysis. These patterns may be better reproduced, which can be highlighted instead by

294 adopting a limited number of more specialized common copulas ~~specialized in reproducing~~
 295 specific behaviors.

296 By limiting the search to a subset of copula functions, comprising only purely symmetric
 297 or purely asymmetric tail behaviors, more interesting results are obtained, as summarized by the
 298 frequency plot in Fig. 4. The grid cells where symmetric tail behavior copulas are selected as
 299 optimal are about 55% of the domain (see Fig. 4b), with a predominance of Student-t copula but
 300 also with a non-negligible fraction of cells (23%) where the Gaussian (symmetric and without tail-
 301 dependences) is chosen (see Fig. 4a). The remaining grid cells are almost equally split between
 302 upper and lower tail-dependences, with Gumbel (and its rotated counterpart, Gumbel 180) as the
 303 most selected among the asymmetric options.

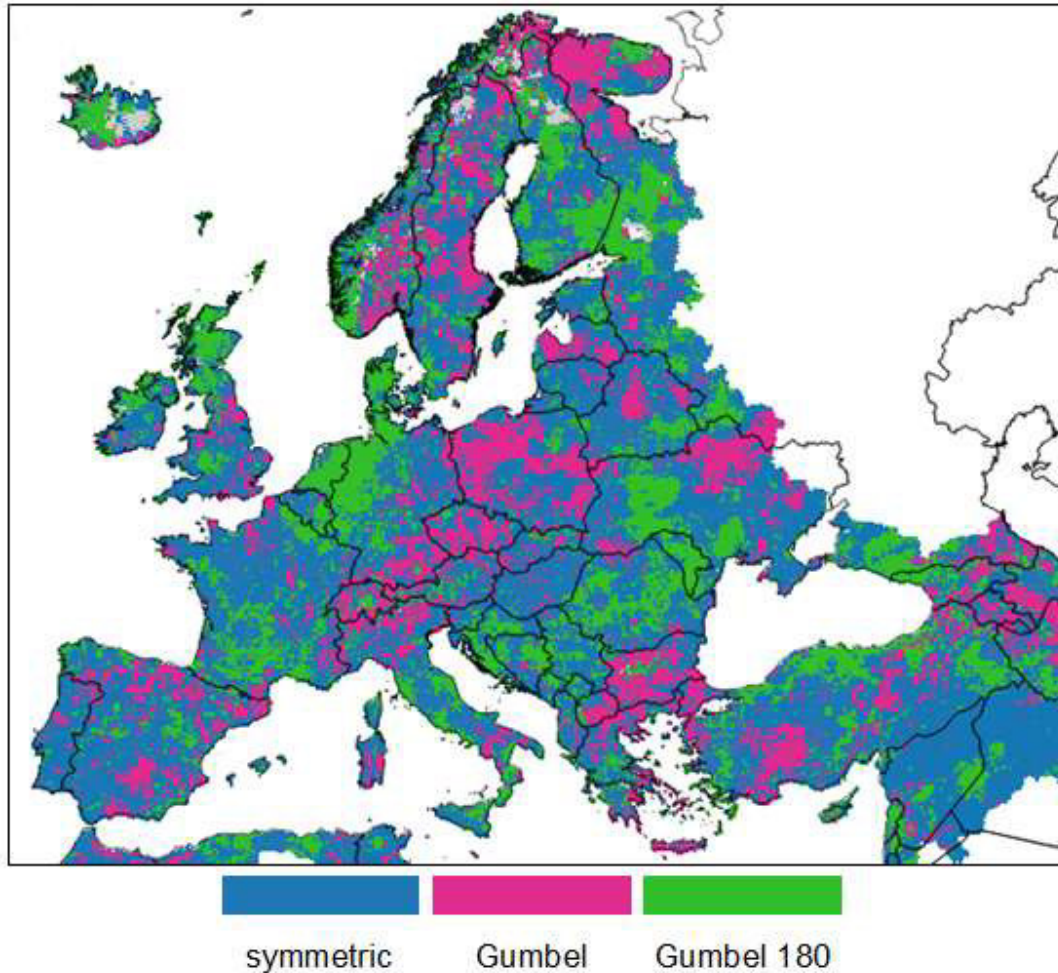
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305 **Fig. 4.** Frequency of the optimal copulas based on the minimum AIC. The barplot in panel a)
 306 shows the frequency offer for each copula, while increas the box in panel b) reports a synthetic compact
 307 description of the subdivision of the entire domain among the 4 most frequent copulas.
 308
 309

310 The spatial distribution of these optimal copulas (Fig. 5) confirms mostly agree with of the
 311 patterns observed in Fig. 3, supporting the findings as a converge of evidence on the TD spatial
 312 distribution of TD coefficients of the regions where either of the two TD is predominant. In
 313 addition, this result further confirmsation that a rather limited range of simple copula functions is
 314 able to capture the overall dynamics of dependence between precipitation and soil moisture over
 315 the entire European domain. Despite the observed spatial clusters in the obtained optimal copulas,
 316 the overall patterns observed in Fig. 5 are still rather noisy and may be difficult to interpret. This
 317 erratic behavior can be partially explained by the fact that different copulas may perform quite

318 similarly over some grid cells, hence the AIC of the optimal copula (AIC_{\min}) may not differ
319 significantly from the AIC of other functions.

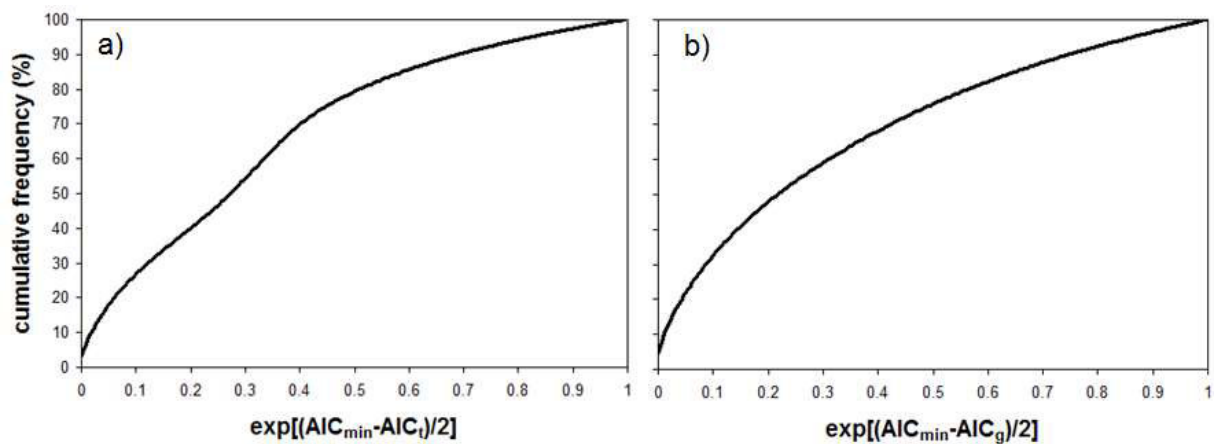


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321 **Fig. 5.** Spatial distribution of the optimal copulas obtained by minimizing the AIC. The symmetric
322 tail behavior class includes both Gaussian and Student-t copulas.
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324 To further investigate this hypothesis, we evaluated the possibility to replace the optimal
325 copulas with either a Student-t or a Gumbel (direct and rotated) over the entire domain. The
326 Gaussian copula was excluded from this analysis under the assumption that the no tail-dependence
327 of the Gaussian can be adequately reproduced by the Student-t with a small enough tail-
328 dependence. The plots in Fig. 6 reports the relative likelihood for the Student-t (panel a) and
329 Gumbel families (panel b) compared to the locally selected optimal copulas. Low values of this
330 metric correspond to conditions where the optimal copula cannot be replaced by the alternative
331 function (being either the Student-t or the Gumbel).

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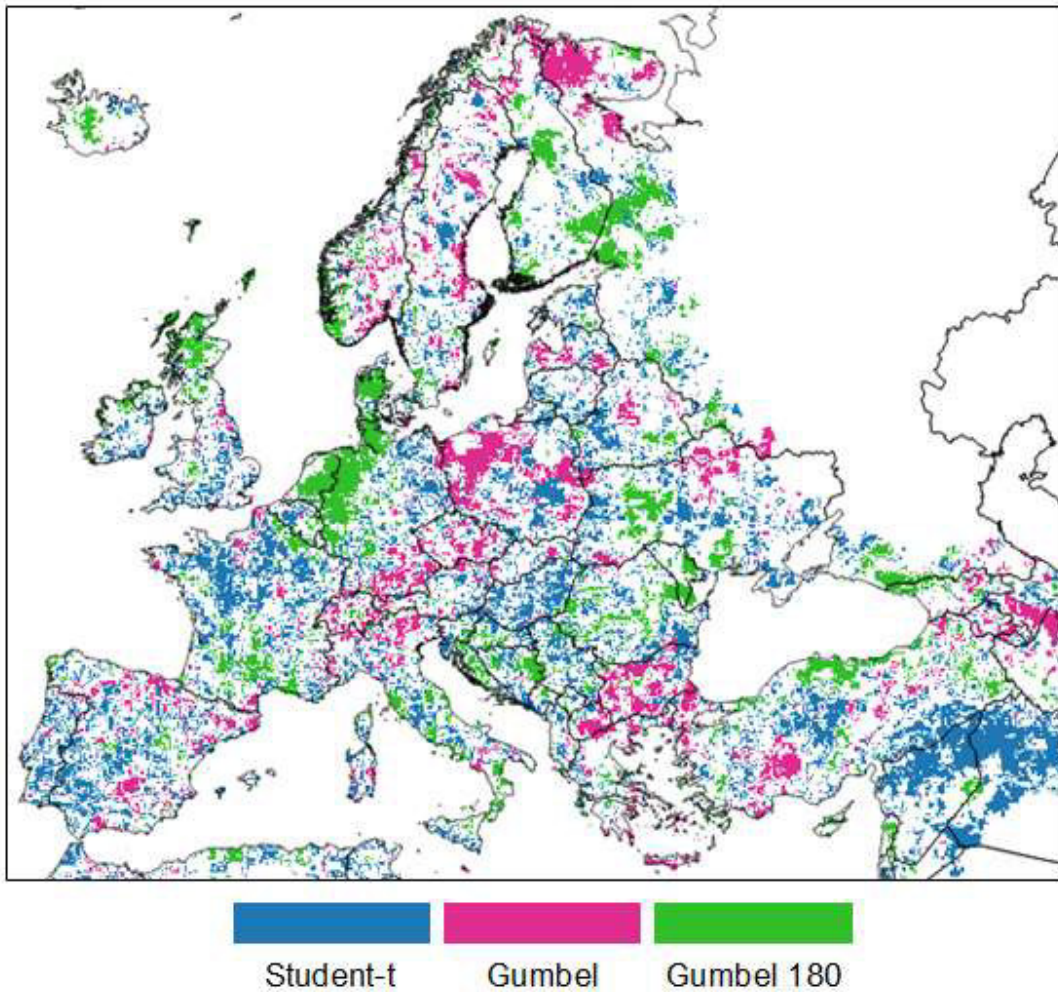
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334 **Fig. 6.** Frequency analysis of the relative likelihood computed between the optimal AIC (AIC_{\min})
335 and: a) Student-t (AIC_t), or b) Gumbel (AIC_g) families. The grid cells where either the Student-t
336 or the Gumbel was already the optimal solution were excluded from the respective frequency
337 analysis.

338

339 The results in Fig. 6 show that, if we assume a relative likelihood of 0.1 as a threshold to
340 detect a statistically significant difference, the Student-t cannot reasonably replace the local
341 optimal copula in about 18% of the entire domain (Fig. 6a), whereas this fraction is about 17% for
342 the Gumbel family (Fig. 6b). ~~It is also possible to observed that how~~ ~~It emerges that~~ the Gumbel
343 family is the optimal ~~once copula~~ in almost the totality (about 99%) of the grid cells where the
344 Student-t is not a suitable replacement of the local optimal, whereas almost only symmetric copulas
345 (63% Student-t and 34% Gaussian) are the optimal functions where the Gumbel family is not a
346 suitable replacement. Overall, these results suggest that the selection of the optimal copula is
347 “univocal” (i.e., cannot be reasonably replaced by another function) in about 35% (18+17) of the
348 domain, whereas either the Student-t or the Gumbel families can be adopted in the remain fraction
349 of the domain with similar performances in terms of AIC (and no clear TD behavior). ~~It is worth~~
350 ~~mentioning how t~~ This analysis also confirms the assumption that all the areas where the Gaussian
351 was chosen as optimal copula can be satisfactory modeled also by using the Student-t (i.e. without
352 a statistically significant increase in AIC).

353



354

355 **Fig. 7.** Spatial distribution of the grid cells where the selection of the optimal copula is “univocal”
 356 according to the relative likelihood criterion.
 357

358

359 The “univocal” areas derived from the previous analysis are mapped in Fig. 7, highlighting
 360 some of the more consistent spatial clusters already observed in both Figs. 3 and 5, as well as a
 361 large fraction of cells in northern Europe where a “univocal” optimal copula cannot be selected.
 362 These grid cells with “univocal” copula are used as a starting point for the random forest
 363 classification, given the robustness in their signal, [and the agreement in the outcome of both
 parametric and non-parametric TD behaviors.](#)

364

365 A sample [corresponding to](#) 25% of the “univocal” grid cells (~~corresponding to~~ about 8%
 366 of the entire domain) was used to train the random forest, adopting a number of trees (ntree) of 80
 367 and a single feature randomly sampled at each split (mtry = 1). The training size and [the](#) minimum
 values of hyperparameters were chosen to reduce the problem of overfitting. Among the possible

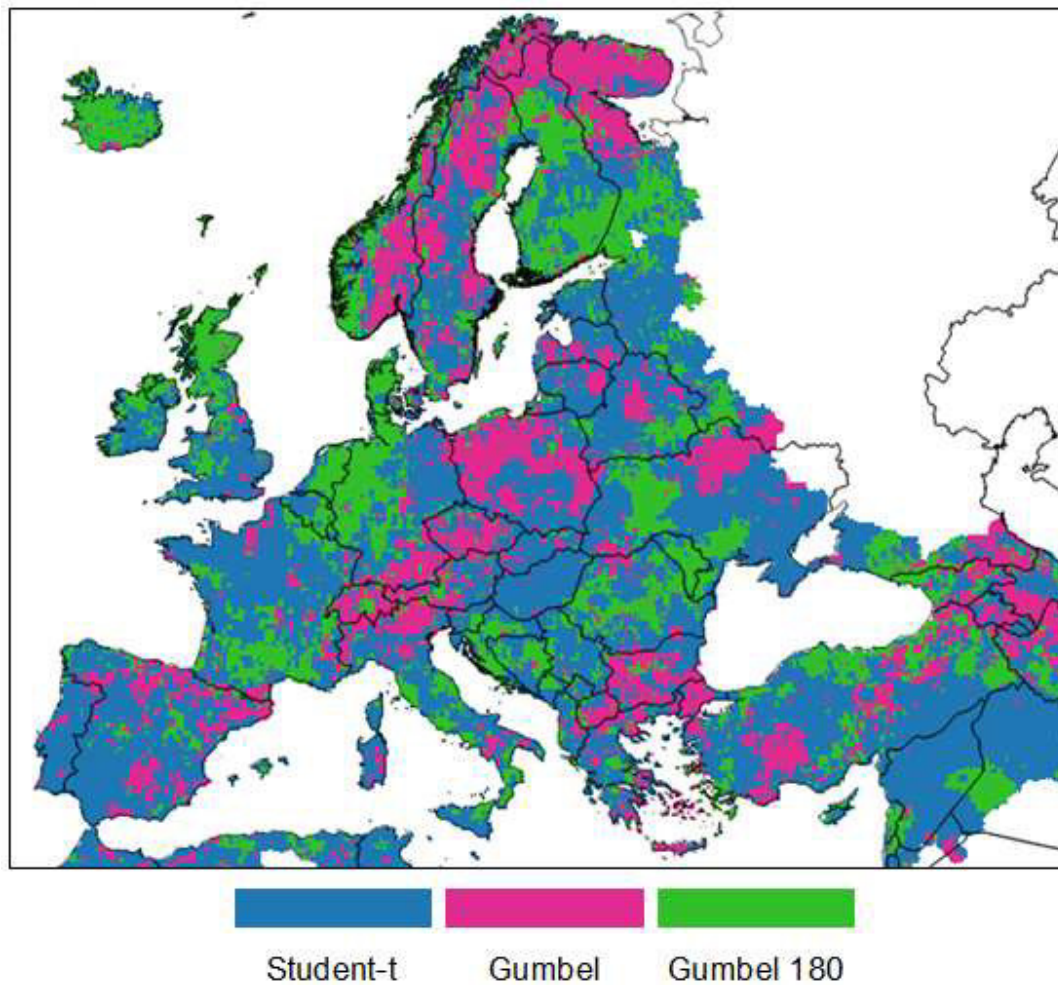
368 features, three variables were selected by analyzing the variable importance plots, as well as the
 369 ease of access: annual average temperature, annual total precipitation, and precipitation
 370 seasonality. The trained classifier was then applied to the testing subset (the remaining 75% of the
 371 “univocal” grid cells) and the outcomes were analyzed by mean of a confusion matrix, which
 372 results are summarized in Table 2. Overall, the obtained classification has a very satisfactory
 373 matching with the test subset, with a general high accuracy (ACC = 0.86) and with all the metrics
 374 pointing toward a significant improving in the performance compared to the reference No-
 375 Information-Rate (NIR) (i.e., small p-values) and a high probability to have correct modeled values
 376 compared to simple chance (i.e., high Cohen’s *K*).

377

378 **Table 2.** Summary of the confusion matrix analysis applied to the trained random forest on the
 379 testing subset.

Accuracy (ACC)	0.86
No-Information-Rate (NIR)	0.50
p-value (ACC > NIR)	$< 2.2 \times 10^{-16}$
McNemar’s test p-value	3.44×10^{-5}
Cohen’s kappa statistic (<i>K</i>)	0.78

380



381

382 **Fig. 8.** Map of the optimal copula as modeled by the trained random forest classifier.

383

384 Finally, the trained classifier was [extended-applied](#) to the entire [domain-dataset](#) to obtain a
 385 classification of the [entire](#)-European domain in term of the expected optimal copula and [the](#)
 386 [corresponding](#) TD behavior. This map, reported in Fig. 8, [shows bears](#)-a strong resemblance to
 387 both the empirically-derived map in Fig. 3 and the optimal AIC fitting in Fig. 5. Beside this overall
 388 agreement, some notable discrepancies can be observed over northern Scandinavia and Iceland,
 389 two regions where low Kendall's τ and a small fraction of "univocal" selected copulas were
 390 already identified.

391

392 **4. Discussion**

393 The overarching goal of the study is to investigate the joint probability of two [standardized](#)
 394 variables aiming at capturing agricultural drought conditions, hence the overall agreement between

395 these two quantities is a fundamental prerequisite. ~~A The expected~~ direct relationship between
396 standardized 3-month cumulated precipitation and soil moisture is expected, ~~as since~~ both SPI-3
397 and SMA are similarly-used agricultural drought indices, and this can support the identification of
398 the most suitable set of copula families (Salvadori et al., 2007; Genest et al., 2007). This ~~behavior~~
399 direct relationship is overall confirmed by the positive Kendall's τ values ~~estimated~~observed over
400 most of the domain ($\tau = 0.42 \pm 0.1$). Moderately high correlation values ~~of were observed between~~
401 standardized precipitation and soil moisture were estimated also in other studies. Kwon et al.
402 (2018) ~~reporting~~ Pearson's r values between 0.4 and 0.6 for 55 stations in South Korea, albeit
403 with seasonal patterns; Gaona et al. (2022) ~~reported found~~ similar values over the Ebro basin with
404 both land-surface modeled and satellite soil moisture, and Sepulcre-Cantó et al. (2012) obtained
405 an average value of r of about 0.6 over nine stations across Europe.

406 Sehler et al. (2019) studied the correlation between remote sensing-based precipitation and
407 soil moisture, finding moderate correlation over southern Europe, and a weak (often not
408 significant) correlation in central Europe. However, central Europe is close to the upper limit of
409 the analyzed remote sensing products, which can explain such low performance. Limited
410 correlation even among different soil moisture products has been observed in northern Europe in
411 other studies (Almenda-Martín et al., 2022), confirming the difficulty to model ~~the~~ soil moisture
412 dynamics over this region.

413 The obtained values for the Kendall's τ fall in a somewhat optimal range for the analysis
414 of the joint probability, since ~~they τ values~~ are statistically significant almost everywhere (i.e.,
415 consistency in the produced outcomes of the two indices are to a certain degree consistent) but not
416 too high to make meaningless any joint use of the two datasets (i.e., the two indices are too similar
417 and provide the same information ~~products~~).

418 ~~Even more interesting is~~ The outcome of the tail-dependence analysis is even more
419 interesting, given the role that such metric quantity, ~~(and in particular the low tail)~~, plays in the
420 detection of ~~drought~~ extreme events (and in particular the low-tail for droughts). The TD
421 investigation is sometimes overlooked in the development of multivariate drought analyses indices,
422 where previous studies often focused on optimizing the copula to the local data without analyzing
423 the implicit assumption on the TD, the consistency with the empirical non-parametric TD, and the
424 implications of the associated dependencies ~~for for the modeling of drought conditions~~. ~~Indeed,~~
425 Previous references in the scientific literature for studies on the joint probability of precipitation

426 ~~and soil moisture are rather scarce, and~~ TD is rarely the focus of ~~extensive such~~ analyses ~~or, at~~
427 ~~least, limited to specific areas and/or conditions not to, such an extend as the one results reported~~
428 ~~in this study for the entire Europe, and previous references in the scientific literature for~~
429 ~~precipitation and soil moisture are rather scarce.~~

430 As an example, Manning et al. (2018) performed a very detailed analysis over 11 FluxNet
431 sites in Europe on the role of precipitation and evapotranspiration on soil moisture drought, based
432 on pair copula constructions, but the authors did not provide any indication on which bivariate
433 copula was the optimal one for each site. Kwon et al. (2018) reported that Frank copula was the
434 most frequent optimal choice in their study over South Korea, ~~However, but~~ some clear spatial
435 patterns ~~were also~~ observed in their outcomes ~~were are not discussed~~, with Frank being the selected
436 copula mostly in the central area of the domain, but ~~also with~~ Gumbel and Student-t performing
437 the best in the southern and eastern coasts, respectively.

438 Dash et al. (2019) found Frank (among the Archimedean copulas) working the best for 3-
439 month precipitation and soil moisture over an Indian basin, ~~while~~ Hao and AghaKouchak (2013)
440 highlighted the good performance of Frank and Gumbel in five regions ~~of in~~ California, even if
441 neither Gaussian nor Student-t were considered. In all these applications, no specific
442 considerations on the ~~observed~~ TD behaviors were reported, even if a common trend seems to be
443 the good performance of Frank copula, ~~This which~~ is in contrast with our results, where the Frank
444 was very rarely selected as optimal (less than 1% of the domain). A possible explanation of these
445 results may be our focus on empirical marginal frequencies rather than theoretical ones, given the
446 well-documented increasing uncertainty in parametric fitting in the tails (Farahmand and
447 AghaKouchak, 2015; Laimighofer and Laaha, 2022). As a possible confirmation of this
448 hypothesis, a good performance of Gumbel and Gaussian has been observed over Iran by Bateni
449 et al. (2018), similarly to our results, when a non-parametric form for SPI and SSI (Standardized
450 Soil Moisture Index) was used.

451 The absence of a strict standard procedure to investigate tail-dependence may be another
452 factor affecting the limited focus on the topic ~~of in many studies on multivariate~~ drought
453 ~~studies indices~~. Non-parametric TD has the clear advantage to avoid any alteration of the data due
454 to the fitting procedure, but the outcomes in this study also show a high degree of spatial noise
455 likely due to the intrinsic nature of non-parametric analyses, the large uncertainty in non-
456 parametric methods (Serinaldi et al., 2015), as well as the effects of ~~to~~ the limited sample size

457 ~~which affects the estimates of TD.~~ (For this last issue see also the illustration 3.18 in Salvadori
458 et al. (2007). The threshold used here to define a symmetric behavior, based on a random shuffling
459 of the data, seems to successfully overcome the difficulty to define a self-consistent maximum
460 difference in TD values, but it cannot be seen as a reliable approach to easily identify TD symmetry
461 without the support of further evidence (e.g., by theoretical analyses or other means).

462 In this regard, ~~the~~ fitting of parametric copula functions returns ~~more consistent~~ spatial
463 patterns in ~~our study~~ TD coefficients similar to the ones obtained with the non-parametric approach.
464 ~~However, but evidence on~~ the absence of “univocal” fittings can be observed for large areas, as
465 well as some contrasting results compared to the non-parametric TD especially over northern
466 Europe (areas with low correlation). The grid cells where a given copula clearly outperforms the
467 alternative options is limited to roughly 1/3 of the domain, further stressing the evidence that clear-
468 cut outcomes are difficult to infer from a single methodology. ~~Thus~~ ~~In this regard~~, it seems
469 reasonable to state that infer that only a only a critical concerted analysis of both parametric and
470 non-parametric TDs can return robust practical indications based on a converge of evidence.

471 A clear outcome of our study is the predominance of regions with symmetric tail-
472 dependences coefficients, where the Student-t copula is suitable to reproduce the joint probability
473 of standardized precipitation and soil moisture. An even split of the remaining domain between
474 areas with either lower or upper tail-dependence ~~only~~ is also observed, where the Gumbel copula
475 (in either is direct or 180 rotated forms) is proven to be a suitable option. These results are crucial
476 in defining the role of standardized precipitation and soil moisture datasets in detecting drought
477 events, and to which extent they can work in synergy in a drought monitoring system. ~~In fact,~~
478 ~~w~~While the correlation between the two datasets highlights the extent of their overall agreement,
479 which in this study was somewhat uniform across most of the domain (τ ranging between 0.3 and
480 0.5), very different degrees of tail-consistency can be obtained for similar Kendall’s τ if the TDs
481 differ substantially. Regions with higher LTD will have a higher agreement in the detection of
482 drought extremes compared to the areas with a UTD predominance, hence a low number of false
483 alarm and a higher signal-to-noise ratio may be expected.

484 To further explore this behavior ~~stress on this issue~~, the time series of standardized variables
485 were converted in binary vectors based on a commonly used standardized drought threshold of
486 -1 (corresponding to an empirical frequency of 0.16). On these data, the pair-wise binary
487 correlation coefficient, $\rho(-1)$, was computed separately for the grid cells with LTD and UTD.

Results are shown reported in Fig. 9, for grid-cells with low ($0.1 < \tau \leq 0.4$, panel a) and high ($\tau > 0.4$, panel b) overall correlation, respectively. They show a net increase in the pairwise binary correlation for the grid cells with LTD (of about 0.15 in both cases) compared to the cases with UTD, even if the overall correlation is comparable. This increase in $\rho(-1)$ translates in a stronger agreement in the detection of extremes when a low tail-dependence is observed (i.e., less uncertainty), resulting in a more robust detection of the drought conditions thanks to the concurrency of extreme conditions in both drought indices (i.e., convergence of evidence).

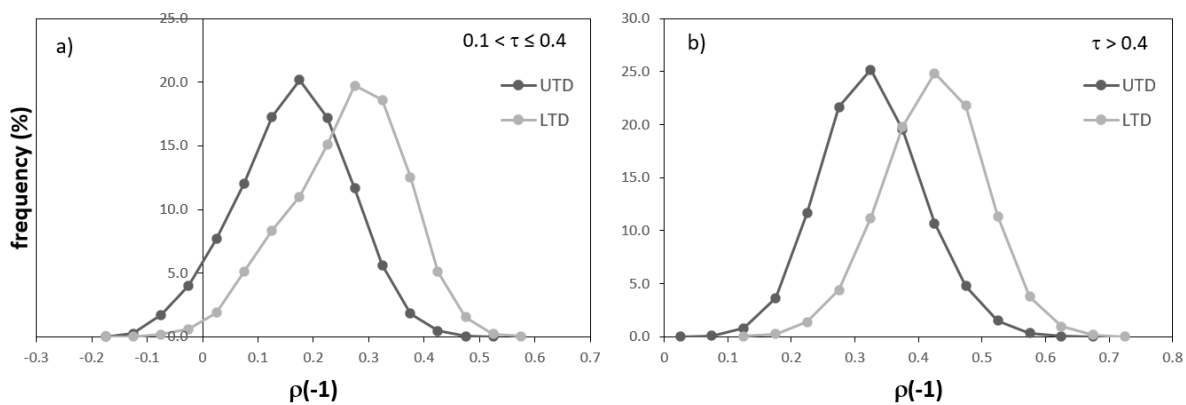


Fig. 9. Frequency distribution of the pairwise binary correlation between standardized precipitation and soil moisture lower than -1, computed separately for grid cells with UTD (dark grey lines) and LTD (light grey lines). Panel a) reports the results for the grid cells with low overall correlation ($0.1 < \tau \leq 0.4$), while panel b) reports the results for the grid cells with high correlation ($\tau > 0.4$).

In this regard, regions such as southern France, northern UK, northern Germany and Denmark, (where a strong LTD is observed, (see Fig. 8)), are appropriate candidates for a robust assessment of agricultural drought conditions based on a joint precipitation-soil moisture index, whereas some regions in central Europe (i.e., Poland, Czechia Republic, Switzerland) may not equally benefit from the use of a joint index due to the lower importance of absence of LTD.

Overall, the fact that the parametric copula fittings confirm most of the non-parametric TD patterns suggests that a parametric approach is suitable for an operational implementation of a precipitation-soil moisture joint drought index over most of Europe. This implies that, and that the proposed procedure, based on the combination of parametric and non-parametric analyses, can be considered as well as a reliable tool to provide meaningful insight on the potential application of joint probability as detector of extreme droughts.

514 ~~Even if, a~~At first glance, it may seem difficult to assign an ~~an~~ [meaningful](#) explanation for the
515 observed spatial patterns in LTD and UTD. ~~;~~ [However,](#) the proven possibility to reasonably
516 reconstruct these spatial patterns with a random forest classifier, starting from only a small sample
517 of robust training data (less than 10% of the domain) and with ~~commonly~~ [available](#) driving
518 features, suggests that the observed clusters are unlikely to be caused only by chance and that
519 hidden structures may be present and ~~may be~~ further explored. This result is encouraging for an
520 extension of the derived ~~approach~~[considerations](#) to other ~~spatial~~ regions of the world.

522 5. Summary and Conclusions

523 The use of combined indices based on copula seems a promising development in the field of
524 drought detection and monitoring. In this study, we analyzed the joint probability of two
525 ~~variables~~[quantities](#) commonly used ~~to—in agricultural derive~~ drought ~~indices~~[analyses](#); ~~the~~
526 [empirical frequencies of](#) 3-month cumulated precipitation and soil moisture. ~~We with a special~~
527 focus on the probabilistic characteristics ~~being that are~~ key for ~~their usage in~~ agricultural drought
528 [analyses](#)[studies](#).

529 The overall agreement in the marginal probability of the two [standardized](#) variables
530 suggests that they are indeed valid candidates for the development of a joint drought index over
531 the European domain. However, an in-depth analysis of the tail-dependence, derived with both
532 non-parametric and parametric approaches, shows some clear spatial patterns, which have direct
533 repercussion on the capability of such data to provide robust [and coherent](#) estimates of ~~the~~ [drought](#)
534 extremes. In this regard, regions such as southern France, northern UK, northern Germany, and
535 Denmark may benefit more from the joint use of the two [standardized](#) variables thanks to the
536 observed strong low tail-dependence ~~(i.e., increasing agreement on the left over tail extremes)~~.
537 The joint dependence of [standardized](#) precipitation and soil moisture is well reproduced ~~by~~ using
538 three ~~commonly used~~ copulas (Student-t, Gumbel and 180 rotated Gumbel), ~~which with~~ spatial
539 patterns [that](#) were successfully reconstructed with a random forest classification, suggesting the
540 presence of a structure in the outcomes not related to chance.

542 **Code availability:** The codes used for this analysis can be provided upon request via the
543 corresponding author.

545 **Data availability:** All the data used in this study can be accessed and retrieved through the
546 European Drought Observatory (EDO) web portal
547 (<https://edo.jrc.ec.europa.eu/gdo/php/index.php?id=2112>).
548

549 **Author contribution:** CC designed the experiments, with inputs from AT and CDM. CC
550 developed the codes and performed the analyses. CC prepared the manuscript, which was
551 expanded and revised by all co-authors.
552

553 **Competing interests:** At least one of the (co-)authors is a member of the editorial board of
554 Hydrology and Earth System Sciences.

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